



Genome-wide association study identifies 48 common genetic variants associated with handedness

Handedness has been extensively studied because of its relationship with language and the over-representation of left-handers in some neurodevelopmental disorders. Using data from the UK Biobank, 23andMe and the International Handedness Consortium, we conducted a genome-wide association meta-analysis of handedness ($N = 1,766,671$). We found 41 loci associated ($P < 5 \times 10^{-8}$) with left-handedness and 7 associated with ambidexterity. Tissue-enrichment analysis implicated the CNS in the aetiology of handedness. Pathways including regulation of microtubules and brain morphology were also highlighted. We found suggestive positive genetic correlations between left-handedness and neuropsychiatric traits, including schizophrenia and bipolar disorder. Furthermore, the genetic correlation between left-handedness and ambidexterity is low ($r_G = 0.26$), which implies that these traits are largely influenced by different genetic mechanisms. Our findings suggest that handedness is highly polygenic and that the genetic variants that predispose to left-handedness may underlie part of the association with some psychiatric disorders.

Handedness refers to the preferential use of one hand over the other. Conversely, ambidexterity refers to the ability to perform the same action equally well with both hands. Hand preference is first observed during gestation as embryos begin to exhibit single arm movements^{1,2}. Across the life span, the consistent use of one hand leads to alterations in the macromorphology and micromorphology of bone³, which results in enduring asymmetries in bone form and density^{4,5}. At the neurological level, handedness is associated with the lateralization of language (the side of the brain involved in language) and other cognitive effects^{6,7}. The prevalence of left-handedness in modern western cultures is approximately 9%⁸ and is greater in males than females⁹. While handedness is conceptually simple, its aetiology and whether it is related to brain and visceral (internal organ) asymmetry is unclear.

Since the mid-1980s, the literature regarding the genetics of handedness and lateralization has been dominated by the right-shift¹⁰ and dextral-chance¹¹ theories. Both theories involve additive biallelic monogenic systems in which an allele at the locus biases an individual towards right-handedness, while the second allele is a null allele that results in the random determination of handedness by fluctuating asymmetry. The allele frequency of the right-shift variant has been estimated at ~43.5%¹⁰, while that of the dextral-chance variant has been estimated at ~20% in populations with a 10% prevalence of left-handedness¹¹. A joint analysis of data from 35 twin studies found that additive genetic factors accounted for 25.5% (95% confidence interval (CI) of 15.7, 29.5%) of the phenotypic variance of handedness¹², which is consistent with predictions of the variance explained under the single gene right-shift and dextral-change models. However, linkage studies^{13–16}, candidate gene and genome-wide association studies (GWAS)^{17–21} have failed to identify any putative major gene for handedness.

Most recently, two large-scale GWAS identified four genomic loci containing common variants of small effect associated with handedness^{20,21}. However, both GWAS failed to replicate signals at the *LRRTM1*, *PCSK6* and the X-linked androgen receptor genes that had previously been reported in smaller genetic association

studies^{17–19}. In this study, we present findings from the world's largest GWAS meta-analysis of handedness to date ($N = 1,766,671$), which combined data from 32 cohorts from the International Handedness Consortium (IHC) ($N = 125,612$), 23andMe ($N = 1,178,877$) and the UK Biobank (UKBB) ($N = 462,182$).

Results

GWAS of left-handedness. Across all studies, the handedness phenotype was assessed by a questionnaire that evaluated either which hand was used for writing or for self-declared handedness. All cohorts were randomly ascertained with respect to handedness. Combining data across the 32 IHC cohorts, 23andMe and UKBB yielded 1,534,836 right-handed and 194,198 left-handed (11.0%) individuals (Supplementary Table 1). After quality control (Methods), the GWAS meta-analysis included 13,346,399 single nucleotide polymorphisms (SNPs; including autosomal and X chromosome SNPs) with a minor allele frequency (MAF) of >0.5%.

The genetic correlations as estimated by bivariate linkage disequilibrium (LD) score regression²² among the results from the UKBB, 23andMe and IHC GWAS were $r_G^{\text{UKBB-23andMe}} = 0.88$ (s.e. = 0.05), $r_G^{\text{UKBB-IHC}} = 0.73$ (s.e. = 0.16) and $r_G^{\text{IHC-23andMe}} = 0.60$ (s.e. = 0.11), which suggests that the three GWAS were capturing many of the same genetic loci for handedness. There was some inflation of the test statistics following meta-analysis ($\lambda_{GC} = 1.22$); however, the intercept from the LD score regression analysis²³ was 1.01. This suggests that the inflation was due to polygenicity rather than bias due to population stratification or duplication of participants across the UKBB, 23andMe and IHC studies.

We identified 41 loci that met the threshold for genome-wide significance ($P < 5 \times 10^{-8}$) (Fig. 1 and Supplementary Table 2). Loci were defined as distinct if independent genome-wide significant signals were separated by at least 1Mb, except for the MHC and 17q21.31 regions (the 17q21.31 region contains a common inversion polymorphism²⁴) for which we only report the lead signals due to the extent of LD across these loci. Summary statistics for the lead variants at genome-wide significant loci

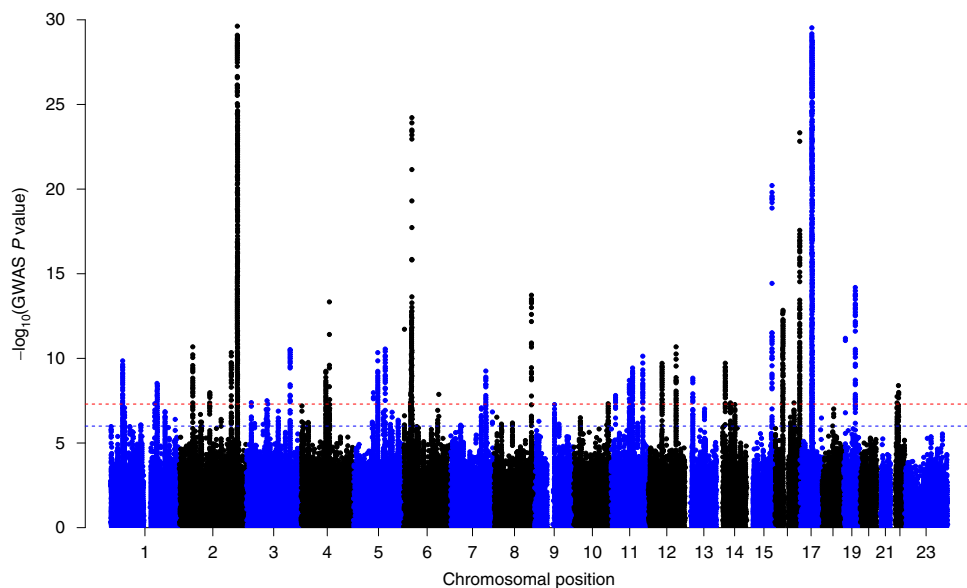


Fig. 1 | Manhattan plot of the left-handedness meta-analysis. Manhattan plots for the left-handedness GWAS meta-analysis ($N=1,534,836$ right-handed versus 194,198 left-handed). Each dot represents a SNP. The red broken line highlights the genome-wide levels of significance threshold ($P < 5 \times 10^{-8}$); the blue broken line shows the threshold for suggestive associations.

