

Propensity for Risk Taking Across the Life Span and Around the Globe



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Psychological Science
2016, Vol. 27(2) 231–243
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DOI: 10.1177/0956797615617811
pss.sagepub.com
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Abstract

Past empirical work suggests that aging is associated with decreases in risk taking. But are such effects universal? Life-history theory suggests that the link between age and risk taking is a function of specific reproductive strategies that can be more or less risky depending on the ecology. We assessed variation in the age-risk curve using World Values Survey data from 77 countries ($N = 147,118$). The results suggest that propensity for risk taking tends to decline across the life span in the vast majority of countries. In addition, there is systematic variation among countries: Countries in which hardship (e.g., high infant mortality) is higher are characterized by higher levels of risk taking and flatter age-risk curves. These findings suggest that hardship may function as a cue to guide life-history strategies. Age-risk relations thus cannot be understood without reference to the demands and affordances of the environment.

Keywords

risk taking, adult development, gender differences, cross-cultural differences, open materials

Received 3/11/15; Revision accepted 10/27/15

How does propensity for risk taking change across the life span and around the world? Several lines of evidence suggest that propensity for risky behavior increases in adolescence, peaks in young adulthood, and declines with aging (Dohmen et al., 2011; Mandal & Roe, 2014; Quetelet, 1842/2013). One paradigmatic example of this progression is the link between age and criminal behavior—one possible manifestation of risk taking that has been well documented since the 19th century (Quetelet, 1842/2013). The relation between age and crime has been replicated in different cohorts and cultures, albeit with significant variation (Steinberg, 2013; Ulmer & Steffensmeier, 2014), but to what extent do cultures vary systematically in age-risk progression?

One influential conception of risk taking is that it serves a functional role (i.e., an adaptation) that may be biologically determined (Mishra, 2014; Sih & Del Giudice, 2012; Wilson & Daly, 1985). In line with this view, propensity for risk taking and associated constructs, such as impulsivity and sensation seeking, have been conceptualized as traits with strong biological underpinnings (Steinberg, 2008; Zuckerman, 2007) that show moderate to high heritability (Anokhin, Golosheykin, Grant, & Heath, 2012; Benjamin

et al., 2012; Bezdjian, Baker, & Tuvblad, 2011) and reliable gender differences (Cross, Copping, & Campbell, 2011; Cross, Cyrenne, & Brown, 2013).

The view that risk taking serves a functional role is best discussed in the context of life-history theory, a framework that addresses how organisms allocate time and energy to tasks and traits so as to maximize their fitness. This framework focuses particularly on how evolutionary forces shape the timing of life events involved in development, growth, and reproduction as a result of ecological characteristics (Kaplan & Gangestad, 2005). According to life-history theory, even universal adaptations “may be limited by sex, life history stage, or circumstance” (Tooby & Cosmides, 1990, p. 393). In other words, life-history strategies, such as reproductive strategies, can be expected to change as a function of ecological circumstances.

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Current views suggest that one can frame human reproductive strategies along a continuum (Ellis et al., 2012): Toward one end, individuals may adopt a slower life-history strategy that focuses on avoiding risks and producing a few high-quality offspring that are likely to survive and reproduce; toward the other end, individuals may adopt faster life strategies that consist of taking risks in the service of promoting mating opportunities, early reproduction, and a greater number of offspring with more variable outcomes. These risky strategies may be particularly adaptive in harsh environments, in which morbidity and mortality are high and individuals have to compete fiercely for resources. In contrast, in rich, predictable environments, a slower reproductive strategy could be more appropriate. In other words, the rationale is that harsh, unpredictable environments may lead individuals to gamble on shorter life spans and earlier reproduction, given that fitness is likely enhanced by breeding early and abundantly rather than wasting resources on promoting one's own (unlikely) survival in such conditions (Ellis et al., 2012; Frankenhuis & de Weerth, 2013). There is indeed empirical evidence of the dependency between reproductive strategies and the harshness or unpredictability of local environments (Belsky, Schlomer, & Ellis, 2012; Simpson, Griskevicius, Kuo, Sung, & Collins, 2012; Wilson & Daly, 1997). Life-history theory also suggests that risky behaviors can be expected to be more prevalent among males, who are more likely than females to face reproductive competition (Ellis et al., 2012).

In sum, local conditions, such as the availability of resources and associated competition, are likely to affect individuals' propensity for risk taking. Ultimately, such factors may play a role in determining the shape of the age-risk relation, and resource scarcity and hardship may lead to longer periods of risk taking across the adult life span. In the present study, we tested whether local conditions of hardship could be used to predict cross-cultural variation in risk taking across the life span.

The Present Study

There has been considerable interest in accounting for similarities and differences in risk taking between cultures and countries (Becker, Dohmen, Enke, & Falk, 2014; Hsee & Weber, 1999; Rieger, Wang, & Hens, 2015; Vieider et al., 2015). However, this past work has not considered the extent to which the propensity to take risks is associated with age across cultures. We aimed to contribute to this effort by investigating the following research questions: Is a universal progression of risk propensity associated with age, such that risk propensity declines from adulthood to old age? Do local characteristics (e.g., exposure to hardship), as well as age and gender differences, account for potential differences in risk taking across cultures?

To answer these questions, we analyzed data from the World Values Survey (<http://www.worldvaluessurvey.org>), which aims to explore people's values and beliefs around the globe. It consists of a series of nationally representative surveys of various countries in which similar questionnaires are used, mostly in face-to-face interviews. We analyzed data collected in the last two independent waves of the survey (World Values Survey Association, 2008, 2014), which included one item we take to measure propensity for risk taking. Data for this item were available for 77 countries (see Fig. 1). In particular, participants were asked to report their similarity to a hypothetical individual: "Adventure and taking risks are important to this person; to have an exciting life" (for details, see Method). This item stems from Schwartz's (2012) Value Survey, which was designed to tap into a set of 10 independent universal values. According to Schwartz, this item captures individuals' need for variety and stimulation to maintain an optimal level of activation, and relates to feelings of excitement, variety seeking, and daringness. We take this item to measure the closely linked constructs of propensity for risk taking and sensation seeking that are empirically and theoretically related. For example, sensation seeking has been characterized as "a trait defined by the *seeking* of varied, novel, complex, and *intense* sensations and experiences, and the willingness to take physical, social, *legal*, and *financial* risks for the sake of such experience" (Zuckerman, 2007, p. 27).

