

1 Title: The genomic determinants of adaptive evolution in a fungal
2 pathogen

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22 fungi

23 **Runing Title:** Adaptive evolution in a fungal plant pathogen

24 **Abstract:**

25 Antagonistic host-pathogen co-evolution is a determining factor in the outcome of
26 infection and shapes genetic diversity at the population level of both partners. While the
27 molecular function of an increasing number of genes involved in pathogenicity is being
28 uncovered, little is known about the molecular bases and genomic impact of hst-pathogen
29 coevolution and rapid adaptation. Here, we apply a population genomic approach to infer
30 genome-wide patterns of selection among thirteen isolates of the fungal pathogen
31 *Zymoseptoria tritici*. Using whole genome alignments, we characterize intragenic
32 polymorphism, and we apply different test statistics based on the distribution of non-
33 synonymous and synonymous polymorphisms (pN/pS) and substitutions (dN/dS) to (1)
34 characterise the selection regime acting on each gene, (2) estimate rates of adaptation and (3)
35 identify targets of selection. We correlate our estimates with different genome variables to
36 identify the main determinants of past and ongoing adaptive evolution, as well as purifying and
37 balancing selection. We report a negative relationship between pN/pS and fine-scale
38 recombination rate and a strong positive correlation between the rate of adaptive non-
39 synonymous substitutions (ω_a) and recombination rate. This result suggests a pervasive role of
40 Hill-Robertson interference even in a species with an exceptionally high recombination rate
41 (60 cM/Mb). Moreover, we report that the genome-wide fraction of adaptive non-synonymous
42 substitutions (α) is ~ 44%, however in genes encoding determinants of pathogenicity we find a
43 mean value of alpha ~ 68% demonstrating a considerably faster rate of adaptive evolution in
44 this class of genes. We identify 787 candidate genes under balancing selection with an
45 enrichment of genes involved in secondary metabolism and host infection, but not predicted
46 effectors. This suggests that different classes of pathogenicity-related genes evolve according
47 to distinct selection regimes. Overall our study shows that sexual recombination is a main
48 driver of genome evolution in this pathogen.

49 Introduction

50 Antagonistic host-pathogen interactions drive co-evolutionary dynamics between
51 pathogens and their hosts. Signatures of selection in genomes inform about mechanisms of
52 evolution and identify targets of selection at interacting loci (Möller & Stukenbrock 2017). In
53 general, genome studies of microbial pathogens have focused on rapidly evolving genes
54 involved in pathogenicity, such as “effector” genes encoding proteins that interfere with host
55 defenses and may determine host range of the pathogen (Lo Presti *et al.* 2015). Effector genes
56 of fungal pathogens have frequently been found to associate with repetitive DNA and it is
57 proposed that repeat rich genome compartments provide particularly favorable environments
58 for rapid evolution of new virulence specificities (e.g. (Ma *et al.* 2010; Spanu *et al.* 2010;
59 Klosterman *et al.* 2011; Daverdin *et al.* 2012)). Repetitive DNA may locally increase mutation
60 rate and contribute to gene duplications and structural variation among alleles. Yet little is
61 known about the factors that shape genome evolution in fungal pathogens, in particular the
62 interplay of mutation, natural selection, genetic drift and, for sexually reproducing species,
63 recombination along the genome of fast evolving pathogens.

64 Evolution of genes involved in the antagonistic interaction with the host can be
65 driven by positive selection whereby new alleles recurrently replace existing alleles in response
66 to allelic changes in the host, a scenario termed arms race evolution (Van Valen 1973; Tellier *et al.*
67 *al.* 2014). Variation in pathogenicity related genes can also be maintained by balancing
68 selection, a trench-warfare scenario where a set of alleles are maintained in the population over
69 long evolutionary times (Stahl *et al.* 1999). In plants, balancing selection has been described as
70 a main driver of evolution in genes encoding resistance proteins (e.g. (Tian *et al.* 2002; Huard-
71 Chauveau *et al.* 2013)), however the importance of balancing selection in pathogen genomes is
72 less understood.

73 Population genomic data reflect signatures of past and on-going selection acting on
74 the organism. While past signatures of selection can be related to ecological specialization, on-

75 going positive selection reflects local adaptation in the existing population. In plant pathogens,
76 signatures of on-going selection can reflect host-pathogen arms race or trench warfare
77 evolution as well as adaptations to other local environmental conditions, in agricultural
78 systems notably fungicide treatments (Hayes *et al.* 2015; Delmas *et al.* 2017). Rapid adaptation
79 is fueled primarily by large effective populations sizes as well as high recombination and
80 mutation rates that promote the emergence, spread and fixation of new advantageous alleles. In
81 research of Eukaryote pathogen genome evolution, most studies have used genome scans to
82 detect outlier genes and genomic regions. Based on the finding of high variability in specific
83 genome compartments it has been proposed that plant pathogens represent exceptional outliers
84 in terms of evolutionary rates (Raffaele & Kamoun 2012; Upson *et al.* 2018). Nevertheless,
85 quantitative measures of evolution in pathogen genomes are missing to test this hypothesis.

86 In this study we have addressed the impact of selection on genome evolution in a
87 fungal plant pathogen, *Zymoseptoria tritici*. *Z. tritici* infects wheat and reproduces by the
88 production of asexual spores in infected leaf tissues and by sexual recombination between
89 isolates of opposite mating type (Waalwijk *et al.* 2002). Previous studies based on mating
90 experiments and population genomic data have reported exceptional high recombination rates
91 in this species (~60 cM/Mb), including intragenic recombination hotspots that underline the
92 putative key role of recombination in evolution of this species (Croll *et al.* 2015; Stukenbrock
93 & Dutheil 2017). The genome of *Z. tritici* consists of thirteen core and several accessory
94 chromosomes. The latter are present at variable frequencies in different individuals,
95 constituting a particular case of karyotypic polymorphism (Goodwin *et al.* 2011). The
96 accessory chromosomes comprise repeat rich, heterochromatic DNA with a low gene content
97 and they encode traits with quantitative effects on virulence (Grandaubert *et al.* 2015;
98 Schotanus *et al.* 2015; Habig *et al.* 2017). We have previously shown that evolutionary rates on
99 these chromosomes are particularly high suggesting that genes on the accessory chromosomes
100 in general evolve under less selective constraints (Stukenbrock *et al.* 2010). Previous studies
101 based on comparative population genomic analyses of *Z. tritici* and two closely related species,

102 *Zymoseptoria pseudotritici* and *Zymoseptoria ardabiliae* used genome-wide estimates of non-
103 synonymous and synonymous divergence to identify past species-specific signatures of
104 selection in the wheat pathogen (Stukenbrock *et al.* 2010, 2011). Functional characterization of
105 some of these genes revealed amino acid substitutions important for *in planta* development and
106 asexual spore formation in the wheat-adapted pathogen, and thereby confirmed the use of
107 evolutionary predictions to identify functionally relevant traits (Poppe *et al.* 2015).

108 We here apply a population genomics approach to infer genome-wide signals of
109 natural selection, including purifying, positive and balancing selection among thirteen isolates
110 of *Z. tritici* collected from bread wheat in Europe and the Middle East. We specifically ask to
111 which extent recombination contributes to adaptive evolution in a sexual pathogen. Our
112 analyses based on more than 1.4 million single nucleotide polymorphisms (SNPs) and
113 including 700,000 coding sites allow us to identify past and on-going signatures of selection in
114 the genome of *Z. tritici*. Our analyses reveal a strong importance of recombination in gene
115 evolution, for both positive and negative selection, and a particularly high rate of adaptive
116 substitutions in genes encoding putative effectors. On the other hand, balancing selection is
117 more prevalent in genes located in repeat-rich parts of the genome implying that transposable
118 elements also contribute to the maintenance of genetic variation. Overall, our analyses
119 underline the potential of rapid adaptation of virulence related traits in this important
120 agricultural pathogen.

121 **Results and Discussion**

122 **Population structure of *Z. tritici* correlates with geographical origin**

123 We generated a population genomic dataset of thirteen *Z. tritici* isolates obtained
124 from different field populations in Europe and Iran (Table S1). Given the high extent of
125 structural variation in genomes of *Z. tritici* isolates, we *de novo* assembled and aligned the
126 thirteen genomes. After filtering (see Material and Methods), the resulting multiple genome

127 alignment of ~27 Mb (Table S2) comprised a total of 1,489,362 SNPs of which approximately
128 50% locate in protein coding regions. The SNP data was used to compute the overall genetic
129 diversity of the sample showing a mean value of $\pi = 0.022$ per site. Importantly, the multiple
130 genome alignment of *de novo* assembled genomes is a priori exempt of paralogous sequences.

131 We first used the genomic data to investigate the relationship of the *Z. tritici* isolates.
132 We assessed population genetic structure by analyzing the ancestral recombination graph of the
133 thirteen genomes. To this end, we slid 10 kb windows along the multiple genome alignment,
134 and estimated the genealogy for each window. The resulting 1,850 trees were combined into a
135 super tree (Fig. 1A). If the sample of genomes is taken from a panmictic population with
136 recombination, the super tree is expected to be a star tree. However, here we observe at least
137 two clusters, one comprising all European isolates and the other comprising two isolates
138 collected in Iran (Fig. 1A). We further investigated population structure using the program
139 ADMIXTURE (Alexander *et al.* 2009) and found the strongest support for a model with two
140 ancestral populations supporting the tree-based clusters of European and Iranian isolates (Fig.
141 1B). This pattern is consistent with some extent of geographical barriers preventing gene flow
142 between European and Middle East *Z. tritici* populations, and possibly local adaptation to
143 distinct host genotypes, i.e. wheat cultivars.

