# bbcontacts - Supplementary information 

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## S1 Supplementary methods

## S1.1 Datasets

BetaSheet916 (Cheng and Baldi, 2005) consists of 916 protein chains with an available X-ray structure of resolution below $2.5 \AA$. These chains contain $31,638 \beta$-residue contacts distributed into 4519 antiparallel $\beta$-strand contacts, 2214 parallel $\beta$-strand contacts and 1429 isolated $\beta$-bridges.
BetaSheet1452 (Savojardo et al., 2013) was built from the structures deposited in the Protein Data Bank after 2004 but using a procedure similar to the BetaSheet 916 building procedure. BetaSheet1452 involves $56,552 \beta$-residue contacts distributed into 3937 antiparallel $\beta$-strand contacts, 7892 parallel $\beta$-strand contacts and 2412 isolated $\beta$-bridges.

To build our training dataset, we extracted all CATH domains that did not belong to any of the fold groups identified in the test datasets in CATH v3.5. This set of 22,563 domains belonging to 864 fold groups was then filtered to reduce redundancy. For this purpose, we used the pdbfilter.pl script from the HH-suite (Remmert et al., 2011) with parameters -cov 0 -e 0.01 -id 0 (no sequence identity restriction for filtering, but the minimum Evalue between any two representative sequences is 0.01 and no minimum coverage was applied when discarding redundant sequences). Among the 1482 PDB domains in this redundancy-filtered dataset, 943 domains containing $\beta$-contacts form our training dataset (867 X-ray structures with resolution below $3.5 \AA$ and 76 NMR structures). These 943 domains contain 19,339 $\beta$-contacts: 2511 parallel $\beta$-contacts, 16,041 antiparallel $\beta$-contacts and $787 \beta$-bridges.
Because not all chains in BetaSheet916 and BetaSheet1452 were fully annotated in CATH v3.5, there might remain some redundancy between the training dataset and the test dataset. We verified that the results for BetaSheet916 and BetaSheet1452 did not deteriorate when the dataset was restricted to the subset of each dataset containing all chains fully annotated in CATH v3.5 (and thus non-redundant with the training dataset) (see section S2.1 below and Figure S1). In Figures S21, S22, S23, S24 and S25, we also show results for the training dataset and the test dataset BetaSheet1452.
Because bbcontacts relies on correlated mutations and thus predicts side-chain and not backbone contacts, the positions involved in $\beta$-bulges were adjusted to reflect the expected pattern: for a $\beta$-bulge between res1 and res2 (in one strand) and resX (in the other strand), all three side-chains must point in the same direction with respect to the plane formed by the $\beta$-sheet.

## S1.2 Data used for HMM training

To build the multiple sequence alignments, we started from sequences based on the ATOM records of the PDB files: this allowed us to have perfect matching between the structure,
the DSSP assignment and the first sequence of the MSA for each protein in the training dataset, while the two test datasets were unaffected by this choice as by construction they contain only proteins with no backbone interruption (Cheng and Baldi, 2005; Savojardo et al., 2013).

We first ran HHblits v2.0.15 (Remmert et al., 2011) against the uniprot20 database (dated March 2013), with options -all -maxfilt 100000 -realign_max 100000 -B 100000 -Z 100000, thus avoiding any filtering in order to retrieve as many homologous sequences as possible. We then performed a filtering step using HHfilter with options -id 90 -neff 15 -qsc -30 (each alignment is filtered down to $90 \%$ sequence identity).

The distribution of the number of sequences in the MSAs for the training dataset and the two test datasets is provided in Figure S2.

The secondary structure predictions were obtained with PSIPRED (Jones, 1999), as implemented in the addss.pl script from the HH-suite (Remmert et al., 2011). This means that the MSAs were first filtered down to Neff $\leq 7$ and that the procedure included fine-tuning of the secondary structure predictions with psipass2.

Direct coupling predictions were obtained with CCMpred (Seemayer et al., 2014) run with the default options, including initial sequence reweighting and final post-processing using the average-product correction (Dunn et al., 2008). The MSAs were not filtered to remove columns or rows with many gaps.

When building MSAs of reduced diversity for the training dataset, rather than sampling from the alignment at random, we ran HHfilter (Remmert et al., 2011) with different values of the qsc parameter, describing the entropy per column in the MSA. We tried different qsc values through a dichotomic search, until the filtered alignment contained the number of sequences expected for a given $\eta \in\{0.05,0.1,0.2, \ldots, 1.0,1.2\}$. For an initial MSA diversity value of $\eta_{0}$, diversity-filtered alignments can be obtained for each $\eta<\eta_{0}$.

The respective numbers of domains and numbers of $\beta$-contacts in each diversity-filtered dataset are given in Table S1.

Table S1: Number of domains (\#domains), number of parallel residue-residue $\beta$-contacts (\#parallel) and number of antiparallel residue-residue $\beta$-contacts (\#antiparallel) in each diversity-filtered dataset

| $\eta$ | 0.05 | 0.1 | 0.2 | 0.3 | 0.4 | 0.5 | 0.6 | 0.7 | 0.8 | 0.9 | 1.0 | 1.2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| \#domains | 805 | 621 | 384 | 235 | 155 | 103 | 68 | 48 | 35 | 23 | 18 | 10 |
| \#parallel | 2144 | 1606 | 1148 | 588 | 287 | 186 | 102 | 57 | 49 | 33 | 23 | 9 |
| \#antiparallel | 11888 | 8470 | 4549 | 2601 | 1675 | 962 | 656 | 409 | 281 | 208 | 160 | 84 |

## S1.3 HMM parameters

The HMM emission probability $e_{z}(i, j)$ for a given HMM state $z$ and a given position $(i, j)$ contains a product of two terms: one term based on couplings $e c_{z}(i, j)$ and one term based on secondary structure $\operatorname{ess}_{z}(i, j)$.

## S1.3.1 Coupling-based emissions.

The coupling-based part of the emission probability at position $(i, j)$ was expressed as the product of three odds-ratios relative to the background: one for the central coupling at position $(i, j)$ and one for each of the two couplings at the positions adjacent to $(i, j)$ that belong to the secondary diagonals of the pattern. This can be written as follows for the case of parallel $\beta$-strands, for any HMM state $z$ apart from the "start" and "end" states:

$$
\begin{equation*}
e c_{z}(i, j)=\frac{p\left(x_{i, j} \mid z, i, j\right)}{p\left(x_{i, j} \mid b g, i, j\right)} \frac{p\left(x_{i+1, j-1} \mid z, i, j\right)}{p\left(x_{i+1, j-1} \mid b g, i, j\right)} \frac{p\left(x_{i-1, j+1} \mid z, i, j\right)}{p\left(x_{i-1, j+1} \mid b g, i, j\right)} \tag{1}
\end{equation*}
$$

where $x_{i, j}$ denotes the coupling value at position $(i, j)$ and $b g$ denotes the background.
Still for the case of parallel $\beta$-strands, we fitted only 3 distributions:

- $p\left(x_{i, j} \mid z, i, j\right)$ for the main diagonal of the pattern,
- $p\left(x_{i+1, j-1} \mid z, i, j\right)=p\left(x_{i-1, j+1} \mid z, i, j\right)$ for the secondary diagonals of the pattern,
- $p\left(x_{i, j} \mid b g, i, j\right)=p\left(x_{i+1, j-1} \mid b g, i, j\right)=p\left(x_{i-1, j+1} \mid b g, i, j\right)$ for the background.

The case of antiparallel $\beta$-strands was treated similarly, except that the positions on the secondary diagonals are $(i+1, j+1)$ and $(i-1, j-1)$.

To fit the distributions, the diversity-filtered alignments were used. After centering each coupling distribution at zero by subtracting a shift parameter $x_{0}$, we fitted it using a combination of two transformed Gamma functions, one for positive coupling values $x_{+}$ and one for negative coupling values $x_{-}$. For instance, in the case of the main diagonal fit:

$$
p\left(x_{i, j} \mid z, i, j\right)= \begin{cases}f_{+}\left(x_{i, j}-x_{0}\right) & \text { if } x_{i, j} \geq x_{0}  \tag{2}\\ f_{-}\left(-x_{i, j}+x_{0}\right) & \text { if } x_{i, j}<x_{0}\end{cases}
$$

where

$$
\begin{align*}
& f_{+}: x_{+} \mapsto w_{+} \frac{b_{+}^{-\frac{1}{\alpha_{+}}}}{\Gamma\left(\frac{1}{\alpha_{+}}\right)} \frac{\alpha_{+}}{\beta_{+}} \exp \left(-\frac{\left(\frac{\left|x_{+}\right|}{\beta_{+}}\right)^{\alpha_{+}}}{b_{+}\left[1+\left(\frac{\left|x_{+}\right|}{\beta_{+}}\right)^{\alpha_{+}}\right]}\right)  \tag{3}\\
& f_{-}: x_{-} \mapsto\left(1-w_{+}\right) \frac{b_{-}-\frac{1}{\alpha_{-}}}{\Gamma\left(\frac{1}{\alpha_{-}}\right)} \alpha_{-} \exp \left(-\frac{\left|x_{-}\right|^{\alpha_{-}}}{b_{-}}\right) \tag{4}
\end{align*}
$$

To describe each of the transformed Gamma fits for a given value of $\eta$, we thus need 7 parameters: the shift $x_{0}$ needed to center the coupling distribution, the relative weight of the positive and negative sides $w_{+}$, plus $b_{-}$and $\alpha_{-}$for negative couplings and $b_{+}, \alpha_{+}$and $\beta_{+}$for positive couplings.

To describe the dependency of the coupling distributions on $\eta$, we expressed each parameter as a function of the alignment diversity $\eta$ : first, the shift was fitted as a quadratic function of $\eta$ by linear regression and the fitted shift was subtracted from all coupling values. All remaining parameters were expressed as linear functions of $\eta$.

The optimization was performed using a global maximum likelihood estimation with the BFGS algorithm, using multiple initialization values for the parameters ( 32 for the background fit and over 1,000 for the signal fits). We performed each round of optimization using the training data (predicted coupling matrices) generated for all 12 values of $\eta$. The training was performed after local background correction of the coupling matrices with $S=10$, but the resulting parameters were found to be very similar when the matrices without local background correction were used for training. We picked the fit with the maximum likelihood over all runs. The final fits are illustrated in Figure S4.

The final number of parameters for the coupling-based emission probabilities is thus 90: 2 (parallel/antiparallel) times 3 (background, main-diagonal signal, secondary-diagonal signal) times 15 ( 6 parameters with a linear dependency on $\eta$ and one parameter with a quadratic dependency).

Because the coupling densities are fitted on cases with alignments from a limited range of $\eta$ values, and because the parameters are expressed as linear or quadratic functions of $\eta$, the formulae obtained for each parameter as a function of $\eta$ do not give acceptable results when $\eta$ is either too small or too large (for instance, the $b$ parameters can become negative, which is unacceptable). On the other hand, we observed that for very low values of $\eta$, the coupling distributions derived from true $\beta$-contact patterns are almost superimposed with the background distributions (the noise overcomes the signal) and for high values of $\eta$, the distributions do not change much. Therefore, we set boundaries on the values of $\eta$ as follows:

- if $\eta<0.022$, only the secondary structure part of the emission probabilities is used (we assume that there is no coupling signal so that the odds-ratios in the couplingbased emissions are always equal to 1 ),
- if $\eta>1.98$, we set $\eta=1.98$ for calculating the coupling-based emissions.

In addition, because the fits are not perfect representations of the coupling densities, and because the range of coupling values observed in the training dataset is limited, the fits can display unexpected behaviors for relatively low or relatively high coupling values. Typically, for low couplings, $p\left(x_{i, j} \mid z\right) / p\left(x_{i, j} \mid b g\right)$ can become larger than 1 , which is unexpected. For high couplings, $p\left(x_{i, j} \mid z\right) / p\left(x_{i, j} \mid b g\right)$ can become smaller than 1 , which is also unexpected. To avoid such problems, we set (for both the main-diagonal and secondary-diagonal oddsratios):

$$
\left\{\begin{array}{l}
\frac{p\left(x_{i, j} \mid z\right)}{p\left(x_{i, j} \mid b g\right)}=1 \text { if } \frac{p\left(x_{i, j} \mid z\right)}{p\left(x_{, j} \mid b g\right)}>1 \text { and } x_{i, j}<x_{0}(b g) \\
\frac{p\left(x_{i, j} \mid z\right)}{p\left(x_{i, j} \mid b g\right)}=1 \text { if } \frac{p\left(x_{i, j} \mid z\right)}{p\left(x_{i, j} \mid b g\right)}<1 \text { and } x_{i, j}>x_{0}(\text { maindiag })+0.1
\end{array}\right.
$$

where $x_{0}(b g)$ is the fitted shift for the background distribution and $x_{0}$ (maindiag) is the fitted shift for the main diagonal positions. This has a very minor effect on the vast majority of our results but makes bbcontacts more robust.

## S1.3.2 Secondary-structure-based emissions.

Here, we distinguished between different HMM states (denoted by $z$ ) because they exhibit a different behaviour with respect to observed secondary structure predictions.

We tested two different versions of the secondary-structure-based emissions $\operatorname{ess}_{z}(i, j)$.
The first version (hereafter called "non-conditional" for simplicity) was based on the probability $p\left(\sigma_{i}, \sigma_{j} \mid z\right)$ of observing a pair of secondary structure states $\left(\sigma_{i}, \sigma_{j}\right)$ in state $z$ at position $(i, j)$ :

$$
\begin{equation*}
e s s_{z}(i, j)=\frac{p\left(\sigma_{i}, \sigma_{j} \mid z\right)}{p\left(\sigma_{i}, \sigma_{j} \mid b g\right)} \tag{5}
\end{equation*}
$$

The second version (hereafter called "conditional") was based on the probability $p\left(\sigma_{i}, \sigma_{j} \mid \sigma_{i_{\mathrm{prev}}}, \sigma_{j_{\mathrm{prev}}}, z\right)$ of observing a pair of secondary structure states $\left(\sigma_{i}, \sigma_{j}\right)$ in state $z$ at position $(i, j)$ given that we additionally observed secondary structure states $\left(\sigma_{i_{\text {prev }}}, \sigma_{j_{\text {prev }}}\right)$ at the previous position ( $i_{\text {prev }}, j_{\text {prev }}$ ), where

$$
\left(i_{\text {prev }}, j_{\text {prev }}\right)= \begin{cases}(i-1, j-1) & \text { for the parallel case }  \tag{6}\\ (i-1, j+1) & \text { for the antiparallel case }\end{cases}
$$

In order to reduce the number of parameters necessary to describe the secondary-structure-based "conditional" emission probabilities, we used the following factorization:

$$
\begin{equation*}
p\left(\sigma_{i}, \sigma_{j} \mid \sigma_{i_{\mathrm{prev}}}, \sigma_{j_{\mathrm{prev}}}, z\right)=p\left(\sigma_{i} \mid \sigma_{i_{\mathrm{prev}}}, z\right) p\left(\sigma_{j} \mid \sigma_{j_{\mathrm{prev}}}, z\right) \tag{7}
\end{equation*}
$$

We verified that the numerical difference between the initial and the factorized form was small.