Our analytic strategy proceeded in two steps. First, we analyzed the link between age and propensity for risk taking across the 77 countries. Second, we examined the extent to which a measure of exposure to hardship (i.e., a composite index capturing economic and social hardship through measures of gross domestic product per capita, homicide rate, and income inequality, among others) could account for cross-country variation in the pattern of the propensity for risk taking across the life span. The rationale for the latter analysis was to test the expectation that countries in which individuals are most exposed to hardship are likely to show higher levels of risk taking and longer periods of risk taking across the adult life span.

Method

Participants and procedure

We used data from the last two independent waves, Waves 5 and 6, of the World Values Survey (World Values Survey Association, 2008, 2014). We based our analysis on a balanced sample that consisted of respondents with valid answers on the risk item as well as all covariates of interest. Our final sample comprised 147,118 individuals

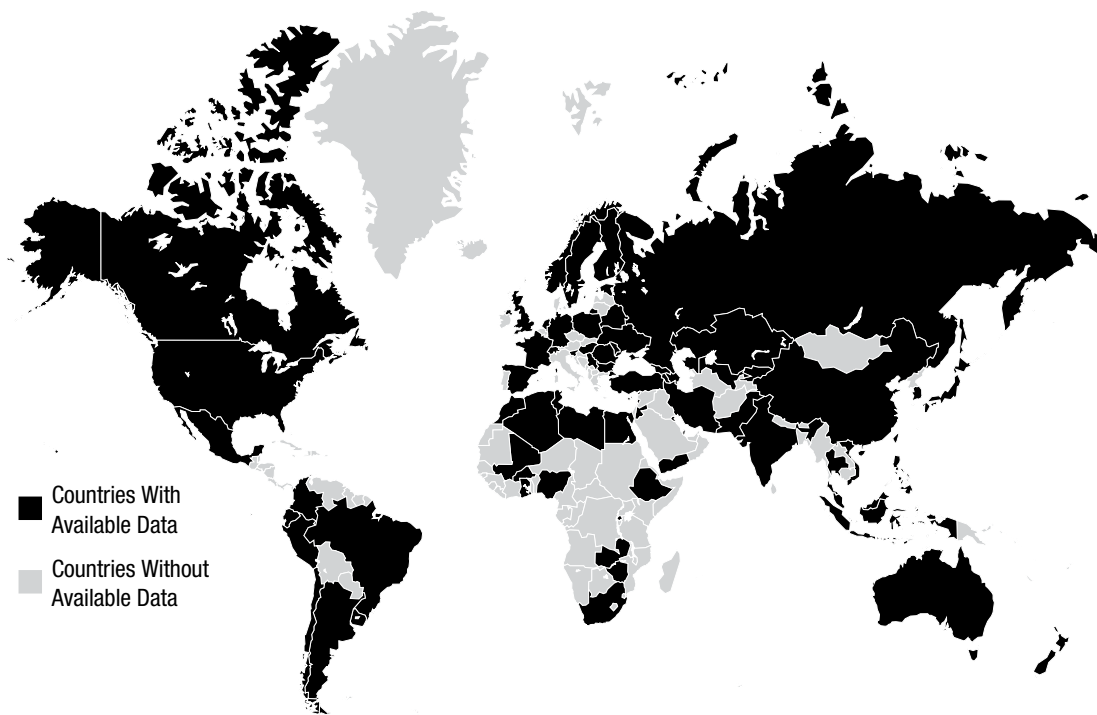


Fig. 1. World map showing countries for which data on the measure of propensity for risk taking were available from the World Values Survey.

(76,617 females, 52%; age range = 15–99 years). The countries included in the analysis cover the full range of global variation, from very poor to very rich countries, in all of the world's major cultural zones.

Measures

Propensity for risk taking and demographic covariates. Each respondent in the World Values Survey (World Values Survey Association, 2008, 2014) heard the following information:

Now I will briefly describe some people. Using this card, would you please indicate for each description whether that person is very much like you, like you, somewhat like you, a little like you, not like you, or not at all like you?

Respondents were then asked to rate a number of statements, including the following statement about adventure, excitement, and risk taking: “Adventure and taking risks are important to this person; to have an exciting life.” Respondents rated the statements using a 6-point scale (1 = *very much like me*, 6 = *not at all like me*).

In our analyses, we reversed the scale of the item such that the highest value (6) represented the highest propensity for risk taking and the lowest value (1) represented the lowest propensity for risk taking. We also

considered a number of demographic variables from the survey, including age and gender, education, marital status, parental status, and current occupational status. These variables represent (a) important indicators of human capital and (b) life-cycle phases that have been hypothesized to influence risk taking (Dohmen et al., 2011; Wilson & Daly, 1985).

Hardship index. To capture exposure to hardship in each country, we considered a number of indicators that could plausibly capture adversity and economic and social strife: homicide rate, gross domestic product, income inequality, infant mortality, life expectancy at birth, and gender equality (as indexed by the ratio of males to females receiving primary education). We manually compiled data concerning these indicators from the World Health Organization (homicide rate; World Health Organization, 2015), the World Bank (gender equality; World Bank, 2015), and the U.S. Central Intelligence Agency (gross domestic product, income inequality, infant mortality, life expectancy at birth; U.S. Central Intelligence Agency, 2015). By and large, the single indicators were significantly correlated. To obtain a single index representing exposure to hardship in each country, we *z*-standardized all of the indicators and used appropriate transformations (i.e., log transform); some of the indicators required reverse coding. Each indicator had missing data; the number of countries with missing data

ranged from one (1% of the sample) to nine (12% of the sample). We imputed the missing values with the median of each indicator so that we could use all countries and indicators in our analyses. Overall, the standardized and transformed hardship indicators were reasonably consistent, Cronbach's $\alpha = .86$. Consequently, we obtained a single hardship index by averaging all six z -standardized and transformed indicators. The specific data sources, as well as our procedure and the intercorrelations among indicators, are described in detail in the Supplemental Tables and Figures in the Supplemental Material available online.