are presented in Table 1 along with the gene nearest to the lead SNP. A description of the putative functions of the nearest gene is included in Supplementary Table 2. Conditional analyses identified nine additional independent SNPs at genome-wide significance near the lead SNPs on chromosomes 2q, 6p, 16q and 17q (Supplementary Table 3). Interestingly, the list of genome-wide significant associations included multiple variants close to genes involved in microtubule formation or regulation (that is, *MAP2*, *TUBB*, *TUBB3*, *NDRG1*, *TUBB4A*, *TUBA1B*, *BUB3* and *TTC28*). A phenome-wide association scan (PheWAS) of the lead SNPs using GWAS summary data from 1,349 traits revealed that 28 out of the 41 lead SNPs have previously been associated with other complex traits (Supplementary Table 4). Among these results, we highlight that the rs6224, rs13107325 and rs45527431 variants have previously been associated with schizophrenia at genome-wide levels of significance ($P < 5 \times 10^{-8}$). Alleles at these loci had the same direction of effect (that is, those that increased the odds of left-handedness also increased the risk of schizophrenia). Furthermore, we found that seven variants associated with left-handedness were also associated with educational attainment; however, the direction of effect of these SNPs on left-handedness and educational attainment was not consistent. Future colocalization analyses are needed to assess whether the same SNPs associated with handedness also affect these other traits or whether the pattern of signals are more likely due to LD with another causal variant.

To identify the most likely tissues and pathways underlying the GWAS signals, we used DEPICT²⁵ and MAGMA²⁶. Results from both the DEPICT and MAGMA tissue-enrichment analyses implicated (false discovery rate (FDR) $< 5\%$) the central nervous system, including brain tissues such as the hippocampus and cerebrum (Table 2 and Supplementary Table 5), which is consistent with the hypothesis that handedness is primarily a neurological trait. We observed significant evidence for pathways involved in left-handedness (FDR $< 5\%$), including regulation of microtubules and axons, as well as neurogenesis and morphology regulation of the cerebral cortex and hippocampus (Table 3).

We then performed gene-based analyses using gene-expression prediction models of brain tissues using S-MultiXcan to

identify additional loci²⁷. In total, we tested the association between the predicted expression of 14,501 genes in brain tissues and left-handedness. In addition to detecting significant associations ($P < 3.44 \times 10^{-6}$) of genes within the loci identified during the meta-analysis, we observed an association between left-handedness and the predicted expression of *AMIGO1* ($P = 2.82 \times 10^{-7}$), a gene involved in the growth and fasciculation of neurites from cultured hippocampal neurons and may also be involved in the myelination of developing neural axons²⁸. Supplementary Table 6 shows the significant associations from the S-MultiXcan analysis.

We also applied the summary-data-based Mendelian randomization (SMR) approach using expression quantitative trait loci (eQTL) data from PsychENCODE^{29,30} and from a meta-analysis of eQTL data from brain tissues³¹ that included results from GTEx³², CMC³³ and ROSEMAP³⁴ to identify additional loci and to pinpoint genes behind the GWAS associations. Through this approach, we were unable to identify genes outside the loci from our main GWAS; however, we were able to implicate the *NMT1*, *TUBA1C*, *FES*, *CENPBD1* and *BCR* genes as candidates underlying some of our genome-wide significant associations. Supplementary Table 7 shows all the statistically significant associations from the SMR analysis.

Multiple studies have reported that left-handedness and ambidexterity are more prevalent in males than in females⁹. Consistent with this observation, we found that 11.9% of male participants in the IHC cohorts reported being left-handed or ambidextrous compared with only 9.3% of females (odds ratio (OR) = 1.31, 95% CI = 1.25–1.38, $P < 2.2 \times 10^{-16}$) (Supplementary Table 8). Similarly, in the UKBB data, 10.5% of males and 9.9% of females were left-handed (OR = 1.07, 95% CI = 1.05–1.09, $P = 1.87 \times 10^{-11}$), and in 23andMe, 15.6% of males and 12.6% of females were left-handed (OR = 1.28, 95% CI = 1.26–1.30, $P < 2.2 \times 10^{-16}$). Sex differences in ambidexterity were also apparent in the UKBB and 23andMe cohorts (these data were not available for the IHC cohorts). In the UKBB data, 2% of males and 1.30% of females reported being ambidextrous (OR = 1.55, 95% CI = 1.47–1.62, $P < 2.2 \times 10^{-16}$), while in 23andMe, 3.45% of males and 2.61% of females were ambidextrous (OR = 1.33, 95% CI = 1.28–1.37, $P < 2.2 \times 10^{-16}$). Birth year had a small but significant effect on left-handedness, with individuals who were born more recently being more likely

Table 1 | Loci associated with left-handedness after a meta-analysis of 23andMe, UKBB and IHC data

