

1 Cooperative phenotype predicts economic conservatism, policy views, and political party  
2 support

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## Abstract

14 Decades of research suggest that our political differences are best captured by two  
15 dimensions of political ideology: economic and social conservatism. The dual evolutionary  
16 framework of political ideology predicts that these dimensions should be related to variation  
17 in general preferences for cooperation and group conformity. Here, we show that, controlling  
18 for a host of demographic covariates, a general cooperative preference captured by a suite of  
19 incentivised economic games (the “cooperative phenotype”) is indeed negatively correlated  
20 with two widely-used measures of economic conservatism — Social Dominance Orientation  
21 and Schwartz’s altruistic vs. self-enhancement values. The cooperative phenotype also  
22 predicts political party support and economically progressive views on political issues like  
23 income redistribution, welfare, taxation, and environmentalism. By contrast, a second  
24 “norm-enforcing punishment” dimension of economic game behaviour, expected to be a proxy  
25 for social conservatism and group conformity, showed no reliable relationship with political  
26 ideology. These findings reveal how general social preferences that evolved to help us  
27 navigate the challenges of group living continue to shape our political differences even today.

28 *Keywords:* cooperation, punishment, behavioural economics, political ideology,  
29 economic conservatism

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32 support

33 Humans differ profoundly in their views on political issues like income redistribution,  
34 taxation, welfare, military spending, and criminal justice. With increasing levels of political  
35 partisanship<sup>1,2</sup>, understanding the sources of this variation is more pressing than ever. One  
36 popular explanation for political differences is that people vary along a unidimensional  
37 liberal-conservative spectrum of political ideology<sup>3</sup>. This model is widely referred to in both  
38 the popular media and scientific literature<sup>4</sup>. However, attitudinal variation in the modern  
39 electorate reveals that the liberal-conservative spectrum fails to capture the full diversity of  
40 human political differences. For example, libertarians in the United States support  
41 conservative free market and pro-business policies, but often simultaneously hold socially  
42 progressive views on abortion, same-sex marriage, and gender roles<sup>5</sup>. In Europe, people are  
43 increasingly supporting political parties that promote the welfare state (a “liberal” policy)  
44 but also aim to limit immigration (a “conservative” policy)<sup>6</sup>. Examples like these show that  
45 a single liberal-conservative spectrum misses important features of the political landscape.

46 In contrast to the unidimensional view, decades of interdisciplinary work indicate that  
47 political ideology is best described along two dimensions<sup>4,7-11</sup>. The first dimension, often  
48 referred to as economic conservatism, captures views on issues like taxation, welfare,  
49 capitalism, and big business<sup>9</sup>. The second dimension, often referred to as social conservatism,  
50 captures views on issues like military spending, patriotism, sexual morality, and criminal  
51 justice<sup>9</sup>. These two dimensions have repeatedly emerged from independent lines of research  
52 in political psychology<sup>8</sup>, moral psychology<sup>12</sup>, and cross-cultural psychology<sup>7</sup>, though they are  
53 often given different labels in different fields<sup>4,8</sup>. While economic and social conservatism tend  
54 to positively covary in Western societies<sup>8</sup>, they are distinct<sup>8</sup> and are negatively correlated  
55 with one another in many societies around the globe<sup>13</sup>.

56 Although previous research has repeatedly identified two dimensions of political

57 ideology, less work has examined the essential nature of the two dimensions or asked why  
58 this particular two-dimensional structure organises political attitudes. An evolutionary  
59 approach has the potential to provide answers at this ultimate (rather than proximate<sup>14</sup>)  
60 level of analysis<sup>15</sup>. Evolutionary approaches to human political ideology are supported by  
61 evidence that the two dimensions of ideology are heritable<sup>16,17</sup>, found across cultures<sup>7</sup>, and  
62 correlated with physiological and neurological differences<sup>18</sup>. These approaches broadly define  
63 politics as the process of dealing with the conflicts of interest that arise from human group  
64 living<sup>19-21</sup>, and align with Harold Lasswell's famous definition of politics as the process of  
65 deciding "who gets what, when, and how"<sup>22</sup>. Defined in this way, politics is not a modern  
66 phenomenon. Political tensions have characterised human groups throughout our  
67 evolutionary history, and have precursors in our group-living primate relatives<sup>23</sup>.

68 We recently proposed two evolutionary foundations of political ideology<sup>4</sup>. In this dual  
69 evolutionary framework, we argue that the two dimensions of political ideology reflect two  
70 key steps in the evolution of human group living<sup>24-28</sup>. In the first key step, humans  
71 developed more egalitarian sharing preferences and began to cooperate across wider  
72 interdependent networks. In the second key step, humans created cultural markers to  
73 identify members of their group, and began adhering to and enforcing group-wide social  
74 norms to enhance group viability. These two fundamental challenges of group living favoured  
75 general preferences for cooperation and group conformity in ancestral humans, transitioning  
76 human group living from the small kin networks characteristic of great apes to relatively  
77 egalitarian, culturally-bound hunter-gatherer communities<sup>26,29,30</sup>.

78 This dual evolutionary framework hypothesises that variation in general preferences for  
79 cooperation and group conformity underlie the two dimensions of political ideology in  
80 modern humans<sup>4</sup>. Functional variation in general preferences can be maintained either by  
81 balancing selection on fitness trade-offs, creating enduring heritable individual differences, or  
82 by behavioural plasticity in response to local socio-ecological conditions<sup>4</sup>. This emergent  
83 variation in general preferences is predicted to produce differences in political opinion. Under

84 the dual evolutionary framework, an increased general preference for cooperation beyond  
85 close kin results in greater support for economically progressive policies that promote  
86 large-scale cooperation, such as income redistribution, taxation, welfare, and  
87 pro-environmentalism. By contrast, an increased general preference for group conformity  
88 results in greater support for socially conservative policies that promote in-group conformity  
89 and norm enforcement, such as military spending, capital punishment, and promoting  
90 traditional religious values.

91 If general preferences do underlie differences in political opinion, such differences  
92 should be manifest in social preferences identified by behavioural economics. Behavioural  
93 economic games allow researchers to study variation in enduring social preferences<sup>31</sup>.  
94 Experimental games model the basic payoff matrices of social interactions, abstracting away  
95 from the complexity of the real world to capture variation in “pure” preferences. Decades of  
96 work using experimental economic games have revealed that humans are not selfish  
97 payoff-maximisers, as predicted by economic theory, but incorporate the payoffs of others  
98 into their cooperative decision-making<sup>32–36</sup>. This research also highlights wide variation in  
99 social preferences between individuals<sup>37–40</sup>.

100 Previous research has studied the covariation between social preferences in economic  
101 games and political ideology. Studies linking gameplay to unidimensional political party  
102 support have produced mixed results, with some studies finding that supporters of  
103 left-leaning parties are more cooperative in economic games<sup>41–44</sup> and other studies finding no  
104 relationship between gameplay and political party support<sup>45,46</sup>. However, a number of studies  
105 find support for an association between cooperative preferences and the economic dimension  
106 of ideology<sup>47,48</sup>. A recent meta-analysis of the links between personality and prosocial  
107 behaviour found that cooperative behaviour in a number of social dilemma games that allow  
108 exploitation of others was negatively correlated with Social Dominance Orientation (SDO; a  
109 measure of support for inequality and hierarchy, widely used as a proxy for economic  
110 conservatism)<sup>49</sup>. In contrast, other research has failed to find a relationship between SDO

111 and cooperative behaviour in the Prisoner's Dilemma Game, in which two players choose  
112 whether to cooperate or defect<sup>46,47</sup>, or the Stag Hunt Game, in which two players must  
113 coordinate on the same strategy<sup>47</sup>. These findings suggest that the economic dimension of  
114 ideology may relate to particular affordances of individual games (e.g., exploitation<sup>49</sup>) rather  
115 than to a general cooperative preference that applies across situations with different payoff  
116 structures. Furthermore, research exploring the relationship between the social dimension of  
117 ideology and general conformist and norm-enforcing preferences remains limited.

118 To overcome these limitations, we build on previous work that identified general social  
119 preferences using a battery of economic games. Peysakhovich et al.<sup>50</sup> asked participants to  
120 play three cooperation games and three punishment games. The cooperation games were the  
121 Dictator Game (in which players can share their endowment with another player), the Trust  
122 Game (in which players can transfer money to another player who may or may not return  
123 some of the multiplied amount), and the Public Goods Game (in which multiple players can  
124 choose to contribute to a shared public good). The punishment games were the Ultimatum  
125 Game (in which players reject unfair offers), the Second-Party Punishment Game (in which  
126 players punish defection in a Prisoner's Dilemma), and the Third-Party Punishment Game  
127 (in which players punish stealing behaviour as an impartial observer). Factor analysis  
128 revealed that the cooperative decisions all positively covaried and could be described by a  
129 single underlying latent variable, dubbed the "cooperative phenotype". The study also  
130 reported a positive correlation between the cooperative phenotype and a single item  
131 measuring support for "an increase in taxes if it were used to help the less well off in society"  
132 (an economically progressive view), but the authors did not analyse potential links with  
133 political ideology. While subsequent research has replicated the cooperative phenotype factor  
134 structure<sup>47,51-53</sup> across several cultures<sup>54</sup>, this work has not systematically examined the  
135 association between the cooperative phenotype and economic conservatism, policy views, and  
136 political party support. In addition, Peysakhovich et al.<sup>50</sup> distinguished the cooperative  
137 phenotype from a second latent variable that emerged from the punishment games, labelled

138 “norm-enforcing punishment”<sup>55,56</sup>. Under the dual evolutionary framework of political  
139 ideology, this latent variable could plausibly be linked to social conservatism.

140 Here, we extend this prior work by combining an expanded set of economic games with  
141 comprehensive survey data to test whether and to what extent the cooperative phenotype  
142 covaries with widely-used measures of economic conservatism, as well as policy views and  
143 political party support. In addition, we examine the relationship between the second latent  
144 variable (norm-enforcing punishment) and social conservatism. In our pre-registered study  
145 (<https://osf.io/dwx8g/>), we recruited a diverse nationally-representative New Zealand sample  
146 ( $n = 926$ ) from the ongoing longitudinal New Zealand Attitudes and Values Study<sup>57</sup> (see  
147 Methods and Figure S1 for sample characteristics). We asked these participants to play a  
148 suite of eight real-time anonymous one-shot economic games measuring both cooperation  
149 and punishment (see Methods for detailed descriptions of games). Six of these games were  
150 identical to those used in Peysakhovich et al.<sup>50</sup>, and two were additional coordination games:  
151 a version of the Stag Hunt Game, in which multiple players must coordinate on the same  
152 strategy, preferably the payoff-dominant equilibrium; and the Stag Hunt Game with  
153 Punishment, in which players may punish others for not choosing the strategy commensurate  
154 with the payoff-dominant equilibrium. We added these two games to study whether the  
155 general preferences for cooperation and punishment extended to coordination problems that  
156 do not involve exploitation of others and are arguably better models for the real-world  
157 cooperative dilemmas faced by our human ancestors<sup>26</sup>. Participants played the online  
158 economic games simultaneously with other participants across New Zealand. Gameplay was  
159 incentivised with a fixed \$20 NZD show-up fee plus a bonus payment of between \$10-35 ( $M$   
160  $= \$25.17$ ,  $SD = \$2.47$ ) depending on participants’ decisions, resulting in \$41,826 overall  
161 spent on participant payment.