144 **Recombination contributes to high rates of adaptive evolution in *Z. tritici***

145 We next aimed to obtain a quantitative assessment of adaptive substitutions in the
146 genome of *Z. tritici*. To this end, we first estimated the non-synonymous and synonymous
147 divergence dN and dS using a genome alignment of *Z. tritici* and its sister species *Z.*
148 *ardabiliae*. Furthermore, we used the *Z. tritici* SNP data to compute the unfolded site frequency
149 spectrum (SFS) of synonymous and non-synonymous sites using *Z. ardabiliae* as outgroup.
150 The synonymous nucleotide diversity was on average over all genes 0.054, reflecting the high
151 diversity in this species. By contrasting divergence and polymorphism data, we estimated the
152 parameters α (proportion of adaptive non-synonymous substitutions, dN_a / dN) and ω_a

153 (proportion of the dN / dS ratio that is attributable to adaptive mutations, dN_a / dS). The SFS is
154 strongly affected by demography and the presence of slightly deleterious mutations segregating
155 at low frequencies (Eyre-Walker & Keightley 2007). State-of-the-art statistical methods
156 account for the latter by modeling the distribution of fitness effects (DFE) of mutations
157 (Gossmann *et al.* 2010; Galtier 2016). Potential confounding demographic factors such as
158 variable population size, population structure and linked selection are accounted for by fitting
159 additional parameters to accommodate deviations from a constant size neutral model of
160 evolution. This generic correction assumes that these factors affect both synonymous and non-
161 synonymous mutations equivalently.

162 We estimated α as well as ω_a , the rate of adaptive substitutions, using four distinct
163 DFE models accounting for mutations with both slightly deleterious and beneficial effects (see
164 Materials and Methods) and found that the Gamma-Exponential model best fitted our data
165 (Table 1) in agreement with studies from animals (Galtier 2016). This suggests the existence of
166 slightly deleterious, as well as slightly beneficial segregating mutations in the genome of *Z.*
167 *tritici* (Table 1). The estimates provide an α value of 35% as a genome average, and an ω_a value
168 of 0.044. Both values are in the range of what is observed for Mammals (with the exception of
169 Primates) but considerably higher than estimates from plants (Gossmann *et al.* 2010; Galtier
170 2016). In candidate effector genes, however, the rate of mutations fixed by selection is more
171 than twice as high as in non-effector genes (ω_a equal to 0.120 vs. 0.048, Table 1), with 60% of
172 non-synonymous substitutions in these genes inferred to be adaptive. This average estimate is
173 close to the highest values reported in animals (Galtier 2016), and reflects the strong selective
174 pressure acting on these genes. We note that estimates of α and ω_a are slightly higher when
175 only non-effector genes are used (6,639 genes) than when using the complete gene set (6,767
176 genes), a small difference that likely results from sampling variance. In order to assess the
177 significance of the observed differences between effector and non-effector genes while
178 accounting for the difference in gene numbers in both categories (128 and 6,639 genes,
179 respectively), we performed a bootstrap analysis where we estimated α and ω_a in random

180 samples of 128 genes in each category. The results of 100 resamples are shown on Fig. 2,
181 revealing a highly significant difference between the two distributions (Wilcoxon test, p-value
182 $< 2.2 \cdot 10^{-16}$) and confirming the significantly higher rate of adaptation in effector genes.

183 We hypothesized that recombination could be an important driver of adaptive
184 evolution in *Z. tritici*. Previous inference of recombination maps in *Z. tritici* based on
185 experimental crosses and population genomic data have revealed exceptionally high rates of
186 recombination in this species (Croll *et al.* 2015; Stukenbrock & Dutheil 2017). To assess the
187 role of recombination in adaptive evolution of *Z. tritici* we used the recombination maps
188 generated in these previous studies. Genetic maps resulting from crossing experiments allow
189 inference of the recombination rate r (measured as cM / Mb), but are limited in resolution.
190 Conversely, linkage disequilibrium-based maps generated from population genomic data offer
191 an improved resolution, but only allow inference of $\rho = 4 \cdot N_e \cdot r$, where N_e is the effective
192 population size. As such, ρ is a proxy for r that is affected by both selection and demography.
193 We clustered all analyzed genes according to their r and ρ values, and estimated α and ω_a for
194 each case using the Gamma-Exponential distribution of fitness effects. In order to assess the
195 variance of our estimates and their robustness to the sampled genes, we further conducted a
196 bootstrap analysis where we sampled genes in each category 100 times. We report a significant
197 positive correlation between α (averaged over 100 bootstrap replicates) and r (Kendall's tau =
198 0.31, p-value = 0.004354) and ω_a and r (Kendall's tau = 0.31, p-value = 0.006041). We note
199 that similar correlations are observed when ρ is used instead of r , or when effector genes are
200 discarded (Supplementary Data). These results suggest that a higher recombination rate favors
201 the fixation of adaptive mutations, as expected under a Hill-Robertson interference scenario,
202 where selected mutations reduce the effective population size at linked loci (Hill & Robertson
203 1966; Marais & Charlesworth 2003).

204 We further explored the relationship between α , ω_a and r . We fitted four models:
205 linear (as in (Campos *et al.* 2014)), power law, curvilinear (as in (Castellano *et al.* 2016)), and
206 logarithmic (see Materials and Methods). While we find a higher support for the logarithmic

207 model (Fig. 3), the effect is very weak and our data does not allow further estimation of the
208 asymptotic value (Castellano *et al.* 2016). When using ρ instead of r , the curvilinear model is
209 preferred when all genes are considered, but the power law offers a better fit when effectors are
210 excluded (Supplementary Material).

211 In summary, our results represent a quantitative assessment of adaptive evolution in
212 the genome of a fungal pathogen and reveal a strong role of recombination on adaptation (Fig.
213 3). The exceptionally high rate of adaptation in effector genes (Fig. 2) likely reflects arms race
214 evolution driven by the antagonistic interaction of *Z. tritici* and its host.

215 **Local rates of recombination are correlated with the strength of purifying** 216 **selection revealing pervasive background selection**

217 We next addressed the genome-wide strength of purifying selection in protein coding
218 genes of *Z. tritici* using the ratio of non-synonymous to synonymous polymorphisms (pN / pS
219 ratio). We computed pN and pS for each gene as the average pairwise heterozygosity
220 (Romiguier *et al.* 2014; Ellegren & Galtier 2016). To investigate which genome parameters
221 impact the strength of purifying selection in *Z. tritici*, we compared the pN / pS ratio for each
222 gene to 1) the mean gene expression, 2) the GC content at third codon positions (GC3), 3) the
223 protein length, 4) the local recombination rate, 5) the density in protein coding sites and 6) the
224 density in transposable elements. We used *Z. tritici* gene expression data from early host
225 colonization (four days after spore inoculation on leaves of seedlings of a susceptible wheat
226 host) and *in vitro* growth (Kellner *et al.* 2014). Recombination rates were averaged in 20 kb
227 windows and recombination was analysed independently as r (Croll *et al.* 2015) and ρ
228 (Stukenbrock & Dutheil 2017). We further considered whether the gene 7) is an effector
229 candidate and 8) is located on an accessory chromosome (Fig. 4). We restricted our analysis to
230 genes for which pN / pS could be computed (6,627 genes, see Materials and Methods) and for
231 which pN / pS was estimated to be < 1 (6,621 genes).

232 We identify several variables that significantly impact the strength of purifying
233 selection (summarized in Table 2). Mean gene expression and GC at third codon position have
234 the strongest effect on pN / pS (Fig. 4A and 4B), displaying highly significant negative
235 correlations (Kendall's tau = -0.369 and -0.237 respectively, p-values < $2.2 \cdot 10^{-16}$ in both cases).
236 Consistent with this observation, studies in yeast and bacteria have previously documented a
237 strong impact of expression levels on gene evolution whereby highly expressed genes are more
238 conserved reflected as lower pN / pS values (Drummond *et al.* 2006; Liao *et al.* 2006). GC3
239 and mean gene expression are intrinsically highly correlated (Kendall's tau = 0.222, p-value <
240 $2.2 \cdot 10^{-16}$), possibly reflecting biases in codon usage whereby optimal codons are GC-rich at
241 their third position (Fig. S1), as also observed in other organisms (Duret & Mouchiroud 1999).
242 An alternative explanation for the effect of the GC content on pN / pS could be a possible
243 indirect effect of recombination as we also observe a positive correlation of GC3 and the
244 recombination rate (Kendall's tau = 0.097, p-value < $2.2 \cdot 10^{-16}$). A similar correlation of
245 recombination and GC3 is found in other organisms (Duret 2002). In *Saccharomyces*
246 *cerevisiae* this correlation has been explained by the impact of biased gene conversion on
247 sequences evolution (Birdsell 2002). However, a thorough search for signatures of GC-biased
248 gene conversion did not find any pervasive effect of this phenomenon in *Z. tritici* (Stukenbrock
249 & Dutheil 2017). The relationship between pN / pS and GC3 is therefore more likely a by-
250 product of the correlation with gene expression.

251 Protein size is slightly positively correlated with pN / pS (Table 2), although the
252 effect is due to very short proteins being more conserved and the observed effect disappears when
253 testing only proteins with > 100 amino acids (excluding 348 proteins out of 6,621, Kendall's
254 tau = 0.012, p-value = 0.1443, Fig. 4C). Gene density, estimated in a 50 kb regions centered on
255 the gene (see Material and Methods), does not have a significant effect on pN / pS (Kendall's
256 tau = 0.0067, p-value = 0.4127, Fig. 4D). The genome of *Z. tritici* is compact and uniform in
257 terms of gene localization, and the distribution of gene density is almost normal with a median
258 around 54%. This likely explains that gene density does not have an impact on strength of

259 purifying selection. Conversely, we observe a significant negative correlation between pN / pS
260 and recombination rate r (Kendall's tau = -0.031, p-value = $1.85 \cdot 10^{-4}$, Fig. 4E) or ρ (Kendall's
261 tau = -0.039, p-value = $1.52 \cdot 10^{-6}$, Fig. 4F and Fig. 5). These results are in agreement with a
262 model of background selection, where purifying selection at linked loci with low
263 recombination rates reduces the local effective population size, therefore reducing the efficacy
264 of selection and allowing slightly deleterious mutations to spread more frequently than at loci
265 with high recombination rates (Charlesworth *et al.* 1993; Nordborg *et al.* 1996).