Statistical analysis

We opted to use a linear regression approach to model the dependent variable, propensity for risk taking. First, however, we normalized it to have a mean of 50 and a standard deviation of 10 (i.e., T score), which is a common approach when using single, ordinal variables in linear regression models (Stevenson & Wolfers, 2008). Note, however, that the results were identical for the linear regression and the ordinal logistic regression, but the latter are less straightforward to depict and interpret (see Supplemental Tables and Figures).

We estimated the effects of different independent variables on the risk measure using mixed-effects linear regression in R (R Development Core Team, 2013). Specifically, we used the function `lmer` in the `lme4` package (Bates, Maechler, Bolker, & Walker, 2014), and we obtained p values for each effect on the basis of Satterthwaite's approximation using the package `lmerTest` (Kuznetsova, Brockhoff, & Christensen, 2015). Note that our modeling approach was particularly suited to our research question because it considered fixed effects (i.e., across country, average) of age and gender but also random effects (i.e., country specific) of these factors. This approach permitted us to assess the effect of ecological (i.e., country) characteristics on age and gender differences in propensity for risk taking and to control for other potentially relevant demographic covariates, such as marital status and education.

We report models estimating linear effects of age and gender but no age-by-gender interactions because models including covariates as well as quadratic effects of age or interactions with gender (or both) failed to converge. This finding suggests that the models that included interactions do not provide an appropriate description of the data. In all our analyses, we used age as a continuous variable and binary or dichotomized predictors to simplify coefficient estimation and interpretation, leading to the following additional predictors: gender (0 = female, 1 = male), marital status (0 = not married, 1 = married), parental status (0 = no children, 1 = children), education

(0 = no or incomplete primary education, 1 = primary education or higher), and occupational status (0 = not currently employed, 1 = currently employed). We compared different regression models using log-likelihood tests.

Results

We conducted a number of mixed-effects regression models with propensity for risk taking as the dependent variable. Table 1 presents the fixed-effects coefficients of all relevant models. We first compared a baseline model that did not consider any predictors (intercept-only model; not shown) with Model 1, which included age and gender as predictors. The significantly better fit of Model 1 relative to baseline suggests that age and gender are important predictors that contribute to explaining a

Table 1. Estimated Fixed-Effects Coefficients From the Mixed-Effects Regression Models of Propensity for Risk Taking Across the 77 Countries

Model and predictor	<i>b</i>	<i>SE</i>	<i>T</i> score	<i>p</i> value
Model 1: Age + gender				
Intercept	51.46	0.29	174.99	< .001
Age	-1.98	0.11	-18.15	< .001
Gender	-2.33	0.12	-20.10	< .001
Model 2: Age + gender + demographic covariates				
Intercept	52.09	0.27	192.40	< .001
Age	-1.43	0.09	-16.19	< .001
Gender	-2.17	0.12	-17.82	< .001
Parental status	-1.32	0.13	-10.27	< .001
Marital status	-0.85	0.10	-8.59	< .001
Occupational status	0.16	0.08	2.06	.04
Education	0.79	0.12	6.66	< .001
Model 3: Age + gender + demographic covariates + hardship				
Intercept	52.10	0.26	198.17	< .001
Age	-1.42	0.07	-19.22	< .001
Gender	-2.16	0.11	-18.83	< .001
Parental status	-1.32	0.13	-10.29	< .001
Marital status	-0.85	0.10	-8.70	< .001
Occupational status	0.16	0.08	2.02	.05
Education	0.79	0.12	6.83	< .001
Hardship	0.67	0.33	2.04	.04
Hardship × Age	0.48	0.10	4.97	< .001
Hardship × Gender	0.29	0.14	2.03	.05

Note: Variables were coded as follows—gender: 0 = male, 1 = female; parental status: 0 = no children, 1 = children; marital status: 0 = unmarried, 1 = married; occupational status: 0 = unemployed, 1 = employed; and education: 0 = no or incomplete primary education, 1 = primary education or higher.

significant amount of variance in propensity for risk taking, $\chi^2(7, N = 147,118) = 9,293, p < .001$. Table 1 shows that propensity for risk taking tended to decrease as a function of age and was lower for females than for males. We also ran Model 2, which included additional covariates of interest (i.e., education, parental status, marital status, and occupational status). Model 2 provided a significant improvement in fit relative to Model 1, $\chi^2(26, N = 147,118) = 1,524, p < .001$, but the results in Table 2 show that the main effects of age and gender remained after the inclusion of the additional demographic predictors.

But to what extent are life-span reductions in propensity for risk taking universal? Figure 2 plots the aggregate results as well as the country-specific effects of age and gender as estimated from Model 2. The pattern of reduction in propensity for risk taking across the life span, as well as the increased propensity for risk taking of males relative to females, was replicated in the vast majority of countries.

Despite the commonalities across countries, Figure 2 also highlights considerable variance in propensity for risk taking; it steeply declined with age in most countries but there are exceptions, such as Nigeria or Mali. As expected from the predictions of life-history theory, variation between countries in propensity for risk taking was associated with local characteristics as captured by our hardship index. Model 3, which included the hardship index as a covariate as well as interactions of hardship with age and gender, provided an additional improvement in fit relative to Model 2, $\chi^2(3, N = 147,118) = 124,283, p < .001$. Moreover, the results suggest that hardship was related to intercept differences in propensity for risk taking, as well as the age and gender effects identified in the previous models.