162 We pre-registered three hypotheses. First, we hypothesised that a two-factor structure  
163 of cooperation and punishment would emerge from this expanded suite of games, replicating  
164 and extending previous work. Second, we hypothesised that SDO would negatively predict

165 variation in the cooperation factor. Third, we hypothesised that, to the extent that  
166 punishment in our economic games reflects norm-enforcement, Right Wing Authoritarianism  
167 (RWA; a widely-used measure of social conservatism) would positively predict variation in  
168 the punishment factor. Beyond these pre-registered hypotheses, we took advantage of the  
169 wealth of survey data from the New Zealand Attitudes and Values Study to further explore  
170 the relationships between gameplay and additional measures of political ideology, policy  
171 views, and political party support.

172

## Results

173 To test the first hypothesis that gameplay would produce a two-factor structure, we  
174 focused on patterns of covariation among the economic games with pre-registered  
175 correlational, principal components, and confirmatory factor analyses. Pairwise Spearman's  
176 rank correlations between game decisions are visualised in Figure 1a. Replicating previous  
177 work<sup>50</sup>, we found significant positive correlations between most cooperation decisions ( $r_s =$   
178  $0.07-0.33$ ), significant positive correlations between all punishment decisions ( $r_s = 0.10-0.49$ ),  
179 and non-significant or negative correlations between most cooperation and punishment  
180 decisions. Principal components analyses suggested that these patterns of covariation could  
181 be explained by two underlying factors: cooperation and punishment. When including only  
182 games from previous work, principal components analysis supported a two-factor solution.  
183 Using orthogonal varimax rotation and retaining factors with eigenvalues above one (Figure  
184 S2a), this analysis revealed that cooperative behaviour loads highly onto one factor and  
185 punitive behaviour loads highly onto a second distinct factor (Figure 1b). Together, these  
186 factors explained 44% of the variance in game decisions. This two-factor structure held when  
187 adding our novel coordination games, with the Stag Hunt Game loading onto the  
188 cooperation factor and the Stag Hunt Game with Punishment loading onto the punishment  
189 factor (Figures 1c and S2b; 41% variance explained).

190

In addition to the data-driven statistics reported above, the two-factor structure of

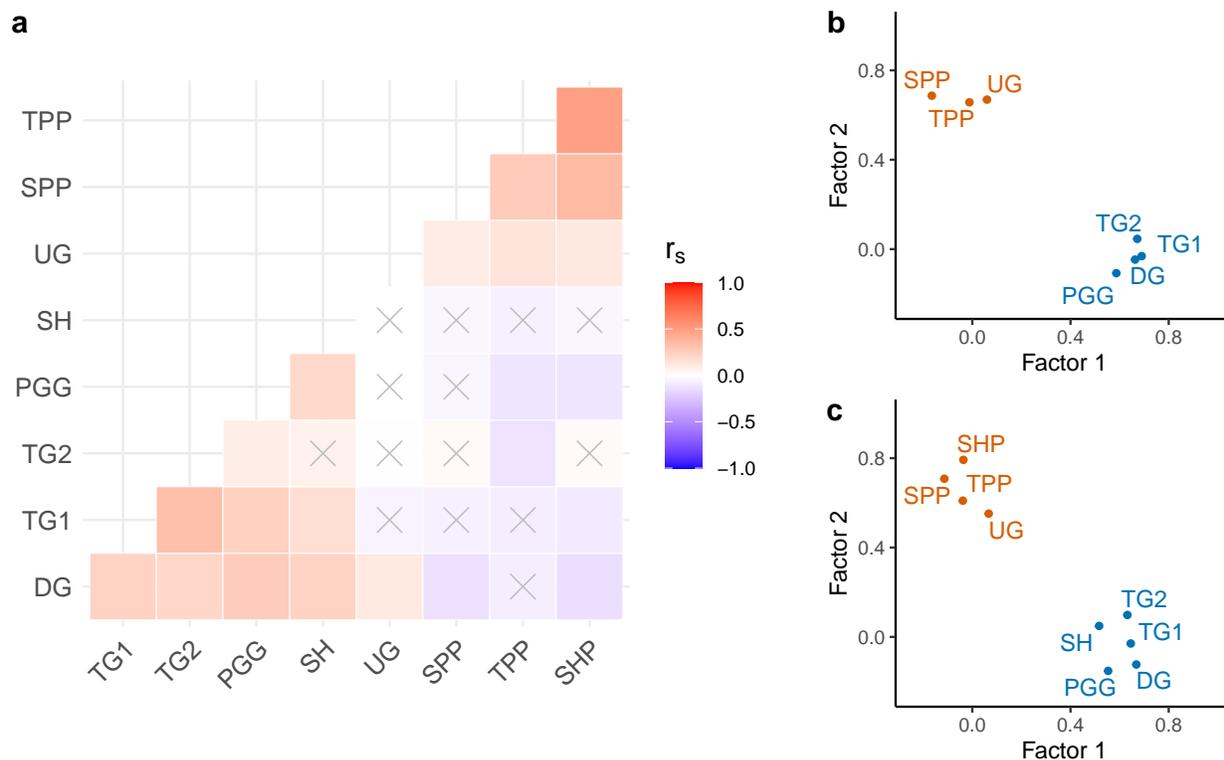
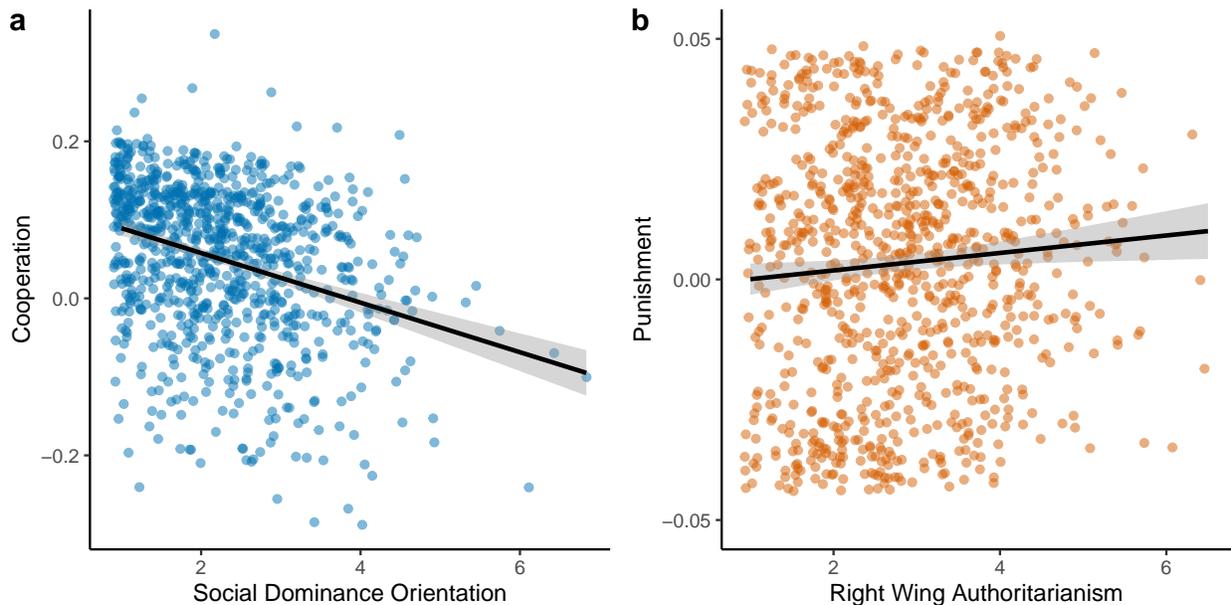


Figure 1. Factor structure of economic games. (a) Pairwise Spearman's rank correlations for game decisions. Grey crosses indicate non-significant correlations with Benjamin-Hochberg corrected  $p$ -values above .05. (b) Factor loadings from a principal components analysis with only the game decisions from previous work<sup>50</sup>. (c) Factor loadings from an extended principal components analysis including coordination games. DG = Dictator Game, TG1 = Trust Game (Give), TG2 = Trust Game (Return), PGG = Public Goods Game, SH = Stag Hunt Game, UG = Ultimatum Game (Minimum Acceptable Offer), TPP = Third-Party Punishment Game (Punish), SPP = Second-Party Punishment Game (Punish Defector), SHP = Stag Hunt Game with Punishment (Punish Defector).

191 cooperation and punishment was also supported by confirmatory factor analyses. As with  
192 our principal components analysis, we initially restricted our confirmatory factor model to  
193 include only the games used in previous work<sup>50</sup>. Controlling for game comprehension, we  
194 loaded the Dictator Game, Trust Game (Give), Trust Game (Return), and Public Goods  
195 Game onto the cooperation latent variable, and loaded the Ultimatum Game (Minimum  
196 Acceptable Offer), Third-Party Punishment Game (Punish), and Second-Party Punishment  
197 Game (Punish Defector) onto the punishment latent variable. According to established fit  
198 statistic cutoffs<sup>58,59</sup>, this model fitted the data well, with a Root Mean Square Error of  
199 Approximation (RMSEA) value of 0.06 and a Standardised Root Mean Square Residual  
200 (SRMR) value of 0.04. All indicators had significantly positive loadings, and the two latent  
201 variables were significantly negatively correlated with one another ( $r = -0.20$ ,  $p = .023$ ;  
202 Figure S3). Model fit improved further when additionally loading the Stag Hunt Game onto  
203 the cooperation latent variable and loading the Stag Hunt Game with Punishment (Punish  
204 Defector) onto the punishment latent variable (RMSEA = 0.04; SRMR = 0.04; Figure S4).