266 Background selection is notably expected to be stronger in regions of higher density
267 of coding sites, implying that the negative correlation between recombination and pN / pS
268 should be higher in gene-dense regions. To further test this effect of coding site density, we
269 split our gene set in two subsets, whether the density of coding sites was below (low-density
270 set) or above (high-density set) the median. For the low-density set, we report a marginally
271 significant negative correlation between r and pN / pS (Kendall's tau = -0.022, p-value =
272 0.05856), while the correlation is stronger and significant for the high-density set (Kendall's
273 tau = -0.039, p-value = $7.89 \cdot 10^{-04}$). These results provide evidence that background selection
274 impacts the genome of *Z. tritici*, and support a central role of recombination in the removal of
275 non-adaptive mutations in the genome of *Z. tritici* consistent with patterns described in other
276 species such as *Drosophila melanogaster* (Campos *et al.* 2014). We note, however, that the
277 effect of recombination is smaller than the one of functional variables such as mean expression.
278 This is likely related, we hypothesize, to the globally high level of recombination and reduced
279 linkage throughout the genome of *Z. tritici*.

280 One part of the *Z. tritici* genome where recombination is low is the accessory
281 chromosomes (Stukenbrock & Dutheil 2017). This low recombination rate is reflected in our
282 estimates of purifying selection. We find a significantly higher pN / pS ratios in genes located
283 on accessory chromosomes (Wilcoxon rank test, p-value = 0.0162 , Fig. 4G), and on the right
284 arm of chromosome 7 (Fig. 5), a genomic region predicted to be an ancestral accessory
285 chromosome fused with a core chromosome (Schotanus *et al.* 2015). Accessory chromosomes

286 have a reduced effective population size due to their presence/absence variation among
287 individuals, resulting in a reduced efficacy of selection and a higher pN / pS ratio on these
288 chromosomes..

289 Finally, we compared the strength of purifying selection of effector and non-effector
290 genes, and find that genes predicted to encode effector proteins have a significantly higher pN /
291 pS ratio compared to other genes (Wilcoxon rank test, p-value = $1.5 \cdot 10^{-14}$, Fig. 4H). We
292 speculate that this pattern is due to the fast evolution through positive selection of this
293 particular category of genes, and the higher pN / pS ratio in these genes reflects the fixation of
294 slightly deleterious mutations by linkage.

295 **Detection of on-going balancing selection in *Z. tritici* identifies candidate** 296 **pathogenicity factors**

297 The recurrent interaction with different host genotypes can confer the maintenance of
298 multiple alleles at selected loci in the pathogen population. To identify specific sites and genes
299 in the *Z. tritici* genome showing signatures of balancing selection, we fitted models of codon
300 sequence evolution as implemented in the CodeML program of the PAML package to detect
301 genes with significant signatures of balancing selection using likelihood ratio tests (Yang
302 2007). Two models are typically compared: a model with sites evolving only under neutrality
303 or purifying selection, with an ω ratio (non-synonymous rate of polymorphisms / synonymous
304 rate of polymorphisms) equal or below one (neutral model with purifying selection), and a
305 model allowing for some sites to evolve under positive selection with a ω ratio above one
306 (positive selection model). A likelihood ratio test (LRT) is then used to test for the occurrence
307 of positive selection. When applied to population data, sites with $\omega > 1$ can reflect balancing
308 selection (Anisimova *et al.* 2001). We fitted codon models for genes present in at least three
309 isolates (83% of genes located on core chromosomes and 31% of genes on the accessory
310 chromosomes, Table S3). After correcting for multiple testing, we identified a final set of 787

311 genes (including 24 on the accessory chromosomes) evolving under balancing selection (false
312 discovery rate < 0.01, Table S3).

313 As selection tests based on codon model comparison were previously shown to
314 potentially suffer from an inflated false discovery rate (FDR) in the presence of recombination
315 (Anisimova *et al.* 2003), we conducted simulations with parameters reflecting the
316 characteristics of our data set (see Material and Methods). In agreement with previous results,
317 we report an increase in FDR in the presence of recombination within the gene (Fig. 6). While
318 our results appear to be relatively independent of the level of diversity in the alignment, we
319 report a strong effect of the number of sites in the alignment: for a given number of
320 recombination events, we see a higher FDR in long compared to short genes. This suggests that
321 the recombination rate, which is lower in longer genes for a given number of recombination
322 events, is not the only determinant of erroneous rejection of the null model. Large alignments,
323 on the contrary, carry more statistical signal (Anisimova *et al.* 2001) and can lead to strong
324 support of the wrong model in case an incorrect tree is provided, an effect that is independent
325 of the actual number of recombination events. In agreement with this hypothesis, we see that
326 the FDR also increases with the number of sequences in the alignment (Fig. 6). To assess the
327 extent of false discovery in our analysis, we sorted genes for which all 13 individuals were
328 present (6,627 genes) according to their corresponding protein lengths: more than 100, 500,
329 1,000 and 2,000 amino acids, respectively. We find that the proportion of genes significantly
330 rejecting the null hypothesis of no positive selection is systematically higher than the
331 maximum observed FDR for the corresponding length (Fig. 6). This suggests that false
332 discovery due to recombination does not explain all of our candidates and that the selection test
333 was able to capture biological signal.

334 To further assess the gene-specific selection regime, we computed Tajima's D for
335 each gene in our data set. We find that the distribution of Tajima's D is globally shifted toward
336 negative values (Fig. 7), as expected for coding sequences under purifying selection.
337 Furthermore, *Z. tritici* was previously found to have undergone population expansion since the

338 domestication of wheat and speciation of the pathogen (Stukenbrock *et al.* 2007). This
339 population expansion also contributed to the overall negative Tajima's D values. We further
340 observe that there is no significant differences between genes encoding predicted effector
341 proteins and others (Wilcoxon rank test, p-value = 0.1597). Genes predicted to be under
342 balancing selection by PAML, however, display significantly higher Tajimas' D values
343 (Wilcoxon rank test, p-value < $2.2 \cdot 10^{-16}$), a typical signature of balancing selection (Tajima
344 1989). However, given the population structure that we observe, it cannot be excluded that for
345 some of these genes, the signal of the LRT results from population differentiation. In such case,
346 the detected gene could be under positive selection resulting from local adaptation. In the
347 following, we further investigate the genome distribution and biological function of genes
348 detected by PAML.

349 In several pathogen genomes, rapidly evolving genes are found clustered in
350 particular genomic environments often associated with repetitive sequences (e.g (Raffaele *et*
351 *al.* 2010; Dutheil *et al.* 2016)). To address whether the same pattern is found in *Z. tritici*, we
352 assessed the spatial distribution of genes under balancing selection along chromosomes (see
353 Materials and Methods). Of the 787 positively selected genes, 240 are located within a distance
354 of 5 kb from each other and thereby form 108 clusters containing two to four genes with
355 signatures of positive selection. At the genome scale, clusters containing two or three genes do
356 not show a significant pattern as the same pattern can be obtained by randomly distributing the
357 positively selected genes across the genome. However, there are two significant clusters (p-
358 value = $1.8 \cdot 10^{-3}$) containing four and eight genes, respectively. The clusters comprising four
359 genes is located in a 24 kb region of chromosome 5, and the cluster comprising eight genes in a
360 31 kb region of chromosome 9. For genes in both clusters no functional relevance can be
361 assigned, but the clusters represent interesting candidates for future functional studies.

362 **The genomic determinants of balancing selection in *Z. tritici***

363 In order to test which factor drives the occurrence of balancing selection in the *Z.*
364 *tritici* genome, we fitted (generalized) linear models. We assessed the impact of 1) the
365 recombination rate, 2) the density in protein coding sites, 3) the density in transposable
366 elements (TE), 4) whether the gene is predicted to encode an effector protein, and 5) the mean
367 gene expression (see Material and Methods). We fitted a binary logistic model, where the
368 response variable is whether a gene is predicted to be under positive balancing selection by
369 PAML. We find that positive selection is less likely at highly expressed genes (Table 3), an
370 effect that, we hypothesize, is due to highly expressed genes being on average more
371 constrained (see above).

372 We find a significant effect of effector encoding genes (Table 3), that is, effector
373 genes are, intriguingly, less likely to be under balancing selection than non-effector encoding
374 genes. As described above, we find that effector genes tend to undergo a higher rate of adaptive
375 evolution, and our analyses thereby suggest that distinct categories of genes are evolving under
376 arms race (recurrent selective sweeps) and trench-warfare (balancing selection) scenarios in the
377 pathogen genome. The rate of recombination has a weak but significant positive effect on the
378 occurrence of positive selection, *i.e.* genes detected to be under balancing selection are more
379 frequent in highly recombining regions (Table 3). This effect can be interpreted as a better
380 efficacy of selection in highly recombining regions, but can also be due to an increased false
381 discovery rate in the presence of recombination. In order to disentangle the two hypotheses, we
382 fitted a similar linear model with Tajima's D as a response variable and find consistent results
383 where recombination rate has a significant positive effect on Tajima's D (Table 3).