CHR	BP	SNP	Gene	EA	NEA	EAF	Z	OR ^a	P	Direction
1	44172458	rs34550543	ST3GAL3	T	C	0.41	6.42	1.02	1.4 × 10 ⁻¹⁰	+++
1	160398240	rs66513715	VANGL2	D	I	0.20	-5.46	0.97	4.77 × 10 ⁻⁸	-??
1	169112399	rs10081960	NME7	C	G	0.60	-5.93	0.98	2.96 × 10 ⁻⁹	---
2	48624007	rs4953572	FOXP2	A	G	0.66	6.70	1.02	2.11 × 10 ⁻¹¹	+++
2	109954066	rs4676276	SH3RF3	A	C	0.52	5.72	1.02	1.06 × 10 ⁻⁸	+++
2	187522750	rs13006483	ITGAV	T	G	0.28	6.59	1.03	4.51 × 10 ⁻¹¹	+++
2	210300731	rs62213410	MAP2	A	T	0.71	-11.45	0.96	2.37 × 10 ⁻³⁰	---
3	18167162	rs1398651	SATB1	A	T	0.56	5.49	1.02	4.11 × 10 ⁻⁸	+++
3	74246260	rs201072423	CNTN3	D	I	0.51	5.53	1.02	3.17 × 10 ⁻⁸	+++
3	77574555	rs62251113	ROBO2	A	C	0.37	5.45	1.02	4.94 × 10 ⁻⁸	+++
3	158017859	rs1526194	RSRC1	T	C	0.58	-6.65	0.98	3.02 × 10 ⁻¹¹	---
4	89910701	rs28658282:T	FAM13A	T	C	0.10	-6.20	0.96	5.77 × 10 ⁻¹⁰	-??
4	103188709	rs13107325	SLC39A8	T	C	0.08	7.54	1.06	4.62 × 10 ⁻¹⁴	+++
5	71890187	rs246628	LINC02056	C	G	0.41	5.72	1.02	1.06 × 10 ⁻⁸	+++
5	87825490	rs2194028	TMEM161B-AS1	T	C	0.34	6.59	1.02	4.52 × 10 ⁻¹¹	+++
5	114471109	rs1422070	TRIM36	A	C	0.60	-6.65	0.98	2.85 × 10 ⁻¹¹	---
6	3143866	rs35551703	BPHL	A	G	0.04	-7.04	0.94	1.93 × 10 ⁻¹²	---
6	26599509	rs45527431	ABT1	A	G	0.91	5.86	1.04	4.63 × 10 ⁻⁹	++?
6	30688427	rs3132584	TUBB	T	G	0.21	-10.31	0.95	6.12 × 10 ⁻²⁵	---
6	127643791	rs148342778:GTA	ECHDC1	D	I	0.65	-5.68	0.97	1.32 × 10 ⁻⁸	-??
7	127268806	rs806188	PAX4	T	C	0.32	6.20	1.02	5.65 × 10 ⁻¹⁰	+++
8	134274226	rs2233324	NDRG1	C	G	0.16	-7.66	0.96	1.86 × 10 ⁻¹⁴	---
10	124992505	rs12414988	BUB3	A	G	0.21	5.46	1.02	4.75 × 10 ⁻⁸	+++
11	16474017	rs1000565	SOX6	A	G	0.60	5.65	1.02	1.56 × 10 ⁻⁸	+++
11	66173400	rs11227478	NPAS4	A	G	0.21	-6.00	0.98	1.97 × 10 ⁻⁹	---
11	77531890	rs11820337	RSF1	T	C	0.35	6.27	1.02	3.64 × 10 ⁻¹⁰	+++
11	115081563	rs9645660	CADM1	T	C	0.52	-6.51	0.98	7.43 × 10 ⁻¹¹	---
12	49539892	rs11168884	TUBA1B	T	C	0.34	-6.37	0.98	1.96 × 10 ⁻¹⁰	---
12	100324975	rs7132513:G	ANKS1B	C	G	0.61	6.70	1.03	2.1 × 10 ⁻¹¹	+??
13	27294638	rs9581731	WASF3	T	C	0.71	-6.05	0.98	1.49 × 10 ⁻⁹	---
14	29628115	rs8016028	AL133166.1	T	C	0.81	-6.37	0.97	1.92 × 10 ⁻¹⁰	---
14	48430794	rs8012503	LINC00648	C	G	0.88	5.48	1.03	4.27 × 10 ⁻⁸	+++
15	91423543	rs6224	FURIN	T	G	0.47	-9.39	0.97	6.16 × 10 ⁻²¹	---
16	28828834	rs62036618	ATXN2L	A	C	0.61	-7.39	0.98	1.43 × 10 ⁻¹³	--?
16	69224615	rs1424114	SNTB2	T	C	0.35	-5.48	0.98	4.21 × 10 ⁻⁸	---
16	89991599	rs4550447	TUBB3	C	G	0.12	10.12	1.06	4.67 × 10 ⁻²⁴	+++
17	43757450	rs55974014	CRHR1	A	C	0.21	-11.43	0.95	2.97 × 10 ⁻³⁰	---
19	6499231	rs66479618	TUBB4A	T	C	0.20	-6.87	0.97	6.48 × 10 ⁻¹²	--+
19	42439263	rs112737242	RABAC1	D	I	0.35	-7.80	0.97	6.4 × 10 ⁻¹⁵	--?
22	23663848	rs4822384	BCR	T	G	0.39	-5.63	0.98	1.82 × 10 ⁻⁸	---
22	28628209	rs5762532	TTC28	T	C	0.59	-5.88	0.98	4.04 × 10 ⁻⁹	---

BP, base pair positions based on the GRCh37-hg19 human genome assembly; CHR, chromosome; EA, effect allele; EAF, effect allele frequency; NEA, non-effect allele; P, meta-analysis P value; Z, Z-statistic. The direction of effects is shown in the following order: 23andMe, UKBB and IHC. ^aOR corresponds to that derived from the 23andMe results. Where the SNP is missing in a cohort, a question mark is indicated in the Direction column.

to be left-handed (OR=1.008 per year, 95% CI=1.007–1.009, $P < 2.2 \times 10^{-16}$).

The differences in prevalence between males and females and a previously reported association between the X-linked androgen receptor gene and handedness¹⁷ could reflect the involvement of

hormone-related genes in handedness aetiology. We therefore carried out a sex-stratified GWAS of handedness in the UKBB data using left-handed individuals as cases and right-handed individuals as controls; however, we did not identify any genome-wide significant loci. Despite this, the point estimate of the genetic correlation between

Table 2 | Results of the tissue-enrichment analysis for left-handedness (DEPICT)

Tissue	Group	P value
Corpus striatum	Nervous system	0.00000673
Basal ganglia	Nervous system	0.0000148
Hippocampus	Nervous system	0.0000227
Central nervous system	Nervous system	0.0000327
Brain	Nervous system	0.0000357
Telencephalon	Nervous system	0.0000398
Parahippocampal gyrus	Nervous system	0.0000414
Entorhinal cortex	Nervous system	0.0000414
Limbic system	Nervous system	0.0000423
Cerebrum	Nervous system	0.0000427
Prosencephalon	Nervous system	0.0000487
Temporal lobe	Nervous system	0.0000587
Cerebral cortex	Nervous system	0.0000668
Parietal lobe	Nervous system	0.000305
Mesencephalon	Nervous system	0.000843
Occipital lobe	Nervous system	0.00131
Visual cortex	Nervous system	0.00156
Brain stem	Nervous system	0.00166

Only results with FDR values <5% are shown.

male handedness and female handedness computed using LD score regression was lower than unity but not significantly different from one ($r_G = 0.77$ (s.e. = 0.12), $P = 0.055$).

Associations with previously reported candidate genes. All loci identified in recent GWAS of handedness from the UKBB^{20,21} were replicated in our study. However, we found no evidence of association between left-handedness and genes and genetic variants reported in other prior studies. The SNPs rs1446109, rs1007371 and rs723524 in the *LRRTM1* locus reported by Francks et al.¹⁸ did not reach nominal significance in any of the analyses performed ($P > 0.05$). Similarly, the SNP rs11855415 reported by Scerri et al.¹⁹ as associated with left-handedness in individuals with dyslexia did not show evidence of association ($P > 0.05$). Furthermore, we investigated whether the 27 genes exhibiting asymmetric expression in early development of the cerebral cortex described by Sun et al.³⁵ were associated with handedness in our S-MultiXcan analyses. Only 11 out of the 27 asymmetry genes were available in our analysis, and after adjusting the results for multiple testing, we did not observe any significant association (Supplementary Table 9). In a more recent study, Ocklenburg and colleagues³⁶ list 74 genes displaying asymmetric expression in cervical and anterior thoracic spinal cord segments of five human fetuses. In total, 43 out of the 74 genes were in our S-MultiXcan analyses, of which only *HIST1H4C* was statistically significant after correcting for multiple testing ($P = 2.2 \times 10^{-4}$) (Supplementary Table 10).

Heritability of left-handedness and genetic correlations with other traits. Previous twin studies have estimated the heritability of left-handedness as around 25%¹². In the present study, we employed LD score regression, genome-based restricted maximum likelihood (REML) analysis, as implemented in BOLT-LMM, and maximum likelihood analysis of identity by descent (IBD) sharing in close relatives³⁷ to provide complementary estimates of SNP heritability and total heritability that relied on a different set of assumptions to the classical twin model. Using GWAS summary statistics from our study and LD score regression, we estimated that the variance

explained by SNPs was 3.45% (s.e. = 0.17%) on the liability scale, assuming the prevalence of left-handedness is 10% (Table 4). Using genotypic data from the UKBB study (and age and sex as covariates) and genome-based REML analysis, we also obtained low estimates of the SNP heritability (5.87%, s.e. = 2.21%). Due to the large disparity between estimates of heritability from twin studies and the lower estimates of SNP heritability from the above approaches, we estimated the heritability of handedness using autosomal IBD information from closely related individuals³⁷ in the UKBB data (estimated genome-wide IBD > 8%). We partitioned the phenotypic variance into additive genetic effects (A), shared environmental effects (C) and individual environmental effects (E) (Methods). We estimated that additive genetic effects explained 11.9% (95% CI = 7.2–17.7) of the phenotypic variance in handedness, while shared environmental effects and individual environment effects accounted for 4.6% (95% CI = 0–9.0) and 83.6% (95% CI = 75.2–85.6) of the variance in liability, respectively (Table 4). Dropping (C) from the model did not significantly worsen the fit of the model ($P = 0.29$). The estimate from the A+E model was 19.7% (95% CI = 13.6–25.7) for additive genetic effects, which overlapped with those from twin studies.