205 Having confirmed the existence of a two-factor structure underlying gameplay, we next  
206 fitted pre-registered structural equation models to test the relationships between the two  
207 latent variables and political ideology. We simultaneously regressed the cooperation and  
208 punishment latent variables onto mean scores of six Likert-scale items for SDO (Cronbach's  
209  $\alpha = 0.82$ ) and RWA ( $\alpha = 0.74$ ; see Methods for details on SDO and RWA items). In line  
210 with our second hypothesis, we found that SDO significantly negatively predicted the  
211 cooperation latent variable with a small-to-medium effect size (standardised  $\beta = -0.24$ ,  
212 unstandardised  $b = -0.03$ , 95% CI [-0.05, -0.02],  $p < .001$ , semi-partial  $r = 0.24$ ; Figure 2a).  
213 SDO also significantly positively predicted the punishment latent variable with a small effect  
214 size, which we did not hypothesise ( $\beta = 0.10$ ,  $b = 0.003$  [0.000, 0.006],  $p = .029$ ,  $r = 0.11$ ).  
215 In contrast with our third hypothesis, we found no significant relationship between RWA and  
216 the punishment latent variable ( $\beta = 0.04$ ,  $b = 0.001$  [-0.001, 0.003],  $p = .301$ ,  $r = 0.06$ ;  
217 Figure 2b). RWA was also unrelated to the cooperation latent variable ( $\beta = 0.00$ ,  $b = 0.00$

218  $[-0.01, 0.01]$ ,  $p = .955$ ,  $r = 0.02$ ). This pattern of results held when additionally controlling  
 219 for age, gender, ethnicity, education, and religiosity, though the relationship between SDO  
 220 and the punishment latent variable was attenuated when controlling for socio-economic  
 221 status and local deprivation (Table S1).



*Figure 2. Relationships between model-predicted latent variable scores for cooperation and punishment and the two dimensions of political ideology. (a) Social Dominance Orientation (mean score) is negatively related to the cooperation latent variable. (b) Right Wing Authoritarianism (mean score) is unrelated to the punishment latent variable. Lines are predictions from linear regressions, shaded areas are 95% confidence intervals.*

222 To ensure that our pattern of results generalised beyond SDO and RWA, we replicated  
 223 the above analyses with another widely-used measure of the two dimensions of political  
 224 ideology: Schwartz's values<sup>60</sup>. Schwartz argues that basic personal values vary along two  
 225 dimensions of altruism vs. self-enhancement and openness vs. conservation<sup>60</sup>. These reflect  
 226 the economic and social dimensions of political ideology, respectively<sup>4,8</sup>. When we replace  
 227 SDO and RWA with Schwartz's values, we find the same pattern of results. Controlling for  
 228 demographics, altruistic values significantly positively predicted the cooperation latent  
 229 variable ( $\beta = 0.14$ ,  $b = 0.02$   $[0.01, 0.04]$ ,  $p = .006$ ,  $r = 0.14$ ), while self-enhancement values

230 significantly negatively predicted the cooperation latent variable ( $\beta = -0.16$ ,  $b = -0.03$  [-0.05,  
231 -0.01],  $p = .003$ ,  $r = 0.15$ ). Altruistic and self-enhancement values were unrelated to the  
232 punishment latent variable. In contrast, neither openness values ( $\beta = 0.06$ ,  $b = 0.002$  [-0.001,  
233 0.006],  $p = .193$ ,  $r = 0.06$ ) nor conservation values ( $\beta = 0.04$ ,  $b = 0.003$  [-0.004, 0.009],  $p =$   
234  $.390$ ,  $r = 0.05$ ) predicted the punishment latent variable. Openness and conservation values  
235 were unrelated to the cooperation latent variable.

236 To further investigate the relationships between political ideology and gameplay, we  
237 conducted a series of exploratory linear regressions with mean SDO and RWA scores as  
238 outcome variables and individual game decisions as predictors, controlling for demographics  
239 and game comprehension. Strikingly, every cooperative decision was negatively associated  
240 with SDO (Figure 3). Holding all other variables constant, individuals were predicted to  
241 have higher SDO scores if they gave less in the Dictator Game, did not give in the Trust  
242 Game, returned less in the Trust Game, contributed less in the Public Goods Game, did not  
243 coordinate in the Stag Hunt Game, offered less in the Ultimatum Game, stole in the  
244 Third-Party Punishment Game, defected in the Prisoner's Dilemma Game, and did not  
245 coordinate in the Stag Hunt Game with Punishment. The effect sizes for these relationships  
246 were small, but consistent across all games (semi-partial  $r = 0.06$ -0.11). SDO was also  
247 positively related to some punishment decisions, including anti-social punishment  
248 (punishment of cooperators) in the Second-Party Punishment Game and Stag Hunt Game  
249 with Punishment, and punishment of defectors in the latter game. In contrast, fewer  
250 individual game decisions predicted RWA (Figure 4). Holding all other variables constant,  
251 individuals were predicted to have higher RWA scores if they anti-socially punished  
252 cooperators in the Second-Party Punishment Game and the Stag Hunt Game with  
253 Punishment, and if they stole in the Third-Party Punishment Game.

254 The cooperation latent variable also reliably predicted economically progressive views  
255 across a wide range of political issues (Figure 5). Controlling for the punishment latent  
256 variable, game comprehension, and demographics, exploratory structural equation modelling

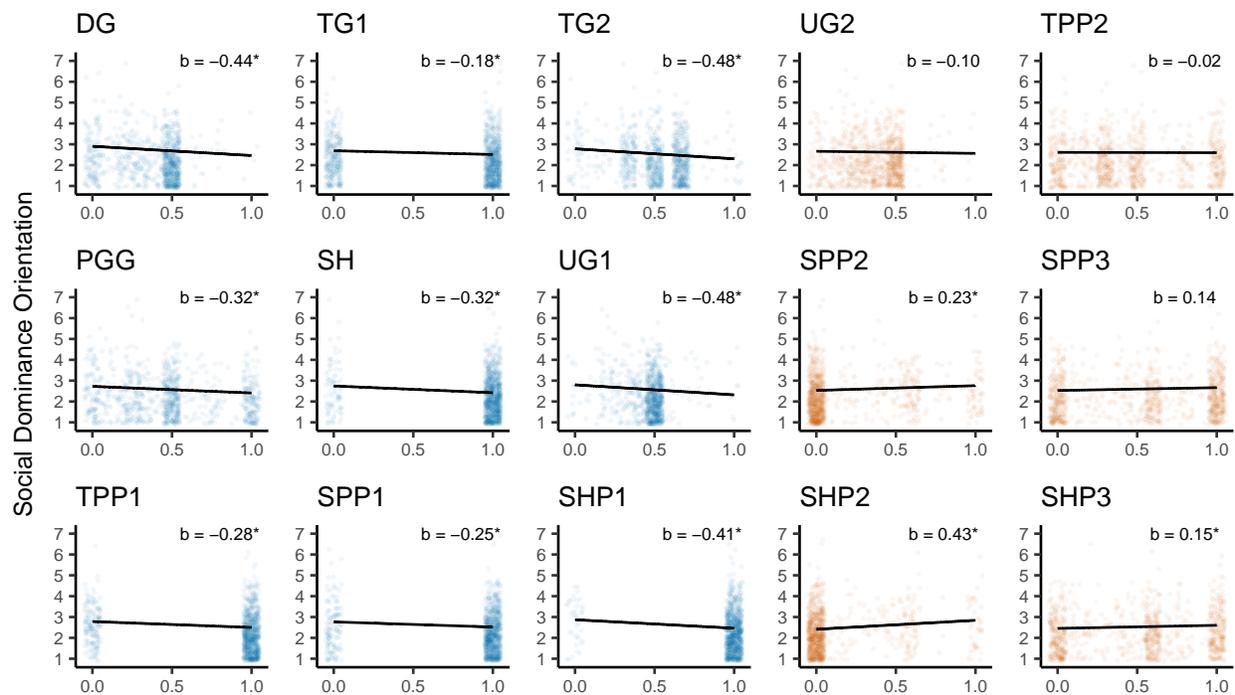


Figure 3. Individual game decisions predicting Social Dominance Orientation. Lines are predictions from linear regressions, controlling for Right Wing Authoritarianism, game comprehension, and demographics. Blue plots are cooperation decisions, orange plots are punishment decisions. Numbers are unstandardised coefficients. \* $p < 0.05$ . DG = Dictator Game, TG1 = Trust Game (Give), TG2 = Trust Game (Return), PGG = Public Goods Game, SH = Stag Hunt Game, UG1 = Ultimatum Game (Offer), TPP1 = Third-Party Punishment Game (Steal), SPP1 = Second-Party Punishment Game (Cooperate), SHP1 = Stag Hunt Game with Punishment (Coordinate), UG2 = Ultimatum Game (Minimum Acceptable Offer), TPP2 = Third-Party Punishment Game (Punish), SPP2 = Second-Party Punishment Game (Punish Cooperator), SPP3 = Second-Party Punishment Game (Punish Defector), SHP2 = Stag Hunt Game with Punishment (Punish Coordinator), SHP3 = Stag Hunt Game with Punishment (Punish Defector).