384 We further extended our analyses of the genes predicted to be under balancing
385 selection by characterizing known protein domains. From the 787 candidate genes, 602 of the
386 encoded proteins (76.5%) have an *in silico* attributed function or harbor known protein
387 domains. We conducted a PFAM domain enrichment analyses and identified 21 significantly
388 enriched domains (FDR \leq 0.05) (Table S4). These domains can be grouped into different

389 categories based on their associated molecular function and include carbohydrate-active
390 enzymes (CAZymes), polyketide synthases (PKS), non-ribosomal peptide synthetases (NRPS),
391 and cellular transporters (Table S4). Several of these categories are relevant for the pathogen to
392 interact with its host. For example, CAZymes are proteins involved in the break down of
393 glycosidic bonds contained in plant cell walls (André *et al.* 2014), and PKS and NRPS are
394 multimodular enzymes involved in the biosynthesis of secondary metabolites of which many
395 also have been shown to be involved in virulence of plant pathogenic fungi (Howlett 2006).
396 Among the positively selected *Z. tritici* genes, we also searched for homologs of known
397 virulence factors described in other plant pathogens. The PFAM domain PF14856 corresponds
398 to the mature part of a virulence factor Ecp2 described in the tomato fungal pathogen
399 *Cladosporium fulvum* (Van den Ackerveken *et al.* 1993). Ecp2 has been described in several
400 other plant pathogens (Stergiopoulos *et al.* 2010) and has three homologs in *Z. tritici*. We find
401 that two of these homologs comprise sites under positive selection supporting a virulence
402 related role of this gene also in this wheat pathogen.

403 **Signatures of past selection in *Z. tritici* and related species**

404 In a previous study the evolutionary history of *Z. tritici* was inferred using a whole
405 genome coalescence analyses (Stukenbrock *et al.* 2011). We showed that divergence of *Z.*
406 *tritici* and its sister species *Z. pseudotritici* occurred recently and likely coincides with the
407 onset of wheat domestication and thereby specialization of *Z. tritici* to a new host. We
408 hypothesize that genes important for the colonization of distinct hosts have been under
409 selection during the divergence of *Zymoseptoria* species. To infer signatures of past selection
410 we applied the branch model implemented in the program package PAML to estimate the
411 branch-specific dN / dS ratio for core *Zymoseptoria* genes (present in four species *Z. tritici*, *Z.*
412 *pseudotritici*, *Z. ardabiliae* and *Z. brevis*) (Yang & Nielsen 1998; Grandaubert *et al.* 2015). The
413 branch-specific dN / dS ratios reflect the proportion of non-synonymous to synonymous
414 substitutions accumulated in each branch of the *Zymoseptoria* phylogeny. Our analyses

415 identified 47 genes with a dN /dS ratio > 1 on the *Z. tritici* branch indicative of an increased
416 non-synonymous divergence, 54 genes in *Z. pseudotritici*, 60 genes in *Z. brevis* and 15 genes in
417 *Z. ardabiliae* (Table 4). Based on their putative function in host-pathogen interactions, we
418 hypothesized that some positively selected genes encode secreted proteins and putative
419 effectors. In order to test this hypothesis, we fitted linear models with branch specific dN / dS
420 ratios as response variables and whether the corresponding gene family encodes an effector or
421 not in *Z. tritici* as explanatory variable. For all four extant species and the common ancestor of
422 *Z. tritici* and *Z. pseudotritici*, we find a significantly higher dN / dS ratio for genes encoding
423 predicted effector proteins (Table 4), in agreement with the general observation that effector-
424 encoding genes are fast-evolving. For *Z. tritici* only we report that effector-encoding genes are
425 more likely to have a dN / dS > 1. For *Z. ardabiliae*, we only find 15 genes under positive
426 selection, and none in effector-encoding candidate genes.

427 We next performed a PFAM domain enrichment analysis for the positively selected
428 genes with a predicted function to address if some functional domains are significantly
429 enriched in this set of genes (FDR < 0.05) (Table S5). The majority of genes encode proteins of
430 unknown function, however among the functionally characterized proteins we find a gene
431 encoding a regulator of chromosome condensation that was previously described to be
432 functionally relevant for virulence in *Z. tritici* (Poppe *et al.* 2015). Our analyses reveal an
433 enrichment of genes encoding proteinases, one gene encoding Lysin motifs (LysM) already
434 described as an effector in *Z. tritici* and other fungal pathogens (de Jonge *et al.* 2010; Marshall
435 *et al.* 2011), and one gene encoding a CFEM domain. Cystein-rich CFEM domains have been
436 described in G-protein-coupled receptors in the rice blast pathogen *Magnaporthe oryzae*
437 playing an important role in pathogenicity (Kulkarni *et al.* 2005). These genes provide
438 interesting candidates for further studies of molecular determinants of host specificity in *Z.*
439 *tritici*.

440 **Conclusions**

441 In this study we have used the fungal wheat pathogen *Z. tritici* to assess the patterns
442 of selection acting on the genome, both qualitatively and quantitatively. We measure the rate of
443 adaptation using models of distributions of fitness effects, and with polymorphism and
444 divergence models of codon sequence evolution, we provide evidence for signatures of both
445 positive and balancing selection in protein coding genes, as expected under arms race and
446 trench warfare scenarios, respectively. The rate of adaptive substitutions in the plant pathogen
447 is similar to estimates in animal species and considerably faster than corresponding estimates
448 in plants. Notably, rates of evolution in genes encoding effector proteins are more than twice
449 as fast as the genome average. Furthermore, our results suggest widespread occurrence of
450 linked selection (both Hill-Robertson interference and background selection), as both the rate
451 of adaptation and the strength of negative selection correlate with the recombination rate.
452 Finally, we infer signatures of balancing selection and find an enrichment of genes encoding
453 pathogenicity related functions - but not effector proteins - among detected genes. Our results
454 thereby demonstrate that different categories of genes evolve under arms race and trench-
455 warfare selection in this pathogen. Interestingly, we show that signatures of positive selection
456 and balancing selection do not correlate with the presence of transposable elements as
457 predicted in other studies. These results highlight the fundamental role of recombination and
458 sexual reproduction in adaptive processes of rapidly evolving organisms.

459 **Materials and Methods**

460 **Re-sequencing, assembly and alignment of *Z. tritici* isolates**

461 In this study we used a geographical collection of thirteen field isolates of *Z. tritici*
462 isolated from infected leaves of bread wheat (*Triticum aestivum*) (Table S1). Genome data of
463 three isolates, including the reference isolate IPO323, were published previously (Goodwin *et*
464 *al.* 2011; Stukenbrock *et al.* 2011). For the remaining ten isolates full genomes were

465 sequenced. DNA extraction was performed as previously described (Stukenbrock *et al.* 2011).
466 Library preparation and paired end sequencing using an Illumina HiSeq2000 platform were
467 conducted at Aros, Skejby, Denmark. Sequence data of the ten isolates has been deposited
468 under the NCBI BioProject IDs PRJNA312067. We used SOAPdenovo2 (Luo *et al.* 2012) to
469 construct de novo genome assemblies for each isolate independently. For each genome, the k-
470 mer value maximizing the weighted median (N50) of contigs and scaffolds was selected.

471 Prior to generating a multiple genome alignment, we pre-processed the individual
472 genomes of the thirteen *Z. tritici* isolates. First, we masked repetitive sequences using a library
473 of 497 repeat families identified de novo in four *Zymoseptoria* species (Grandaubert *et al.*
474 2015). Repeats were soft-masked using the program RepeatMasker (option -xsmall) to retained
475 information of repeat sites in the alignment (A.F.A. Smit, R. Hubley & P. Green RepeatMasker
476 at <http://repeatmasker.org>). Second, we filtered the genome assemblies to contain only contigs
477 with a length \geq 1 kb. Multiple genome alignments were generated by the MULTIZ program
478 using the LASTZ pairwise aligner from the Threaded Blockset Aligner (TBA) package
479 (Blanchette *et al.* 2004). The alignment was projected on the IPO323 reference genome using
480 the maf_project program from the TBA package.

481 **Inference of population structure**

482 In order to infer population structure, we generated genealogies of the thirteen
483 isolates using the multiple genome alignment. We used the MafFilter program (Dutheil *et al.*
484 2014) to compute pairwise distance matrices using maximum likelihood under a Kimura 2
485 parameter model in 10 kb sliding windows along the chromosomes of the reference genome.
486 For each window, a BioNJ tree was reconstructed from the distance matrices. The resulting
487 1,850 genealogies were used to build a super tree using SDM for the generation of a distance
488 supermatrix (Crisuolo *et al.* 2006) and FastME was used to infer a consensus tree (Lefort *et*
489 *al.* 2015). We used the program ADMIXTURE (Alexander *et al.* 2009) (software using the —
490 haploid='*' option) to estimate the number of ancestral populations based on single nucleotide

491 polymorphism data. Filtered biallelic variants were exported as PLINK files using MafFilter
492 (Dutheil *et al.* 2014). A cross-validation procedure was conducted as described in the manual
493 of the ADMIXTURE in order to determine the optimal number of partitions. We further
494 assessed the effect of linkage by removing SNPs with an R2 value higher than 0.1 with any
495 other SNP in windows of 50 SNPs, slid by 10 SNPs, using PLINK (Chang *et al.* 2015). This
496 filtering did not affect the conclusion of the cross-validation procedure.

497 **Prediction of effector candidates**

498 Gene models from the *Z. tritici* reference strains (Grandaubert *et al.* 2015) were used
499 to predict proteins targeted for secretion using SignalP (Petersen *et al.* 2011). Genes predicted
500 to encode a secreted protein were further submitted to effector prediction using the EffectorP
501 software (Sperschneider *et al.* 2016).

502 **Estimation of rates of adaptation**

503 Based on the coordinates of each predicted gene model in the reference genome
504 IPO323 (Goodwin *et al.* 2011; Grandaubert *et al.* 2015), exons were extracted from the
505 multiple genome alignment of *Z. tritici* isolates using MafFilter (Dutheil *et al.* 2014). Complete
506 coding sequences (CDS) were concatenated to generate individual alignments of all
507 orthologous CDS. If one or more exons were not extracted from the alignment due to missing
508 information, the gene was discarded from further analyses. Each complete CDS alignment was
509 filtered according to the following criteria: (i) CDS were discarded if they contained more than
510 5% gaps in one or more individuals, (ii) CDS with premature stop codon were likewise
511 deleted, and (iii) only alignments comprising three or more CDS were kept. In some cases, due
512 to indels in the genome alignment, the codon phasing of some genes was lost. This issue was
513 overcome by refining the CDS alignment using the codon-based multiple alignment program
514 MACSE (Ranwez *et al.* 2011). The final data set contains 9,412 gene alignments, among which
515 7,040 contain a sequence for all 13 isolates. We further created a data set containing an

516 outgroup sequence, taken from *Z. ardabiliae*, leading to 6,767 alignments with all 13 isolates
517 together with the outgroup sequence.