The substantive interpretation of the hardship effects is better captured in Figure 3, which depicts the relation between the random coefficients for each country (i.e., intercepts, age, gender) from the model without hardship as a predictor (Model 2). Figure 3 presents the zero-order correlations between the model coefficients for each country and the hardship index. The hardship index was significantly correlated to the intercept of propensity for risk taking in each country, $r = .24, p < .03$, as well as with the age coefficients, $r = .56, p < .001$, and gender coefficients, $r = .40, p < .001$. For example, Figure 3 shows that the higher the hardship experienced in each country, the closer the age coefficient was to 0, which represents a flat propensity-for-risk-taking curve across the adult life span.

In summary, we found that harsher environments were associated with increased propensity for risk taking in young adults, smaller gender differences, and smaller differences in propensity for risk taking between younger and older individuals (i.e., a flattening of the age-risk

curve). Overall, this finding suggests that ecologically dire circumstances may reduce differences in propensity for risk taking between younger and older individuals.

Discussion

We analyzed data from a large-scale survey of 77 countries to test whether the typical age-risk progression, which peaks in young adulthood and declines with increased age, represents a pancultural regularity. We found that the overwhelming majority of countries show the typical age-risk pattern, but there is significant variation in the relation between age and propensity for risk taking. Crucially, we found that an index of hardship in each country is significantly associated with the shape of the age-risk function: Hardship is associated with flatter age-risk curves and thus with smaller differences between younger and older age groups and between males and females. In other words, ecologies with scarce resources and therefore heightened competition may lead to increased propensity for risk taking regardless of age and gender. Our work matches expectations from life-history theory that associate ecological characteristics with life-span development of traits and reproductive strategies (Ellis et al., 2012; Mishra, 2014; Wilson & Daly, 1985).

Our work has connections to the broader debate about universals in life-span personality development; exploration and risk taking may represent important facets of such development. Whereas some researchers emphasize universals (McCrae et al., 2000), others emphasize the importance of normative as well as idiosyncratic life events in shaping personality; these events are likely to vary across cultures and individuals (Roberts, Wood, & Smith, 2005). Previous research indicates that there are indeed reliable age differences in personality development, such as decreases in openness to new experiences or increases in conscientiousness with increasing age, that have been replicated across samples and cultures (Roberts, Walton, & Viechtbauer, 2006). Nevertheless, considerable variation in personality development across cultures is explained by differences in timing of normative life events (Bleidorn et al., 2013). For example, Bleidorn et al. showed that cultures with an earlier onset of adult-role responsibilities, such as starting employment and parenting, were marked by relatively early personality maturation. Our results, which demonstrate a default progression of propensity for risk taking modified by ecological circumstances, are in line with the findings of Bleidorn et al. and with associated theories positing that personality development is a product, at least in part, of experience with the characteristics of local ecologies. Viewed more generally, our work contributes to understanding the causes underlying cultural variation (Weber & Hsee, 1999) and resonates with calls

Table 2. Demographics and Results From Model 2 by Country

Country	ISO Country Code	n	Female (%)	Mean age	Age range	Hardship index	Intercept	Age	Gender	b				Education
										Parental status	Marital status	Occupational status	Education	
Algeria	DA	1,101	50	37.14	18-81	0.24	53.41	-2.20	-1.21	-3.73	-1.17	0.20	1.67	
Andorra	AD	1,001	50	40.63	18-88	-0.81	53.54	-1.47	-2.70	-1.79	-1.49	0.31	0.98	
Argentina	AR	978	53	42.26	18-88	0.03	49.43	-1.29	-2.35	-2.36	-1.30	0.07	2.05	
Armenia	AM	1,087	66	46.47	18-85	-0.73	52.91	-1.97	-3.05	-2.06	-1.08	0.18	0.88	
Australia	AU	2,367	55	50.29	18-95	-0.90	51.89	-1.98	-3.25	-1.34	-1.23	0.40	0.72	
Azerbaijan	AZ	1,002	50	41.13	18-85	0.05	52.46	-1.58	-4.58	-1.09	-1.18	-0.16	0.55	
Bahrain	BH	1,190	45	39.23	18-72	-0.79	54.97	0.05	-1.01	0.37	0.08	0.59	0.27	
Belarus	BY	1,518	56	44.33	18-86	-0.43	51.83	-2.45	-3.33	-1.92	-0.82	-0.50	0.92	
Brazil	BR	2,922	60	41.32	18-93	0.84	49.19	-1.23	-2.52	-1.52	-1.46	-0.15	0.60	
Bulgaria	BG	935	54	46.88	18-84	0.00	50.19	-1.45	-1.79	-1.19	-1.07	0.74	2.01	
Burkina Faso	BF	1,250	47	34.12	16-94	1.29	53.50	-0.40	-2.28	-0.71	-0.37	-0.08	0.61	
Canada	CA	2,094	58	48.19	16-94	-0.82	53.07	-1.97	-3.19	-1.31	-1.65	0.88	-0.03	
Chile	CL	1,874	53	43.26	18-85	0.10	52.69	-2.07	-2.14	-2.46	-0.69	0.27	1.26	
China	CN	3,872	53	44.20	18-75	-0.03	49.64	-1.79	-1.62	-2.09	-1.01	-0.22	1.17	
Colombia	CO	1,487	50	40.34	18-82	0.69	52.42	-1.19	-3.07	-2.03	-0.84	0.07	0.36	
Cyprus	CY	2,020	52	41.87	17-91	-0.51	53.75	-2.17	-2.90	-1.25	-1.82	0.35	1.29	
Ecuador	EC	1,201	52	39.78	18-97	0.35	54.21	-1.44	-1.45	-0.31	-0.27	0.04	0.88	
Egypt	EG	4,549	64	40.86	18-99	0.20	47.82	-0.53	-0.67	-2.21	-0.39	0.72	1.18	
Estonia	EE	1,507	56	48.25	18-93	-0.37	50.81	-2.26	-2.43	-1.99	-1.14	0.14	0.98	
Ethiopia	ET	1,479	48	29.94	16-76	1.37	52.78	-0.63	-0.22	-1.54	-0.67	1.29	0.02	
Finland	FI	1,004	52	47.52	17-87	-0.90	50.11	-1.69	-2.13	-1.27	-1.34	0.15	0.19	
France	FR	993	52	47.05	18-92	-0.93	52.17	-1.88	-3.12	-2.04	-1.52	-0.07	0.45	
Georgia	GE	2,637	54	45.06	18-91	0.20	51.27	-1.18	-3.76	-0.44	-1.15	0.36	1.06	
Germany	DE	3,968	53	49.85	17-95	-1.01	48.95	-2.28	-2.78	-1.43	-1.41	0.18	1.16	
Ghana	GH	3,065	50	32.38	16-90	0.66	56.57	-0.70	-1.24	-0.01	0.01	0.19	0.16	
Gibraltar	GI	1,181	47	36.58	18-83	-0.43	50.70	-1.34	-1.75	-1.32	-0.10	-0.23	1.30	
Great Britain	GB	1,006	51	45.67	15-94	-0.74	54.11	-1.61	-3.68	-1.56	-1.57	0.16	-0.45	
Hungary	HU	999	53	45.54	18-91	-0.67	49.69	-2.37	-2.44	-1.66	-1.08	0.14	1.41	
India	IN	2,759	39	40.38	18-90	0.34	54.56	-0.30	-0.77	0.59	-0.19	0.38	0.28	
Indonesia	ID	1,922	47	35.75	15-84	0.22	54.96	-0.90	-2.52	-1.00	-0.55	0.64	0.51	
Iran	IR	2,545	50	32.72	16-90	0.27	52.40	-1.76	-1.33	-1.98	-1.15	0.68	0.60	
Japan	JP	3,019	52	49.32	18-80	-1.07	46.60	-0.55	-2.30	-0.81	-0.95	-0.23	-0.33	
Jordan	JO	1,193	50	39.71	18-84	0.13	53.93	-1.75	-2.24	-1.52	-0.81	0.21	0.57	
Kazakhstan	KZ	1,500	60	40.02	18-88	-0.02	49.62	-1.37	-1.83	-1.71	-0.84	0.45	0.59	
Kuwait	KW	1,165	36	36.56	16-79	-0.40	55.03	-2.89	-2.21	-1.31	-1.26	-0.41	-0.15	
Kyrgyzstan	KG	1,487	51	38.70	18-89	0.51	51.67	-0.35	-1.08	-0.73	0.14	-0.13	1.76	
Lebanon	LB	1,155	51	38.09	18-82	0.11	55.51	-1.56	-2.15	-0.68	-0.40	0.07	-0.06	
Libya	LY	1,963	48	38.13	18-78	-0.02	52.94	-1.99	-2.78	-1.22	-1.16	0.36	0.57	