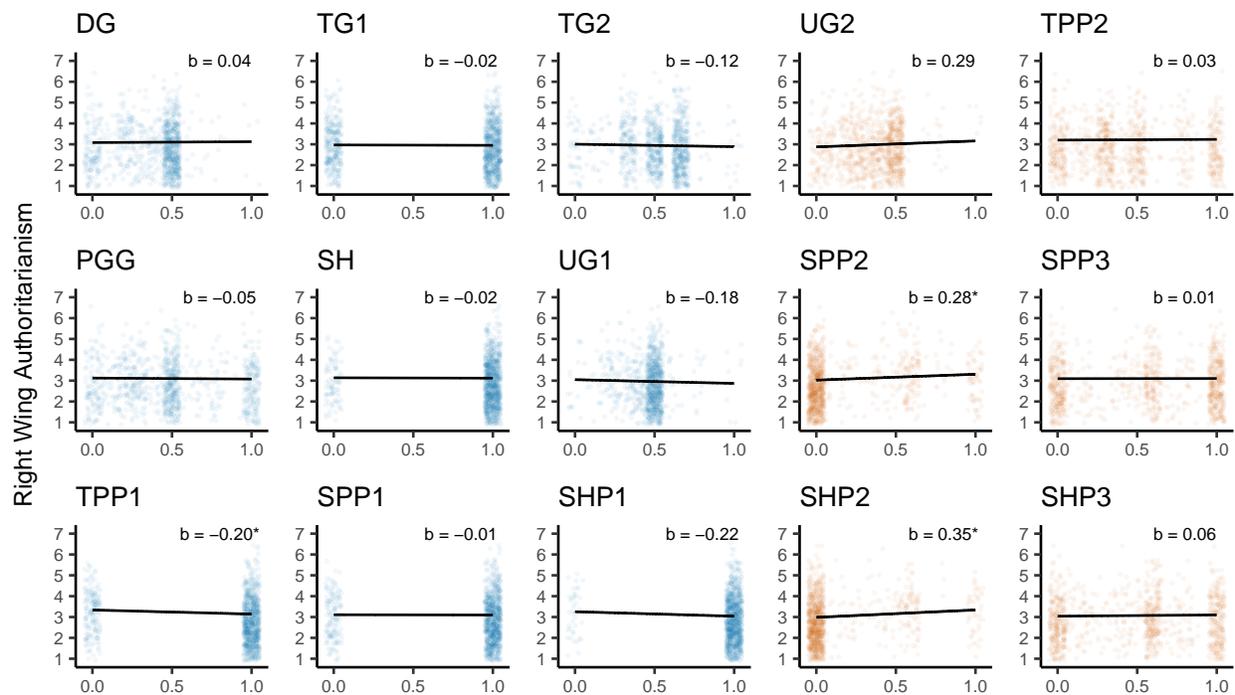


Figure 4. Individual game decisions predicting Right Wing Authoritarianism. Lines are predictions from linear regressions, controlling for Social Dominance Orientation, game comprehension, and demographics. Blue plots are cooperation decisions, orange plots are punishment decisions. Numbers are unstandardised coefficients. \* $p < 0.05$ . DG = Dictator Game, TG1 = Trust Game (Give), TG2 = Trust Game (Return), PGG = Public Goods Game, SH = Stag Hunt Game, UG1 = Ultimatum Game (Offer), TPP1 = Third-Party Punishment Game (Steal), SPP1 = Second-Party Punishment Game (Cooperate), SHP1 = Stag Hunt Game with Punishment (Coordinate), UG2 = Ultimatum Game (Minimum Acceptable Offer), TPP2 = Third-Party Punishment Game (Punish), SPP2 = Second-Party Punishment Game (Punish Cooperator), SPP3 = Second-Party Punishment Game (Punish Defector), SHP2 = Stag Hunt Game with Punishment (Punish Coordinator), SHP3 = Stag Hunt Game with Punishment (Punish Defector).

revealed that the cooperation latent variable positively predicted preferences for income redistribution ( $b = 1.07$  [0.41, 1.72],  $p = .001$ ,  $r = 0.11$ ), willingness to make sacrifices for the environment ( $b = 1.21$  [0.45, 1.96],  $p = .002$ ,  $r = 0.12$ ), endorsement of increased payments for those receiving Jobseeker Support ( $b = 1.36$  [0.60, 2.12],  $p < .001$ ,  $r = 0.15$ ), and endorsement of increased payments for those receiving Sole Parent Support ( $b = 1.42$  [0.66, 2.18],  $p < .001$ ,  $r = 0.15$ ). Cooperation was also negatively related to support for an economically conservative “flat tax” where everyone pays the same percentage of tax on their income regardless of their wealth ( $b = -0.97$  [-1.69, -0.24],  $p = .009$ ,  $r = 0.11$ ) and the belief that people would be less motivated to work hard if incomes were equal ( $b = -1.50$  [-2.17, -0.82],  $p < .001$ ,  $r = 0.16$ ). Though these effect sizes are small, the cooperation latent variable explains a comparable proportion of variance to other demographic variables (Figure S5). In contrast, the punishment latent variable did not predict socially conservative views like support for same-sex marriage ( $b = -0.16$  [-3.08, 2.75],  $p = .913$ ,  $r = 0.00$ ), support for euthanasia ( $b = 1.22$  [-1.56, 4.01],  $p = .388$ ,  $r = 0.04$ ), or support for abortion ( $b = -0.37$  [-3.24, 2.50],  $p = .800$ ,  $r = 0.00$ ).

Finally, the cooperation latent variable predicted political party support. In an exploratory Bayesian multinomial regression analysis, we used the model-predicted cooperation and punishment latent variable scores to predict reported support for the main four political parties in New Zealand: the New Zealand National Party, the New Zealand Labour Party, the Green Party of Aotearoa New Zealand, and the New Zealand First Party. Adding the cooperation latent variable to a null model with only demographic predictors improved model fit (difference in expected log pointwise predictive density = 4.18,  $SE = 1.32$ ). The effect of cooperation on political party support, controlling for demographics and the punishment latent variable, is visualised in Figure 6. Individuals higher in the cooperation latent variable were less likely to support National (New Zealand’s primary conservative party;  $b = -0.74$ , 95% credible interval [-1.59, 0.12]) and were more likely to support the Greens ( $b = 0.66$  [-0.20, 1.51]), although cooperation did not predict support for

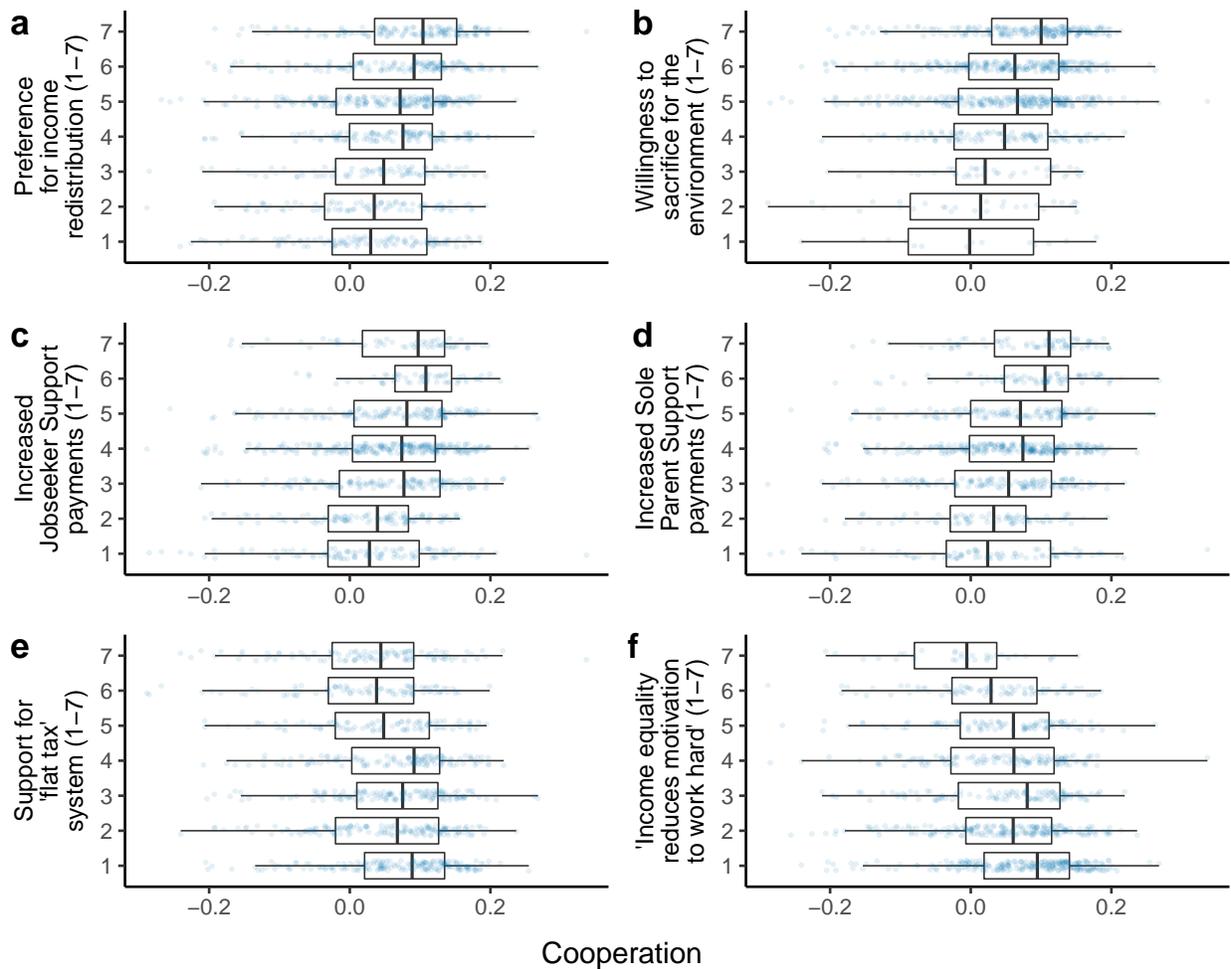


Figure 5. The cooperation latent variable predicts a host of economically progressive views. For every Likert scale, 7 indicates increased support for the policy or agreement with the statement, whereas 1 indicates reduced support for the policy or disagreement with the statement. Boxplots visualise the median, inter-quartile range, and range with outliers removed ( $\pm 1.5 \times$  inter-quartile range).

284 Labour or NZ First. In contrast, adding the punishment latent variable did not improve  
 285 model fit over the null model (difference in expected log pointwise predictive density = 0.03,  
 286  $SE = 0.12$ ).

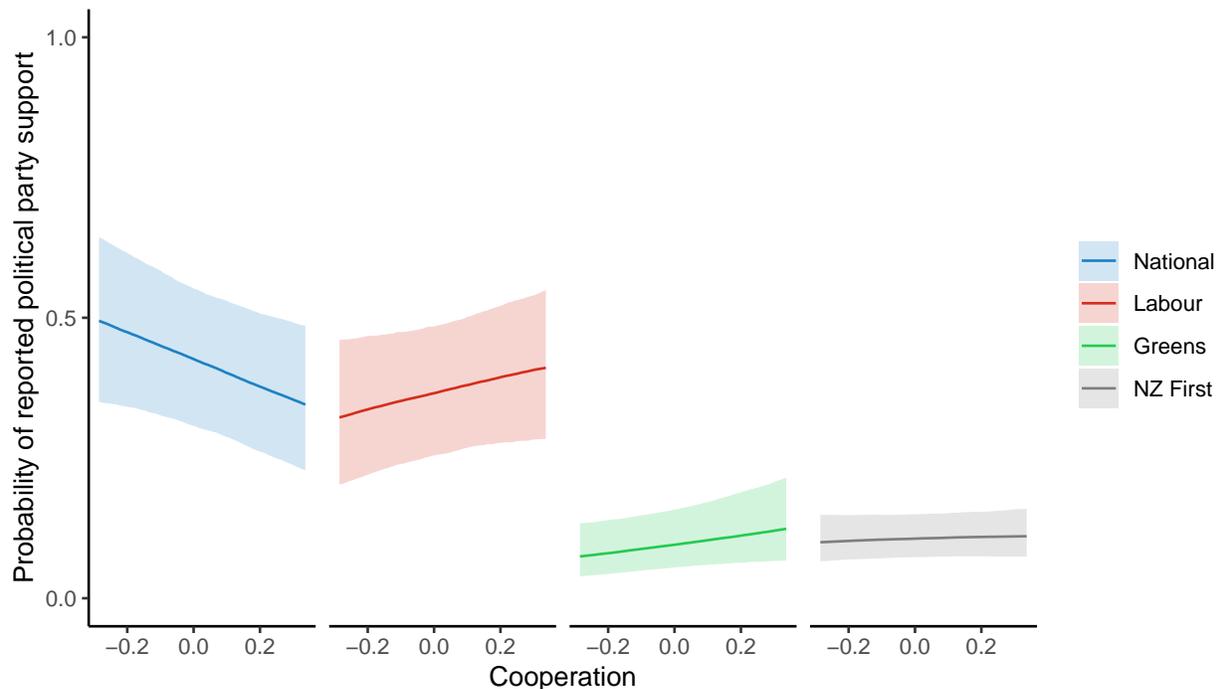


Figure 6. The cooperation latent variable predicts reported political party support. Individuals higher in the cooperation latent variable were less likely to support National and more likely to support the Greens. Lines are median posterior multinomial regression lines, shaded areas are 95% credible intervals.