518 The CDS alignment with outgroup was used to infer the synonymous and non-
519 synonymous divergence based on the rate of synonymous and non-synonymous substitutions.
520 The synonymous and non-synonymous unfolded site frequency spectra (SFS) were computed,
521 using the outgroup sequence to reconstruct the ancestral allele. To do so, we first reconstructed
522 a BioNJ tree for each gene and fitted a codon model of evolution using maximum likelihood.
523 Then ancestral state was inferred using the marginal reconstruction procedure of Yang (Yang *et*
524 *al.* 1995). All calculations were performed using the BppPopStats program from the Bio++
525 Program Suite (Guéguen *et al.* 2013). We used the Grapes program in order to estimate the
526 distribution of fitness effects from the SFS and compute a genome wide estimate of α and ω_a ,
527 the proportion of mutations fixed by selection and the rate of adaptive substitutions
528 respectively (Galtier 2016). The following models were fitted and compared using Akaike's
529 information criterion: Neutral, Gamma, Gamma-Exponential, Displaced Gamma, Scaled Beta
530 and Bessel K. Analyses were conducted on the complete set of gene alignments, as well as on
531 sub-datasets sorted according to whether the individual genes encoded a predicted effector
532 protein or not. We further stratified our data set according to the local recombination rate,
533 computed in 20 kb windows, using both the previously published genetic maps (Croll *et al.*
534 2015) and population estimates from patterns of linkage disequilibrium (Stukenbrock &
535 Dutheil 2017). We discretized the observed distributions of both r and ρ in 41 and 45
536 categories, respectively, using the cut2 command from the Hmisc R package in order to have
537 similar number of genes in each category (comprising between 247 and 258 genes for ρ , and
538 between 67 and 1,323 genes for r , the largest value being obtained for genes with $r = 0$). For
539 each gene sets, 100 bootstrap replicates were generated by sampling genes randomly in each
540 category. Genes in each replicate were concatenated and the Grapes program run with the
541 GammaExpo distribution of fitness effect (Galtier 2016). For each recombination category, the

542 mean estimates of α and ω_a , as well as the standard error over the 100 replicates, were
543 computed.

544 **Genome-wide analysis of selection patterns**

545 We inferred the strength of purifying selection by computing the pN / pS ratio for
546 each gene. Average pairwise synonymous (π_S) and non-synonymous (π_N) nucleotide diversity
547 were computed for each genes, and divided by the average number of synonymous (NS) and
548 non-synonymous (NN) positions, respectively, in order to compute the pN / pS ratio as ($\pi_N /$
549 NN) / (π_S / NS). We compared the strength of purifying selection of each gene to several
550 variables, after discarding 6 genes with pN / pS greater than one, as they might be under
551 positive selection. Local recombination rate in 20 kb windows was obtained from Croll et al
552 (Croll *et al.* 2015), and averaged over the two crosses. Population recombination rates (ρ) in
553 the same 20 kb windows were computed as in (Stukenbrock & Dutheil 2017). Each gene was
554 assigned a recombination rate based on the window(s) it overlap with, using a weighted
555 average in case it overlap with multiple windows. Local protein coding site and TE densities
556 were computed as the proportion of coding sites in a window starting x kb upstream and
557 ending x kb downstream each gene. We compared different estimations for x = 10, 20, 50 or
558 100 kb (see Supplementary Data). For the density of coding sites, we find very little influence
559 of the window size, with a unimodal distribution around ~50%. We therefore selected the
560 intermediate x = 50 kb. The density of TEs showed a large pick at 0 for low values of x. We
561 therefore selected x = 100 kb in order to get a unimodal distribution. GC content at third codon
562 position (GC3) and protein length were also recorded. Expression levels were calculated from
563 (Kellner *et al.* 2014). The mean expression level was computed as the maximum value
564 observed for the gene in axenic culture or plant infection, each averaged over three biological
565 replicates. Genes located on accessory chromosomes were labeled as “dispensable”.
566 Correlation and distribution comparison of the pN / pS ratio with each explanatory variable

567 were performed using rank-based tests (Kendall correlation and Wilcoxon test), as
568 implemented in the R statistical software.

569 **Estimation of codon usage in *Z. tritici***

570 We selected the 10% *Z. tritici* most expressed genes and computed the relative
571 synonymous codon usage of every codons (Sharp *et al.* 1986). Analyses were conducted using
572 the 'uco' function of the seqinr package for R (Charif *et al.* 2005).

573 **Model of codon sequence evolution**

574 We used all 9,412 filtered CDS alignments to reconstruct genealogies for the
575 individual genes using PhyML (model HKY85) (Guindon & Gascuel 2003). To investigate
576 patterns of selection and infer the role of positive selection on adaptive gene evolution, the
577 program CodeML from the PAML package was used (Yang 2007) with the filtered multiple
578 CDS alignments and the corresponding phylogenetic trees as inputs. CodeML allows inference
579 of selection and evolutionary rates by calculating the parameter ω , the ratio of non-
580 synonymous to synonymous rates (dN/dS) for each gene. More specifically, we compared site
581 models that allow ω to vary among codons in the protein (Nielsen & Yang 1998). The models
582 used in this study include the nearly neutral (M1a), positive selection (M2a), beta& ω (M8) and
583 beta& $\omega=1$ (M8a) models. A likelihood ratio test (LRT) was used to compare the fit of null
584 models and alternative models, and the significance of the LRT statistic was determined using
585 a χ^2 distribution. The first LRT tests for the occurrence of sites under positive selection by
586 comparing the M1a and M2a models. In the model M1a sites can be under purifying selection
587 ($0 < \omega < 1$) and evolve by neutral evolution ($\omega = 1$) while the M2a model allows for some sites
588 to be under positive selection ($\omega > 1$). The second LRT compares the M8a and M8 models,
589 where in M8 a discretized beta distribution for ω (limited to the interval [0,1]) and an
590 additional category of sites with $\omega_s > 1$. M8a is obtained by constraining $\omega_s > 1$ setting
591 (Swanson *et al.* 2003). By allowing for a wider range of strength of purifying selection, the M8

592 models are more biologically realistic. They may suffer, however, of the same issue than the
593 M7-M8 LRT, which was shown to display an increased false discovery rate compared to the
594 M1a-M2a comparison (Anisimova *et al.* 2001). We corrected for multiple testing and a false
595 discovery rate of 1% was used for the detection of genes under positive selection (Benjamini &
596 Hochberg 1995). Only genes significant for both tests were considered as genes evolving under
597 positive selection (787 out of 9,412 genes analyzed).

598 To address divergent adaptation, we compared gene evolution among four closely
599 related *Zymoseptoria* species. In a previous study we defined the core proteome of *Z. tritici*, *Z.*
600 *ardabiliae*, *Z. brevis* and *Z. pseudotritici* comprising 7,786 orthologous genes (Grandaubert *et*
601 *al.* 2015). We generated alignments of the corresponding coding sequences using the MACSE
602 sequence aligner (Ranwez *et al.* 2011) and used CodeML with a branch model that allows ω to
603 vary among branches of the phylogeny (Yang & Nielsen 1998). As input we applied a non-
604 rooted tree of the four *Zymoseptoria* species as published in (Stukenbrock *et al.* 2012). Branch
605 lengths were re-estimated for each gene by CodeML.

606 **Simulation with recombination**

607 We used the coalevol program in order to simulate codon alignments in the presence
608 of recombination (Arenas & Posada 2014). We used a haploid effective population size of
609 10,000 (option -e10000 1), a one year generation time (option -/1), one parameter for relative
610 transition vs. transversion rate, set to 2 (option -v1 2), a Goldman-Yang model of codon
611 evolution, with 4 omega classes, in equal proportion and set to 0, 0.33, 0.66 and 1.0,
612 respectively (option -m2 4 0.0 0.25 0.33 0.25 0.66 0.25 1). Two mutation rates were tested,
613 $1 \cdot 10^{-5}$ and $1 \cdot 10^{-6}$ (option -u). One set of simulations was conducted without recombination (-r
614 0.0), and for others a fixed number of recombination events was used, equal to 1, 2, 5, 10, 30
615 or 50 (option -w). Protein length was set to 100, 500, 1,000 or 2,000 codon, for 13 and 30
616 sequence (-s option). Thirty replicates were generated for each parameter combination, and a
617 single phylogenetic tree was inferred using maximum likelihood on the resulting nucleotide

618 aligned, with identical parameters to the real data analysis. M1a and M2a models were then
619 fitted using the estimated tree as input with CodeML. CodeML output was parsed using
620 BioPython (Cock *et al.* 2009).

621 **Functional enrichment analysis**

622 PFAM domains were extracted from Interproscan results from (Grandaubert *et al.*
623 2015). Only domain hits with e-values lower than 1.10^{-5} were considered resulting in 10,026
624 domains present in 7,343 genes. Enrichment tests were performed based on contingency tables,
625 counting the number of genes containing the domain and the number of genes which do not
626 contain it, for both the complete proteome and a given set of candidates to test. A χ^2 test was
627 performed to assess significance.

628 **Gene cluster analysis**

629 To analyze the distribution of genes under positive selection, we considered two
630 genes separated by less than 5,000 bp to be clustered and assessed the probability of such
631 clusters under a random distribution of genes along the chromosomes. To do so on a genome-
632 wide scale, we calculated the probability to obtain clusters encompassing from two to ten
633 genes under positive selection when these genes are randomly distributed across all gene
634 coordinates. Based on 10,000 random permutations, it appeared that only clusters containing
635 more than three genes were significant at the 5% level.