The "conditional" secondary-structure-based emissions can thus be expressed as:

$$
\begin{equation*}
e s s_{z}(i, j)=\frac{p\left(\sigma_{i} \mid \sigma_{i_{\mathrm{prev}}}, z\right)}{p\left(\sigma_{i} \mid \sigma_{i_{\mathrm{prev}}}, b g\right)} \frac{p\left(\sigma_{j} \mid \sigma_{j_{\mathrm{prev}}}, z\right)}{p\left(\sigma_{j} \mid \sigma_{j_{\mathrm{prev}}}, b g\right)} \tag{8}
\end{equation*}
$$

We also added pseudocounts derived from the non-conditional probability distribution to the conditional probabilities, i.e. we replaced

$$
\begin{equation*}
p\left(\sigma_{i} \mid \sigma_{i_{\mathrm{prev}}}, z\right)=\frac{N_{\sigma_{i}, \sigma_{\mathrm{iprev}}, z},}{N_{\sigma_{\mathrm{iprev}}, z}} \tag{9}
\end{equation*}
$$

by a term including $N_{0}$ counts from the non-conditional distribution. The resulting emission probabilities are:

$$
\begin{equation*}
e s s_{z}(i, j)=\frac{\frac{N_{\sigma_{i}, \sigma_{i_{\mathrm{prev}}}, z}+N_{0} \frac{N_{\sigma_{i}, z}}{N_{z}}}{N_{0}+N_{\sigma_{i_{\mathrm{prev}}}, z}}}{\frac{N_{\sigma_{i}, \sigma_{i_{\mathrm{prev}}}, b g}+N_{0} \frac{N_{\sigma_{i}, b g}}{N_{b g}}}{N_{0}+N_{\sigma_{i_{\mathrm{prev}}}, b g}}} \frac{\frac{N_{\sigma_{j}, \sigma_{j_{\mathrm{prev}}}, z}+N_{0} \frac{N_{\sigma_{j}, z}}{N_{z}}}{N_{0}+N_{\sigma_{j_{\mathrm{prev}}}, z}}}{\frac{N_{\sigma_{j}, \sigma_{j_{\mathrm{prev}}}, b g}+N_{0} \frac{N_{\sigma_{j}, b g}}{N_{b g}}}{N_{0}+N_{\sigma_{j_{\mathrm{prev}}}, b g}}} \tag{10}
\end{equation*}
$$

where the $N_{\ldots}$.. terms represent counts observed in the training dataset (e.g. $N_{\sigma_{i}, z}$ is the number of counts observed for secondary structure state $\sigma_{i}$ in state $z$ and $N_{\sigma_{i}, \sigma_{i_{\text {prev }}}, z}$ is the number of counts observed for secondary structure state $\sigma_{i}$ and previous secondary state $\sigma_{i_{\text {prev }}}$ in state $z$ ).

The number of pseudocounts $N_{0}$ was optimized on the training dataset, as illustrated in Figure 57 .
Because we found that the secondary structure states for the coupling matrix cells situated immediately before and immediately after a $\beta$-strand interaction contain information about the likelihood to start and end this interaction, there is also a secondary-structurebased emission term in the Viterbi initialization step ("start" state) and in the Viterbi termination step ("end" state). The term in the initialization step is always a non-conditional probability, for a clean termination of the chain-rule based product of emissions:

$$
p\left(\sigma_{i_{\text {end }}} \mid \sigma_{i_{\text {end }}-1}, z\right) * \ldots * p\left(\sigma_{i} \mid \sigma_{i_{\mathrm{prev}}}, z\right) * \ldots * p\left(\sigma_{i_{\mathrm{start}}+1} \mid \sigma_{i_{\mathrm{start}}}, z\right) * p\left(\sigma_{i_{\mathrm{start}}} \mid z\right)
$$

This initialization term can also be seen as a prior based on secondary structure.
For each situation (DSSP-based emissions or PSIPRED-based emissions), the number of parameters is therefore 415: the number of pseudocounts $N_{0} ; 216$ parameters for the
non-conditional probabilities ( 2 directions times 12 states ( 11 HMM states plus the background) times 9 possible combinations of ( $\left.\sigma_{i}, \sigma_{j}\right)$ ); and 198 parameters for the conditional probabilities ( 2 directions times 11 states ( 11 HMM states minus the start state plus the background) times 9 possible combinations of $\left(\sigma_{i}, \sigma_{i_{\text {prev }}}\right)$ ). Note that for the DSSP-based case, many of these emission parameters are 0 or 1 since $\beta-\beta$ contacts can only be detected between residues assigned as "E" by DSSP.

## S1.3.3 Prior probability distribution depending on sequence separation.

We introduced a prior for starting a $\beta$-strand interaction depending on the sequence separation between the first pair of interacting residues. The prior contains explicit probabilities to have a contact starting at a sequence separation of up to 12 ; the probability is then modeled as a linear function of the sequence separation between 13 and 20 and as an exponentially decreasing function starting from a sequence separation of 21 :

$$
\operatorname{prior}(i-j)= \begin{cases}p_{i-j} & \text { if }|i-j| \leq 12  \tag{11}\\ l_{1}+|i-j| * l_{2} & \text { if } 13 \leq|i-j| \leq 20 \\ e_{1}+e_{2} * e^{\frac{-i-j \mid}{e_{3}}} & \text { if }|i-j| \geq 21\end{cases}
$$

The corresponding 17 parameters $\left(p_{1}, p_{2}, \ldots, p_{12}, l_{1}, l_{2}, e_{1}, e_{2}\right.$ and $\left.e_{3}\right)$ were trained independently for parallel and antiparallel $\beta$-contacts. In addition, for the DSSP-based predictions, the training was limited to regions of the coupling matrix where both residues belong to a $\beta$-strand, because no HMM path can be detected outside of these regions. The fitted parameters are shown in Figure 55 .
We also introduced constraints to prevent decoding in regions of the coupling matrix too close to the diagonal. For this purpose, we always "mask" a region around the diagonal, i.e. we set the emission probabilities to 0 for all positions $(i, j)$ in this region and for all states. This region contains all pairs of positions with a sequence separation of up to (and including) 1 for antiparallel contacts and 6 for parallel contacts.

## S1.4 HMM decoding

The local Viterbi algorithm consists of four major steps: initialization, recursion, termination and back-tracing.

In the initialization step, the Viterbi variables $V[i, j$, start $]$ are initialized for all positions $(i, j)$ in the coupling matrix. Because the coupling matrices are symmetric, only positions where $i>j$ can receive non-zero Viterbi scores $V[i, j, z]$ for $z \notin\{$ start, end $\}$. To make the implementation easier, when going from the start state to the first state in a $\beta$-contact, we take a step in the coupling matrix:

$$
\begin{cases}(i, j) \rightarrow(i+1, j+1) & \text { for the parallel case } \\ (i, j) \rightarrow(i+1, j-1) & \text { for the antiparallel case }\end{cases}
$$

This means that $V[i, j$, start $]$ also has to be initialized for $i \in\{-1,0, \ldots, L\}$ and $j=-1$ for the parallel case and for $i=j$ for the antiparallel case (where $L$ is the protein length).
In principle, all positions should receive an initial probability of 1 , but the priors described above (secondary structure prior and prior depending on sequence separation) are also applied during this initialization step:
$V[i, j$, start $]= \begin{cases}\operatorname{prior}(i-j) * p\left(\sigma_{i_{\text {start }}} \mid \operatorname{start}\right) * p\left(\sigma_{j_{\text {start }} \mid} \mid \text { start }\right) & \text { for the parallel case } \\ \operatorname{prior}(i-j+2) * p\left(\sigma_{i_{\text {start }}} \mid \text { start }\right) * p\left(\sigma_{j_{\text {start }}} \mid s t a r t\right) & \text { for the antiparallel case }\end{cases}$

The reason for having $\operatorname{prior}(i-j+2)$ in the antiparallel case is that we take a step in the coupling matrix when going from the start state to the first state in a $\beta$-contact, so the sequence separation (measuring the distance to the diagonal in the coupling matrix) is unchanged for the parallel case and increased by 2 for the antiparallel case.

In the recursion step, all probabilities $V[i, j, z]$ for $z \notin\{$ start, end $\}$ are calculated using the transition and emission probabilities:

$$
\begin{equation*}
V[i, j, z]=e_{z}(i, j) \max _{k}\left(V\left[i_{\text {prev }}, j_{\text {prev }}, k\right] * t[k][z]\right) \tag{13}
\end{equation*}
$$

where $k$ is any of the HMM states apart from the end state, $t[k][z]$ is the transition probability from state $k$ to state $z$, and the previous Viterbi score $V\left[i_{\text {prev }}, j_{\text {prev }}, k\right]$ is taken from position

$$
\left(i_{\text {prev }}, j_{\text {prev }}\right)= \begin{cases}(i-1, j-1) & \text { in the parallel case if } z \notin\left\{\text { bulgei }_{1}, \text { bulgei } i_{2}, \text { bulge } j_{1}, \text { bulge }_{2}\right\} \\ (i, j-1) & \text { in the parallel case if } z \in\left\{\text { bulgei } i_{1}, \text { bulgei } i_{2}\right\} \\ (i-1, j) & \text { in the parallel case if } z \in\left\{\text { bulge } j_{1}, \text { bulge } j_{2}\right\} \\ (i-1, j+1) & \text { in the antiparallel case if } z \notin\left\{\text { bulgei } i_{1}, \text { bulgei }{ }_{2},{\text { bulge } \left.j_{1}, \text { bulge } j_{2}\right\}}\right. \\ (i, j+1) & \text { in the antiparallel case if } z \in\left\{\text { bulgei } i_{1}, \text { bulgei }_{2}\right\} \\ (i-1, j) & \text { in the antiparallel case if } z \in\left\{\text { bulge } j_{1}, \text { bulge }_{2}\right\}\end{cases}
$$

In the termination step, the $V[i, j$, end $]$ probabilities are calculated for all positions where $i>j$. Like with the start state, to make the implementation easier, we take a step in the coupling matrix between the last $\beta$-contact and the end state, so that the $(i, j)$ position corresponding to the end state is not part of the final path. The formula for $V[i, j, e n d]$ includes the classical maximum over states where the path can end and a secondary-structure-based emission term, as mentioned above (section S1.3.2):

$$
\begin{equation*}
V[i, j, \text { end }]=e s s_{\mathrm{end}}(i, j) \max _{k}\left(V\left[i_{\mathrm{prev}}, j_{\mathrm{prev}}, k\right] * t[k][\mathrm{end}]\right) \tag{14}
\end{equation*}
$$

In the recursion and termination steps, pointers are used to keep track of the most likely sequences of states.

The initialization, recursion and termination steps of the Viterbi decoding are performed separately for the parallel and antiparallel directions, but all Viterbi scores are then merged before the fourth and final back-tracing step. In the back-tracing step, all $V[i, j$, end $]$ probabilities (Viterbi scores) are sorted in decreasing order. The first (most likely) path, corresponding to the highest $V[i, j$, end $]$ probability, is retrieved by back-tracing through the saved pointers and saved. Then, we cross out a region corresponding to a "corridor" around this path (we cross out all residue pairs belonging to the path, plus all residue pairs within $\pm 3$ residues of those belonging to the path) in the Viterbi matrix corresponding to the direction of the path (parallel or antiparallel), i.e. we do not take into account any more Viterbi probabilities for this region and this direction. This avoids retrieving many suboptimal versions of a contact between the same $\beta$-strands. The next path that does not contain any crossed-out residues is then saved and a region around this path is crossed-out.

We proceed iteratively in this manner until we reach a given Viterbi score threshold. This threshold is chosen to be low enough that the precision-recall curve shows only a precision drop after this threshold and no more gain in recall, but not too low for computational efficiency. It is adjusted depending on the parameters used to run bbcontacts (DSSP or PSIPRED-based predictions, PSM triggering).

For numerical stability, all probabilities are expressed in logarithmic space.

## S1.4.1 Prediction-shortening mode (PSM)

In practice, we apply a decrease by 0.3 per PSM iteration in all log-scale transition probabilities, except for the transitions to the end state that are used to maintain the sum of transition probabilities leaving any state equal to 1 . If paths exceeding a length of 50 are predicted, then the decrease in all log-scale transition probabilities is 0.6 per iteration to speed-up the PSM process. A maximum of 20 iterations is also set to limit the runtime of bbcontacts when PSM gets triggered.

## S1.5 Evaluation

Residue-level evaluation is straightforward in all cases: a pair of residues predicted as a $\beta$-contact (i.e. belonging to one of the accepted paths) is counted as a true positive if it is actually a $\beta$-contact (defined by DSSP) and as a false positive if it is not. False negatives are all the true $\beta$-contacts which have not been predicted above a given Viterbi threshold. If bbcontacts predicts a contact that actually corresponds to a $\beta$-bridge as part of a parallel or antiparallel path, then the contact is counted as a true positive at the residue level and at the orientation-independent strand level.

Strand-level evaluation is only straightforward for DSSP-based results, because in this case a given predicted path will contain residues belonging to exactly one strand on each side. For PSIPRED-based results, strand-level evaluation is performed in the following manner. If a predicted path contains interactions between residues belonging to more than one pair of $\beta$-strands, then each pair of strands predicted to be in contact is counted in the strand-level evaluation. If a predicted path contains (on one or both sides) only residues that are not part of a $\beta$-strand, then this path is counted as a false positive in the strand-level evaluation. Finally, if a predicted path contains on both sides a mixture of residues contained in $\beta$-strands and other residues, the interactions between residues that are not part of a strand are ignored in the strand-level evaluation. Because PSIPREDbased strand-level evaluation is based on these additional criteria, it is provided only in an indicative manner and the residue-level evaluation forms the solid basis for comparison between different versions of our method.

## S2 Supplementary results

## S2.1 Verification that the results are not affected by any potential redundancy between training and test datasets

Because not all domains from all protein chains contained in the test datasets BetaSheet916 and BetaSheet 1452 were annotated in CATH v3.5, we need to make sure that the bbcontacts performance is not over-estimated due to over-training. For this purpose, we evaluated bbcontacts (DSSP-based predictions and PSIPRED-based predictions without and with PSM) on the subsets of the test datasets which are fully annotated in CATH v3.5 (and thus non-redundant with the training dataset), i.e. all protein chains in each test dataset for which all domains are annotated in CATH v3.5.

These subsets contain 873 out of 916 chains for BetaSheet 916 and 403 out of 1452 chains for BetaSheet1452. The difference in the proportion of annotated chains between the two test datasets comes from the fact that BetaSheet916 was published in 2005, while BetaSheet1452 was built from PDB structures deposited later than May 2004, so that many structures in BetaSheet1452 are too recent to have been annotated in CATH v3.5.
Because we built the training dataset by taking domains not belonging to any of the folds (CATH Topologies) observed in the annotated protein chains of the test datasets, we are sure that these subsets do not have any redundancy with the training dataset.

In Figure S1, we see that the bbcontacts performance on the subset of BetaSheet916 is almost identical to the performance on the full test dataset; the final recall for PSIPREDbased predictions is even slightly higher. The bbcontacts performance on the subset of BetaSheet1452 is slightly better than the performance on the full dataset for both DSSPbased and PSIPRED-based predictions.

Therefore, we can be confident that the method is not over-trained, because when we remove all chains that are potentially redundant with the training dataset from the evaluation, the performance of bbcontacts is maintained or even slightly increased.

## S2.2 Results for the training dataset

It must be noted that the training dataset is rather different in composition from the test datasets: because it is built from all CATH v3.5 annotated domains not contained in the BetaSheet916 and BetaSheet1452 datasets, it contains many protein domains with few $\beta$-residues, low resolution or missing residues.

## S2.2.1 Influence of the number of pseudocounts in the secondary-structure-based emissions

We looked at the influence of the number of pseudocounts from the non-conditional distribution added to the conditional probabilities for the definition of the secondary-structurebased emissions.

The results are displayed in Figure S7. A number of pseudocounts of 10,000 was chosen as it gives the best precision-recall compromise on the training dataset. A number of pseudocounts of 100,000 gives similar results, with slightly higher initial precision and slightly lower final recall.