287

## Discussion

288 The correlations, principal components analyses, and confirmatory factor analyses we  
 289 report show that behaviour across a large suite of economic games can be captured by two  
 290 underlying general preferences: cooperation and punishment. This finding replicates previous  
 291 work<sup>50,54</sup> and demonstrates that the cooperative phenotype generalises to an  
 292 English-speaking sample outside the United States and Europe. In addition, the unique  
 293 combination of substantial monetary stakes, real-time interactive gameplay, and additional

294 coordination games deployed in our study indicates that the cooperative phenotype reliably  
295 emerges with methodologies that may more reliably track real-world social interactions and  
296 with games that do not involve exploitation of others<sup>49</sup>.

297 As predicted by the dual evolutionary framework of political ideology<sup>4</sup>, we found that  
298 the cooperative phenotype captured by our economic games negatively covaried with two  
299 widely-used measures of economic conservatism: Social Dominance Orientation and  
300 Schwartz's altruistic vs. self-enhancement values. This builds upon previous studies  
301 identifying negative correlations between SDO and cooperative behaviour<sup>43,47-49</sup> and between  
302 altruistic values and cooperative behaviour<sup>61,62</sup>. The small-to-medium effect size for the  
303 relationship between SDO and the general cooperative preference (semi-partial  $r = 0.24$ ) is  
304 comparable to the effect size found in a recent meta-analysis of personality traits and  
305 economic game behaviour<sup>49</sup>. Our results suggest that previous correlations between measures  
306 of economic conservatism and gameplay have emerged because of an underlying relationship  
307 between economic conservatism and a general cooperative preference, rather than because of  
308 idiosyncratic features of particular conservatism measures or particular games.

309 The negative association between SDO and the cooperative phenotype suggests that  
310 the SDO self-report scale is capturing variation in interpersonal cooperation vs. competition.  
311 This claim is at odds with the popular conceptualisation of SDO as a measure of  
312 group-based (not interpersonal) dominance<sup>63-65</sup> and the explicit reference to groups in SDO  
313 scale items (see Methods). We acknowledge that SDO does predict group-based  
314 discrimination and prejudice<sup>8</sup>. However, SDO is also positively correlated with narcissism  
315 and Machiavellianism<sup>63,64,66</sup> and negatively correlated with agreeableness<sup>67</sup>, all of which are  
316 moderators of interpersonal social interactions<sup>66,68</sup>. Individuals higher in SDO also prefer to  
317 keep money for themselves rather than benefit the ingroup in order to harm a real, salient  
318 outgroup<sup>48</sup>, further supporting the claim that SDO reflects interpersonal rather than  
319 group-based dominance. Those studying SDO will need to adjust their conceptualisation of  
320 the measure to accommodate these findings.

321 When we separately explored individual economic games, we found negative  
322 correlations between SDO and cooperation for every cooperative decision (Figure 3). A  
323 recent study hypothesised that SDO should be specifically related to economic game  
324 behaviour that involves exploitation of others (i.e., increasing one's own outcome at another's  
325 cost)<sup>49</sup>. In this study, we have shown that SDO is indeed related to exploitative behaviour,  
326 such as returning less in the Trust Game, but it is also related to cooperative behaviour in  
327 games that do not involve a conflict of interest and so do not allow for exploitation of others,  
328 such as coordinating to mutual benefit in the Stag Hunt Game. This provides further  
329 evidence that economic conservatism is negatively associated with a general cooperative  
330 preference that applies across many different situations with different interaction payoff  
331 structures, rather than a narrow preference that applies in some situations but not others.

332 Further exploratory analyses revealed that the cooperative phenotype was related to  
333 economically progressive views across a wide range of policy issues, including  
334 environmentalism, welfare, taxation, and income redistribution. While the effect sizes for  
335 these associations were small, the cooperation latent variable explained a comparable  
336 proportion of variance to other demographic variables known to predict economically  
337 progressive policy views, such as age, gender, and education<sup>9,69</sup> (Figure S5), and the effects  
338 persisted when controlling for these variables. We do not believe that the associations  
339 between the cooperative phenotype and economically progressive views simply arose from  
340 people mapping their prior political beliefs onto the economic games, since the relationship  
341 held for games with different structures and the game descriptions were abstract with no  
342 political framing. Instead, we believe that variation in the cooperative phenotype predicts  
343 these particular policy views because they all reflect a willingness to cooperate in real-world  
344 large-scale social dilemmas. Taxation can be seen as cooperating in a public goods dilemma,  
345 where people pay individual costs that provide benefits to all (e.g., public infrastructure,  
346 education, healthcare). Making personal sacrifices to protect the environment can be seen as  
347 cooperating in a commons dilemma, where people forgo self-interest by refusing to deplete

348 common environmental resources. Under this view, all other things being equal, the higher  
349 the cooperative phenotype, the more willing people will be to cooperate in real-world  
350 large-scale social dilemmas like these and thus support economically progressive policies.

351 We also found that the cooperative phenotype predicted political party support.  
352 Specifically, individuals with higher levels of the cooperation latent variable were less likely  
353 to support the National Party and more likely to support the Green Party, a party that  
354 promotes environmental protection and progressive taxation. This contrasts with studies  
355 finding no association between gameplay and support for Republicans or Democrats in the  
356 United States<sup>45,46</sup>, but aligns with research showing greater cooperative preferences among  
357 supporters of left-leaning parties in Europe<sup>41-44</sup>. Previous research in New Zealand has  
358 identified key demographic predictors of political party support, such as age, gender,  
359 education, ethnicity, and socio-economic status<sup>70</sup>. For example, individuals in highly  
360 deprived areas of New Zealand are more likely to support Labour, Greens, and NZ First,  
361 presumably because these political parties best serve the economic interests of individuals  
362 low in socio-economic status. However, we have shown that the cooperation latent variable  
363 predicts party support even after controlling for all of these variables. This suggests that, in  
364 contrast to approaches that highlight self-interest as the main determinant of political  
365 attitudes<sup>71</sup>, the cooperative phenotype underlying the economic dimension of ideology  
366 predicts political party support independently of self-interest concerns.

367 We failed to find the predicted association between the punishment latent variable and  
368 RWA or between the punishment latent variable and Schwartz's openness vs. conservation  
369 values, nor did we find any link between the punishment latent variable and policy views or  
370 political party support. One explanation for this lack of relationship is that the punishment  
371 games initially used by Peysakhovich et al.<sup>50</sup> may not be capturing a general group  
372 conformist norm-enforcing preference as expected. Instead, it has been suggested that  
373 punishment in economic games could provide a competitive function, increasing individuals'  
374 relative payoffs over others<sup>72,73</sup>. In hindsight, there are several reasons to favour a

375 competitive-punishment account over a normative-punishment account in this study. First,  
376 this study used one-shot games with anonymous others in which there was no possibility for  
377 signalling, reputation building, or future behaviour modification, all supposedly essential  
378 factors behind the success of norm-enforcing punishment to enhance cooperation. Evidence  
379 suggests that one needs a relatively long horizon for norm-enforcing punishments to increase  
380 cooperation beyond what would prevail in the absence of such punishments<sup>74,75</sup>. In our  
381 one-shot anonymous games, then, punishment may be a last-ditch attempt to equalise  
382 payoffs rather than enforce norms to shape later behaviour. Second, we unexpectedly found  
383 a small positive relationship between SDO and the punishment latent variable, indicating  
384 that punishment in our games may be related to interpersonal competition rather than  
385 norm-enforcement. Third, an additional analysis revealed that, after controlling for  
386 demographics and the cooperation latent variable, the punishment latent variable was  
387 positively correlated with a view of the world as “a dog-eat-dog world where you have to be  
388 ruthless at times” ( $b = 2.97 [0.14, 5.79]$ ,  $p = .040$ ). This positive relationship between the  
389 punishment latent variable and a competitive worldview in our study suggests that  
390 punishment is reflecting a competitive motive, rather than a norm-enforcing motive, at least  
391 in this instance. Nevertheless, the competitive-punishment account is difficult to formally  
392 test against the normative-punishment account with our data as we did not empirically  
393 measure any norms<sup>76</sup> or allow for repeated interactions in our study.

394       There remain several open questions regarding the dual evolutionary framework of  
395 political ideology. First, if the punishment games from Peysakhovich et al.<sup>50</sup> were not  
396 sufficient to measure a general group conformist preference, are there more suitable  
397 behavioural measures for the social dimension of political ideology? Future studies could  
398 make use of repeated games with the opportunity for signalling and reputation building<sup>55</sup>,  
399 empirically-measured group norms<sup>76</sup>, or other measures of norm-adherence to evaluate the  
400 hypothesised general preference for group conformity. Second, while this study has provided  
401 a snapshot of general preferences and ideology at a particular point in time, the dual

402 evolutionary framework also hypothesises that changes in general preferences should predict  
403 subsequent changes in political views. For example, Universal Basic Income<sup>77</sup>, an  
404 economically progressive welfare program that aims to provide unconditional regular  
405 payments to citizens, has garnered increased support during the COVID-19 pandemic<sup>78</sup>. One  
406 explanation for this increased support is that people tend to cooperate more with one  
407 another during crises<sup>79</sup>, and this increased cooperative preference results in more  
408 economically progressive views. Future longitudinal studies will be required to formally test  
409 causal pathways like these. Third, it remains unclear to what extent our findings apply to  
410 non-WEIRD<sup>80</sup> and small-scale societies. If the two dimensions of political ideology indeed  
411 emerge from general cooperative and group conformist preferences that were selected in  
412 ancestral humans, then the relationship between general preferences and political ideology  
413 should be cross-culturally universal. Future anthropological studies should employ similar  
414 behavioural measures in small-scale societies to predict locally-relevant political attitudes.