(continued)

Table 2. (continued)

Country	ISO Country Code	n	Female (%)	Mean age	Age range	Hardship index	Intercept	<i>b</i>					
								Age	Gender	Parental status	Marital status	Occupational status	Education
Malaysia	MY	2,497	49	36.09	15–80	0.30	49.79	-0.73	-1.50	-1.13	-0.17	0.00	1.04
Mali	ML	1,187	52	36.91	16–95	1.74	55.03	-0.44	-0.87	-1.05	0.38	-0.15	0.20
Mexico	MX	3,490	50	38.14	18–93	0.32	52.16	-1.14	-2.75	-2.40	-1.35	0.16	-0.18
Moldova	MD	1,028	53	42.62	18–86	0.19	50.48	-1.80	-1.18	-2.27	-0.86	0.20	0.43
Morocco	MA	1,978	49	37.48	18–87	0.30	52.41	-1.22	-2.55	-1.55	-0.96	0.15	1.18
Netherlands	NL	2,757	52	50.98	15–90	-1.02	49.97	-1.69	-3.09	-0.92	-1.46	0.01	0.66
New Zealand	NZ	739	58	50.53	18–90	-0.73	53.04	-2.21	-2.20	-1.58	-1.18	0.06	0.39
Nigeria	NG	1,759	50	31.22	18–98	1.53	56.60	-0.11	-0.38	-0.70	0.57	0.46	1.85
Norway	NO	1,015	50	45.64	18–79	-1.25	52.78	-2.12	-1.33	-1.62	-0.90	-0.08	1.12
Pakistan	PK	1,176	48	34.40	18–85	1.02	56.03	-0.23	-1.49	-0.46	-0.01	-0.60	0.66
Palestine	PS	968	51	36.32	18–86	-0.08	51.65	-1.84	-2.34	-1.00	-0.93	0.59	1.35
Peru	PE	2,575	50	38.08	18–89	0.38	50.66	-0.94	-2.87	-2.02	-0.96	-0.28	0.34
Philippines	PH	1,198	50	42.69	18–87	0.56	55.63	-0.57	-0.95	-0.49	-0.50	0.18	0.63
Poland	PL	1,926	52	46.77	18–92	-0.60	53.56	-1.43	-2.59	-2.12	-0.60	0.19	1.52
Qatar	QA	1,051	54	37.75	18–93	-0.13	52.90	-1.10	-3.08	-0.85	-0.96	0.52	0.49
Romania	RO	2,943	55	47.96	18–97	-0.38	48.91	-1.57	-2.39	-1.49	-0.93	0.58	1.47
Russia	RU	3,640	53	42.12	16–91	0.08	51.12	-1.27	-2.20	-0.94	-1.24	0.67	1.59
Rwanda	RW	2,573	49	34.31	16–90	0.97	53.30	-0.43	-1.02	0.30	-0.60	0.03	-0.24
Serbia	RS	1,099	49	42.68	18–87	-0.33	51.62	-1.92	-1.79	-2.96	-1.47	0.49	1.27
Singapore	SG	1,924	55	41.83	18–89	-0.86	54.41	-0.97	-1.78	-1.23	-0.49	0.07	0.39
Slovenia	SI	2,032	55	47.56	18–94	-0.99	52.01	-2.69	-2.79	-1.96	-1.88	0.39	1.01
South Africa	ZA	6,370	50	37.59	16–94	1.74	53.26	-0.96	-0.90	-0.18	-0.18	0.72	2.19
South Korea	KR	2,357	50	42.28	19–91	-0.71	53.69	-1.85	-1.42	-1.01	-0.57	0.11	0.48
Spain	ES	2,319	51	46.24	18–99	-0.85	52.79	-1.92	-1.54	-1.63	-1.24	0.26	0.81
Sweden	SE	2,133	51	48.05	18–85	-1.17	52.10	-1.80	-2.65	-1.70	-1.24	-0.25	0.72
Switzerland	CH	1,219	55	52.30	18–86	-1.11	50.43	-1.41	-3.63	-0.64	-1.71	0.32	0.63
Taiwan	TW	2,415	50	44.31	18–85	-0.54	47.70	-1.95	-2.33	-1.54	-1.03	-0.21	0.80
Thailand	TH	2,597	50	45.50	17–88	0.19	51.85	-0.59	-0.90	-1.13	0.25	-0.35	0.82
Trinidad	TT	1,976	55	44.18	18–94	0.46	52.05	-1.74	-2.57	-0.20	-0.81	0.30	1.03
Tunisia	TN	1,096	45	37.33	18–87	0.07	51.64	-2.73	-2.79	-2.11	-1.06	-0.20	0.88
Turkey	TR	2,875	50	37.33	18–86	0.05	52.25	-1.23	-1.67	-1.53	-0.73	0.32	1.66
Ukraine	UA	2,439	62	45.43	18–90	-0.18	50.54	-2.15	-2.59	-1.82	-0.88	0.31	0.71
Uruguay	UY	1,938	54	45.66	18–97	0.15	49.10	-1.61	-1.87	-1.87	-1.03	-0.14	0.37
United States	US	2,163	52	49.16	18–93	-0.26	51.98	-2.21	-2.64	-0.88	-1.17	0.37	0.11
Uzbekistan	UZ	1,424	61	39.09	18–89	0.27	52.73	-0.82	-2.42	-0.46	-0.70	0.36	0.90
Vietnam	VN	1,406	48	40.36	18–86	0.28	49.42	-0.96	-2.71	-0.78	-0.49	-0.40	-0.25
Yemen	YE	922	50	35.09	18–90	1.27	48.08	-1.37	-2.29	-2.06	-1.03	0.01	0.90
Zambia	ZM	1,419	49	29.72	16–80	1.57	53.37	-0.95	-1.82	-0.34	0.48	-0.42	0.96