## S2.2.2 Choice of Viterbi score threshold for F1-score evaluation

The threshold for calculating F1-scores on the test datasets was chosen as the Viterbi score giving the maximum residue-level F1-score on the training dataset. The evolution of the

F1-score when including predictions with decreasing Viterbi score is shown in Figure S16 for the DSSP-based predictions and the PSIPRED-based predictions without and with PSM. A vertical line marks the chosen threshold: 1.7 for DSSP-based predictions and -1.6 for PSIPRED-based predictions.

## S2.2.3 Contribution of the different terms in bbcontacts and final results

The precision-recall plots shown in the main text for BetaSheet916 (main Figures 3, 4a and 4 b ) are shown for the training dataset in Figures S21 and S22.

These figures show that the trends, choices and conclusions discussed in the main text for BetaSheet 916 results also hold for results obtained on the training dataset. They also show that the performance of bbcontacts on the training dataset is not higher than on the test datasets, which is a sign that our method is not overfitted.

## S2.3 Additional results for BetaSheet916

Comparison between Figure 3a in the main text and Figure S9(a) shows that the effect of changing the secondary-structure-based emissions is very different for DSSP and PSIPRED-based results. For the DSSP case, the probabilities for HMM (non-background) states are essentially unaffected by any of the changes, as they simply reflect the fact that $\beta$-contacts can only occur between two $\beta$-residues. On the other hand, background probabilities are strongly affected by the change from non-conditional to conditional. Adding pseudocounts from the non-conditional distribution to the conditional probabilities almost does not affect the background, as discussed above, and this explains why the blue and purple lines are superimposed in Figure S9(a). For PSIPRED-based results, adding pseudocounts from the non-conditional distribution to the conditional probabilities is a good compromise that improves the performance of bbcontacts.

Comparison between Figure 3b in the main text and Figure $\mathrm{S9}$ (b) also shows a difference in the impact of local background correction on DSSP-based results compared to PSIPRED-based results. The DSSP-based predictions are only impacted by darker regions if the corresponding couplings occur between two $\beta$-strands, while in the PSIPRED-based case, strong couplings cause the coupling-based emission probabilities to overtake the secondary-structure-based emissions, so that $\beta$-contacts can be predicted even in a region where the secondary structure composition is highly unfavorable.
For an easier comparison with the results from previous papers, Tables S2 and S3 provide recall, precision and F1-scores at the residue level and at the strand level, on the BetaSheet916 dataset. The "SS source" column specifies whether true or predicted secondary structure was used as an input. In these tables, the results for bbcontacts are given for a Viterbi score threshold corresponding to the threshold giving the maximum residue-level F1-score on the training dataset (see above, section S2.2.2). The results for all methods except bbcontacts are taken from Savojardo et al. (2013). In Table S3, the column " $\mathrm{F} 1 \geq 70 \%$ " shows the percentage of chains in the BetaSheet916 dataset that have an F1-score higher than $70 \%$ at the strand level (correct $\beta$-strand pairing).

## S2.3.1 Comparison of bbcontacts with BCov* and CMM*

BCov (Savojardo et al., 2013) uses PSICOV (Jones et al., 2012) to generate direct coupling matrices, but it has been shown that pseudo-likelihood-based methods give better precision (Kamisetty et al., 2013). CMM (Burkoff et al., 2013) uses a different correlated mutation measure which has not been assessed in terms of general contact prediction performance.

Table S2: Residue-level performance on the BetaSheet916 dataset
(the largest value in each column is highlighted in bold)

| Method | SS source | Recall (\%) | Precision (\%) | F1-score (\%) |
| :---: | :---: | :---: | :---: | :---: |
| bbcontacts | PSIPRED | 47.3 | 54.8 | 50.7 |
| bbcontacts+PSM | PSIPRED | 47.2 | 55.0 | 50.8 |
| bbcontacts | DSSP | $\mathbf{6 0 . 9}$ | $\mathbf{6 9 . 4}$ | $\mathbf{6 4 . 8}$ |
| BCov6 | DSSP | 43.9 | 42.4 | 43.1 |
| BCov | DSSP | 42.4 | 40.9 | 41.6 |
| CMM | DSSP | 44.0 | 44.0 | 44.0 |
| MLN-2S | DSSP | 42.7 | 47.3 | 44.9 |
| MLN | DSSP | 39.3 | 46.1 | 42.4 |
| BetaPro | DSSP | 44.1 | 38.0 | 40.8 |

Table S3: Strand-level performance (correct $\beta$-strand pairing) on BetaSheet916 (the largest value in each column is highlighted in bold)

| Method | SS source | Recall (\%) | Precision (\%) | F1-score (\%) | F1 $\geq 70 \%$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| bbcontacts | PSIPRED | 48.2 | 81.1 | 60.5 | 39.7 |
| bbcontacts+PSM | PSIPRED | 48.3 | 79.8 | 60.1 | 39.5 |
| bbcontacts | DSSP | 57.6 | $\mathbf{8 3 . 7}$ | $\mathbf{6 8 . 3}$ | $\mathbf{5 5 . 9}$ |
| BCov | DSSP | $\mathbf{6 2 . 0}$ | 59.5 | 60.7 | 44.2 |
| CMM | DSSP | 55.0 | 61.0 | 58.0 | 35.0 |
| MLN-2S | DSSP | 59.8 | 58.4 | 59.1 | 36.2 |
| MLN | DSSP | 55.5 | 59.8 | 57.6 | 33.7 |
| BetaPro | DSSP | 59.7 | 53.1 | 56.2 | 31.7 |

However, almost $80 \%$ of the original CMM alignments for BetaSheet916 have less than 1000 sequences, as opposed to $27 \%$ for the original BCov alignments and $34 \%$ for the alignments used in the present paper (compare Figure 1 in Burkoff et al. (2013) with Figure 3 in Savojardo et al. (2013) and Figure S2 in the present paper). Therefore, we can expect than by using better contact predictions as an input, the performance of BCov and CMM should improve.

A comparison of the DSSP-based residue-level performance of bbcontacts, BCov, CMM, BCov* and CMM* is shown in Figure S13. BCov* and CMM* correspond to using coupling matrices predicted with CCMpred as well as DSSP assignments as inputs to the $\beta$-topology prediction algorithms of BCov and CMM.

For $\mathrm{BCov}^{*}$ and $\mathrm{CMM}^{*}$, we do not apply local background correction to the coupling matrices, because local background correction does not contribute a lot to DSSP-based results (see Figure S9). The default BCov parameters are used, in particular the minimum sequence separation of 6 for parallel strand pairing. For $\mathrm{CMM}^{*}$, the recommended parameters that were used in the original publication (Burkoff et al., 2013) are used for sampling ( 50 resets, 1 million samples for each reset).

Figure S13 shows that BCov* displays intermediate results between BCov and bbcontacts. The precision-recall curve for bbcontacts displays a better robustness for highconfidence contacts than $\mathrm{CMM}^{*}$, as the bbcontacts precision remains above $80 \%$ for recall up to almost $50 \%$. However, the $\mathrm{CMM}^{*}$ precision-recall curve displays a higher initial precision than the bbcontacts curve. This can be explained by several factors. First,

CMM was developed to make use of strong topological constraints for $\beta$-strand interactions, but some of these constraints rely heavily on the availability of the exact $\beta$-strand positions and we thus decided not to include them in bbcontacts. This is the case for instance of the specific treatment of DSSP E (or B) assignments of length 1, which always correspond to $\beta$-bridges, and of the explicit modelling of the number of residues with no $\beta$-partners at the end of $\beta$-strands (Burkoff et al., 2013). In addition, the CMM output contains a probability for each pair of $\beta$-residues to be in contact, while in bbcontacts, the final score is given to a path containing several $\beta$-contacts. Thus, contrary to CMM, bbcontacts cannot distinguish between central pairs of $\beta$-residues in a strand-strand contact and (less confident) pairs of $\beta$-residues close to strand extremities.

## S2.4 Results for BetaSheet1452

The precision-recall plots shown in the main text for BetaSheet916 (main Figures 3, 4a and 4b) are shown for the BetaSheet1452 test dataset in Figures $\operatorname{S23}$ and S24. In addition, the F1-scores for individual test cases depending on the number of MSA sequences are shown in Figure 525.
These figures show that the trends, choices and conclusions discussed in the main text for BetaSheet 916 results also hold for results obtained on BetaSheet1452. In particular, comparison with the results from previous methods BCov and CMM (obtained from Savojardo et al. (2013)) also shows that bbcontacts performs much better than these previous methods when using the DSSP assignment, and the residue-level precision and recall reached by bbcontacts when using PSIPRED predictions are higher than the precision and recall of BCov and CMM when these methods use DSSP assignments.

## S3 Supplementary figures

(a)


Figure S1: Comparison between residue-level performance evaluated on the full test dataset or on a subset of the test dataset for which all domains are annotated in CATH v3.5, thus making sure that there is no redundancy between this subset and our training dataset. (a) Test dataset BetaSheet916. The fully annotated subset contains 873 out of 916 chains (95\%). (b) Test dataset BetaSheet1452. The fully annotated subset contains 403 out of 1452 chains ( $28 \%$ ).
Further discussion of these results is provided in section S2.1.
(a)

(b)

(c)

Training dataset


Figure S2: Cumulative distribution of the number of sequences in the MSAs (filtered at $90 \%$ sequence identity) for (a) test dataset BetaSheet916, (b) test dataset BetaSheet1452 and (c) training dataset. In each plot, the intersection of the curve with the red vertical line marks the proportion of alignments with less than 1000 sequences: $34 \%$ for BetaSheet 916 , $25 \%$ for BetaSheet 1452 and $66 \%$ in the training dataset.


Figure S3: Illustration of the effect of the local background correction procedure with $S=10$ on two cases, both belonging to the BetaSheet916 dataset.
(a) Protein chain 1gygB (370 residues): the initial coupling matrix shows a darker region in the top-right corner; the local background correction has a strong effect on the predicted couplings; in particular, the coupling values for the top-right corner get strongly reduced. (b) Protein chain 1 p 9 yA ( 117 residues): the local background correction has a very mild effect on the values of the predicted couplings and does not change the overall appearance of the contact map; the visible patterns are not affected by the local background correction.


Figure S4: Coupling distribution densities and corresponding fits, for parallel (left) and antiparallel (right) $\beta$-contacts, for $\sqrt{N} / L \in\{0.1,0.3,0.5\}$


Figure S5: True distributions of the probability for starting an interaction between two $\beta$-strands, depending on sequence separation, and fits used for the prior depending on sequence separation. The two panels use two different scales, with the left panel focused on the large probabilities associated with antiparallel $\beta$-contacts at short sequence separation and the right panel focused on larger sequence separations. The fits contain three regions: up to a sequence separation of 12 , there is an explicit probability for each sequence separation; between 13 and 20, the fit is a linear function of sequence separation; starting at a sequence separation of 21 , the fit is an exponentially decreasing function of sequence separation.
This plot corresponds to the DSSP case (where the probabilities are calculated only for regions containing exclusively $\beta$-residues). In the PSIPRED case, the distributions have a similar shape but the probabilities are much lower because they are normalized over all possible pairs of residues within a protein.


Figure S6: An example where prediction-shortening mode (PSM) is necessary. The two contacts between the yellow and light-orange $\beta$-strands and between the green and dark-orange $\beta$-strands form patterns (highlighted in red boxes) that are very close in the coupling matrix and almost aligned. Therefore, without PSM, only one path is detected, which leads to the prediction of several false positive residue-residue contacts between the green-to-yellow linker (around positions 75-80) and the light-orange-to-dark-orange linker (around positions 105-110). When PSM is triggered, it shortens the predictions until two separate paths are detected.


Figure S7: Influence of the number of pseudocounts added from the non-conditional distribution to the conditional distribution, evaluated on the training dataset, using no prior depending on sequence separation, no local background correction of the coupling matrices and with PSM turned off. (left) DSSP-based predictions: all conditional distributions without pseudocounts or with any number of pseudocounts are superimposed. (right) PSIPRED-based predictions.
(a)


PSIPRED-based predictions


Figure S8: Influence of different model parameters on the residue-level performance of bbcontacts on the BetaSheet 916 dataset, using PSIPRED predictions as an input and using the reference version of bbcontacts mentioned in the main text (local background correction of the coupling matrices with $S=10$, conditional secondary-structure-based emissions with 10,000 pseudocounts).
(a) Influence of the prior depending on sequence separation.
(b) Influence of the signal coming from the secondary diagonals of the patterns on test dataset BetaSheet916: runs "without secondary diagonal signal" (dashed lines) contain only signal from the main diagonal of the pattern, while runs "with secondary diagonal signal" (solid lines) contain signal from both the main and the secondary diagonals. We also test the influence of PSM (predictions without PSM in red, with PSM in brown).


Figure S9: Influence of different model parameters on the residue-level performance of bbcontacts on BetaSheet916, using DSSP assignments as input for secondary structure.
(a) Influence of the type of secondary-structure-based emission probabilities on the residuelevel performance of bbcontacts: non-conditional (red), conditional (blue), conditional with 10,000 pseudocounts (purple). The blue and purple lines are superimposed.
(b) Effect of local background correction applied to coupling matrices, for different values of $S$.
(a)


Figure S10: Influence of different model parameters on the residue-level performance of bbcontacts on the BetaSheet916 dataset, using DSSP assignments as an input and using the reference version of bbcontacts mentioned in the main text (local background correction of the coupling matrices with $S=10$, conditional secondary-structure-based emissions with 10,000 pseudocounts).
(a) Influence of the prior depending on sequence separation.
(b) Influence of the signal coming from the secondary diagonals of the patterns on test dataset BetaSheet916: runs "without secondary diagonal signal" (dashed lines) contain only signal from the main diagonal of the pattern, while runs "with secondary diagonal signal" (solid lines) contain signal from both the main and the secondary diagonals. We also test the influence of PSM (predictions without PSM in red, with PSM in brown).

Strand-level predictions (correct orientation)


Figure S11: Final strand-level performance of bbcontacts on the BetaSheet916 dataset, compared to previous methods, when testing not only for correct pairing of $\beta$-strands but also for correct orientation. For all methods apart from bbcontacts, the results are obtained by multiplying the strand-level precision and recall by the percentage of correct directions provided in Savojardo et al. (2013). This result is not available for CMM.


Figure S12: Examples of predicted contact maps for three PDB chains from the BetaSheet916 dataset. Each of the panels is built like main Figure 1, panels (a-b). On the left: CCMpred coupling matrix (upper-left) and coarse distance matrix (lower-right). On the right: $\beta-\beta$ contacts predicted by bbcontacts using predicted secondary structure (upper-left) and coarse distance matrix (lower-right). The Viterbi score of the local alignment is the confidence value. The true $\beta-\beta$ contacts (annotated by DSSP) are shown as open circles. (a) 1iguB ( $\eta=0.09$ ). (b) $1 \mathrm{jerA}(\eta=0.29)$. (c) 2acyA ( $\eta=0.49$ ).