415 In sum, the dual evolutionary framework of political ideology predicts that differences  
416 in economic and social conservatism are produced by underlying variation in general  
417 preferences for cooperation and group conformity. In support of this framework, we have  
418 shown that a general preference for cooperation can be estimated from a suite of behavioural  
419 economic games, and this cooperative phenotype is related to economically progressive  
420 attitudes, policy views, and political party support.

## 421 **Methods**

### 422 **Participants**

423 **Power analysis.** The minimum size of our sample was determined by conducting a  
424 power analysis on previous data<sup>50</sup>. We used the smallest significant correlation between  
425 economic game play and real-world cooperation in that study as the effect size ( $r = 0.15$ ).  
426 G\*Power<sup>81</sup> suggested a sample of 571 participants to detect this correlation effect size with

427 statistical power of 0.95. This sample size is also sufficient to conduct exploratory factor  
428 analysis<sup>59</sup>. We aimed to sample 1000 participants, substantially above this amount, and  
429 eventually report data from 926 participants.

430 **Participants and sampling.** Participants for this study were sampled from the  
431 ongoing longitudinal New Zealand Attitudes and Values Study<sup>57</sup>. This participant pool is a  
432 representative sample of the New Zealand population, contacted through random draws from  
433 the New Zealand electoral roll. We included participants in our sample frame who: had  
434 completed Wave 4 of the study ( $n = 12,189$ ); had also completed Wave 9 and/or Wave 10 ( $n$   
435  $= 8,095$ ); had not subsequently withdrawn from the study at the time of sampling ( $n =$   
436  $7,833$ ); had consistently indicated at Wave 9 and 10 that they would be willing to participate  
437 in an additional online study ( $n = 4,181$ ); had a valid email address ( $n = 4,040$ ); were living  
438 in New Zealand ( $n = 3,955$ ); were younger than 70 at the time of sampling ( $n = 3,374$ ); and  
439 had a valid cell or landline number ( $n = 3,345$ ). Of these 3,345 participants, we attempted  
440 to contact 3,063 about an additional study involving “economic decision-making in groups”.  
441 We managed to successfully contact 2,731 about the study.

442 Participants were contacted initially by phone and then, if they agreed to participate,  
443 over email in the days leading up to their allocated study session. 1,805 participants either  
444 dropped out of the study entirely (before or during the session), spent less than 10 minutes  
445 or more than 50 minutes on the games, or did not have SDO or RWA scores for Wave 9.  
446 This resulted in a final sample of 926 participants (630 females; age  $M = 50.5$  years,  $SD =$   
447  $12$  years; see Figure S1 for further sample characteristics).

## 448 **Materials**

449 **New Zealand Attitudes and Values Survey measures.** Main variables were  
450 taken from Wave 9 of the New Zealand Attitudes and Values Study<sup>57</sup>. These included SDO,  
451 RWA, age, gender, ethnicity, education level, socio-economic status, local deprivation,  
452 religiosity, and reported political party support. SDO and RWA scores were both mean

453 averages of six separate Likert-scale items (1 = strongly disagree, 7 = strongly agree). For  
454 SDO, the individual items<sup>65</sup> were: “It is OK if some groups have more of a chance in life  
455 than others”; “Inferior groups should stay in their place”; “To get ahead in life, it is  
456 sometimes okay to step on other groups”; “We should have increased social equality”  
457 (reversed); “It would be good if groups could be equal” (reversed); and “We should do what  
458 we can to equalise conditions for different groups” (reversed). For RWA, the individual  
459 items<sup>82</sup> were: “It is always better to trust the judgment of the proper authorities in  
460 government and religion than to listen to the noisy rabble-rousers in our society who are  
461 trying to create doubt in people’s minds”; “It would be best for everyone if the proper  
462 authorities censored magazines so that people could not get their hands on trashy and  
463 disgusting material”; “Our country will be destroyed some day if we do not smash the  
464 perversions eating away at our moral fibre and traditional beliefs”; “People should pay less  
465 attention to The Bible and other old traditional forms of religious guidance, and instead  
466 develop their own personal standards of what is moral and immoral” (reversed); “Atheists  
467 and others who have rebelled against established religions are no doubt every bit as good  
468 and virtuous as those who attend church regularly” (reversed); and “Some of the best people  
469 in our country are those who are challenging our government, criticizing religion, and  
470 ignoring the ‘normal way’ things are supposed to be done” (reversed). Additional items on  
471 Schwartz’s values, policy views, and competitive worldview were taken from Waves 5, 6, and  
472 8 of the New Zealand Attitudes and Values Study.

473 **Battery of economic games.** All eight economic games were conducted online  
474 using oTree software<sup>83</sup>. In an attempt to replicate previous work, these games are largely  
475 identical to those from Peysakhovich et al.<sup>50</sup>. All games involved one-shot decisions between  
476 multiple players. The strategy method was used to elicit responses in all possible roles.  
477 Participants played for points, which were converted to New Zealand dollars (1 point =  
478 \$0.035). Code for these games can be found at <https://osf.io/dwx8g/>.

479 We included three cooperation games in which individuals decided whether to pay a

480 personal cost to benefit another player:

- 481 • *Dictator Game.* Player A is given 100 points. They must decide how many of these  
482 points to transfer to Player B. Player A keeps the remaining points. Player B is passive  
483 in the interaction.
- 484 • *Trust Game.* Players A and B both start with 50 points. First, Player A decides  
485 whether or not to transfer all 50 points to Player B, in the knowledge that the  
486 transferred amount will be tripled to 150 points. If Player A transfers, Player B now  
487 has 200 points. Player B must then decide to transfer 0 - 150 points back to Player A.
- 488 • *Public Goods Game.* Four players begin with 100 points each. They can contribute 0 -  
489 100 points into a shared group project. All four decisions are made simultaneously, and  
490 then the amount in the group project is doubled and distributed evenly between all  
491 four players. Each player ends the game with their share from the group project, plus  
492 the points they initially refrained from contributing.

493 We also included three punishment games, in which individuals decide whether to pay  
494 a personal cost to punish another player for their decisions:

- 495 • *Ultimatum Game.* Player A starts with 100 points, and Player B starts with nothing.  
496 Player A must decide how many points to transfer to Player B. However, Player B  
497 simultaneously specifies their “minimum acceptable offer”: namely, the lowest transfer  
498 from Player A that they will accept. If Player A’s transfer amount is lower than this  
499 minimum acceptable offer, both players end the game with 0 points. Otherwise, Player  
500 B receives the transfer amount, and Player A keeps the remaining points.
- 501 • *Third-Party Punishment Game.* Players A, B, and C all start with 100 points. First,  
502 Player A decides whether to “take” from Player B. If Player A takes, Player B loses 50  
503 points and Player A gains 30 points (taking is inefficient). If Player A takes, Player C  
504 can then pay 0 - 20 points to remove points from Player A. Each paid point removes 5  
505 points from Player A. Player B is passive in the interaction.

- 506 • *Second-Party Punishment Game.* Players A and B start with 100 points. This game  
507 has two stages: the transfer stage, and the penalty stage. In the transfer stage, each  
508 player decides whether to transfer 30 points to the other player. Any transferred points  
509 are doubled before the other player receives them. Decisions are made simultaneously.  
510 The transfer stage follows the payoff matrix of a Prisoner's Dilemma. Then, in the  
511 penalty phase, both players can pay 0 - 10 points to remove points from the other  
512 player, depending on their decision in the transfer stage. Each paid point removes 5  
513 points from the other player.

514 We replaced the destructive All-Pay Auction Game used in previous work<sup>50</sup> with two  
515 coordination games: a Stag-Hunt Game, and a Stag-Hunt Game with Punishment. These  
516 games are kept as identical to the Public Goods Game and Second-Party Punishment Game  
517 (respectively) as possible:

- 518 • *Stag-Hunt Game.* Four players begin with 50 points each. They can either contribute  
519 30 points into a shared group project or contribute nothing. Any points in the group  
520 project will be doubled and distributed evenly between all players, *but only if all*  
521 *players contribute.* Otherwise, the points in the group project will be lost. Decisions  
522 are made simultaneously. Each player ends the game with their share from the group  
523 project, plus the points they initially refrained from contributing.
- 524 • *Stag-Hunt Game with Punishment.* Players A and B start with 100 points. This game  
525 has two stages: the transfer stage, and the penalty stage. In the transfer stage, each  
526 player decides whether to transfer 30 points to a group project. Any points in the  
527 group project will be doubled and distributed evenly between both players, *but only if*  
528 *both players contribute.* Otherwise, the points in the group project will be lost.  
529 Decisions are made simultaneously. The transfer stage follows the payoff matrix of a  
530 Stag Hunt. Then, in the penalty phase, both players can pay 0 - 10 points to remove  
531 points from the other player, depending on their decision in the transfer stage. Each

532           paid point removes 5 points from the other player.

### 533 **Procedure**

534           Data collection for the study was conducted between 18th February and 25th July 2019  
535 with weekly staggered recruitment. After expressing interest in an initial phone call,  
536 participants were emailed further information about the study. In the email, they were asked  
537 to complete a Qualtrics survey to specify their availability the following week. This survey  
538 removed any ineligible participants who (1) did not have adequate access to the Internet, (2)  
539 did not have a quiet place to participate in the study, and (3) did not have a New Zealand  
540 bank account (for payment purposes).

541           Study sessions contained between 14 and 97 participants, and were conducted on  
542 midweek evenings (between 6pm and 8pm). Participants knew that they were playing with  
543 others recruited from the New Zealand Attitudes and Values Study, but were not aware of  
544 how many people were present in any particular session. At the time of their session,  
545 participants were emailed with a session-wide oTree link. On the first screen, they were  
546 required to enter their unique participant label. This was followed by a consent form, where  
547 they were informed of their confidentiality, right to withdraw, payment, and ethical approval.  
548 If they agreed to the consent form, participants then read information about the economic  
549 games. They learned that they would be playing for points that would be later converted to  
550 currency, that they would be matched in real-time with other participants from the New  
551 Zealand Attitudes and Values Study, and that every task would count towards their final  
552 bonus payment.

553           The eight economic games were then presented in a random order. For each game,  
554 participants read the instructions and answered a comprehension question about the game.  
555 Participants then provided responses for all possible roles in the game, following the strategy  
556 method. After all the games, participants entered a waiting lobby where they waited for  
557 other participants in the session to complete the games. Once all participants in the session

558 arrived at the waiting lobby, the software calculated payoffs for each game by randomly  
559 matching individuals within a session. The accumulated payoffs across all eight games  
560 determined the final overall payoff. Once matching was complete, participants saw a payoff  
561 screen which summarised the results of each individual game and how much they had earned  
562 from the study.