636 **Association between positive selection and effector-encoding genes**

637 To test whether effector-encoding genes are more likely to be under positive
638 selection, we fitted linear models with (1) dN / dS and (2) dN / dS > 1 as response variables,
639 and whether the gene was predicted to encode an effector protein in *Z. tritici* as an explanatory
640 variable. Models were fitted independently for each branch of the four species phylogeny. A
641 binary logistic regression was fitted in order to predict the occurrence of genes under positive

642 selection. For model (1), residues were normalized using a Box-Cox transform as implemented
643 in the MASS package for the R statistical software. An ordinary least square fit was then
644 obtained using the ols function of the rms package for R (Harrell 2015), using the robcov
645 function to obtain robust estimates of the size effects and associated p-values. For model (2),
646 the lrm function of the rms package was used to fit the binary logistic regression model,
647 together with the robcov function to get robust estimates.

648 **Authors' contributions**

649 JD and EHS conceived and planned the experiments. JG and JD established the
650 computational framework and analyzed the data. All authors contributed to the interpretation
651 of data and wrote the manuscript. All authors read and approved the final manuscript.

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658 **Competing interests:**

659 The authors declare that they have no competing interests.

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661 Tables

662 **Table 1:** Estimates of the proportion of adaptive mutation (α) under various models of
663 distribution of fitness effects.

| Data set | Model | Nb. parameters | Log likelihood | AIC | α | ω_a |
|---------------|----------------|----------------|----------------|----------------|--------------|--------------|
| All genes | Neutral | 16 | -1111.199 | 2254.398 | 0.253 | 0.031 |
| All genes | Gamma | 17 | -286.142 | 606.284 | 0.464 | 0.058 |
| All genes | GammaExpo | 19 | -229.425 | 496.850 | 0.352 | 0.044 |
| All genes | DisplacedGamma | 18 | -286.142 | 608.284 | 0.464 | 0.058 |
| All genes | ScaledBeta | 18 | -273.334 | 582.669 | 0.485 | 0.060 |
| All genes | BesselK | 19 | -294.387 | 626.774 | 0.514 | 0.064 |
| Non-effectors | Neutral | 16 | -1104.260 | 2240.519 | 0.252 | 0.031 |
| Non-effectors | Gamma | 17 | -286.821 | 607.643 | 0.463 | 0.057 |
| Non-effectors | GammaExpo | 19 | -230.221 | 498.442 | 0.388 | 0.048 |
| Non-effectors | DisplacedGamma | 18 | -286.821 | 609.643 | 0.463 | 0.057 |
| Non-effectors | ScaledBeta | 18 | -273.687 | 583.375 | 0.458 | 0.057 |
| Non-effectors | BesselK | 19 | -296.728 | 631.456 | 0.513 | 0.063 |
| Effectors | Neutral | 16 | -101.036 | 234.071 | 0.307 | 0.062 |
| Effectors | Gamma | 17 | -94.255 | 222.510 | 0.485 | 0.097 |
| Effectors | GammaExpo | 19 | -87.332 | 212.664 | 0.666 | 0.134 |
| Effectors | DisplacedGamma | 18 | -94.684 | 225.369 | 0.492 | 0.099 |
| Effectors | ScaledBeta | 18 | -86.845 | 209.689 | 0.600 | 0.120 |
| 664 Effectors | BesselK | 19 | -87.416 | 212.832 | 0.507 | 0.102 |

665 AIC: Akaike's information criterion. α : proportion of adaptive substitutions, ω_a : rate of
666 adaptive substitutions. Values in bold indicate the best model fit for each gene set.

667

668 **Table 2:** Correlation of pN / pS with genomic factors.

| Variable | Effect | P value |
|--|---------|----------|
| GC3 | -0.2372 | <2.2E-16 |
| Expression | -0.3689 | <2.2E-16 |
| Protein size | 0.0227 | 0.0056 |
| Density of protein coding sites | 0.0067 | 0.4127 |
| Recombination rate (cM / Mb) | -0.0309 | 1.85E-04 |
| Population recombination rate (4.Ne.r) | -0.0394 | 1.52E-06 |
| Density of TEs | -0.0030 | 0.7197 |
| Effector | 0.0845 | 1.05E-14 |
| 669 Dispensable chromosome | 0.2590 | 0.0162 |

670 Effects and p-values are calculated using Kendall's correlation of ranks. GC3: GC-content at
 671 third codon positions. TEs: transposable elements.

672 **Table 3:** Genomic factors affecting the occurrence of balancing selection

| Variable | LRT result | | Tajima's D | |
|------------------------------|------------|---------|------------|---------|
| | Coef. | P value | Coef. | P value |
| Intercept | -1.3428 | <0.0001 | -1.4128 | <0.0001 |
| Recombination rate (cM / Mb) | 0.0012 | 0.0006 | 0.0011 | <0.0001 |
| Density of TEs | -0.9250 | 0.0621 | 0.0504 | 0.6465 |
| Coding site density | -0.2991 | 0.5259 | -0.0471 | 0.6768 |
| Effector | -0.9621 | 0.0156 | -0.1164 | 0.1239 |
| 673 Expression | -0.3341 | <0.0001 | -0.0193 | 0.0213 |

674 LRT result: likelihood ratio test obtained from PAML analysis. Coef.: model coefficient. TEs:
 675 transposable elements.

676 **Table 4:** Difference of dN / dS ratios between genes predicted to encode an effector protein or
 677 not, in different branches of the species tree.

| Species | Positive selection | | dN / dS (OLS) | | dN / dS > 1 (BLR) | |
|--|--------------------|-----------|---------------|---------|-------------------|---------|
| | Total | Effectors | Coef. | P-value | Coef. | P-value |
| <i>Z. tritici</i> | 47 | 5 | 0.4281 | 0.0000 | 2.0427 | 0.0002 |
| <i>Z. pseudotritici</i> | 54 | 4 | 0.3319 | 0.0001 | 0.9290 | 0.2027 |
| <i>Z. tritici – pseudotritici</i> ancestor | 1149 | 41 | 1.0968 | 0.0000 | 0.3649 | 0.2592 |
| <i>Z. brevis</i> | 60 | 3 | 0.4593 | 0.0000 | 0.7492 | 0.3026 |
| 678 <i>Z. ardabiliae</i> | 15 | 0 | 0.3173 | 0.0000 | -5.0216 | 0.0000 |

679 OLS: ordinary least square. BLR: binary logistic regression. Coef. : coefficient in the linear
 680 model.

681 **Figures**

682 **Fig. 1:** Population structure of the thirteen *Z. tritici* isolates. A) Consensus super tree of the
683 thirteen isolates based on 1,850 genealogies estimated in 10 kb sliding windows along the
684 multiple genome alignment. This tree suggests the grouping of the isolates into two
685 populations originating from Europe and Iran. B) Based on SNP data, the program
686 ADMIXTURE estimated that the best separation of the isolates is also in two populations
687 (k=2). However, the use of k=3 highlighted a German sub-population within the European
688 isolates.

689 **Fig. 2:** Comparison of the estimates of A) the proportion of adaptive substitution α , and B) the
690 rate of adaptive substitution, ω_a for genes predicted to encode effector proteins or not.
691 Histograms (white bars), kernel density plots and box-and-whiskers charts are computed over
692 100 bootstrap replicates in each case (see Material and Methods).

693 **Fig. 3:** Estimates of A) the proportion of adaptive substitution α , and B) the rate of adaptive
694 substitution, ω_a as a function of the recombination rate (r). Each point and bars represent the
695 mean estimate and corresponding standard error for one recombination category over 100
696 bootstrap replicates. Four models were fitted (colored curved) and corresponding Akaike's
697 information criterion values are indicated in the right margin. Inset plots represent the same
698 data with a logarithmic scale, the b value was set to the corresponding estimate in the third
699 model. Confidence intervals have been omitted for clarity.

700 **Fig. 4:** Correlation of the strength of purifying selection with several genomic factors. The
701 intensity of purifying selection is measured by the pN / pS ratio. Points represent median
702 values and error bar the first and third quartiles of the distributions. A-F: x-axis were
703 discretized in categories with equal point densities for clarity of visualization. Lines represent
704 first, median and third quantile regression on non-discretized data.

705 **Fig. 5:** Patterns of selection along the genome of *Z. tritici*. Recombination rate, population
706 recombination rate, pN / pS ratio and density of coding sites (CDS) are plotted in windows of
707 100 kb along the thirteen essential chromosomes.

708 **Fig. 6:** Effect of recombination on the inference of positive selection. False discovery rate, as
709 estimated from simulations under a model with neutral and purifying selection only, was
710 plotted as a function of the number of recombination events (x-axis), length of the alignment
711 (coloured lines), mutation rate and sample size (panels). Horizontal dash lines show the
712 discovery rate of the real data for distinct minimum gene lengths.

713 **Fig. 7:** Distribution of Tajima's D for different gene categories. Kernel densities were fitted to
714 the distribution of each gene's Tajima's D (x-axis and color scale), sorted per category
715 (detected to be under balancing selection, predicted to encode an effector protein, predicted not
716 to encode an effector protein, all genes).

717 **Supplementary material**

718 **Table S1:** Summary table of isolates used in this study and genome assembly statistics.