We investigated the genetic correlation between left-handedness and 1,349 complex traits using LD score regression as implemented in the Complex-Traits Genetics Virtual Lab (CTG-VL)³⁸. We did not observe any genetic correlations at FDR < 5% beside handedness itself. However, we observed a general inflation of P values across the traits (that is, the expected number of traits with genetic correlations at $P < 0.05$ under the null hypothesis of no association was 67.45, whereas we observed 102 traits with $P < 0.05$). We also observed suggestive positive correlations with neurological and psychiatric traits, including schizophrenia ($P = 0.005663$), bipolar disorder ($P = 0.0023$), intracranial volume ($P = 0.01205$) and educational attainment ($P = 0.001772$), and negative correlations with mean pallidum volume ($P = 0.01124$) (Supplementary Table 11).

GWAS of ambidexterity. We carried out a separate GWAS of ambidexterity with the UKBB and 23andMe data using ambidextrous individuals as cases ($N = 37,637$; ~2% of the total sample) and right-handed individuals as controls ($N = 1,422,823$). This meta-analysis included 12,493,443 autosomal and X chromosome SNPs with a MAF > 0.5%.

Similar to the left-handedness GWAS, before the meta-analysis, we computed the genetic correlation between the UKBB ambidexterity GWAS and the 23andMe GWAS. The estimate of the genetic correlation was $r_G = 1$ (s.e. = 0.15), which indicates that both GWAS were capturing the same genetic loci. After the meta-analysis, we identified seven loci with $P < 5 \times 10^{-8}$ (Fig. 2). Table 5 displays the summary statistics for the lead SNPs at these loci along with the closest gene. Full summary statistics and descriptions of the nearest gene for these loci are included in Supplementary Table 12. There was some overlap between genome-wide significant SNPs associated with left-handedness and ambidexterity. A total of 16 out of the 41 SNPs associated with left-handedness displayed a nominal significant association with ambidexterity ($P < 0.05$), 15 of which were also in the same direction of effect (Supplementary Table 13). Conditional analyses did not identify further independent signals at genome-wide levels of significance. PheWAS revealed that the lead SNPs have been implicated in anthropometric traits and blood biomarkers (Supplementary Table 14).

The DEPICT analysis did not identify any tissue or pathway at FDR < 5%. However, the MAGMA tissue-enrichment analysis highlighted all the brain tissues tested (FDR < 5%), including brain cerebellar hemisphere and the cerebellum (Supplementary Table 15). The MAGMA pathway analysis identified 16 pathways (FDR < 5%), including regulation of cell size, basal dendrite, postsynaptic cytosol, among others that are hard to interpret such as pulmonary valve morphogenesis and development (Supplementary Table 16).

Table 3 | Results from pathway-enrichment analysis for left-handedness (DEPICT and MAGMA)

Pathway ID	Pathway description	P value
DEPICT		
MP:0000788	Abnormal cerebral cortex morphology	0.00000185
MP:0000807	Abnormal hippocampus morphology	0.00000187
GO:0008017	Microtubule binding	0.00000195
MP:0004275	Abnormal postnatal subventricular zone morphology	0.00000549
ENSG00000137285	TUBB2B subnetwork	0.00000571
MP:0000812	Abnormal dentate gyrus morphology	0.00000843
ENSG00000206211	ENSG00000206211 subnetwork	0.0000128
ENSG00000206283	PFDN6 subnetwork	0.0000128
ENSG00000204220	PFDN6 subnetwork	0.0000128
GO:0005874	Microtubule	0.0000140
GO:0021543	Pallium development	0.0000456
MP:0000790	Abnormal stratification in the cerebral cortex	0.0000500
ENSG00000147601	TERF1 subnetwork	0.0000596
GO:0015631	Tubulin binding	0.0000600
REACTOME apoptotic execution phase	REACTOME apoptotic execution phase	0.0000665
GO:0021987	Cerebral cortex development	0.0000696
ENSG00000182901	RGS7 subnetwork	0.0000731
ENSG00000106105	GARS subnetwork	0.0000738
GO:0007409	Axonogenesis	0.0000902
REACTOME: apoptotic cleavage of cellular proteins	REACTOME: apoptotic cleavage of cellular proteins	0.000108
MAGMA		
REACTOME: CRMPs in SEMA3A signalling	CRMPs in SEMA3A signalling	4.13×10^{-8}
REACTOME: axon guidance	Axon guidance	3.13×10^{-6}
GO: neurogenesis	Neurogenesis	3.53×10^{-6}
REACTOME: semaphorin interactions	Semaphorin interactions	9.84×10^{-6}
Matzuk: preovulatory follicle	Preovulatory follicle	1.26×10^{-5}
REACTOME: SEMA3A-PAK-dependent axon repulsion	SEMA3A-PAK-dependent axon repulsion	1.73×10^{-5}
GO: regulation of non-motile cilium assembly	Regulation of non-motile cilium assembly	3.01×10^{-5}

Only results with FDR <5% are shown.

A S-MultiXcan analysis based on the association between predicted gene expression in brain tissues and ambidexterity identified the genes *QTRTD1*, *TMEM215*, *RPL41* and *RAB40C* in addition to those loci identified during the GWAS (Supplementary Table 17). The SMR analysis pinpointed *TUBA1C* and *CYP51A1* as potentially behind the GWAS associations on chromosomes 12 and 7, respectively (Supplementary Table 18).

Heritability of ambidexterity and genetic correlations. The number of ambidextrous individuals in the UKBB was not enough to precisely estimate heritability using maximum likelihood analysis of IBD sharing in close relatives. However, the SNP heritability of ambidexterity on the liability scale (1% prevalence) estimated through LD score regression and REML implemented in BOLT-LMM was higher than that observed for left-handedness ($h^2_g = 0.12$ (s.e. = 0.007) and $h^2_g = 0.15$ (s.e. = 0.014)).

We estimated the genetic correlation between ambidexterity and a catalogue of 1,349 traits with GWAS summary statistics. Our analyses revealed 575 genetic correlations at FDR < 5%. Among the strongest correlations were positive genetic correlations between ambidexterity and traits related to pain and injuries, and body mass index, and a negative genetic correlation with educational

attainment (Supplementary Table 19). Interestingly, the genetic correlation between our left-handedness meta-analysis and our ambidexterity meta-analysis was only moderate ($r_G = 0.24$, s.e. = 0.03), which suggests that there are divergent genetic aetiologies.