563         Session sizes were often not strict multiples of four necessary for random matching,  
564 either due to unforeseen dropouts or uneven sign-ups. To remedy this, simulated players were  
565 used to round up uneven session sizes. These simulated players chose responses in the games  
566 based on the median responses from previous research<sup>50</sup>. Participants were made aware that  
567 they may be matched with simulated players, but only on the payoff screen *after* all game  
568 decisions had been made: “In the rare event that we could not find a participant to match  
569 you with, we have instead matched you with average decisions based on previous research.”

570         Participants were paid a fixed \$20 NZD show-up fee, plus a bonus payment of between  
571 \$10 - 35 ( $M = \$25.17$ ,  $SD = \$2.47$ ) depending on the decisions of themselves and others. In  
572 total, we spent \$41,826 on participant reimbursements. Name and bank account details were  
573 collected at the end of the study. For security purposes, this information was encrypted  
574 while stored online and later decrypted on a local computer for payment. After providing  
575 this information, participants specified if they would like to be contacted again about similar  
576 studies and then re-entered their unique participant label.

577         On average, participants took 22 minutes to complete the eight economic games ( $SD =$   
578  $7.39$  mins, range = 6 - 47 mins). Participants had a 55 minute limit to complete the games.  
579 As the study involved real-time matching between participants, participants could not pause  
580 their completion of the questionnaire and return to it later, as other participants would be  
581 left waiting to match with them. If participants took longer than 55 minutes, the software  
582 skipped them ahead to the waiting lobby and treated them as if they were simulated players  
583 (i.e., submitting median responses from past research). Timeouts were still paid the \$20

584 show-up fee, but no bonus payment.

## 585 **Statistical analysis**

586 Our pre-registered analyses proceeded in two stages (<https://osf.io/dwx8g/>). In the  
587 first stage, we focused on the data from the economic games. Following previous research<sup>50</sup>,  
588 we limited our analyses to game decisions which were not followed by a subsequent  
589 punishment stage: so-called “uncoerced” decisions. These decisions were scaled to vary  
590 between 0 and 1 for all analyses. We used a comprehensive set of statistical tests  
591 (correlations, principal components analyses, and confirmatory factor analyses) to determine  
592 the factor structure of variation in these decisions. We expected that two factors would  
593 emerge from and be supported by the data: cooperation and punishment. In the second  
594 stage, we used structural equation modelling to determine the extent to which these  
595 cooperation and punishment latent variables could be predicted by SDO and RWA,  
596 respectively. We also conducted several additional exploratory analyses that were not  
597 pre-registered: (1) structural equation models predicting the cooperation and punishment  
598 latent variables from Schwartz’s values, (2) regressions predicting SDO and RWA from  
599 individual game decisions, (3) structural equation models predicting policy views from the  
600 cooperation and punishment latent variables, and (4) Bayesian multinomial regressions  
601 predicting reported political party support from the cooperation and punishment latent  
602 variables. For Bayesian models, we used normally distributed priors ( $M = 0$ ,  $SD = 0.5$ ) for  
603 all parameters, and all models converged normally ( $\hat{R} = 1$ ).

604 All analyses were conducted in R Version 3.6.2<sup>84</sup>. We used the *psych*<sup>85</sup> package for  
605 correlations and principal components analyses, the *lavaan*<sup>86</sup> package for confirmatory factor  
606 analyses and structural equation modelling, and the *brms*<sup>87</sup> package for Bayesian  
607 multinomial regressions. Figures were created with the *ggplot2*<sup>88</sup> and *cowplot*<sup>89</sup> packages,  
608 and the manuscript was generated with the *drake*<sup>90</sup> and *papaja*<sup>91</sup> packages.

### Competing Interests

609

610 We declare that none of the authors have competing financial or non-financial interests.

### Data Availability

611

612 A copy of the anonymous data reported in each New Zealand Attitudes and Values  
613 Study publication is available from Professor Chris Sibley (c.sibley@auckland.ac.nz) upon  
614 request from appropriately qualified researchers. Such data will be provided with the explicit  
615 understanding that it is used solely for the purposes of replicating or otherwise checking the  
616 validity of analyses reported in scientific papers analysing New Zealand Attitudes and Values  
617 Study data.

### Code Availability

618

619 Python code for the economic games and R code for the statistical analyses, figures,  
620 and manuscript generation are publicly available at <https://osf.io/dwx8g/>.

### Ethics Statement

621

622 Ethical approval for this study was granted by the University of Auckland Human  
623 Participants Ethics Committee (ref: 021666).

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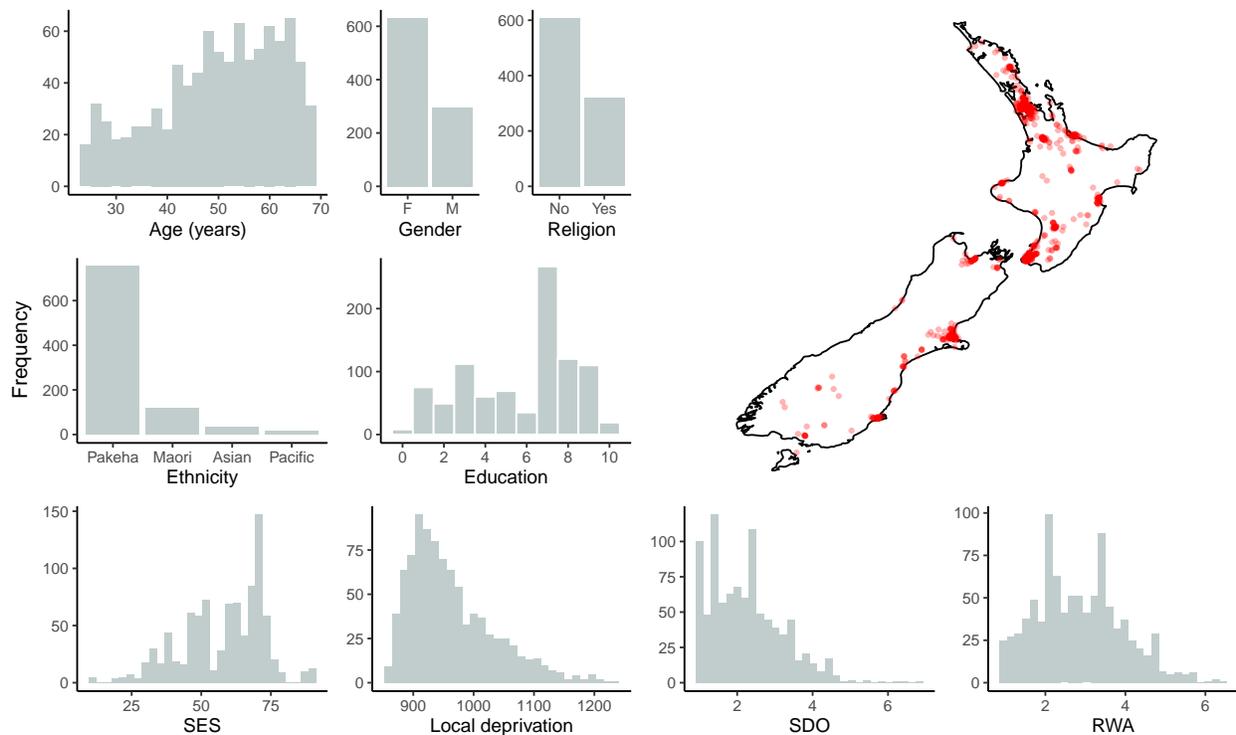
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## Supplementary Materials



*Figure S1. Final sample characteristics (n = 926).* Education is measured by NZREG, an ordinal scale measuring level of educational attainment (range = 0-10). Socio-economic status (SES) is measured by NZSEI, an ordinal rank measure based on occupational categories (range = 10-90). Local deprivation is measured by NZDEP, an ordinal rank measure of local deprivation for each meshblock in New Zealand (range = 840-1550). Social Dominance Orientation (SDO) and Right Wing Authoritarianism (RWA) are mean scores from six Likert scales (range = 1-7). The map of New Zealand shows meshblock locations for participants.