Residue-level predictions


Figure S13: Residue-level performance of bbcontacts on the BetaSheet916 dataset, when using DSSP assignments, compared to the original BCov and CMM results and two new reference points BCov* and CMM*. BCov* and CMM* correspond to a situation where DSSP assignments and couplings predicted with CCMpred are used as an input to the $\beta$-contact prediction algorithms from BCov and CMM. Note that both algorithms require the DSSP-assigned secondary structure. Additional discussion of these results is provided in section S2.3.1.
(a)


Figure S14: Residue-level performance for individual test cases in BetaSheet916 as a function of the number of sequences $N$ in the alignment (filtered at $90 \%$ sequence identity). (a) Performance expressed as the bbcontacts F1-score (calculated on results above a threshold chosen as the Viterbi score giving the maximum residue-level F1-score on the training dataset). (b) Performance expressed as the bbcontacts precision at $40 \%$ recall (i.e. for each test case, all predictions up to $40 \%$ recall are taken into account when calculating precision). For each panel: (left) DSSP-based predictions, (right) PSIPRED-based predictions.

## (a)


(b)


Figure S15: Residue-level performance for individual test cases in BetaSheet916 as a function of the CCMpred precision for $L / 5$ predictions ( $L$ being the length of the protein). (a) Performance expressed as the bbcontacts F1-score (calculated on results above a threshold chosen as the Viterbi score giving the maximum residue-level F1-score on the training dataset). (b) Performance expressed as the bbcontacts precision at $40 \%$ recall (i.e. for each test case, all predictions up to $40 \%$ recall are taken into account when calculating precision). For each panel: (left) DSSP-based predictions, (right) PSIPRED-based predictions.


Figure S16: Evolution of the F1-score on the training dataset when predictions with decreasing Viterbi score are progressively added to the evaluation.
The Viterbi score giving the maximum F1-value is marked by a vertical line and was chosen as a threshold to calculate F1-values on the test datasets.


Figure S17: Residue-level performance of bbcontacts on the BetaSheet916 dataset, compared to CCMpred (Seemayer et al., 2014) and PhyCMAP (Wang and Xu, 2013) baselines obtained by restricting the predictions to (a) DSSP-assigned $\beta$-strand regions and (b) $\beta$ strand regions predicted by PSIPRED. As in main Figure 4c, the false positive predictions with sequence separation smaller than 6 are excluded for CCMpred and PhyCMAP. Three types of evaluation are used: standard and evaluation with $6 \AA$ tolerance (as in main Figure 4 c ) and evaluation with $8 \AA$ tolerance (i.e. all false positives that have a $\mathrm{C} \beta$ distance lower than $8 \AA$ are excluded from the set of false positives) (dotted lines).


Figure S18: Residue-level performance of bbcontacts compared to CCMpred (Seemayer et al. 2014), PhyCMAP (Wang and Xu, 2013) and PconsC2 (Skwark et al., 2014) baselines obtained by restricting the predictions to (a) DSSP-assigned $\beta$-strand regions and (b) $\beta$ strand regions predicted by PSIPRED. This plot contains only results for the subset of the BetaSheet916 dataset for which PconsC2 predictions were obtained (Supplementary Dataset S2). As in main Figure 4c, the false positive predictions with sequence separation smaller than 6 are excluded for CCMpred, PhyCMAP and PconsC2. For clarity, in this plot only the standard evaluation is used.


Figure S19: Residue-level performance of bbcontacts compared to CCMpred (Seemayer et al., 2014), PhyCMAP (Wang and Xu, 2013) and PconsC2 (Skwark et al., 2014) baselines obtained by restricting the predictions to (a) DSSP-assigned $\beta$-strand regions and (b) $\beta$ strand regions predicted by PSIPRED. This plot contains only results for the subset of the BetaSheet916 dataset for which PconsC2 predictions were obtained (Supplementary Dataset S2). As in main Figure 4c, the false positive predictions with sequence separation smaller than 6 are excluded for CCMpred, PhyCMAP and PconsC2. For clarity, in this plot only the $6 \AA$ tolerance and $8 \AA$ tolerance evaluations are used.


Figure S20: Runtimes depending on protein length for all cases in test datasets (a) BetaSheet916 and (b) BetaSheet1452. Because BetaSheet1452 contains much larger protein chains, the scales are different between (a) and (b).
When PSM is enabled, points are colored according to the number of PSM iterations effectively done while running bbcontacts.
The few points that have a runtime lower than the general trend in all plots correspond to cases where $\eta<0.022$, in which case no coupling-based emissions are calculated or used for the predictions.


Figure S21: Influence of different model parameters on the residue-level results for the training dataset.
(a) Influence of the type of secondary-structure-based emission probabilities on the residuelevel performance of bbcontacts: non-conditional (red), conditional (blue), conditional with 10,000 pseudocounts (purple). (left) DSSP-based predictions: the blue and purple lines are superimposed. (right) PSIPRED-based predictions.
(b) Effect of local background correction applied to coupling matrices, for different values of $S$. (left) DSSP-based predictions. (right) PSIPRED-based predictions.
(a)

Strand-level predictions

(b)

Residue-level predictions


Figure S22: Performance of bbcontacts on the training dataset.
(a) Strand-level performance (correct pairing of $\beta$-strands).
(b) Residue-level performance.
(a)


(b)



Figure S23: Influence of different model parameters on the residue-level results for test dataset BetaSheet1452.
(a) Influence of the type of secondary-structure-based emission probabilities on the residuelevel performance of bbcontacts: non-conditional (red), conditional (blue), conditional with 10,000 pseudocounts (purple). (left) DSSP-based predictions: the blue and purple lines are superimposed. (right) PSIPRED-based predictions.
(b) Effect of local background correction applied to coupling matrices, for different values of $S$. (left) DSSP-based predictions. (right) PSIPRED-based predictions.
(a)

Strand-level predictions

(b)

Residue-level predictions


Figure S24: Performance of bbcontacts on dataset BetaSheet1452 and comparison with previous methods.
(a) Strand-level performance (correct pairing of $\beta$-strands).
(b) Residue-level performance.

Results for CMM and BCov on the test dataset BetaSheet1452 are taken from Savojardo et al. (2013).


Figure S25: Residue-level performance for individual test cases in BetaSheet1452, expressed as the F1-score (calculated on results above a threshold chosen as the Viterbi score giving the maximum residue-level F1-score on the training dataset) as a function of the number of sequences $N$ in the alignment (filtered at $90 \%$ sequence identity). (left) DSSP-based predictions, (right) PSIPRED-based predictions.

## S4 Supplementary datasets

## S4.1 Supplementary dataset S1: training dataset (943 domains)

The four columns are the CATH domain identifier, the length of the domain L, the resolution of the PDB structure and the number of sequences N in the HHblits alignment.

| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 3nirA00 | 48 | 0.48 | 61 |
| 2b97A00 | 70 | 0.75 | 98 |
| 1mc2A00 | 122 | 0.85 | 797 |
| 1j0pA00 | 108 | 0.91 | 358 |
| 1vbwA00 | 68 | 0.93 | 414 |
| 3judA00 | 144 | 0.98 | 1315 |
| 1gkmA01 | 193 | 1.00 | 1373 |
| 1gkmA02 | 312 | 1.00 | 1192 |
| 3nvsA02 | 205 | 1.02 | 4369 |
| 2v3iA00 | 433 | 1.05 | 1043 |
| 2hbwA02 | 146 | 1.05 | 4753 |
| 1n62B03 | 166 | 1.09 | 3636 |
| 1n62B02 | 145 | 1.09 | 2556 |
| 1n62B05 | 97 | 1.09 | 3299 |
| 1n62B04 | 254 | 1.09 | 3845 |
| 1g8tA00 | 241 | 1.10 | 1293 |
| 2aibA00 | 98 | 1.10 | 143 |
| 1t8kA00 | 77 | 1.10 | 8302 |
| 3bvxA01 | 382 | 1.10 | 1242 |
| 3a8gA00 | 195 | 1.11 | 262 |
| 3ci3A01 | 177 | 1.11 | 1245 |
| 1sauA01 | 44 | 1.12 | 526 |
| 2r01A02 | 42 | 1.15 | 111 |
| 2awkA00 | 224 | 1.15 | 166 |
| 2ciwA00 | 298 | 1.15 | 383 |
| 1hbnA01 | 99 | 1.16 | 58 |
| 1hbnB02 | 295 | 1.16 | 52 |
| 3essA00 | 199 | 1.19 | 5 |
| 2bmoA01 | 308 | 1.20 | 1506 |
| 3og2A03 | 89 | 1.20 | 166 |
| 3og2A02 | 184 | 1.20 | 187 |
| 1w6sA00 | 595 | 1.20 | 3302 |
| 1jetA02 | 124 | 1.20 | 10531 |
| 1ymtA00 | 235 | 1.20 | 2037 |
| 3qvpA03 | 302 | 1.20 | 3542 |
| 3molB00 | 174 | 1.20 | 51 |
| 1vk1A02 | 122 | 1.20 | 18 |
| 1vr7A00 | 119 | 1.20 | 668 |
| 1vk1A01 | 101 | 1.20 | 2910 |
| 1jetA03 | 216 | 1.20 | 10832 |
| 3qvpA02 | 64 | 1.20 | 2381 |
| 1w6sB00 | 72 | 1.20 | 31 |
| 2iayA00 | 110 | 1.20 | 97 |
| 2xi8A00 | 66 | 1.21 | 21375 |
| 3moeA02 | 108 | 1.25 | 504 |
| 2wlvA00 | 144 | 1.25 | 352 |
| 3m |  |  |  |
| 3bs |  |  |  |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1jniA00 | 62 | 1.25 | 293 |
| 3moeA01 | 226 | 1.25 | 476 |
| 3moeA03 | 272 | 1.25 | 534 |
| 3fciA00 | 223 | 1.27 | 1619 |
| 1qksA02 | 432 | 1.28 | 3381 |
| 1vlbA06 | 126 | 1.28 | 3245 |
| 1vlbA05 | 93 | 1.28 | 3323 |
| 1gk9A01 | 148 | 1.30 | 920 |
| 2prvB00 | 150 | 1.30 | 643 |
| 1eu1A02 | 245 | 1.30 | 5980 |
| 1gk9B02 | 73 | 1.30 | 895 |
| 1rutX01 | 78 | 1.30 | 2635 |
| 3eojA00 | 358 | 1.30 | 11 |
| 3fegA02 | 262 | 1.30 | 5078 |
| 1oqvA00 | 171 | 1.30 | 29 |
| 1gk9B03 | 161 | 1.30 | 467 |
| 2nr7A00 | 193 | 1.30 | 448 |
| 1eu1A03 | 87 | 1.30 | 683 |
| 3i33A04 | 84 | 1.30 | 4612 |
| 1vp8A00 | 183 | 1.30 | 97 |
| 1o9iA02 | 63 | 1.33 | 142 |
| 2qikA02 | 154 | 1.35 | 1597 |
| 3pfgA02 | 59 | 1.35 | 60 |
| 1pinA01 | 32 | 1.35 | 1970 |
| 1gppA00 | 217 | 1.35 | 260 |
| 1ijyA00 | 122 | 1.35 | 576 |
| 2fxuA03 | 92 | 1.35 | 1956 |
| 3bxuA00 | 71 | 1.35 | 544 |
| 2gkpA00 | 161 | 1.35 | 65 |
| 3p0bA01 | 407 | 1.35 | 1143 |
| 1ouwA00 | 148 | 1.37 | 658 |
| 1f1eA00 | 147 | 1.37 | 400 |
| 1s9uA00 | 198 | 1.38 | 935 |
| 2i3fA00 | 206 | 1.38 | 503 |
| 1v30A00 | 118 | 1.40 | 1104 |
| 1tzpB00 | 236 | 1.40 | 272 |
| 1pbjA00 | 116 | 1.40 | 18497 |
| 1s2oA02 | 71 | 1.40 | 293 |
| 1l6rA02 | 64 | 1.40 | 120 |
| 2fsqA00 | 224 | 1.40 | 82 |
| 1ie9A00 | 255 | 1.40 | 2012 |
| 1f8nA03 | 89 | 1.40 | 312 |
| 1ygeA05 | 349 | 1.40 | 671 |
| 1yc5A02 | 84 | 1.40 | 2763 |
| 3iisM00 | 151 | 1.40 | 19 |
| 2ra9A02 | 72 | 1.40 | 304 |
| 12 |  |  |  |
| 193 |  |  |  |