Discussion

We carried out the largest genetic study of handedness to date. Our GWAS and SNP heritability analyses conclusively demonstrate that handedness is a polygenic trait, with multiple genetic variants that implicate multiple biological pathways each increasing the odds of being left-handed or ambidextrous by a small amount. We identified 41 left-handedness and 7 ambidexterity loci that reached genome-wide significance. Our findings are in contrast to the single-gene right-shift¹⁰ and dextral-chance¹¹ hypotheses, whereby the causal genes are hypothesized to account for the heritability of handedness. If these large-effect variants do exist, they should have been detected by our GWAS meta-analysis, which provided over 90% statistical power (Supplementary Table 20) to detect variants with effect sizes as small as a 5% increase in odds per allele for common variants (MAF > 0.05) at genome-wide significance ($\alpha = 5 \times 10^{-8}$). Instead, the present findings firmly support the hypothesis that handedness, like many other behavioural and

Table 4 | SNP heritability and heritability of left-handedness estimated using a range of different approaches

Data used	Method	h_g^2 (s.e.)—liability scale
IHC meta-analysis (32 studies)	LD score regression	0.031 (0.013)
UKBB left-handed individuals only as cases	LD score regression	0.033 (0.004)
23andMe left-handed individuals only as cases	LD score regression	0.040 (0.002)
Meta-analysis UKBB, 23andMe and IHC	LD score regression	0.035 (0.002)
UKBB left-handed individuals only as cases (males)	LD score regression	0.042 (0.006)
UKBB left-handed individuals only as cases (females)	LD score regression	0.032 (0.005)
UKBB left-handed individuals only as cases	REML (BOLT-LMM)	0.059 (0.003)
Right- versus left-handed 0.08 < IBD < 0.3 relatives (no C) + siblings 0.65 > IBD > 0.35 with C in the model	A+C+E model	A = 0.12 ^a (95% CI = 0.07–0.17), C = 0.045 (95% CI = 0–0.09)
Right- versus left-handed 0.08 < IBD < 0.3 relatives (no C) + siblings 0.65 > IBD > 0.35 without C in the model	A+E model	A = 0.20 (95% CI = 0.14–0.26)
Meta-analysis of twin studies of handedness ¹²	A+C+E model	A = 0.25 (95% CI = 0.157–0.30), C = 0 (95% CI = 0–0.076%)

^aEstimate of narrow-sense heritability (h^2).

neurological traits, is influenced by many variants of small effect and multiple biological pathways.

Using different methods and cohorts, we estimated the SNP heritability (h_g^2) of handedness to be between 3% and 6%. However, by using IBD-based methods applied to siblings and other relative pairs, we estimated the narrow-sense heritability (h^2) to be 11.9% (95% CI = 7.2–17.7). Although this is lower than that obtained from twin studies (25%, 95% CI = 15.7–29.5 (ref. ¹²) and 21%, 95% CI = 11–30 (ref. ³⁹)), the CIs for the estimates overlap. Interestingly, h_g^2 estimates for ambidexterity were larger (12–15%), which suggests that common SNPs tag a higher proportion of variability in liability to ambidexterity than in liability to left-handedness.

Our GWAS meta-analysis of left-handedness identified eight loci close to genes involved in microtubule formation and regulation. An enrichment for microtubule-related pathways was then confirmed by the DEPICT analysis. Microtubules are polymers that form part of the cytoskeleton and are essential in several cellular processes, including intracellular transport, cytoplasmic organization and cell division. With respect to handedness, microtubule proteins play important roles during the development and migration of neurons, plasticity and neurodegenerative processes^{40,41}. The association between handedness and variation in microtubule genes also provides insights into differences in the prevalence of various neuropsychiatric disorders and left-handedness observed in some epidemiological studies^{42,43}. Recent genetic studies have identified mutations in a wide variety of tubulin isotypes and

microtubule-related proteins in many major neurodevelopmental and neurodegenerative diseases^{40,44–46}.

We observed an association between left-handedness and the 17q21.31 locus. A deletion in this locus is known to cause Koolen de Vries syndrome, a disorder characterized by intellectual disability, developmental delay and neurological abnormalities of the corpus callosum, hippocampi and ventricles. Variation in the 17q21 locus, including structural variation, has been associated with schizophrenia⁴⁷, autism^{48,49} and cognition⁵⁰. In addition, based on our PheWAS, the rs55974014 SNP within this locus has been associated with mood swings, neuroticism and educational attainment traits. The rs55974014 SNP is located near several genes with neurological functions, including *CRHR1* and *NSF*. Other SNPs close to this gene have been associated with intelligence⁵¹ and Parkinson's disease⁵². Future colocalization analyses are warranted to assess the veracity of the same variant affecting multiple traits.

The genetic correlation between left-handedness and ambidexterity was low, which suggests that the genetic architecture underlying the two traits is different. Only 15 out of the 41 loci associated with left-handedness were associated with ambidexterity at marginal significance levels or lower ($P < 0.05$). However, tissue- and pathway-enrichment analyses indicated that just as for left-handedness, the central nervous system was implicated. Ambidexterity showed significant genetic correlations with multiple traits, particularly anthropometric and those involving pain and injuries. This suggests that reporting being able to write with both hands may be a result from injuries that led to the use of the other hand or may be due to increased injury risk. Future studies into the genetics of ambidexterity should include detailed phenotyping that considers the reasons leading to hand-use preference.

In contrast, left-handedness was not significantly genetically correlated with other (non-handedness) traits in our study. Given that previous studies have shown that the phenotypic correlation between left-handedness and most traits and diseases is low, it is perhaps unsurprising that the magnitude of most genetic correlations with left-handedness was low also. However, our lack of significant results was also partially a reflection of the large number of statistical tests performed and the conservative testing correction applied when estimating the genetic correlation between left-handedness and other traits. Nevertheless, we observed a clear inflation of the distribution of P values when compared to the null, which indicates that there is likely to be a small degree of genetic overlap between handedness and other traits. Among the suggestive genetic correlations ($P < 0.05$), we observed positive genetic correlations between left-handedness and schizophrenia and bipolar disorder, which is consistent with previous observations of greater atypical hand dominance in patients with schizophrenia and in patients with bipolar disorder^{53,54}.

The present study benefitted from having a large sample size that allowed the detection of dozens of novel variants of small effect on handedness. However, it is worth noting that the genetic correlations derived from GWAS summary statistics of left-handedness in the IHC, 23andMe and UKBB data were high but statistically different from one, potentially affecting the statistical power of the meta-analysis. These differences may have been due to the way data were collected in each of the cohorts. For example, in the UKBB, handedness data were obtained at up to three occasions, while this was not the case for the 23andMe and the IHC cohorts. Furthermore, genetic correlations with the IHC data may have been affected by the IHC including ambidextrous and left-handed individuals as cases.

In summary, we report the world's largest GWAS meta-analysis of handedness. We showed that handedness is polygenic and found evidence that microtubule genes may play an essential role in lateralization. Loci mapped in the present study warrant further exploration of their potential role in neurological development and laterality.