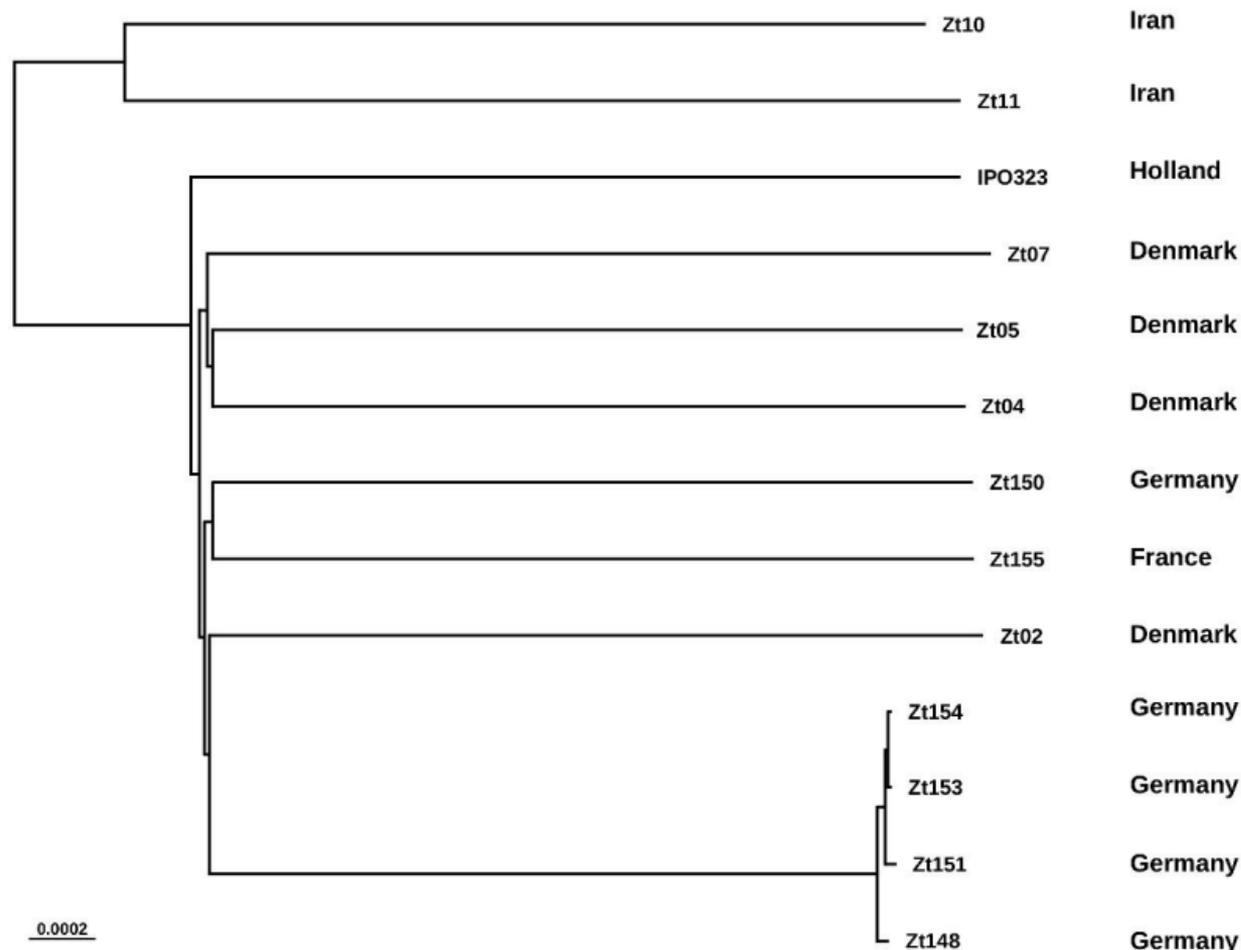
719 **Table S2:** Summary statistics of the multiple genome alignment of thirteen *Z. tritici* genomes.

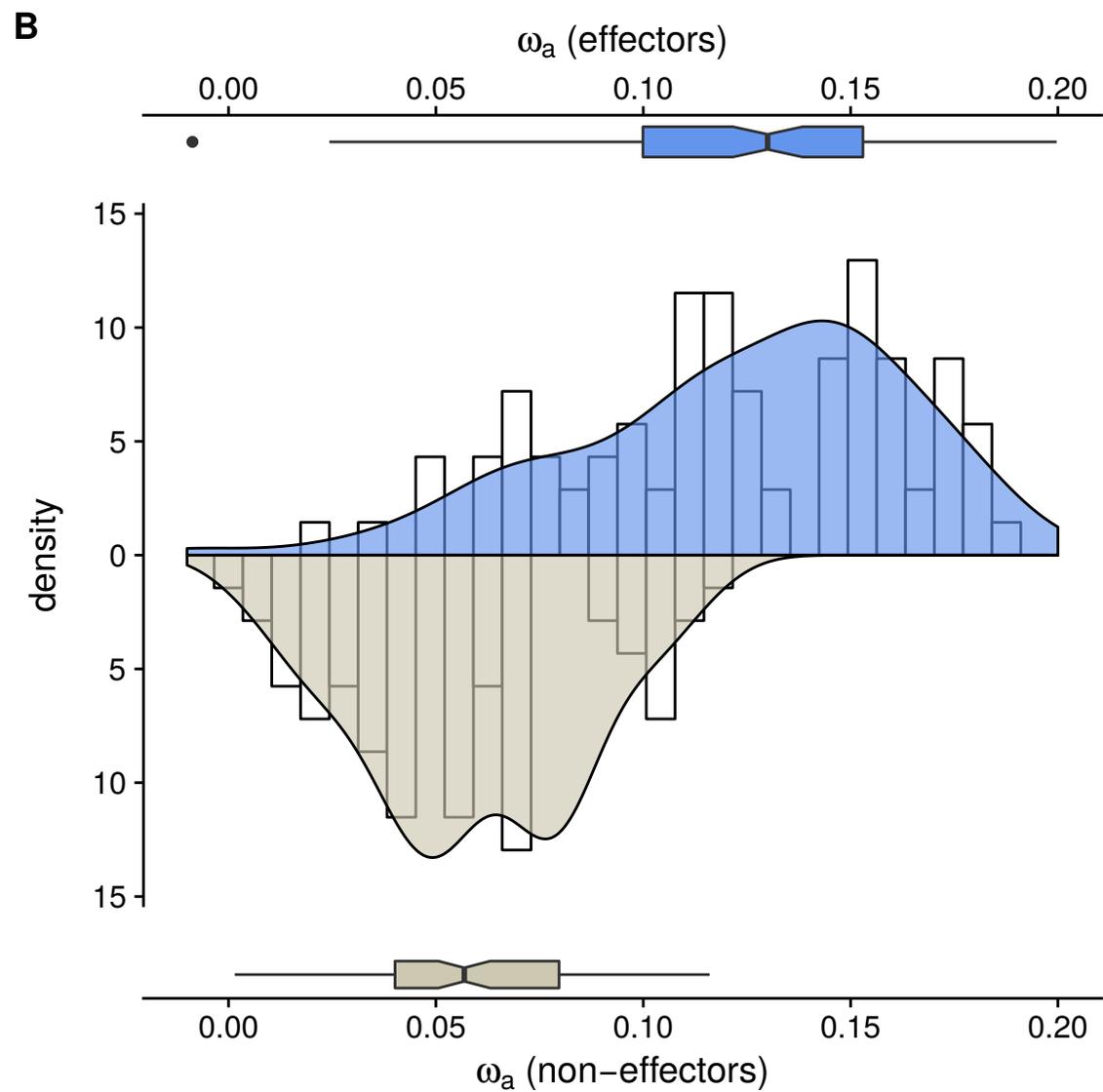
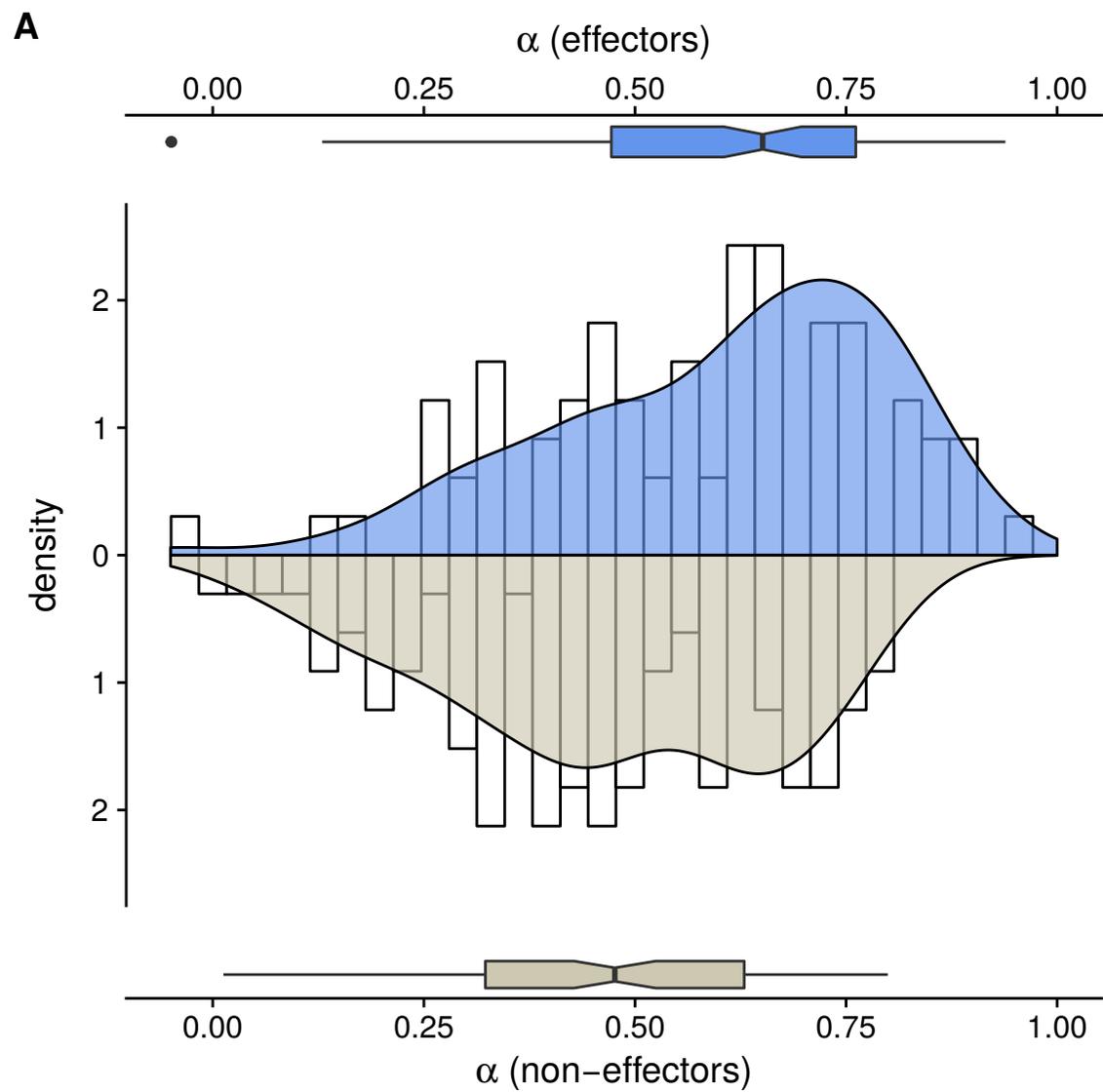
720 **Table S3:** Output of the PAML analysis using codon site models for the 9,412 filtered CDS of
721 *Z. tritici*.