| Domain | L | Resol | N | Domain | L | Resol | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $1 \mathrm{f8nA} 02$ | 100 | 1.40 | 269 | 1wn9A00 | 123 | 1.58 | 13 |
| 2ra9A01 | 54 | 1.40 | 254 | 1k7iA01 | 234 | 1.59 | 3836 |
| 1 tgrA00 | 52 | 1.42 | 251 | 1aopA01 | 166 | 1.60 | 2880 |
| 1pp0B00 | 194 | 1.42 | 13 | 1aopA03 | 143 | 1.60 | 3182 |
| 1m0kA00 | 222 | 1.43 | 946 | $2 \mathrm{olrA02}$ | 79 | 1.60 | 654 |
| 1 bgfA 00 | 124 | 1.45 | 140 | 202kA01 | 243 | 1.60 | 1074 |
| 1ikpA02 | 158 | 1.45 | 3 | $1 \mathrm{ft5A} 00$ | 211 | 1.60 | 839 |
| 2endA00 | 137 | 1.45 | 60 | 1 dd 9 A 01 | 123 | 1.60 | 2085 |
| 3h0nA00 | 182 | 1.45 | 780 | 1150 A 01 | 74 | 1.60 | 1121 |
| 3n2wD00 | 361 | 1.45 | 825 | 1wr8A02 | 69 | 1.60 | 122 |
| 3f6yA01 | 128 | 1.45 | 84 | 2pagA00 | 132 | 1.60 | 221 |
| 1qz5A02 | 22 | 1.45 | 772 | 3ku3A01 | 103 | 1.60 | 268 |
| 1 g 3 pA 01 | 87 | 1.46 | 6 | 1 qgiA 01 | 147 | 1.60 | 73 |
| 1tkeA02 | 102 | 1.46 | 3144 | 1x7dB01 | 164 | 1.60 | 1563 |
| 1 g 3 pA 02 | 104 | 1.46 | 1 | $1 \mathrm{ccwB02}$ | 66 | 1.60 | 67 |
| 1tkeA03 | 58 | 1.46 | 2149 | 1s9rA02 | 78 | 1.60 | 492 |
| 2 fb 6 A 00 | 111 | 1.46 | 113 | 1p5dX03 | 119 | 1.60 | 4263 |
| 3im9A01 | 239 | 1.46 | 7195 | $1 \mathrm{p} 5 \mathrm{dX01}$ | 145 | 1.60 | 4979 |
| 3 mjfA 04 | 96 | 1.47 | 2139 | 1p5dX02 | 78 | 1.60 | 4964 |
| 1rkuA02 | 95 | 1.47 | 102 | $1 \mathrm{kqfC00}$ | 216 | 1.60 | 2844 |
| 1wpuA00 | 147 | 1.48 | 112 | 2ob5A00 | 145 | 1.60 | 707 |
| 3 gvj A04 | 155 | 1.48 | 8 | 1kqfA03 | 266 | 1.60 | 6018 |
| 1 gqiA 03 | 237 | 1.48 | 236 | 1vkiA00 | 165 | 1.60 | 3789 |
| 1vykA00 | 129 | 1.49 | 144 | 1k92A02 | 215 | 1.60 | 1497 |
| 2 qnlA 00 | 159 | 1.50 | 113 | 1hfeL03 | 143 | 1.60 | 1792 |
| 1ui0A00 | 192 | 1.50 | 2604 | 3c8wA01 | 227 | 1.60 | 473 |
| 1 locyA01 | 76 | 1.50 | 758 | 2icuA00 | 207 | 1.60 | 1614 |
| 3c9qA00 | 191 | 1.50 | 113 | 1 aopA02 | 144 | 1.60 | 2216 |
| 3ckcA03 | 103 | 1.50 | 1 | 3os4A00 | 387 | 1.60 | 1777 |
| 1 j 77 A 00 | 199 | 1.50 | 705 | 1dj7A00 | 109 | 1.60 | 217 |
| 1qw2A00 | 97 | 1.50 | 92 | 1rylA00 | 157 | 1.60 | 186 |
| 1 j 10 A 00 | 310 | 1.50 | 308 | 3ku3A02 | 221 | 1.60 | 223 |
| 3mzfA02 | 92 | 1.50 | 1601 | 2 hlyA 00 | 205 | 1.60 | 18 |
| 1inlC02 | 64 | 1.50 | 824 | 1 nln A 00 | 203 | 1.60 | 45 |
| $3 \mathrm{bb0A01}$ | 174 | 1.50 | 140 | $1 \mathrm{vcc} A 00$ | 77 | 1.60 | 18 |
| 3cqlA02 | 58 | 1.50 | 858 | 3079B00 | 105 | 1.60 | 57 |
| 3bb0A02 | 401 | 1.50 | 1047 | 3 gvoA 00 | 342 | 1.60 | 1466 |
| 3 npk A00 | 260 | 1.50 | 5147 | $2 \mathrm{olrA01}$ | 197 | 1.60 | 835 |
| 3bhwA00 | 182 | 1.50 | 159 | $2 \mathrm{olrA03}$ | 251 | 1.60 | 4763 |
| 1g6sA01 | 205 | 1.50 | 4296 | 2v2gD02 | 69 | 1.60 | 2000 |
| 1 locyA02 | 122 | 1.50 | 15 | 3 g 9 mA 00 | 78 | 1.61 | 1600 |
| 1 ofwA01 | 155 | 1.50 | 304 | 2p12A01 | 158 | 1.63 | 217 |
| 2y1qA00 | 137 | 1.50 | 4468 | 1 dw 9 A 02 | 66 | 1.65 | 215 |
| 1h16A00 | 759 | 1.53 | 1172 | 3bs3A00 | 58 | 1.65 | 21444 |
| 1rv9A00 | 242 | 1.53 | 1741 | 2d48A00 | 129 | 1.65 | 53 |
| 3bhqA00 | 194 | 1.54 | 19638 | 1054A01 | 72 | 1.65 | 553 |
| 2jkbA03 | 98 | 1.54 | 32 | 2 yyvB 00 | 224 | 1.65 | 467 |
| 1sz7A00 | 159 | 1.55 | 513 | 1ogoX01 | 203 | 1.65 | 17 |
| 2hhvA02 | 115 | 1.55 | 1982 | 2 fpqA 00 | 414 | 1.65 | 25 |
| 4ubpB00 | 122 | 1.55 | 771 | 1hx6A01 | 229 | 1.65 | 1 |
| 1iuqA02 | 267 | 1.55 | 358 | 2g50A01 | 176 | 1.65 | 2419 |
| 1p9hA00 | 179 | 1.55 | 1296 | $1 \mathrm{dtdB00}$ | 61 | 1.65 | 1 |
| $1 \mathrm{f01A} 01$ | 187 | 1.55 | 1 | 3gbyA00 | 126 | 1.66 | 18013 |
| 1 nc 7 A 00 | 110 | 1.55 | 53 | 2isbA00 | 175 | 1.66 | 885 |
| 2 dxaA 00 | 154 | 1.58 | 3855 | 2aj7A00 | 155 | 1.67 | 397 |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1mw9X03 | 132 | 1.67 | 3616 |
| 1mw9X02 | 163 | 1.67 | 3230 |
| 1mw9X04 | 117 | 1.67 | 3454 |
| 2im9A01 | 116 | 1.67 | 208 |
| $2 \mathrm{r7gA0} 2$ | 158 | 1.67 | 161 |
| 2im9A02 | 147 | 1.67 | 291 |
| 3 kgdA 01 | 240 | 1.68 | 597 |
| $2 \mathrm{j} 8 \mathrm{gA02}$ | 82 | 1.69 | 2526 |
| 1 rxqD 00 | 166 | 1.70 | 1965 |
| $1 \mathrm{pbyC00}$ | 78 | 1.70 | 24 |
| 1qd1A01 | 180 | 1.70 | 303 |
| 2qgmA02 | 68 | 1.70 | 385 |
| 1 vhhA00 | 157 | 1.70 | 91 |
| $1 \mathrm{njhA00}$ | 105 | 1.70 | 73 |
| 1ewfA02 | 276 | 1.70 | 454 |
| 1 vclA 03 | 148 | 1.70 | 3 |
| 2 i 71 A 02 | 145 | 1.70 | 30 |
| 1pg6A00 | 206 | 1.70 | 1103 |
| 2quyA00 | 330 | 1.70 | 1058 |
| 1w27A03 | 122 | 1.70 | 202 |
| 3bi7A01 | 177 | 1.70 | 349 |
| 1 tuoA03 | 117 | 1.70 | 4820 |
| 1tuoA02 | 85 | 1.70 | 4843 |
| 1 ewfA01 | 180 | 1.70 | 346 |
| 2id3A02 | 141 | 1.70 | 5626 |
| 2qgmA01 | 200 | 1.70 | 430 |
| 1wjxA00 | 112 | 1.70 | 1499 |
| 1rh6A00 | 54 | 1.70 | 165 |
| 1mtyG02 | 73 | 1.70 | 12 |
| 2bw0A02 | 102 | 1.70 | 2660 |
| 3kqjA02 | 207 | 1.70 | 4326 |
| 1 kidA00 | 193 | 1.70 | 2364 |
| 1 hp 1 A 02 | 187 | 1.70 | 2835 |
| 3100 A 01 | 62 | 1.70 | 2180 |
| 2a9iA00 | 105 | 1.70 | 240 |
| 1 k 0 rA 01 | 99 | 1.70 | 1222 |
| $1 \mathrm{rwj} A 00$ | 81 | 1.70 | 291 |
| 2axqA03 | 100 | 1.70 | 291 |
| 2ox6D00 | 161 | 1.70 | 10 |
| 2a9dA01 | 246 | 1.70 | 2458 |
| $1 \mathrm{cpqA00}$ | 129 | 1.72 | 529 |
| 2hy5B00 | 132 | 1.72 | 973 |
| 2pbkB00 | 227 | 1.73 | 51 |
| 1uehA00 | 214 | 1.73 | 2744 |
| 3nt1A02 | 511 | 1.73 | 1299 |
| 1 px 5 A 02 | 186 | 1.74 | 182 |
| 1 pcfA 00 | 66 | 1.74 | 279 |
| 1111A02 | 112 | 1.75 | 323 |
| 1y0kA00 | 178 | 1.75 | 10 |
| 1w99A03 | 180 | 1.75 | 156 |
| 1 oi2A02 | 162 | 1.75 | 1124 |
| 1111A03 | 94 | 1.75 | 113 |
| 1 pv 5 A 00 | 254 | 1.75 | 388 |
| 1111A01 | 511 | 1.75 | 2106 |
| 2wbmA01 | 81 | 1.75 | 385 |


| Domain | $\mathbf{L}$ | Resol | N |
| :---: | :---: | :---: | :---: |
| 1u7lA02 | 183 | 1.75 | 221 |
| 3claA00 | 213 | 1.75 | 2450 |
| 1n93X02 | 127 | 1.76 | 11 |
| 2gmqA00 | 96 | 1.76 | 3 |
| 3or1C01 | 40 | 1.76 | 533 |
| 1n93X01 | 208 | 1.76 | 8 |
| 3g0mA00 | 138 | 1.76 | 968 |
| 2qhqB00 | 112 | 1.76 | 167 |
| 2c42A03 | 212 | 1.78 | 2968 |
| 1qkrB00 | 171 | 1.80 | 131 |
| 1f1mA00 | 162 | 1.80 | 198 |
| 2ou6A00 | 180 | 1.80 | 1553 |
| 1fn9A01 | 140 | 1.80 | 13 |
| 1lbuA02 | 129 | 1.80 | 873 |
| 1mugA00 | 165 | 1.80 | 1197 |
| 1lbuA01 | 84 | 1.80 | 4612 |
| 1rzhH01 | 105 | 1.80 | 53 |
| 2p84A01 | 60 | 1.80 | 49 |
| 2q03B00 | 129 | 1.80 | 123 |
| 1v54D00 | 144 | 1.80 | 258 |
| 1orvA01 | 470 | 1.80 | 4976 |
| 1v54A00 | 513 | 1.80 | 4858 |
| 1cmbA00 | 104 | 1.80 | 41 |
| 1kvdB00 | 77 | 1.80 | 2 |
| 1vqqA02 | 120 | 1.80 | 3958 |
| 1a9xA04 | 150 | 1.80 | 3064 |
| 118bB00 | 190 | 1.80 | 716 |
| 1jh6A00 | 181 | 1.80 | 1123 |
| 2sicI00 | 107 | 1.80 | 158 |
| 3canA00 | 158 | 1.80 | 11795 |
| 1vdkA01 | 135 | 1.80 | 1843 |
| 1mwpA00 | 96 | 1.80 | 71 |
| 1fn9A02 | 225 | 1.80 | 4 |
| 3c5nA00 | 231 | 1.80 | 399 |
| 1v54G00 | 83 | 1.80 | 262 |
| 1j09A03 | 116 | 1.80 | 3907 |
| 3fdjA03 | 127 | 1.80 | 1864 |
| 1vdkA02 | 266 | 1.80 | 5553 |
| 1jidA00 | 114 | 1.80 | 371 |
| 2q66A01 | 192 | 1.80 | 374 |
| 1ja1A03 | 126 | 1.80 | 1675 |
| 1lm5B00 | 193 | 1.80 | 245 |
| 1ro7C00 | 240 | 1.80 | 235 |
| 3bq9A01 | 110 | 1.80 | 124 |
| 1d15A02 | 116 | 1.80 | 2 |
| 1v33A01 | 243 | 1.80 | 419 |
| 2z0tA00 | 109 | 1.80 | 168 |
| 2wjrA00 | 204 | 1.80 | 168 |
| 2qh9A00 | 172 | 1.80 | 168 |
| 1nxuA01 | 78 | 1.80 | 1000 |
| 1rzhH02 | 132 | 1.80 | 229 |
| 1yc9A02 | 83 | 1.80 | 3467 |
| 2it9A00 | 115 | 1.80 | 79 |
| 1t07A00 | 75 | 1.80 | 321 |
| 1ow1A00 | 167 | 1.80 | 100 |
| 180 |  |  |  |