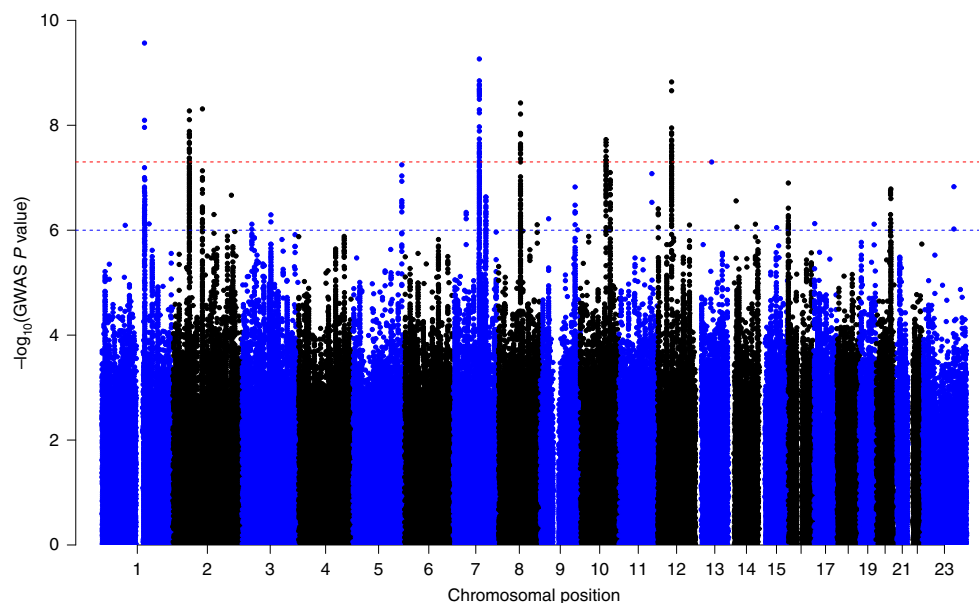


Fig. 2 | Manhattan plot of the ambidexterity meta-analysis. Manhattan plots for the ambidexterity GWAS meta-analysis ($N=1,422,823$ right-handed versus 37,637 ambidextrous individuals). Each dot represents a SNP. The red broken line highlights the genome-wide levels of significance threshold ($P < 5 \times 10^{-8}$); the blue broken line shows the threshold for suggestive associations.

Table 5 | Loci associated with ambidexterity after a meta-analysis of 23andMe and UKBB data

CHR	BP	SNP	Gene	EA	NEA	EAF	Z	OR ^a	P	Direction
1	150317558	rs782122127	<i>PRPF3</i>	D	I	0.19	-6.32	0.88	2.70×10^{-10}	?–
2	58196110	rs2030237	<i>VRK2</i>	A	G	0.58	5.84	1.04	5.29×10^{-9}	++
2	104437850	rs139630683	<i>ACO13727.1</i>	D	I	0.45	5.85	1.05	4.88×10^{-9}	+?
7	91899117	rs2040498	<i>ANKIB1</i>	A	T	0.65	6.21	1.06	5.42×10^{-10}	++
8	77104817	rs10113066	<i>RNU2-54P</i>	T	G	0.51	5.90	1.05	3.74×10^{-9}	++
10	89722731	rs36062478	<i>PTEN</i>	T	C	0.87	-5.62	0.94	1.87×10^{-8}	--
12	49530132	rs35554786	<i>TUBA1B</i>	D	I	0.24	-6.05	0.93	1.49×10^{-9}	-?

The direction of effects is shown in the following order: 23andMe and UKBB. ^aOR corresponds to that derived from the 23andMe results. Where the SNP is missing in a cohort, a question mark is indicated in the Direction column.

Methods

As described below, we performed a meta-analysis on GWAS results from the IHC, the UKBB and 23andMe. Informed consent was provided by all participants. The research was approved by the research ethics committee of each of the individual studies.

Genome-wide association in the UKBB. The UKBB is a large long-term biobank study from the United Kingdom that aims to identify the contribution of genetic and environmental factors to disease. Detailed information on phenotyping and genotyping is presented elsewhere⁵⁵. In brief, the UKBB recruited 502,647 individuals aged 37–76 years across the country and gathered information regarding their health and lifestyle, including handedness, via questionnaires. Genotype data from the UKBB are available for 487,411 participants. Genotypes were imputed by the UKBB against the UK10K reference panel using IMPUTE 2 (ref. ⁵⁶). In addition to the quality control metrics performed centrally by the UKBB⁵⁵, we defined a set of participants of European ancestry by clustering the first two principal components (PCs) derived from the genotype data. Using a K-means algorithm with $K=4$, we identified a group of 463,023 individuals of European ancestry. From this group, 462,182 individuals (250,767 females) provided self-reported data on handedness. In total, 410,677 participants identified themselves as right-handed, 43,859 as left-handed and 7,646 as ambidextrous. The mean birth year of the participants was 1951 (s.d. = 8.04).

We tested 11,498,822 autosomal and X chromosome SNPs with MAF > 0.005 and info score of > 0.4 for associations with handedness using BOLT-LMM, which implements a linear mixed model to account for cryptic relatedness and population structure. Sex and age were included as covariates in all models. We performed

four analyses: (1) right- versus left-handed; (2) right versus ambidextrous; (3) right- versus left-handed (male only); and (4) right- versus left-handed (female only). Analyses of X chromosome genotypes were performed in BOLT-LMM, fitting sex as a covariate and coding the male genotypes as 0/2.

Genome-wide association in 23andMe. All individuals included in the analyses were research participants of the personal genetics company 23andMe, Inc., a private company. The phenotypes, including self-reported handedness, of the research participants were collected via online surveys. DNA extraction and genotyping were performed on saliva samples by the National Genetics Institute (NGI), a CLIA-licensed clinical laboratory and a subsidiary of the Laboratory Corporation of America. Samples were genotyped on one of five genotyping platforms. The v.1 and v.2 platforms were variants of the Illumina HumanHap550+ BeadChip, including about 25,000 custom SNPs selected by 23andMe, with a total of about 560,000 SNPs. The v.3 platform was based on the Illumina OmniExpress+ BeadChip, with custom content to improve the overlap with the v.2 array, with a total of about 950,000 SNPs. The v.4 platform was a fully customized array, including a lower redundancy subset of v.2 and v.3 SNPs with additional coverage of lower-frequency coding variation, and about 570,000 SNPs. The v.5 platform was an Illumina Infinium Global Screening Array (~640,000 SNPs) supplemented with ~50,000 SNPs of custom content. Samples that failed to reach a 98.5% call rate were re-analysed. Individuals whose analyses repeatedly failed were re-contacted by 23andMe customer service to provide additional samples.

For our standard GWAS, we restricted participants to a set of individuals who have a specified ancestry (predominantly European ancestry) determined through an analysis of local ancestry⁵⁷. A maximal set of unrelated individuals was chosen

for each analysis using a segmental IBD estimation algorithm⁵⁸. Individuals were defined as related if they shared more than 700 cM IBD, including regions where the two individuals share either one or both genomic segments of IBD. This level of relatedness (roughly 20% of the genome) corresponds approximately to the minimal expected sharing between first cousins in an outbred population. In total, 1,012,146 individuals included in the analysis identified themselves as right-handed, 136,740 as left-handed and 29,991 as ambidextrous. The mean birth year of the participants was 1972.

We used Minimac3 to impute genotype data against a reference panel consisting of the May 2015 release of the 1000 Genomes Phase 3 haplotypes⁵⁹ and the UK10K imputation reference panel⁶⁰. We computed associations by logistic regression assuming additive allelic effects. We used the imputed dosages rather than best-guess genotypes and included covariates for age, sex, the top five PCs to account for residual population structure and indicators for genotype platforms to account for genotype batch effects. For associations on the X chromosome, male genotypes were coded as if they were homozygous diploid for the observed allele. For quality control of the GWAS results, we removed SNPs with $rsq < 0.3$, $MAF < 0.005$ and available sample size $< 20\%$ of the total sample, as well as SNPs that had strong evidence of a platform batch effect. We also flagged logistic regression results that did not converge due to complete separation, identified by $abs(effect) > 10$ or $s.e. > 10$ on the log odds scale. Two analyses were performed: (1) right- versus left-handed and (2) right versus ambidextrous.