Note: For each country, the coefficient estimate for the intercept was obtained by summing the average coefficient and that country's deviation from the average coefficient. ISO = International Organization for Standardization.

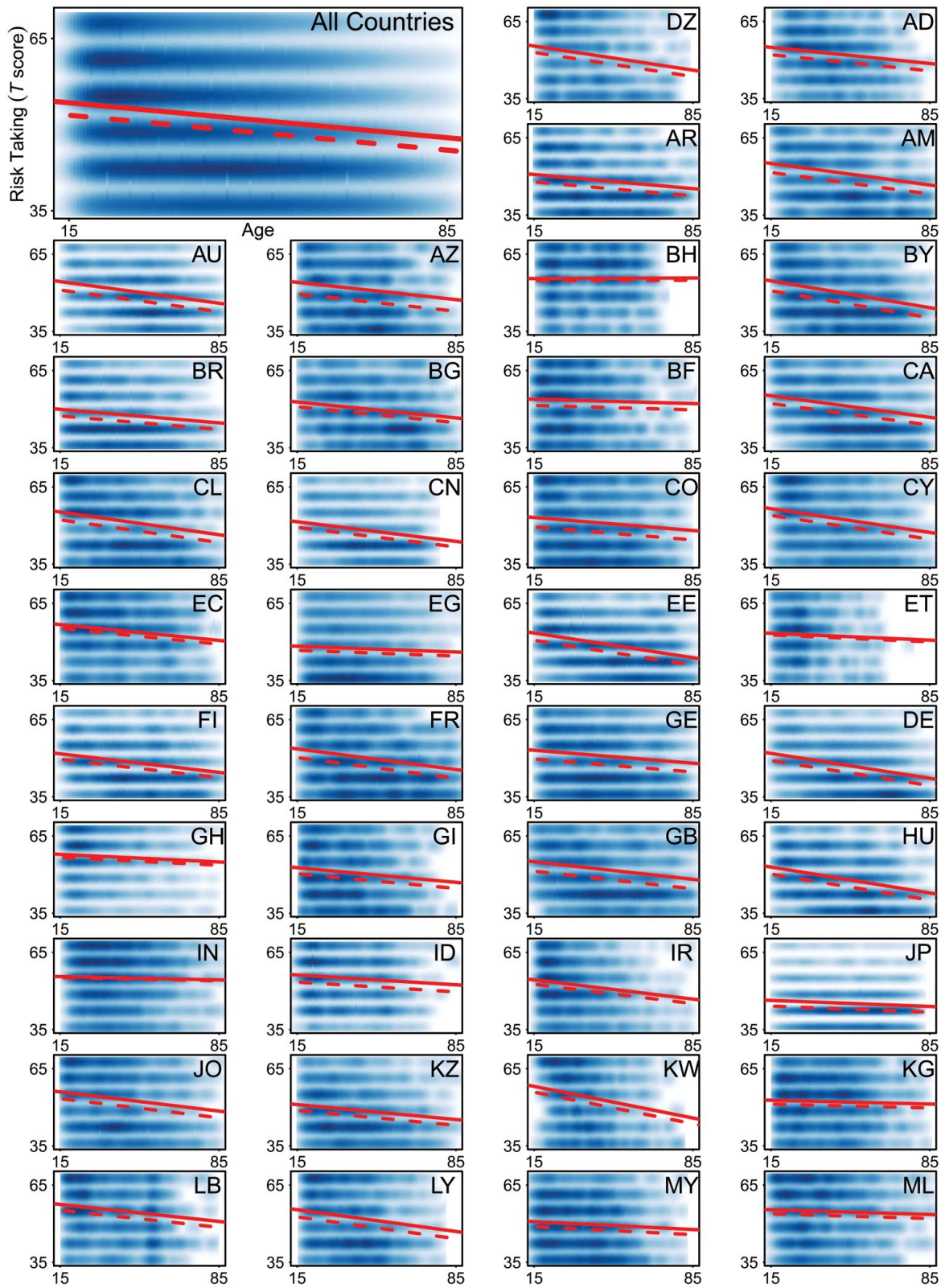


Fig. 2. (continued on next page)