| Domain | L | Resol | N | Domain | L | Resol | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2p1gB01 | 130 | 1.80 | 201 | 1ae9A00 | 171 | 1.90 | 20746 |
| 3 mhjA 00 | 208 | 1.80 | 1432 | 1 u 94 A 02 | 59 | 1.90 | 1731 |
| 1k6kA00 | 142 | 1.80 | 4144 | 1 musA 03 | 110 | 1.90 | 178 |
| 2r6zA01 | 54 | 1.80 | 22 | 2 qx 2 A 00 | 313 | 1.90 | 156 |
| 1 ualA02 | 81 | 1.80 | 1712 | 1 ee 8 A 01 | 120 | 1.90 | 2171 |
| 2y28B00 | 176 | 1.80 | 2494 | 1 ux 6 A 01 | 127 | 1.90 | 669 |
| 1iq4A00 | 179 | 1.80 | 1624 | 1 svbA 03 | 80 | 1.90 | 110 |
| 1 c 96 A 02 | 113 | 1.81 | 3945 | 1 sr 8 A 01 | 39 | 1.90 | 28 |
| 1 c 96 A 04 | 221 | 1.81 | 3657 | $1 \mathrm{aooC03}$ | 176 | 1.90 | 123 |
| 1c96A03 | 175 | 1.81 | 3294 | 1 sr 8 A 02 | 153 | 1.90 | 632 |
| 1 c 96 A 01 | 201 | 1.81 | 2718 | 1 qcsA 02 | 102 | 1.90 | 208 |
| 2yvwA01 | 207 | 1.81 | 4477 | 1lshA04 | 251 | 1.90 | 231 |
| $20 q m B 01$ | 166 | 1.83 | 367 | $3 \mathrm{cx5C00}$ | 385 | 1.90 | 5065 |
| 2w2rA00 | 177 | 1.83 | 13 | 1q16B03 | 77 | 1.90 | 333 |
| 1 hq 0 A 00 | 295 | 1.83 | 22 | 2q5xA00 | 151 | 1.90 | 294 |
| 2pw8I00 | 60 | 1.84 | 10 | 1 rssA 00 | 140 | 1.90 | 1598 |
| $1 \mathrm{u} k \mathrm{kA} 00$ | 129 | 1.85 | 68 | 1iv8A02 | 114 | 1.90 | 411 |
| 2auwB02 | 67 | 1.85 | 749 | 1 nh 2 D 02 | 48 | 1.90 | 149 |
| 1 b 25 A 02 | 178 | 1.85 | 683 | 1 nh 2 C 00 | 50 | 1.90 | 129 |
| 2auwA01 | 82 | 1.85 | 425 | 1 vpsB 00 | 289 | 1.90 | 73 |
| 2qtqB00 | 192 | 1.85 | 23700 | 1at0A00 | 142 | 1.90 | 473 |
| $2 \mathrm{ptrB02}$ | 275 | 1.85 | 5547 | 1dzfA01 | 138 | 1.90 | 244 |
| 1 b 25 A 01 | 209 | 1.85 | 760 | 3 fn 2 A 00 | 94 | 1.90 | 7 |
| 3do6A02 | 113 | 1.85 | 1427 | 1q16A01 | 27 | 1.90 | 147 |
| 1o6aA00 | 85 | 1.85 | 1527 | 2qcvA02 | 41 | 1.90 | 96 |
| 1ssqA01 | 138 | 1.85 | 1743 | 1 epwA02 | 331 | 1.90 | 24 |
| 1 pk A01 | 305 | 1.85 | 241 | 2ww2A02 | 166 | 1.90 | 1256 |
| 1 tolA02 | 73 | 1.85 | 792 | 2g3wA00 | 173 | 1.90 | 312 |
| 2 in 3 A 02 | 123 | 1.85 | 3014 | $11 \mathrm{sh} A 01$ | 263 | 1.90 | 506 |
| 1 b 25 A 03 | 213 | 1.85 | 623 | 1lshA03 | 98 | 1.90 | 174 |
| 3cqbA01 | 85 | 1.86 | 4945 | 3 c 9 fA 02 | 168 | 1.90 | 277 |
| 11 mlA 03 | 63 | 1.86 | 267 | 1 amuA 03 | 81 | 1.90 | 24720 |
| 2wjnC02 | 156 | 1.86 | 75 | $1 \mathrm{jdh} A 00$ | 508 | 1.90 | 5373 |
| 2wjnC01 | 138 | 1.86 | 82 | 1 svbA 01 | 136 | 1.90 | 143 |
| 11 mlA 04 | 99 | 1.86 | 164 | 1 e 4 fT 02 | 65 | 1.90 | 961 |
| 1jb7B00 | 216 | 1.86 | 6 | 1 musA 02 | 276 | 1.90 | 316 |
| 11 mlA 02 | 124 | 1.86 | 603 | 1wteB01 | 147 | 1.90 | 2 |
| 2 gviA 01 | 62 | 1.87 | 16 | $1 \mathrm{r} 4 \mathrm{vA00}$ | 145 | 1.90 | 42 |
| 2ijqA00 | 145 | 1.88 | 455 | 2ck3G02 | 117 | 1.90 | 1686 |
| $3 q 0 i A 02$ | 104 | 1.89 | 2701 | 1wteA02 | 124 | 1.90 | 11 |
| 7 ahlF 01 | 292 | 1.89 | 65 | 3d7aA01 | 136 | 1.90 | 241 |
| 1 t 5 oA 01 | 141 | 1.90 | 1206 | 1 oisA02 | 128 | 1.90 | 204 |
| 2qe9A01 | 158 | 1.90 | 2088 | 2 gukA 00 | 107 | 1.91 | 82 |
| 1 svbA 02 | 83 | 1.90 | 85 | 3bf5A02 | 42 | 1.91 | 1 |
| 2 osoA00 | 152 | 1.90 | 442 | 1 r 0 A 00 | 123 | 1.91 | 1546 |
| 1h6wA01 | 41 | 1.90 | 17 | 3ku8A00 | 134 | 1.93 | 2541 |
| 2o5hA00 | 126 | 1.90 | 75 | 1 kblA 02 | 63 | 1.94 | 829 |
| 1oaoD05 | 130 | 1.90 | 105 | $1 \mathrm{vgj} A 00$ | 181 | 1.94 | 2184 |
| 1dzfA02 | 73 | 1.90 | 309 | 3dt5A00 | 118 | 1.94 | 1 |
| 1 eerA00 | 166 | 1.90 | 43 | 3 dnh 402 | 82 | 1.94 | 445 |
| 315 xA 00 | 101 | 1.90 | 27 | 1b12A02 | 74 | 1.95 | 402 |
| 1 mpxA 02 | 67 | 1.90 | 135 | 1 m 48 B 00 | 126 | 1.95 | 31 |
| $2 \mathrm{dtrA02}$ | 66 | 1.90 | 1377 | 1p7tA01 | 85 | 1.95 | 435 |
| 11slA01 | 56 | 1.90 | 2693 | 1 g 81 A 03 | 92 | 1.95 | 2687 |
| 1i1qA00 | 512 | 1.90 | 4465 | 1 p 7 tB 04 | 132 | 1.95 | 620 |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 2o34A00 | 241 | 1.95 | 867 |
| 1qhdA02 | 227 | 1.95 | 25 |
| 1g8lA04 | 74 | 1.95 | 2466 |
| 1xlyA00 | 224 | 1.95 | 20 |
| 2wb0X02 | 168 | 1.95 | 43 |
| 1beaA00 | 116 | 1.95 | 160 |
| 1p7tB02 | 156 | 1.95 | 463 |
| 3c6kA02 | 56 | 1.95 | 828 |
| 3kjdA01 | 136 | 1.95 | 335 |
| 1g81A02 | 57 | 1.95 | 2386 |
| 2pspA01 | 65 | 1.95 | 414 |
| 1ko7A01 | 129 | 1.95 | 1293 |
| 2gufA01 | 118 | 1.95 | 21282 |
| 1r1hA02 | 374 | 1.95 | 1760 |
| 3c8iA00 | 127 | 1.95 | 41 |
| 2ra8A01 | 74 | 1.95 | 502 |
| 1pucA00 | 101 | 1.95 | 212 |
| 1k7wA01 | 137 | 1.96 | 1687 |
| 1k7wA02 | 228 | 1.96 | 5606 |
| 2oezA01 | 75 | 1.97 | 252 |
| 3nx6A00 | 75 | 1.97 | 1872 |
| 2ny1A00 | 305 | 1.99 | 15373 |
| 2hq4A00 | 158 | 1.99 | 14 |
| 3eupB00 | 200 | 1.99 | 24271 |
| 1he1A00 | 135 | 2.00 | 18 |
| 3cexA00 | 165 | 2.00 | 218 |
| 2hqvA00 | 167 | 2.00 | 392 |
| 1ex0A02 | 319 | 2.00 | 2459 |
| 1kmoA02 | 519 | 2.00 | 24020 |
| 1io1A03 | 95 | 2.00 | 16 |
| 1oe4A00 | 245 | 2.00 | 485 |
| 1qlmA01 | 118 | 2.00 | 173 |
| 1io1A01 | 169 | 2.00 | 2906 |
| 1e7uA04 | 158 | 2.00 | 1070 |
| 1k8kE00 | 173 | 2.00 | 182 |
| 1o22A00 | 146 | 2.00 | 2 |
| 11kiA00 | 172 | 2.00 | 53 |
| 1i7wD00 | 57 | 2.00 | 317 |
| 1xo0A02 | 200 | 2.00 | 19150 |
| 1vkyB01 | 206 | 2.00 | 1606 |
| 1r7lA00 | 103 | 2.00 | 13 |
| 1n7zA02 | 156 | 2.00 | 29 |
| 2e2dC01 | 85 | 2.00 | 127 |
| 1h3nA03 | 57 | 2.00 | 1 |
| 3cngA01 | 34 | 2.00 | 684 |
| 1eg3A01 | 38 | 2.00 | 71 |
| 2qgqB01 | 205 | 2.00 | 10939 |
| 2qv8A00 | 143 | 2.00 | 349 |
| 1e8yA05 | 187 | 2.00 | 1487 |
| 2hnuA00 | 81 | 2.00 | 106 |
| 1fm2B03 | 66 | 2.00 | 818 |
| 2hkuB00 | 182 | 2.00 | 22856 |
| 1nijA02 | 116 | 2.00 | 2454 |
| 1pzxA03 | 122 | 2.00 | 1903 |
| 1g14A00 | 273 | 2.00 | 540 |
| 1 |  |  |  |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| $1 \mathrm{pocA00}$ | 134 | 2.00 | 271 |
| $1 \mathrm{mi8A} 00$ | 141 | 2.00 | 1166 |
| 1 wpbO 02 | 113 | 2.00 | 107 |
| 1o0wA01 | 154 | 2.00 | 3174 |
| 1 qoy A00 | 303 | 2.00 | 29 |
| 2ichA01 | 178 | 2.00 | 372 |
| 2ichB02 | 128 | 2.00 | 402 |
| 1 tvfA02 | 69 | 2.00 | 55 |
| 2qf7A03 | 90 | 2.00 | 568 |
| 2 fytA 02 | 170 | 2.00 | 927 |
| $3 \mathrm{reaC00}$ | 125 | 2.00 | 1047 |
| 2a1kA00 | 215 | 2.00 | 46 |
| 2 ppqA 02 | 212 | 2.00 | 7135 |
| 1 wdjA 00 | 186 | 2.00 | 4314 |
| 1 qlmA 02 | 198 | 2.00 | 155 |
| $1 \mathrm{e} 8 \mathrm{cA01}$ | 99 | 2.00 | 3980 |
| 1 m 3 y A01 | 188 | 2.00 | 139 |
| 1 m 3 y A02 | 215 | 2.00 | 121 |
| 1xkwA01 | 100 | 2.00 | 19500 |
| 1 n 7 zA 01 | 155 | 2.00 | 31 |
| 1 kp 8 A 02 | 94 | 2.00 | 2408 |
| 2 atzA00 | 172 | 2.00 | 32 |
| 1n1bA01 | 200 | 2.00 | 824 |
| 1 fm 2 B 02 | 177 | 2.00 | 816 |
| 1 pujA 02 | 93 | 2.00 | 1469 |
| 1 j 5 uA 01 | 116 | 2.00 | 390 |
| 3kfla02 | 121 | 2.00 | 3513 |
| 1 sx 3 A 01 | 250 | 2.00 | 4971 |
| 1k8kA02 | 30 | 2.00 | 233 |
| 1 d 2 zA 00 | 102 | 2.00 | 148 |
| 1 qakA 01 | 79 | 2.00 | 1382 |
| 2 cvcA 03 | 114 | 2.00 | 404 |
| 3c2qA01 | 86 | 2.00 | 187 |
| 1io1A02 | 131 | 2.00 | 77 |
| 1dvoA00 | 152 | 2.00 | 256 |
| 1 nigA00 | 146 | 2.00 | 3 |
| 2 arzA02 | 88 | 2.00 | 589 |
| $2 \mathrm{oyrA01}$ | 54 | 2.00 | 244 |
| 1olzA02 | 57 | 2.00 | 684 |
| 1ohtA00 | 167 | 2.00 | 2182 |
| 2i5tA00 | 165 | 2.01 | 1585 |
| $2 \mathrm{~h} 5 \mathrm{nC00}$ | 124 | 2.01 | 841 |
| 1 ddgA 02 | 105 | 2.01 | 1415 |
| $2 \mathrm{ii0} 0 \mathrm{~A} 02$ | 242 | 2.02 | 1362 |
| 2ii0A01 | 210 | 2.02 | 1037 |
| 2 nrjA 01 | 323 | 2.03 | 148 |
| 1em9B00 | 141 | 2.05 | 19 |
| 2qnuA00 | 207 | 2.05 | 209 |
| 3bjqJ00 | 292 | 2.05 | 258 |
| 1 y 7 mA 02 | 115 | 2.05 | 3875 |
| 1wlfA02 | 80 | 2.05 | 165 |
| 1uz5A04 | 72 | 2.05 | 2467 |
| $2 \mathrm{g7zA} 02$ | 118 | 2.05 | 1884 |
| 1 i 7 dA 03 | 141 | 2.05 | 3660 |
| 2 gtqA 05 | 324 | 2.05 | 690 |