The IHC. The IHC is a large-scale collaboration between 32 cohorts ($N = 125,612$) with existing GWAS data to identify common genetic variants influencing handedness. Across all studies, the phenotype was collected by a questionnaire that asked either which hand was used for writing or for self-declared handedness. As these two measures are highly ($\sim 95\%$) concordant, not all studies reported on ambidexterity, and around 1–2% of participants reported being able to write with both hands^{12,61}; both left-handed and ambidextrous individuals were classified as cases where this information was available. All cohorts were population samples with respect to handedness, thus combining the data from the 32 studies yielded 13,599 left-handed and 112,013 right-handed individuals (Supplementary Table 1).

All individuals were of self-declared European ancestry (confirmed by genotypic PCA in each cohort). Within each cohort, the genotypic data were imputed to Phase I and II combined HapMap CEU samples (build 36 release 22) with the exception of the Finnish Twin Cohort study, the Health Professionals Follow-Up Study and Nurses' Health Study HPFS/NHS, the Netherlands Twin Registry (NTR), and TOP cohorts, which were imputed to 1000G Phase 3 V5 European population. Within each sample, genome-wide association analyses were conducted for both genotyped and imputed SNPs. The imputed genotypes were analysed using the dosage of an assumed effect allele under an additive model with covariates for year of birth and sex. Supplementary Table 21 shows the imputation and analysis software used in each of the cohorts.

To examine any potential impacts of including ambidextrous individuals as cases on the IHC, we ran a GWAS on the UKBB and 23andMe samples using both phenotypic definitions ((1) left versus right and (2) right versus left + ambidextrous) and computed the genetic correlations between the two analyses. The genetic correlations were $r_G^{23andMe} = 0.95$ (s.e. = 0.003) and $r_G^{UKBB} = 0.98$ (s.e. = 0.006), which suggests that there is only a minor impact of the inclusion of ambidextrous individuals as cases on the GWAS results.

Meta-analysis of IHC, UKBB and 23andMe. A weighted Z-score meta-analysis was conducted with the METAL software⁶² using the summary GWAS statistics from each of the 32 IHC cohorts, UKBB and 23andMe. Given the large discrepancies between the number of cases and controls, we elected to weight each sample by the effective sample size for binary traits, defined as $N_{eff} = 4 / (1/N_{cases} + 1/N_{controls})$.

Before the meta-analysis, quality control thresholds were applied to each of the GWAS results from the individual studies ($r^2 \geq 0.3$, $MAF \geq 0.005$, $P_{HWE} \geq 1 \times 10^{-5}$). We used EasyQC⁶³ to identify and remove SNPs that had allele frequencies that substantially differed from the Haplotype Reference Consortium. In total, up to 13,346,399 SNPs remained for the left-handedness meta-analysis. For the ambidexterity meta-analysis, only 23andMe and the UKBB datasets were used. This meta-analysis included up to 12,493,443 SNPs.

Tissue-expression and pathway analyses. Tissue-expression and pathway analyses were performed using DEPICT (v.1 rel. 194)²⁵ implemented in the CTG-VL (beta 0.1)³⁸ and MAGMA²⁶ implemented in the functional mapping and annotation of genetic associations (FUMA) web application⁶⁴ (accessed on 20 March 2020). DEPICT assesses whether genes in associated loci are highly expressed in any of the 209 medical subject heading (MeSH) tissue and cell-type annotations based on RNA sequencing data from the GTEx project³². Molecular pathways were constructed based on 14,461 gene sets from diverse database and data types, including Gene Ontology, the Kyoto encyclopedia of genes and genomes (KEGG) and REACTOME. As input for DEPICT, we used independent SNPs (based on clumping with a r^2 (LD) between SNPs < 0.05 and 2-Mb windows) with $P < 1 \times 10^{-5}$. MAGMA analyses were performed with the default options of FUMA using data 1000 Genomes Phase 3 for LD reference and GTEx v.8 (ref. 32) for the

tissues-enrichment analysis. Enrichment analyses were performed for 15,483 pathways from MsigDB v.7.0 (ref. 65) and 54 tissues from GTEx. Associations with Benjamini–Hochberg FDR values of $< 5\%$ are reported for both analyses.

Gene-based association analyses. Gene-based association analyses were carried out using S-MultiXcan²⁷ and the SMR method²⁷ implemented in the CTG-VL (beta 0.1)³⁸. S-MultiXcan conducts a test of association between phenotypes and gene-expression levels predicted by data derived from the GTEx project³². In this study, we performed S-MultiXcan using prediction models of all the brain tissues available from the GTEx project. This included amygdala, anterior cingulate cortex BA24, caudate basal ganglia, cerebellar hemisphere, cerebellum, brain cortex, frontal cortex BA9, hippocampus, hypothalamus, nucleus accumbens basal ganglia, putamen basal ganglia, spinal cord cervical c-1 and substantia nigra samples. As a total of 14,501 genes expressed in different brain tissues were tested, the Bonferroni-corrected significance threshold was set at $P = 3.44 \times 10^{-6}$.

SMR conducts a test for pleiotropic associations between the expression level of a gene and a complex trait using eQTL data and GWAS summary statistics. SMR uses a heterogeneity in dependent instruments (HEIDI) test to distinguish pleiotropy from linkage. A rejection ($P < 0.05$) of the null hypothesis (pleiotropy) indicates that the association of the SNP with gene expression and the trait of interest is probably due to linkage of that SNP with two distinct causal SNPs (one for the gene expression and one for the trait). In this study, we performed SMR using eQTL data derived from a meta-analysis of eQTL data from brain tissues (Brain-eMeta)³¹ that included GTEx³², CMC³³ and ROSEMAP³⁴. In addition to this, we carried out SMR analysis using PsychENCODE eQTL summary data from Wang et al.²⁹ and Gandal et al.³⁰. For each of the analyses, we report associations below a Bonferroni-corrected significance threshold based on the number of genes tested that was up to 11,013 genes (that is, $P < 4.54 \times 10^{-6}$).

Genetic correlations. To test whether handedness shares a genetic background with other complex traits with available GWAS summary data, we used CTG-VL³⁸, which implements LD score regression and contains a large database of summary GWAS statistics. In total, we assessed the genetic correlation of left-handedness and ambidexterity with 1,349 different traits. As many of the 1,349 traits tested are correlated between each other, hence not considered independent tests, we adopted a Benjamini–Hochberg FDR of $< 5\%$ to account for multiple testing.

Heritability estimates. To estimate the proportion of phenotypic variance explained by SNPs, we used two statistical methods. REML, implemented in BOLT-LMM, was used to estimate the variance explained by additive effects of genotyped SNPs (h^2_g)⁶⁶. Using a prevalence estimate of 10%, the observed h^2_g was transformed to SNP heritability on an unobserved continuous liability scale⁶⁷. LD score regression was used to estimate the variance explained by all the SNPs using the GWAS summary statistics. Similar to REML, the observed h^2_g was transformed to the liability scale using a prevalence estimate of 10%.