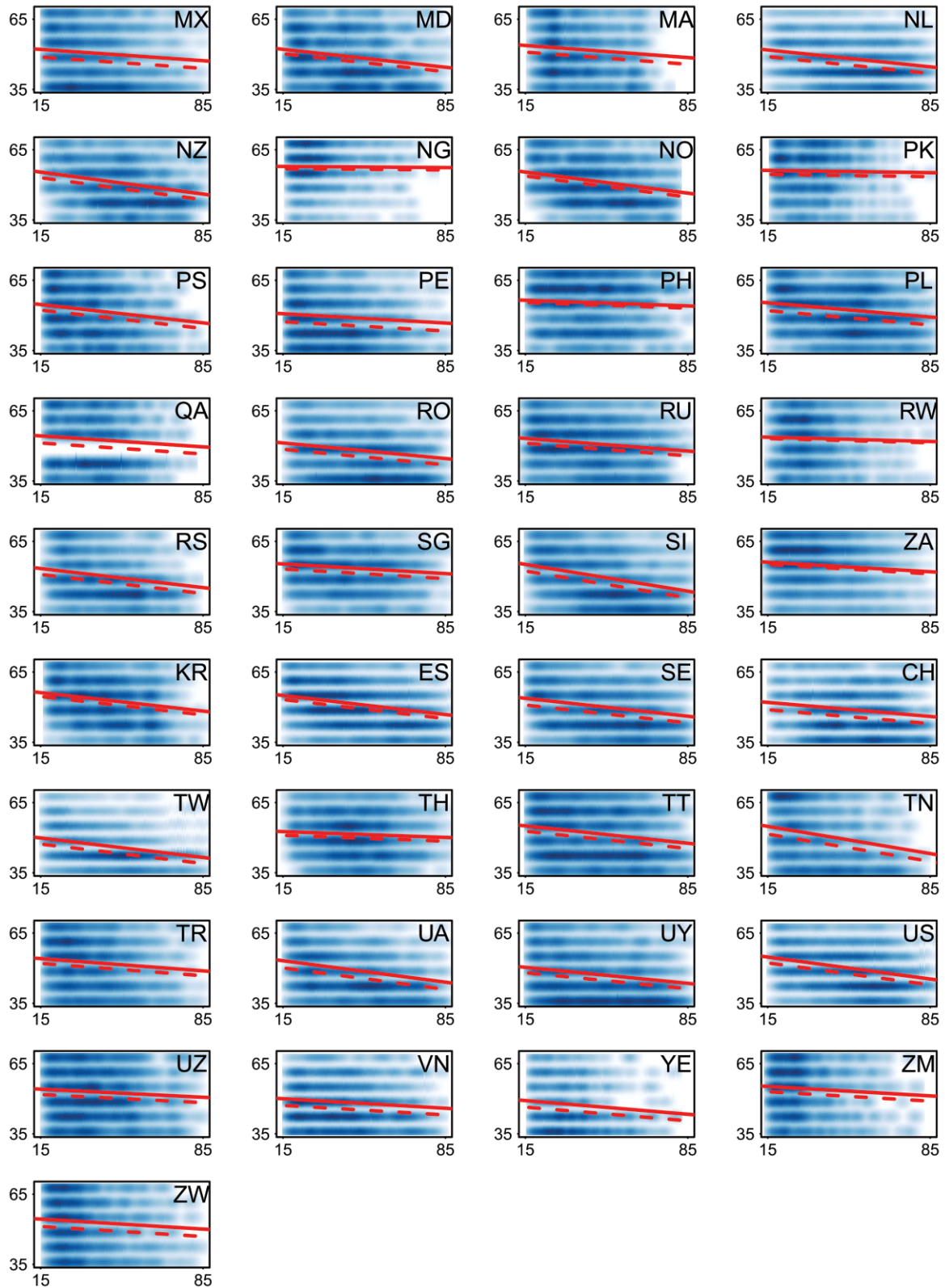


Fig. 2. Density plots of propensity for risk taking as a function of age for all countries combined and for each country separately. The blue background represents the response density; darker colors represent higher densities. Solid lines and dashed lines represent the estimated patterns of propensity for risk taking among males and females, respectively.