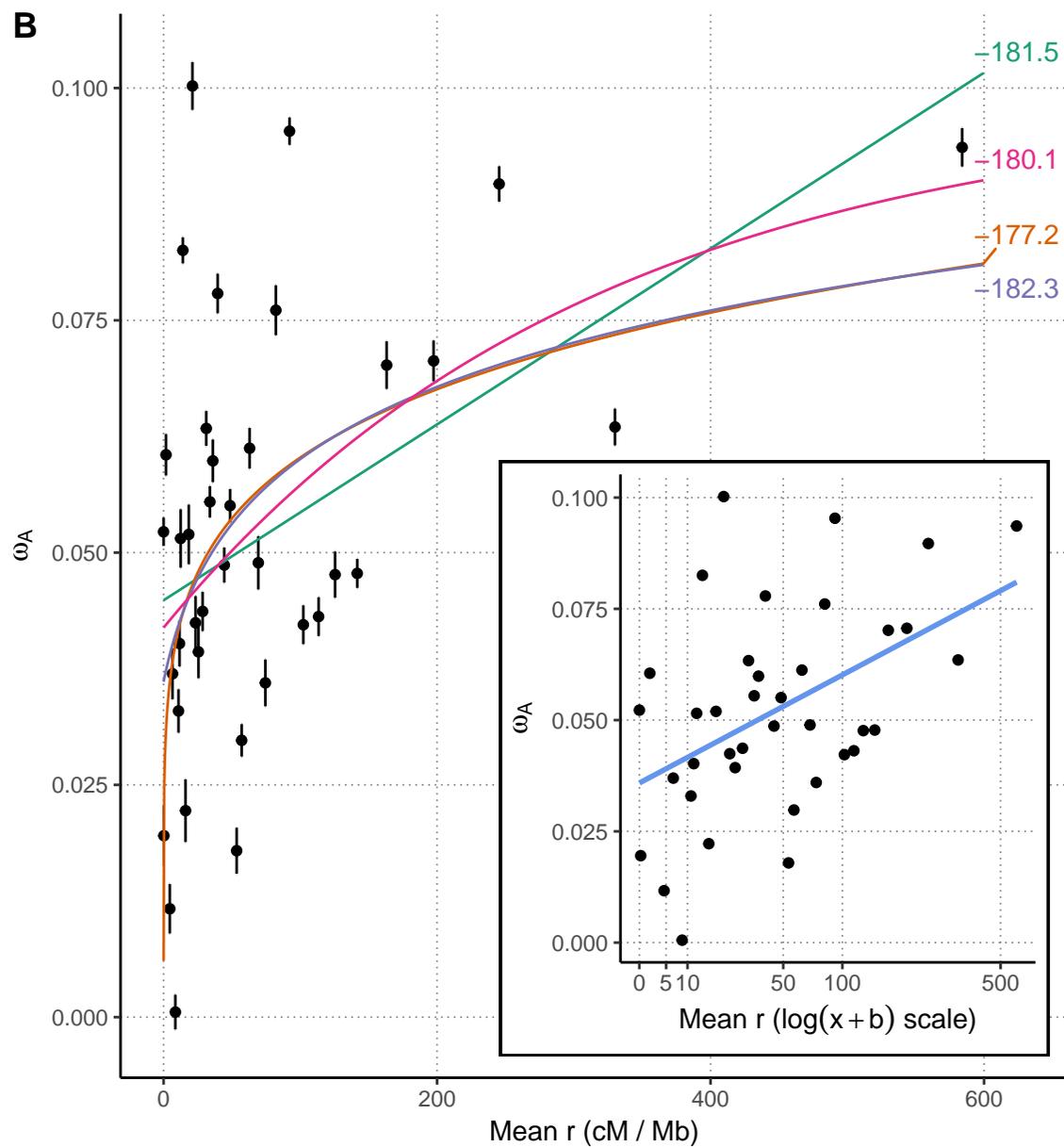
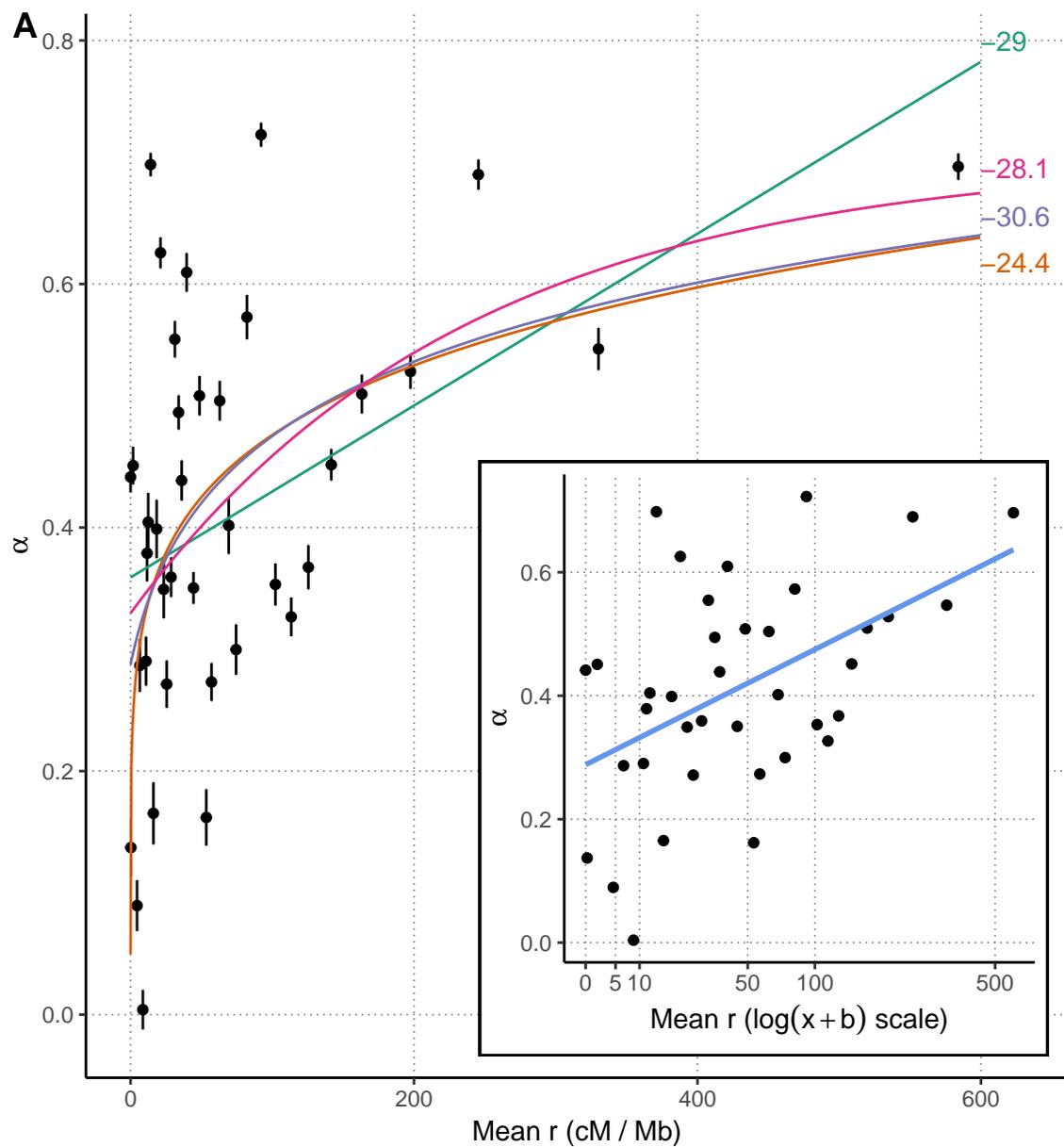
722 **Table S4:** Functional enrichment analysis using PFAM domains for the 787 genes with sites
723 under positive selection in *Z. tritici*.

724 **Table S5:** Functional enrichment analysis using PFAM domains for the genes under positive
725 selection in four *Zymoseptoria* species.

726 **Fig. S1:** Codon usage in *Z. tritici*. Relative synonymous codon usage (RSCU) in the 10% most
727 expressed genes of *Z. tritici*. Codon usage, according to the base type at the third position.

A**B**





Model $a+b.x$ $a.x^b$ $a.\log(x+b)$ $a+b.\exp(-c.x)$

Model $a+b.x$ $a.x^b$ $a.\log(x+b)$ $a+b.\exp(-c.x)$