| Domain | L | Resol | N | Domain | L | Resol | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 y 7 mA 01 | 46 | 2.05 | 8629 | 1scfB00 | 118 | 2.20 | 33 |
| $3 \mathrm{cjrB01}$ | 70 | 2.05 | 933 | 1 eteA00 | 134 | 2.20 | 19 |
| 1udxA01 | 154 | 2.07 | 1760 | 1oxwC00 | 350 | 2.20 | 4917 |
| 2qs7A00 | 130 | 2.09 | 734 | 1 n 7 vA 02 | 123 | 2.20 | 2 |
| 1ppjI00 | 42 | 2.10 | 68 | 1 n 7 vA 03 | 232 | 2.20 | 2 |
| 3 g 4 nA 02 | 173 | 2.10 | 33 | $1 \mathrm{ba3A} 05$ | 53 | 2.20 | 502 |
| $3 \mathrm{fnaB00}$ | 114 | 2.10 | 13986 | 1 qhlA 00 | 203 | 2.20 | 14576 |
| 16vpA00 | 311 | 2.10 | 18 | 1 vq 8 A 03 | 78 | 2.20 | 500 |
| 1g3jD00 | 34 | 2.10 | 25 | 1p2zA02 | 152 | 2.20 | 85 |
| 2fiyA00 | 285 | 2.10 | 364 | 3 hu 3 A 02 | 93 | 2.20 | 360 |
| 1wruA02 | 88 | 2.10 | 147 | 2i06A01 | 214 | 2.20 | 116 |
| 1ewnA00 | 200 | 2.10 | 833 | 1 i 5 pA 03 | 198 | 2.20 | 14 |
| 3fy6D01 | 105 | 2.10 | 7 | 1pfoA01 | 183 | 2.20 | 96 |
| 1 x 9 mA 02 | 169 | 2.10 | 864 | 1 n 7 vA 01 | 177 | 2.20 | 3 |
| 2f2gA00 | 211 | 2.10 | 1226 | $3 \mathrm{mudA01}$ | 119 | 2.20 | 34 |
| 2psbA00 | 287 | 2.10 | 257 | 1p2zA03 | 265 | 2.20 | 79 |
| 3 g 4 nA 01 | 89 | 2.10 | 24 | 2i06A02 | 77 | 2.20 | 71 |
| 1a31A03 | 150 | 2.10 | 550 | 2rhqB01 | 69 | 2.20 | 1708 |
| 1h5wB03 | 45 | 2.10 | 6 | 2 rhqB 04 | 75 | 2.20 | 2265 |
| 3c2bA02 | 146 | 2.10 | 2339 | 2g03A00 | 172 | 2.20 | 2547 |
| 3 g 27 A 01 | 65 | 2.10 | 84 | 1u19A00 | 348 | 2.20 | 29498 |
| 1dq3A01 | 177 | 2.10 | 755 | 1iq8A02 | 70 | 2.20 | 180 |
| 1 mswD 01 | 310 | 2.10 | 349 | 1nmpA01 | 119 | 2.20 | 1902 |
| 1u8bA01 | 69 | 2.10 | 1065 | 1p2zA04 | 208 | 2.20 | 40 |
| 1 qd 6 C 00 | 240 | 2.10 | 350 | 1 p 8 A 00 | 597 | 2.20 | 841 |
| $2 \mathrm{fi0A00}$ | 79 | 2.10 | 559 | 1p2zA01 | 155 | 2.20 | 163 |
| 1okgA03 | 66 | 2.10 | 7 | 2qziA00 | 99 | 2.20 | 59 |
| 2qzbA00 | 145 | 2.10 | 86 | 1kyqA03 | 82 | 2.20 | 92 |
| $1 \mathrm{na6A01}$ | 171 | 2.10 | 35 | $1 \mathrm{gml} \mathrm{A}^{\text {0 }} 0$ | 154 | 2.20 | 2123 |
| 3cddF03 | 60 | 2.10 | 149 | 2rhqB03 | 202 | 2.20 | 2365 |
| $1 \mathrm{jcfA03}$ | 76 | 2.10 | 1346 | 3 bqwA 01 | 210 | 2.20 | 168 |
| 1h3iA01 | 134 | 2.10 | 3672 | 3n1hA00 | 161 | 2.20 | 60 |
| $1 \mathrm{z1nX02}$ | 117 | 2.10 | 312 | $3 \mathrm{ci0J01}$ | 97 | 2.20 | 232 |
| 1wruA01 | 175 | 2.10 | 172 | 1zvpA00 | 128 | 2.20 | 361 |
| 1 accA 01 | 228 | 2.10 | 1124 | 1vq8P01 | 55 | 2.20 | 281 |
| 1fs0G01 | 130 | 2.10 | 2134 | 1pfoA02 | 53 | 2.20 | 43 |
| 3 g 4 nA 03 | 187 | 2.10 | 61 | 1sczA00 | 233 | 2.20 | 3598 |
| 2c36A00 | 274 | 2.11 | 29 | 3bl4A01 | 69 | 2.20 | 14 |
| 2raaA00 | 174 | 2.12 | 2822 | 3mw6F01 | 82 | 2.21 | 249 |
| 1n0uA03 | 107 | 2.12 | 834 | 3cwcA02 | 230 | 2.23 | 1227 |
| 1dvkA00 | 149 | 2.15 | 205 | 1 r 44 A 00 | 202 | 2.25 | 1038 |
| 1rlzA00 | 344 | 2.15 | 556 | 1 ciyA 02 | 196 | 2.25 | 155 |
| 3 csvA 02 | 235 | 2.15 | 2443 | $1 q \mathrm{tqA} 04$ | 109 | 2.25 | 3190 |
| 2 bkkC 02 | 173 | 2.15 | 5095 | $1 q \mathrm{tqA} 03$ | 79 | 2.25 | 1008 |
| 2iojA00 | 117 | 2.15 | 1428 | 1p32C00 | 176 | 2.25 | 342 |
| 2 g 8 yB 01 | 109 | 2.15 | 1037 | 3 dclA 01 | 147 | 2.25 | 55 |
| 3 ar 4 A 01 | 171 | 2.15 | 13464 | 3 dclA 03 | 39 | 2.25 | 49 |
| 1h54A03 | 74 | 2.15 | 1199 | 1xviA02 | 94 | 2.26 | 227 |
| 3 ar 4 A 04 | 244 | 2.15 | 7426 | 2j58A03 | 80 | 2.26 | 2061 |
| 2r19A00 | 135 | 2.16 | 1846 | 2j58A01 | 103 | 2.26 | 2408 |
| 2qyaA01 | 107 | 2.17 | 104 | 2j58A02 | 93 | 2.26 | 2575 |
| 2igsA00 | 211 | 2.17 | 3 | 2wvyA04 | 172 | 2.26 | 1197 |
| 1 k 1 fD 00 | 63 | 2.20 | 16 | 2pw6A00 | 234 | 2.27 | 1392 |
| 1p35C00 | 295 | 2.20 | 11 | 3 bzcA 02 | 260 | 2.27 | 1393 |
| 1b4uB00 | 298 | 2.20 | 1636 | 2o3iA02 | 130 | 2.30 | 185 |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1twfF00 | 84 | 2.30 | 440 |
| 2pifA01 | 138 | 2.30 | 373 |
| 1gd8A00 | 105 | 2.30 | 1345 |
| 1e5rB02 | 84 | 2.30 | 14 |
| 1js8B02 | 105 | 2.30 | 56 |
| 3rk1B02 | 87 | 2.30 | 604 |
| 1cr5B02 | 95 | 2.30 | 199 |
| 2pifA02 | 106 | 2.30 | 364 |
| 2b5uA03 | 98 | 2.30 | 66 |
| 2qm4D01 | 133 | 2.30 | 87 |
| 2vutI00 | 42 | 2.30 | 952 |
| 1inpA01 | 47 | 2.30 | 56 |
| 2o3iA01 | 227 | 2.30 | 324 |
| 1bobA01 | 128 | 2.30 | 209 |
| 3k3fA00 | 332 | 2.30 | 371 |
| 1q3qA02 | 107 | 2.30 | 2044 |
| 1twfB04 | 177 | 2.30 | 738 |
| 111sA00 | 108 | 2.30 | 742 |
| 3cniA00 | 143 | 2.30 | 3344 |
| 1f3mA00 | 70 | 2.30 | 355 |
| 1u2mA00 | 84 | 2.30 | 989 |
| 1jsuC00 | 69 | 2.30 | 214 |
| 2fokB01 | 281 | 2.30 | 16 |
| 3pikA02 | 77 | 2.30 | 3261 |
| 1g31A00 | 107 | 2.30 | 106 |
| 3fggA00 | 138 | 2.30 | 15 |
| 2in5A00 | 190 | 2.30 | 158 |
| 2hr7A02 | 118 | 2.32 | 557 |
| 2auaA02 | 89 | 2.35 | 16 |
| 2auaA01 | 106 | 2.35 | 19 |
| 3cdlB02 | 130 | 2.36 | 2572 |
| 3jyuB01 | 125 | 2.37 | 504 |
| 2gjvA00 | 136 | 2.39 | 43 |
| 115jA03 | 173 | 2.40 | 2662 |
| 1af6A00 | 421 | 2.40 | 311 |
| 1fepA02 | 531 | 2.40 | 24813 |
| 1ax8A00 | 130 | 2.40 | 54 |
| 2gmfA00 | 121 | 2.40 | 28 |
| 115jA02 | 200 | 2.40 | 4001 |
| 1epuA02 | 91 | 2.40 | 552 |
| 1t77A02 | 289 | 2.40 | 674 |
| 1x87B01 | 267 | 2.40 | 509 |
| 115jA04 | 128 | 2.40 | 642 |
| 3b8oA01 | 213 | 2.40 | 222 |
| 1kfqA01 | 207 | 2.40 | 5107 |
| 2a6hC02 | 339 | 2.40 | 1899 |
| 2ph7A01 | 119 | 2.40 | 3 |
| 1yisA02 | 310 | 2.40 | 5706 |
| 2a6hC01 | 211 | 2.40 | 2103 |
| 2fywA02 | 122 | 2.40 | 1822 |
| 2fywC01 | 126 | 2.40 | 1719 |
| 2pusA01 | 334 | 2.40 | 13 |
| 1r8eA02 | 73 | 2.40 | 9656 |
| 1ozjA00 | 126 | 2.40 | 203 |
| 2a6hC03 | 180 | 2.40 | 1254 |
| 2ab |  |  |  |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 2fdoA00 | 89 | 2.40 | 3 |
| 1chkA01 | 143 | 2.40 | 85 |
| 2vpzA04 | 61 | 2.40 | 486 |
| 3kasA03 | 142 | 2.40 | 607 |
| 1n8yC02 | 119 | 2.40 | 426 |
| 2ahxB04 | 138 | 2.40 | 1347 |
| 2a6hC04 | 65 | 2.40 | 1922 |
| 2ftsA04 | 81 | 2.41 | 2454 |
| 3g74C00 | 83 | 2.43 | 105 |
| 1ya5T01 | 84 | 2.44 | 19 |
| 1ei7A00 | 158 | 2.45 | 37 |
| 2h21C01 | 254 | 2.45 | 1791 |
| 3c6mD01 | 35 | 2.45 | 25 |
| 1hk8A00 | 561 | 2.45 | 961 |
| 2h21A02 | 165 | 2.45 | 374 |
| 2ajrA02 | 52 | 2.46 | 765 |
| 1lpbA00 | 85 | 2.46 | 69 |
| 3l4jA05 | 153 | 2.48 | 396 |
| 3l4jA03 | 184 | 2.48 | 1984 |
| 2fpnA01 | 135 | 2.49 | 35 |
| 1j3eA00 | 115 | 2.50 | 106 |
| 2hdiB00 | 103 | 2.50 | 3 |
| 3clqA04 | 160 | 2.50 | 126 |
| 1cjyB02 | 485 | 2.50 | 815 |
| 1khvA01 | 59 | 2.50 | 3 |
| 1ckmA03 | 54 | 2.50 | 5 |
| 1a0pA02 | 180 | 2.50 | 20621 |
| 3bt3A01 | 72 | 2.50 | 4 |
| 2qsdB02 | 78 | 2.50 | 134 |
| 1eg7A03 | 90 | 2.50 | 1630 |
| 1jb0D00 | 138 | 2.50 | 96 |
| 1ztpA01 | 210 | 2.50 | 113 |
| 2re3B03 | 29 | 2.50 | 243 |
| 3clqA02 | 150 | 2.50 | 134 |
| 3c2iA01 | 62 | 2.50 | 261 |
| 1jeyB03 | 102 | 2.50 | 181 |
| 2vqeR00 | 73 | 2.50 | 1239 |
| 1jb0A00 | 740 | 2.50 | 296 |
| 1t11A02 | 166 | 2.50 | 1802 |
| 1td6A02 | 92 | 2.50 | 2 |
| 1uunA02 | 52 | 2.50 | 28 |
| 2q83A02 | 226 | 2.50 | 3287 |
| 1ibvA00 | 81 | 2.50 | 20 |
| 1jeyA03 | 63 | 2.50 | 132 |
| 1jb0L00 | 151 | 2.50 | 135 |
| 1x9yA01 | 170 | 2.50 | 10 |
| 2re3A02 | 69 | 2.50 | 319 |
| 2p62A02 | 93 | 2.50 | 10 |
| 1jeyB02 | 201 | 2.50 | 977 |
| 1jeyA02 | 126 | 2.50 | 394 |
| 1fiqC04 | 156 | 2.50 | 3460 |
| 3bh1A02 | 101 | 2.51 | 192 |
| 3bh1A001 | 239 | 2.51 | 169 |
| 1pc6A00 | 141 | 2.51 | 174 |
| 1otsB00 | 441 | 2.51 | 3138 |
|  |  |  |  |
| 2and |  |  |  |


| Domain | L | Resol | N | Domain | L | Resol | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 fcdA 03 | 74 | 2.53 | 219 | 1jmuB03 | 157 | 2.80 | 18 |
| 1914A00 | 171 | 2.53 | 310 | 1 tlj A 00 | 189 | 2.80 | 266 |
| 2j8sB05 | 223 | 2.54 | 12852 | 1f45B00 | 133 | 2.80 | 40 |
| 2 j 8 sA 408 | 89 | 2.54 | 8524 | $1 \mathrm{fgj} A 01$ | 240 | 2.80 | 523 |
| $2 \mathrm{j} 8 \mathrm{sA04}$ | 96 | 2.54 | 8560 | 1 dkgA02 | 60 | 2.80 | 2183 |
| 3bf0A02 | 98 | 2.55 | 1753 | 2r6gF03 | 88 | 2.80 | 113 |
| 1 w 3 fA 02 | 165 | 2.58 | 207 | 2 glfA 02 | 133 | 2.80 | 1059 |
| 2r6iA01 | 94 | 2.59 | 402 | 2zihB00 | 281 | 2.80 | 324 |
| 1zy8K00 | 44 | 2.59 | 3485 | 1 vfg A 02 | 217 | 2.80 | 3622 |
| 2vglB00 | 579 | 2.59 | 2535 | 2gk9B01 | 157 | 2.80 | 671 |
| 1e50B00 | 130 | 2.60 | 48 | 2r6gF02 | 84 | 2.80 | 175 |
| $2 \mathrm{p} 5 \mathrm{zX02}$ | 97 | 2.60 | 1502 | 3oaeA00 | 421 | 2.80 | 639 |
| 1nktA04 | 177 | 2.60 | 1988 | 2r6gF04 | 235 | 2.80 | 18893 |
| 1 hynR 00 | 300 | 2.60 | 1174 | $1 \mathrm{fgj} A 02$ | 230 | 2.80 | 226 |
| 2r7jA01 | 144 | 2.60 | 15 | 1 xp 4 C 02 | 89 | 2.80 | 1225 |
| 2w9jA00 | 69 | 2.60 | 173 | $1 \mathrm{yewB00}$ | 238 | 2.80 | 419 |
| 1c4zA02 | 81 | 2.60 | 1640 | 1jx7A00 | 114 | 2.80 | 1739 |
| 1 nktA 02 | 122 | 2.60 | 1625 | $1 \mathrm{jmuB02}$ | 124 | 2.80 | 17 |
| 1 k 2 fA 02 | 56 | 2.60 | 385 | 1jmuB01 | 120 | 2.80 | 11 |
| 11 ktA 00 | 104 | 2.60 | 29 | 1 tdjA03 | 161 | 2.80 | 1576 |
| $3 \mathrm{dplC01}$ | 117 | 2.60 | 593 | 3 k 6 e A00 | 139 | 2.81 | 18446 |
| 1 fpsA 00 | 348 | 2.60 | 5360 | 3doaA01 | 153 | 2.81 | 969 |
| 1 sigA00 | 305 | 2.60 | 5213 | 1sxjA03 | 117 | 2.85 | 230 |
| 1 divA 02 | 90 | 2.60 | 1742 | 1 vsgA 02 | 159 | 2.90 | 46 |
| 1jg5A00 | 83 | 2.60 | 44 | 1cjaA01 | 150 | 2.90 | 9 |
| $11 r v A 00$ | 233 | 2.60 | 757 | 3 fgxA 00 | 96 | 2.90 | 67 |
| 3 kicA00 | 520 | 2.60 | 141 | 1 vsgA 01 | 203 | 2.90 | 157 |
| $1 \mathrm{t} 8 \mathrm{sA01}$ | 140 | 2.60 | 260 | $1 \mathrm{nt} 2 \mathrm{B01}$ | 71 | 2.90 | 53 |
| 2nrqA00 | 136 | 2.60 | 226 | 1 gaxA 04 | 50 | 2.90 | 5 |
| 2 bghA 01 | 207 | 2.60 | 1909 | 1 cjaA 02 | 177 | 2.90 | 24 |
| 2 bghA 02 | 206 | 2.60 | 1944 | $1 \mathrm{f02T00}$ | 66 | 2.90 | 11 |
| 2hwjB01 | 118 | 2.61 | 214 | 2 idbA 02 | 113 | 2.90 | 800 |
| 3 cyg A02 | 93 | 2.61 | 494 | 2gy5A03 | 135 | 2.90 | 7080 |
| 1 floC02 | 257 | 2.65 | 9 | 1 jhnA 02 | 146 | 2.90 | 556 |
| 2idgC00 | 159 | 2.69 | 528 | 3bxjA01 | 34 | 3.00 | 2 |
| 1k17A01 | 92 | 2.70 | 1277 | $1 \mathrm{mqsA0} 0$ | 120 | 3.00 | 430 |
| 2 vsgA 02 | 160 | 2.70 | 19 | 2 ijzA 02 | 107 | 3.00 | 1010 |
| 3 bu 2 A 02 | 69 | 2.70 | 177 | 1bo1A02 | 159 | 3.00 | 866 |
| 2vsgA01 | 198 | 2.70 | 128 | 1ltlA01 | 90 | 3.00 | 575 |
| $1 \mathrm{bcpB01}$ | 86 | 2.70 | 4 | 2bpa100 | 426 | 3.00 | 21 |
| 3 b 8 mC 01 | 151 | 2.70 | 218 | 1ldjA05 | 74 | 3.00 | 811 |
| $208 \mathrm{rB02}$ | 151 | 2.70 | 1417 | 1x9nA01 | 270 | 3.00 | 809 |
| 3hhwK02 | 200 | 2.70 | 72 | 2 alaA 01 | 124 | 3.00 | 33 |
| 3hhwK01 | 183 | 2.70 | 84 | 1qvrC01 | 139 | 3.00 | 4088 |
| 1nltA02 | 66 | 2.70 | 2956 | 1tlyA00 | 251 | 3.01 | 276 |
| 2 ijrA 02 | 173 | 2.70 | 41 | 3 fwlA 03 | 85 | 3.09 | 277 |
| 1 c 4 kA 04 | 129 | 2.70 | 1120 | 3 fwlA 04 | 172 | 3.09 | 4377 |
| 3 cslA 01 | 128 | 2.70 | 19919 | 3 h 9 vA 02 | 273 | 3.10 | 227 |
| $1 \mathrm{dt9} 901$ | 105 | 2.70 | 197 | 3 b 8 pB 00 | 173 | 3.10 | 227 |
| 2ha9B00 | 398 | 2.70 | 425 | 1w36B04 | 146 | 3.10 | 736 |
| $1 \mathrm{kf6C00}$ | 130 | 2.70 | 95 | 1e0fI01 | 33 | 3.10 | , |
| 1jroB05 | 94 | 2.70 | 3318 | 2r6fA02 | 154 | 3.20 | 2870 |
| 3 cucA00 | 257 | 2.71 | 3617 | 2zjsY00 | 415 | 3.20 | 1843 |
| 2iahA03 | 543 | 2.73 | 25281 | 2aj2A01 | 92 | 3.21 | 885 |
| 1vz6A01 | 253 | 2.75 | 1344 | 1shyB02 | 49 | 3.22 | 276 |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1v7nY00 | 139 | 3.30 | 40 |
| 1 pw 4 A 02 | 212 | 3.30 | 9546 |
| 1x9pA02 | 297 | 3.30 | 47 |
| 1tx9A00 | 141 | 3.31 | 5 |
| 1 zcdA 00 | 376 | 3.45 | 1126 |
| 1 g 03 A 00 | 134 | NMR | 7 |
| 1 n 91 A 00 | 107 | NMR | 875 |
| 1 hdl A00 | 55 | NMR | 150 |
| $1 \mathrm{mknA00}$ | 59 | NMR | 32 |
| 2jovA01 | 71 | NMR | 272 |
| 1 gccA 00 | 63 | NMR | 1591 |
| 1 auu A00 | 55 | NMR | 685 |
| 1 d 4 uA 00 | 111 | NMR | 223 |
| $1 \mathrm{hyw} A 00$ | 58 | NMR | 60 |
| 1szlA01 | 52 | NMR | 2198 |
| 2jynA01 | 136 | NMR | 165 |
| 1 e 8 pA 00 | 46 | NMR | 69 |
| 1 c ra00 | 94 | NMR | 65 |
| 1z60A00 | 59 | NMR | 245 |
| $1 \mathrm{dqcA00}$ | 73 | NMR | 1200 |
| 1 xn 8 A 00 | 131 | NMR | 39 |
| 106wA02 | 33 | NMR | 1241 |
| 2ez5W01 | 36 | NMR | 1003 |
| $1 \mathrm{hnrA00}$ | 47 | NMR | 702 |
| $1 \mathrm{cz4A} 02$ | 81 | NMR | 282 |
| 1 tbaA00 | 67 | NMR | 44 |
| $1 \mathrm{~d} 1 \mathrm{rA00}$ | 83 | NMR | 985 |
| 2if1A00 | 126 | NMR | 1091 |
| 1wfeA00 | 86 | NMR | 681 |
| $1 \mathrm{nbl} \mathrm{A}^{0} 00$ | 46 | NMR | 63 |
| 2jroA01 | 65 | NMR | 74 |
| 11 v 3 A 00 | 65 | NMR | 570 |
| $1 \mathrm{nd9A00}$ | 49 | NMR | 981 |
| 2hg7A00 | 60 | NMR | 21 |
| $1 \mathrm{nr3A} 00$ | 122 | NMR | 16132 |
| 1imlA00 | 76 | NMR | 2854 |
| 1 x 6 a 01 | 62 | NMR | 2706 |
| 1 j 57 A 00 | 143 | NMR | 86 |
| 2 hfqA 00 | 85 | NMR | 47 |
| 1 n 0 zA 00 | 45 | NMR | 533 |
| 2hg6A00 | 106 | NMR | 20 |
| 1 d 6 gA 00 | 47 | NMR | 37 |
| 1t23A00 | 93 | NMR | 46 |
| 2hh8A00 | 127 | NMR | 50 |
| 1yuaA01 | 64 | NMR | 115 |
| 1ev0A00 | 58 | NMR | 539 |
| 1xu6A00 | 80 | NMR | 86 |
| $1 \mathrm{apj} A 00$ | 74 | NMR | 198 |
| 2jv8A00 | 73 | NMR | 1 |
| 1 jbiA 00 | 100 | NMR | 401 |
| $1 \mathrm{kmxA00}$ | 54 | NMR | 31 |
| 1 q 5 fA 01 | 150 | NMR | 130 |
| 1 l 4 A 01 | 65 | NMR | 274 |
| 1 mkcA 00 | 43 | NMR | 47 |
| $1 \mathrm{rhxA00}$ | 87 | NMR | 393 |