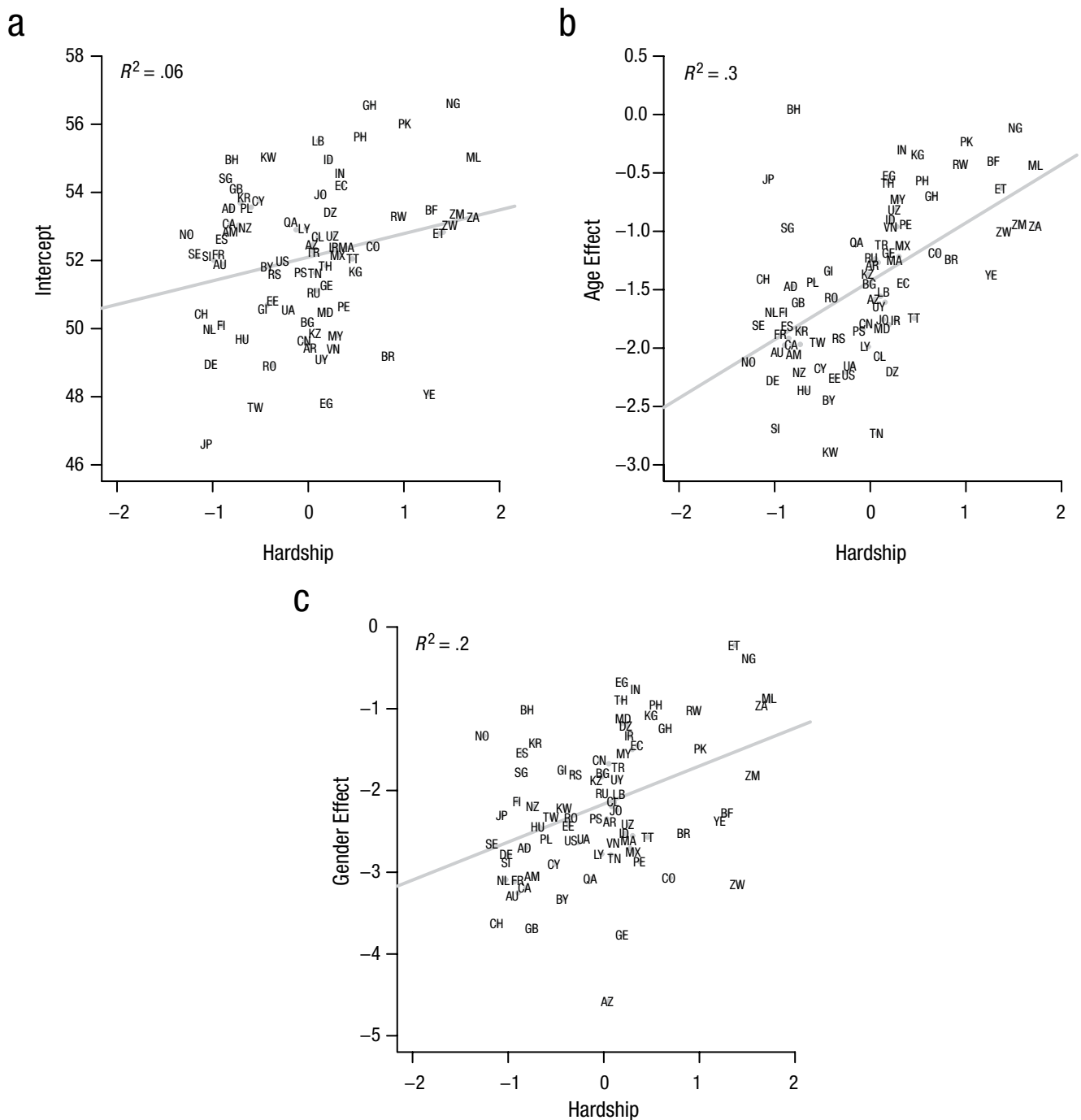


Fig. 3. Scatterplots (with best-fitting regression lines) of the relations between the hardship index and the country-specific (a) intercepts, (b) age-effect estimates, and (c) gender-effect estimates obtained from the mixed-effects regression model in which age and gender (but not hardship) were used to predict risk taking (Model 2). Values on the y-axes in (b) and (c) represent deviations from the mean estimate of the effects of age and gender, respectively. See Table 2 for explanations of the country codes.

for the use of diverse samples and cohorts (Henrich, Heine, & Norenzayan, 2010).

There are a number of limitations associated with the data we report. First, the item “adventure and taking risks are important to this person; to have an exciting life”

likely reflects a number of constructs, including propensity for risk taking, sensation seeking, and impulsivity, which are related but not necessarily identical (Zuckerman, 2007). Research on such constructs suggests that each involves distinct components that may merit

individual investigation (Cross et al., 2013; Mata, Josef, Samanez-Larkin, & Hertwig, 2011; Sharma, Markon, & Clark, 2014). Future work may consider other, more specific measures than these that disentangle potential sub-components of these traits to examine any differential life-span courses between them.

Second, we relied solely on a self-report measure, which may capture current as well as retrospective reports of risk taking and thus cannot assess whether similar patterns would be observed for behavioral measures of risk taking. Past work suggests that there is a correlation, albeit small, between self-reported propensity for risk taking and behavior in economic tasks (Dohmen et al., 2011; Lauriola, Panno, Levin, & Lejuez, 2013; Mishra & Lalumière, 2011). However, the pattern of age differences in behavioral measures of risk is considerably heterogeneous; only a few tasks suggest a decline in propensity for risk taking with increased age (Mata et al., 2011). In future work, researchers will need to systematically assess the link between self-report and behavioral measures of propensity for risk taking and do so across cultures (Rieger & Mata, 2013). Finally, a third limitation of the data is that the meaning of adventure, risk, and excitement is likely to differ between cultures, rendering direct comparisons between countries challenging. Future research may want to relate cultural differences in the perceptions of risk behavior to age differences in propensity for risk taking.

Our work also raises some questions. The monotonic age-related decline in risk taking that we found may not immediately follow from a life-history framework: If risk taking reflects the expected future trajectory of fitness prospects, life-history theory leads to the prediction that very old individuals with increased risk of morbidity and mortality could be more willing to take risks in the hope of immediate successful reproductive efforts, which would lead to a peak in propensity for risk taking in old age (Daly & Wilson, 2005). Some researchers have suggested that the monotonic reduction in risk taking across the adult life span could be explained by the possibility of resource transfers from older adults (i.e., individuals with lower reproductive value) to their offspring with higher reproductive value (Rogers, 1994). Future work that tracks the risk patterns of older individuals in combination with resource transfer behavior and goals could perhaps test this possibility.

In conclusion, age is associated with reduced propensity for risk taking in a quasi-universal fashion. Nevertheless, the considerable variation in the link between age and propensity for risk taking is systematically associated with local hardship. Specifically, high-risk ecologies favor reproductive strategies associated with increased risk taking across the life span and a flattening of the age-risk curve. Age-risk relations appear to

reflect, among other factors, individuals' adjustment to the characteristics of local ecologies and cannot be understood without reference to the demands and affordances of the environment.

Author Contributions

R. Mata developed the study concept, conducted the analyses, and drafted the manuscript. A. K. Josef and R. Hertwig provided critical revisions. All authors approved the final version of the manuscript for submission.

Declaration of Conflicting Interests

The authors declared that they had no conflicts of interest with respect to their authorship or the publication of this article.

Funding

This work was supported by Swiss National Science Foundation Grant 100014-156172.

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