To estimate the narrow-sense heritability, we fit a variance components model to estimate the proportion of phenotypic variance attributable to additive genetic effects (A), shared environmental effects (C) and individual environmental effects (E)⁶⁸. We modelled the genetic sharing between close relative pairs using IBD information (as calculated by the KING software (v.2.1.6)⁶⁸) on 20,277 sibling pairs ($0.65 > IBD > 0.35$) and 49,788 relative pairs with $0.3 > IBD > 0.8$ from the UKBB study to estimate trait heritability. In the model, we also estimated a variance component due to a shared environment (siblings only), which made siblings potentially more similar in terms of handedness, and a unique environmental component, which did not contribute to similarity between relative pairs. Variance components were estimated using maximum likelihood using the OpenMx package⁶⁹.

Reporting Summary. Further information on research design is available in the Nature Research Reporting Summary linked to this article.

Data availability

GWAS summary statistics of the meta-analysis of the UKBB, IHC and 23andMe data for the top 10,000 independent SNPs as well as summary statistics of the meta-analysis between the UKBB and IHC data for all the SNPs are available at <https://evansgroup.di.uq.edu.au/gwas-results.html>. Access to the full summary statistics from the 23andMe sample (for all SNPs) can be obtained by qualified researchers through a data transfer agreement with 23andMe that protects participant privacy. Please contact 23andMe at <https://research.23andme.com/dataset-access> for more information.

Code availability

The code used to perform the meta-analysis will become available on GitHub upon publication.

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Competing interests

G.C.-P., N.E., D.A.H. and J.Y.T. are employees of 23andMe, Inc., and hold stock or stock options in 23andMe. S.H.M., K.S., H.S., S.S. and G.T. are employees of deCODE Genetics/Amgen. M.I.M. is a Wellcome Senior Investigator and a NIHR Senior Investigator. The views expressed in this article are those of the author(s) and not necessarily those of the NHS, the NIHR or the Department of Health. M.I.M. has served on advisory panels for Pfizer, NovoNordisk and Zoe Global; has received honoraria from Merck, Pfizer, NovoNordisk and Eli Lilly; has stock options in Zoe Global; has received research funding from Abbvie, AstraZeneca, Boehringer Ingelheim, Eli Lilly, Janssen, Merck, NovoNordisk, Pfizer, Roche, Sanofi Aventis, Servier and Takeda. As of June 2019, he is an employee of Genentech, and holder of Roche stock. All other authors report no conflicts of interest.

Additional information

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| <input type="checkbox"/> | <input checked="" type="checkbox"/> | A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons |
| <input type="checkbox"/> | <input checked="" type="checkbox"/> | A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient) AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals) |
| <input type="checkbox"/> | <input checked="" type="checkbox"/> | For null hypothesis testing, the test statistic (e.g. F , t , r) with confidence intervals, effect sizes, degrees of freedom and P value noted
<i>Give P values as exact values whenever suitable.</i> |
| <input checked="" type="checkbox"/> | <input type="checkbox"/> | For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings |
| <input checked="" type="checkbox"/> | <input type="checkbox"/> | For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes |
| <input checked="" type="checkbox"/> | <input type="checkbox"/> | Estimates of effect sizes (e.g. Cohen's d , Pearson's r), indicating how they were calculated |

Our web collection on [statistics for biologists](#) contains articles on many of the points above.

Software and code

Policy information about [availability of computer code](#)

Data collection This does not apply to this manuscript as the paper reports a meta-analysis and no data collection was undertaken

Data analysis Supplementary Table 21 show the imputation and analysis software used in each of the cohorts.
The code used to perform the meta-analyses will become available upon publication on GitHub.

For manuscripts utilizing custom algorithms or software that are central to the research but not yet described in published literature, software must be made available to editors and reviewers. We strongly encourage code deposition in a community repository (e.g. GitHub). See the Nature Research [guidelines for submitting code & software](#) for further information.

Data

Policy information about [availability of data](#)

All manuscripts must include a [data availability statement](#). This statement should provide the following information, where applicable:

- Accession codes, unique identifiers, or web links for publicly available datasets
- A list of figures that have associated raw data
- A description of any restrictions on data availability

The code used to perform the meta-analyses will become available upon publication on GitHub. GWAS summary statistics of the Meta-analysis of UK Biobank, IHC and 23andMe data for the top 10,000 independent SNPs will be available upon publication as well as summary statistics of the Meta-analysis between UK Biobank and IHC for all the SNPs. Access to the full summary statistics from the 23andMe sample (for all SNPs) can be obtained by qualified researchers through a data transfer agreement with 23andMe. Please contact 23andMe in <https://research.23andme.com/dataset-access> for more information.

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Life sciences study design

All studies must disclose on these points even when the disclosure is negative.

Sample size	In this study, we present findings from the world's largest GWAS meta-analysis of handedness to date (N = 1,766,671), combining data from 32 cohorts from the International Handedness Consortium (IHC) (N = 125,612), 23andMe (N = 1,178,877) and UK Biobank (N = 462,182). No sample size/power calculations were undertaken.
Data exclusions	At the site level participants were excluded if phenotype, covariate and/or genotypic data were missing or if the genotypic data for an individual failed quality control. All participants were of European ancestry. There were no cohort-level exclusions from the meta-analysis and all cohorts that contributed data were included in the analyses. Prior to meta-analysis, quality control thresholds were applied to each of the GWAS results from the individual studies ($r^2 \geq 0.3$, $MAF \geq 0.005$, $PHWE \geq 1 \times 10^{-5}$). We also removed genetic variants for which the frequency substantially differed from one of the Haplotype Reference Consortium panels (frequency difference > 0.2). We used EasyQC 55 to identify SNPs that had allele frequencies which differed substantially from the Haplotype Reference Consortium. In total, up to 13,346,399 SNPs remained for the left-handedness meta-analysis. For the ambidexterity meta-analysis, only 23andMe and the UK Biobank were used. This meta-analysis included up to 12,493,443 SNPs.
Replication	As the frequency of left-handedness is ~10% in European populations there were no sufficiently large populations with GWAS data available to replicate these findings with sufficient power at the time of analysis. In the ST2 and ST12 we provide the results for both the meta-analysis and the cohort-level results from the UK Biobank, 23andMe and IHC analyses to allow readers to examine the replicability of findings. We used also LD-score genetic correlation analyses to assess the internal replicability of findings at the genome-wide level. The genetic correlations as estimated by bivariate LD-score regression between the results from the UK Biobank, 23andMe and IHC GWAS were high ($rg_{UKB-23andMe} = 0.88$, $s.e. = 0.05$, $rg_{UKB-IHC} = 0.73$, $s.e. = 0.16$, $rg_{IHC-23andMe} = 0.60$, $s.e. = 0.11$).
Randomization	This is not applicable as the paper reports a meta-analysis of genome-wide association analyses for an observed variable with no experimental manipulation.
Blinding	This is not applicable as the paper reports a meta-analysis of genome-wide association analyses for an observed variable with no experimental manipulation.

Reporting for specific materials, systems and methods

We require information from authors about some types of materials, experimental systems and methods used in many studies. Here, indicate whether each material, system or method listed is relevant to your study. If you are not sure if a list item applies to your research, read the appropriate section before selecting a response.

Materials & experimental systems

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Methods

n/a	Included in the study
<input checked="" type="checkbox"/>	<input type="checkbox"/> ChIP-seq
<input checked="" type="checkbox"/>	<input type="checkbox"/> Flow cytometry
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