| Domain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1v9vA01 | 95 | NMR | 69 |
| 1hn6A00 | 110 | NMR | 11 |
| 1wvkA00 | 86 | NMR | 196 |
| 2nwtA00 | 69 | NMR | 158 |
| 2jneA00 | 71 | NMR | 125 |
| 1ghhA00 | 81 | NMR | 208 |
| 2jz6A01 | 50 | NMR | 1109 |
| 1y7jA00 | 40 | NMR | 54 |
| 1g10A00 | 102 | NMR | 142 |
| 1widA00 | 117 | NMR | 698 |
| 1q48A00 | 134 | NMR | 1688 |
| 1wibA00 | 92 | NMR | 1042 |
| 1e8rA00 | 50 | NMR | 61 |
| 1imtA00 | 80 | NMR | 166 |
| 1nynA00 | 111 | NMR | 254 |
| 1ngrA00 | 85 | NMR | 334 |
| 1a1wA00 | 83 | NMR | 256 |
| 1gh9A00 | 71 | NMR | 65 |
| 2gpfA01 | 63 | NMR | 552 |
| 1q60A00 | 99 | NMR | 60 |
| 1v9xA00 | 114 | NMR | 341 |
| 1co4A00 | 42 | NMR | 112 |
| 2joeA01 | 128 | NMR | 124 |
| 1so9A00 | 131 | NMR | 521 |
| 2e6iA00 | 64 | NMR | 165 |
| 1svjA00 | 136 | NMR | 7793 |

## S4.2 Supplementary dataset S2: subset of BetaSheet916 for which PconsC2 predictions could be computed

Because of the high computational cost of running PconsC2, the comparison with bbcontacts and other methods was performed on a subset of the main test set BetaSheet916 containing 186 protein chains. This subset was built in the following manner: for each CATH domain present in BetaSheet916, only the shortest chain in BetaSheet916 containing that CATH domain was retained. All chains containing more than 200 residues were excluded to limit the PconsC2 runtime.
In the following table, the four columns are the PDB chain identifier, the length of the chain $L$, the resolution of the PDB structure and the number of sequences N in the HHblits alignment.

| Chain | $\mathbf{L}$ | Resol | $\mathbf{N}$ |
| :---: | :---: | :---: | :---: |
| 1s5uE | 136 | 1.7 | 7558 |
| 1dhnA | 121 | 1.65 | 1697 |
| 1f3zA | 150 | 1.98 | 2056 |
| 1oo0A | 144 | 1.85 | 114 |
| 1q92A | 195 | 1.4 | 12947 |
| 1qhvA | 195 | 1.51 | 67 |
| 1n2mC | 163 | 1.9 | 176 |
| 1jzaA | 66 | 2.2 | 267 |
| 1fjrA | 188 | 2.3 | 95 |
| 1bxyA | 60 | 1.9 | 1342 |
| 1mk0A | 97 | 1.6 | 553 |
| 1b66A | 138 | 1.9 | 1702 |
| 1mg3F | 125 | 2.4 | 47 |
| 1genA | 200 | 2.15 | 1065 |
| 1h59B | 45 | 2.1 | 134 |
| 1h75A | 76 | 1.7 | 14797 |
| 1squB | 154 | 2.4 | 1344 |
| 1svpB | 160 | 2 | 17 |
| 1ootA | 58 | 1.39 | 6749 |
| 1j2lA | 68 | 1.7 | 1030 |
| 1is1A | 185 | 2.2 | 1773 |
| 1ni0C | 158 | 2.5 | 14 |
| 1hufA | 123 | 2 | 6 |
| 1ugiA | 83 | 1.55 | 1 |
| 1qmyA | 156 | 1.9 | 20 |
| 1btnA | 106 | 2 | 3405 |
| 1g13A | 162 | 2 | 344 |
| 1bx7A | 51 | 1.2 | 11 |
| 1extA | 160 | 1.85 | 976 |
| 1whiA | 122 | 1.5 | 1172 |
| 1lqvB | 173 | 1.6 | 3569 |
| 1b33N | 67 | 2.3 | 277 |
| 1jsgA | 111 | 2.5 | 47 |
| 1fqtA | 109 | 1.6 | 7963 |
| 1d8lB | 140 | 2.5 | 2523 |
| 1icfI | 65 | 2 | 596 |
| 1jatA | 152 | 1.6 | 4048 |
| 1fxkC | 133 | 2.3 | 796 |
| 1hruA | 186 | 2 | 3867 |
|  |  |  |  |


| Chain | $\mathbf{L}$ | Resol | $\mathbf{N}$ |
| :---: | :---: | :---: | :---: |
| 1nvjB | 126 | 2.15 | 1395 |
| 1c9sM | 71 | 1.9 | 77 |
| 1ameA | 66 | 1.65 | 66 |
| 1ir2I | 140 | 1.84 | 547 |
| 1mogA | 67 | 1.7 | 399 |
| 1mkpA | 144 | 2.35 | 5135 |
| 1l5bA | 101 | 2 | 281 |
| 1nz0D | 111 | 1.2 | 1741 |
| 1g2rA | 94 | 1.35 | 988 |
| 1r6jA | 82 | 0.73 | 6730 |
| 1e44A | 84 | 2.4 | 32 |
| 1n5bA | 128 | 2 | 17 |
| 1pugA | 94 | 2.2 | 1566 |
| 1bjpA | 62 | 2.4 | 1872 |
| 1k2dB | 185 | 2.2 | 6262 |
| 1flmA | 122 | 1.3 | 1618 |
| 1nycA | 111 | 1.4 | 3 |
| 1v5xA | 200 | 2 | 1897 |
| 11shB | 174 | 1.9 | 111 |
| 1ex6A | 186 | 2.3 | 12668 |
| 1uslC | 158 | 1.88 | 1684 |
| 1d0dA | 60 | 1.62 | 2 |
| 1c5eA | 95 | 1.1 | 33 |
| 2a8vA | 118 | 2.4 | 1039 |
| 1r94B | 97 | 2.3 | 2850 |
| 1mm9A | 127 | 1.66 | 84 |
| 1e5kA | 188 | 1.35 | 16823 |
| 1iibA | 103 | 1.8 | 1485 |
| 2prdA | 174 | 2 | 1347 |
| 1moxC | 49 | 2.5 | 521 |
| 1feuA | 185 | 2.3 | 1545 |
| 1o7zA | 60 | 1.92 | 889 |
| 1jtgB | 165 | 1.73 | 14 |
| 1js2A | 89 | 1.9 | 155 |
| 1b13A | 54 | 1.5 | 1748 |
| 1gefA | 120 | 2 | 672 |
| 1jj2W | 82 | 2.4 | 444 |
| 1gmxA | 108 | 1.1 | 13191 |
| 1ezgB | 82 | 1.4 | 48 |
|  |  |  |  |


| Chain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1rlhA | 151 | 1.8 | 154 |
| 1jj2E | 172 | 2.4 | 2174 |
| 1jj2L | 194 | 2.4 | 296 |
| 1jj2T | 53 | 2.4 | 354 |
| 1j85A | 156 | 2 | 6797 |
| 1ufhB | 154 | 2.2 | 13970 |
| 1f39B | 101 | 1.9 | 4647 |
| 1ec6A | 87 | 2.4 | 2799 |
| 1ihnA | 113 | 2.2 | 428 |
| 1ew4A | 106 | 1.4 | 496 |
| 1v54F | 98 | 1.8 | 234 |
| 1lm4B | 184 | 1.45 | 3002 |
| 1d1mB | 65 | 2.05 | 2237 |
| 1ptqA | 50 | 1.95 | 1299 |
| 1hdfB | 100 | 2.35 | 146 |
| 1ihfA | 96 | 2.5 | 4009 |
| 1jj2Y | 73 | 2.4 | 328 |
| 1nrzA | 163 | 1.75 | 942 |
| 1go4C | 195 | 2.05 | 399 |
| 1a5kA | 100 | 2.2 | 576 |
| 1oqjA | 90 | 1.55 | 149 |
| 1qysA | 92 | 2.5 | 1 |
| 1bysA | 152 | 2 | 7308 |
| 1kh8A | 125 | 2 | 455 |
| 1kcqA | 103 | 1.65 | 798 |
| 1i8nA | 89 | 2.2 | 1 |
| 1ltiA | 185 | 2.13 | 155 |
| 1qqhA | 144 | 2.1 | 270 |
| 1jhsA | 188 | 1.9 | 197 |
| 3rhnA | 115 | 2.1 | 5129 |
| 1gpqB | 128 | 1.6 | 90 |
| 1gp0A | 133 | 1.4 | 434 |
| 1e6tA | 129 | 2.20 | 4 |
| 1n07B | 155 | 2.45 | 2096 |
| 1ku6B | 61 | 2.5 | 287 |
| 1h4yA | 115 | 1.61 | 6272 |
| 1i4jB | 110 | 1.8 | 1917 |
| 1iwmA | 177 | 1.9 | 441 |
| 1f47B | 144 | 1.95 | 321 |
| 1o5uA | 88 | 1.83 | 3723 |
| 1j3lB | 164 | 2.3 | 1748 |
| 1oqwB | 144 | 2 | 6513 |
| 1rlkA | 116 | 1.95 | 684 |
| 1i59B | 188 | 1.8 | 21133 |
| 1qgwA | 76 | 1.63 | 26 |
| 1fsjB | 134 | 1.8 | 128 |
| 1pqfA | 127 | 2 | 836 |
| 1fmbA | 104 | 1.8 | 3089 |
| 1dg5A | 159 | 2 | 3004 |
| 1ds6B | 179 | 2.35 | 282 |
| 1nplA | 109 | 2 | 2022 |
| 1o13A | 107 | 1.83 | 1654 |
| 1mi0A | 61 | 1.85 | 16 |
| 1fuxB | 164 | 1.81 | 1989 |
| 1udzA | 179 | 1.8 | 6684 |
| 1 |  |  |  |


| Chain | L | Resol | N |
| :---: | :---: | :---: | :---: |
| 1jj2N | 115 | 2.4 | 2230 |
| 118 rA | 101 | 1.65 | 75 |
| 1a73A | 162 | 1.8 | 18 |
| 1 m 8 nA | 120 | 2.45 | 24 |
| 1fx3B | 149 | 2.5 | 561 |
| 1durA | 55 | 2 | 14150 |
| 1b2uD | 90 | 2.1 | 543 |
| 1xxaC | 73 | 2.2 | 855 |
| 1 gwyB | 175 | 1.71 | 51 |
| 1e79H | 131 | 2.4 | 2128 |
| 1 g 3 kA | 173 | 1.9 | 1444 |
| 1 tulA | 102 | 2.2 | 28 |
| $1 \mathrm{k} 9 \mathrm{j} A$ | 130 | 1.9 | 6505 |
| 1agqD | 95 | 1.9 | 316 |
| 1iktA | 115 | 1.75 | 1483 |
| 1 ucrB | 75 | 1.2 | 48 |
| 2 ablA | 163 | 2.5 | 4215 |
| 1ca9A | 191 | 2.3 | 1277 |
| 1 g 6 gA | 127 | 1.6 | 5380 |
| 1 ocuA | 134 | 2.3 | 2515 |
| 1 g 1 bA | 164 | 1.99 | 499 |
| 1ycqA | 88 | 2.3 | 53 |
| 1b78A | 184 | 2.2 | 2381 |
| 1hxrB | 115 | 1.65 | 130 |
| 1d0qA | 102 | 1.71 | 2742 |
| 1vjhA | 120 | 2.1 | 679 |
| 1 kptA | 105 | 1.75 | 46 |
| 1 jyhA | 155 | 1.8 | 2695 |
| 1q9uB | 128 | 1.8 | 871 |
| 1 h 8 pA | 88 | 1.82 | 405 |
| 1hjzB | 192 | 1.7 | 2139 |
| 1 cw 0 A | 155 | 2.3 | 1091 |
| 1 g 5 cC | 169 | 2.1 | 2581 |
| 1pchA | 88 | 1.8 | 2258 |
| 1c2aA | 120 | 1.9 | 177 |
| 1oapA | 108 | 1.93 | 8567 |
| 1uutA | 195 | 2 | 27 |
| 1a9nA | 162 | 2.38 | 21317 |
| 1 rmdA | 116 | 2.1 | 10578 |
| 1no5B | 102 | 1.8 | 3591 |
| 1 di 2 A | 69 | 1.9 | 2899 |
| $1 \mathrm{lj0A}$ | 89 | 2 | 3392 |
| 1kuhA | 132 | 1.6 | 543 |
| 1 n 9 nA | 108 | 2.3 | 4516 |
| 1ub4C | 75 | 1.7 | 1719 |
| 1 fd 4 A | 41 | 1.7 | 191 |
| 1 nfj A | 87 | 2 | 285 |
| 1 ktg A | 137 | 1.8 | 19521 |
| $1 \mathrm{nn7}$ A | 105 | 2.1 | 1428 |
| 1 bylA | 122 | 2.3 | 15658 |
| 1 gmuC | 140 | 1.5 | 535 |
| 1 josA | 100 | 1.7 | 1627 |
| 1 dqoA | 134 | 2.2 | 202 |

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