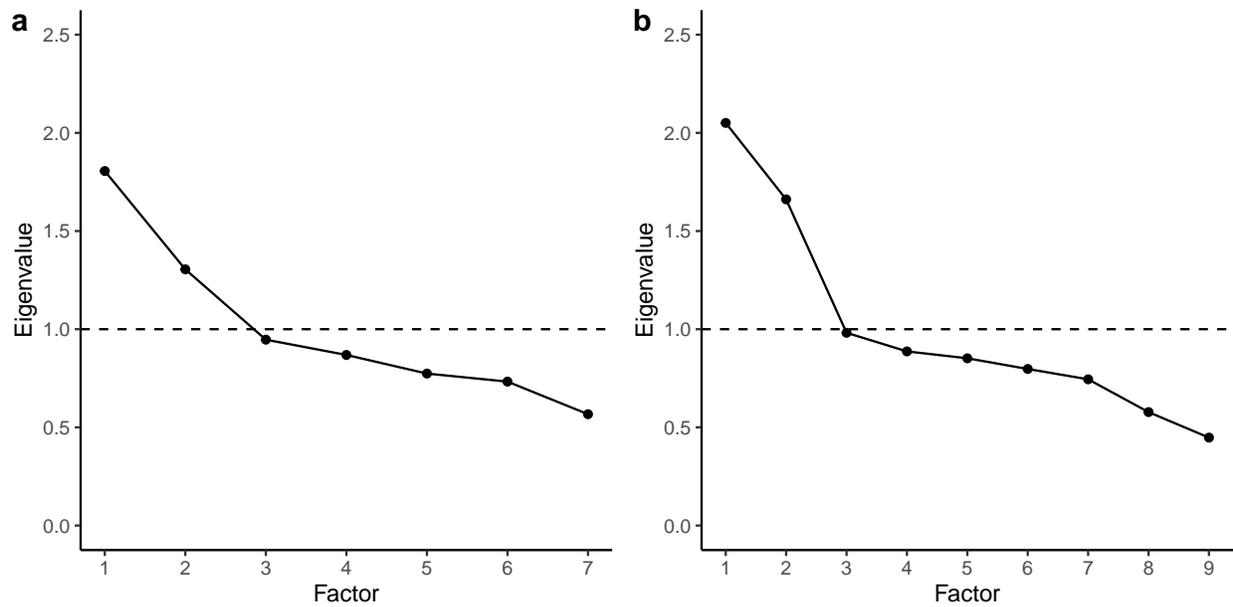
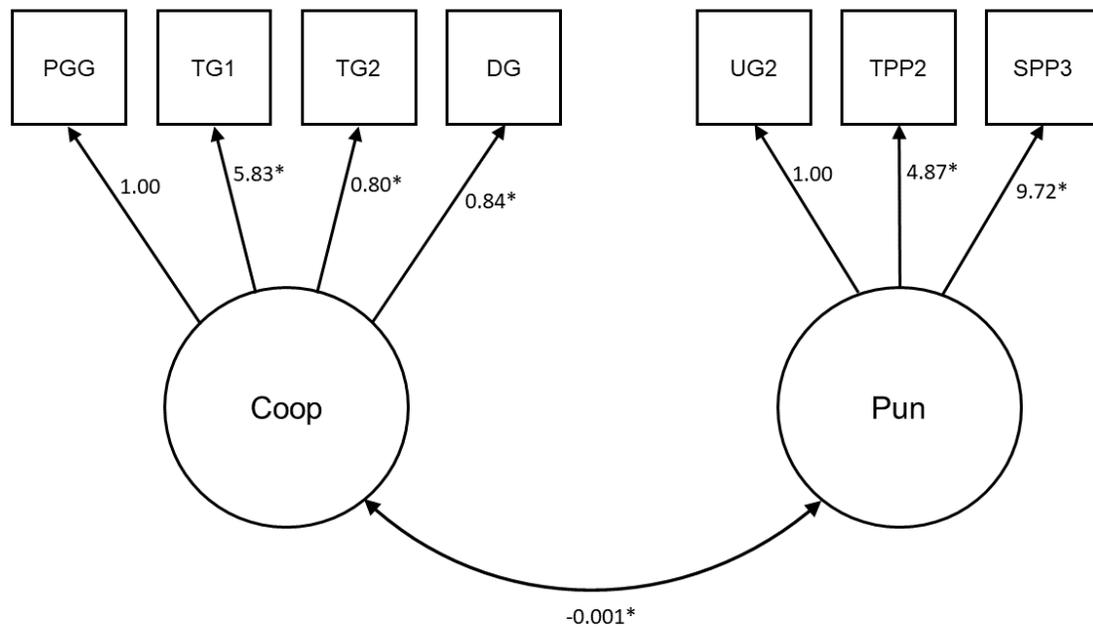
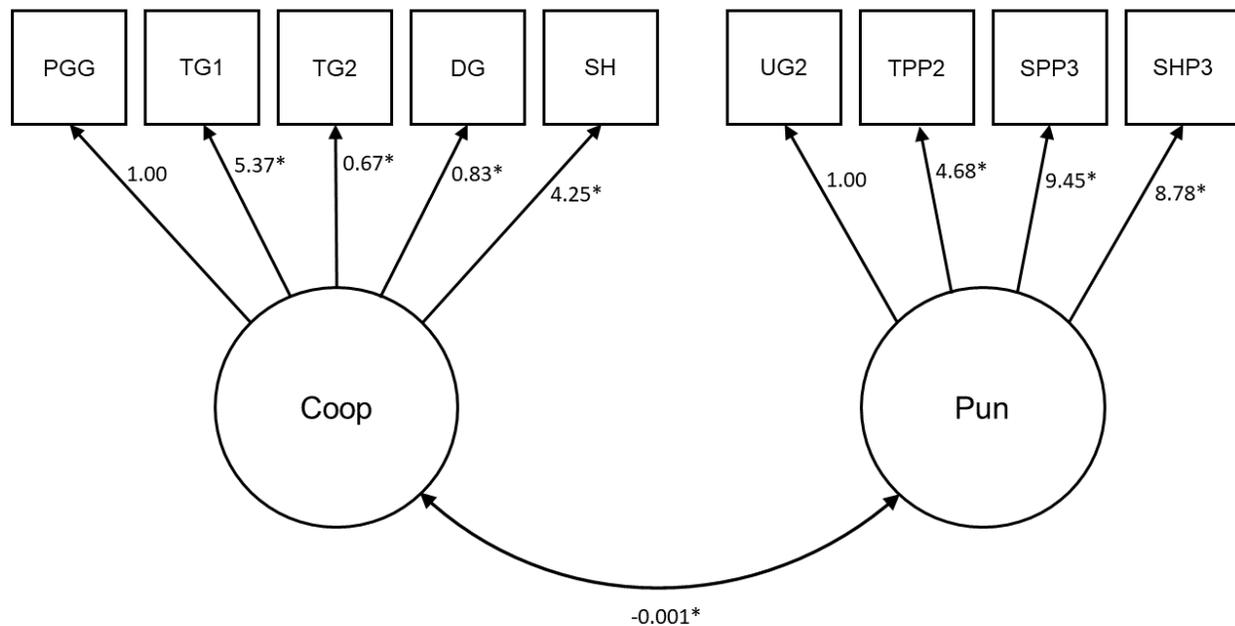


Figure S2. Scree plots for (a) the principal components analysis with only the game decisions from previous work and (b) the extended principal components analysis including coordination games. Principal components analyses initially contained as many factors as there were variables. Both scree plots show that only the first two factors have eigenvalues above 1, indicating a two-factor solution.



*Figure S3. Confirmatory factor model using only the game decisions from previous work.* TG1 is treated as a binary endogenous variable, and paths for PGG and UG2 are constrained to 1. Note that the regression paths controlling for game comprehension are not included in this figure. Numbers are unstandardised coefficients. \* $p < 0.05$ . PGG = Public Goods Game, TG1 = Trust Game (Give), TG2 = Trust Game (Return), DG = Dictator Game, UG2 = Ultimatum Game (Minimum Acceptable Offer), TPP2 = Third-Party Punishment Game (Punish), SPP3 = Second-Party Punishment Game (Punish Defector).



*Figure S4. Confirmatory factor model with additional coordination games. TG1 and SH are treated as binary endogenous variables, and paths for PGG and UG2 are constrained to 1. Note that the regression paths controlling for game comprehension are not included in this figure. Numbers are unstandardised coefficients.  $*p < 0.05$ . PGG = Public Goods Game, TG1 = Trust Game (Give), TG2 = Trust Game (Return), DG = Dictator Game, SH = Stag Hunt Game, UG2 = Ultimatum Game (Minimum Acceptable Offer), TPP2 = Third-Party Punishment Game (Punish), SPP3 = Second-Party Punishment Game (Punish Defector), SHP3 = Stag Hunt Game with Punishment (Punish Defector).*

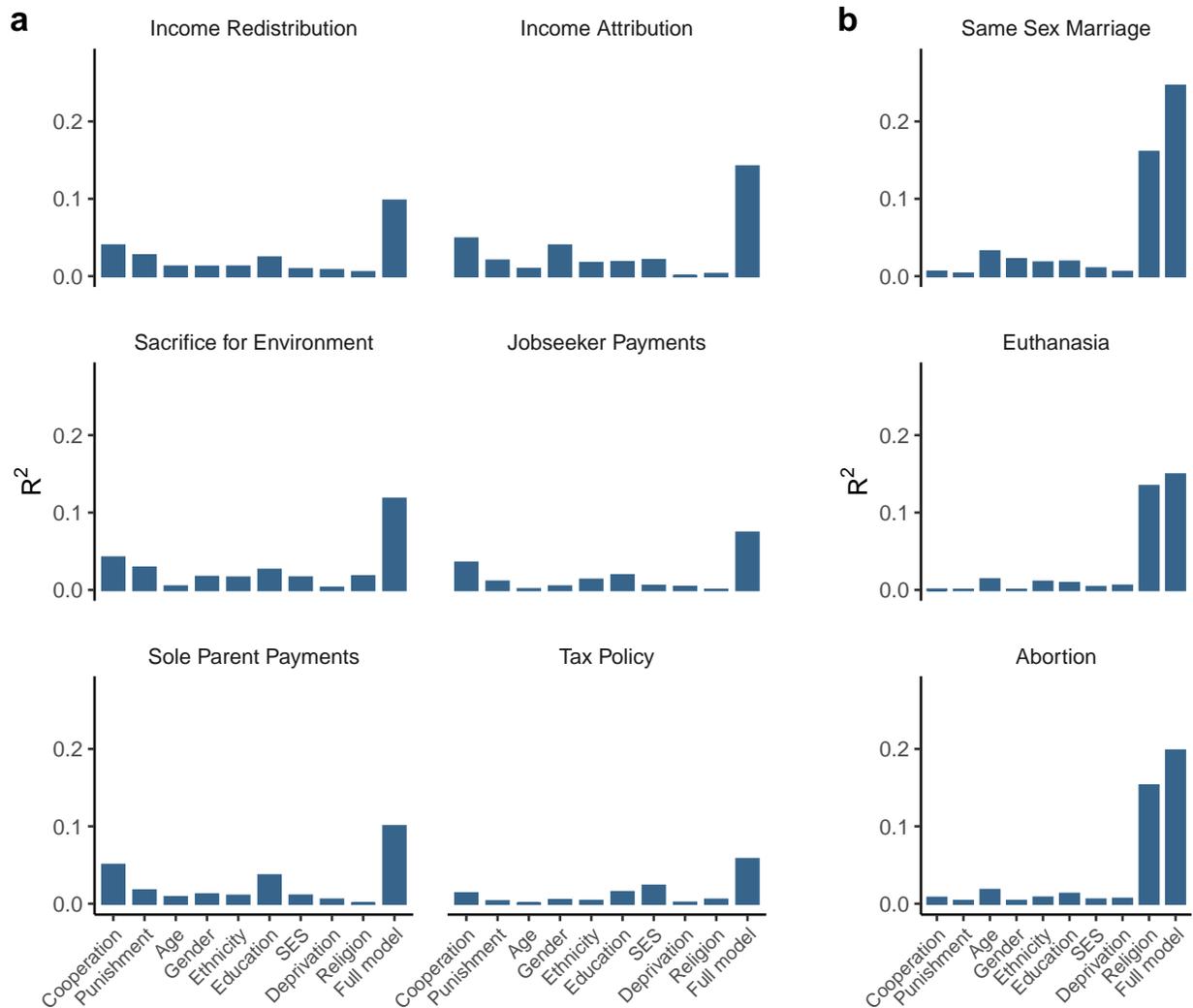


Figure S5. Bar plots for (a) economic and (b) social policy views showing the proportion of variance explained ( $R^2$ ) by the cooperation latent variable, the punishment latent variable, and various demographic controls individually, and their combination in a full model. The cooperation latent variable explains a comparable proportion of variance in policy views to other demographics.

Table S1

*Summary of the regression coefficients predicting the cooperation and punishment latent variables in the full structural equation model.*  
*Reference ethnicity category is Pakeha.*

DV	IV	Estimate	SE	z	p
Cooperation	SDO	-0.029	0.007	-4.453	0.000
	RWA	-0.001	0.006	-0.216	0.829
	Age	-0.001	0.001	-2.312	0.021
	Gender (Male = 1)	0.006	0.012	0.479	0.632
	Ethnicity (Maori)	-0.006	0.016	-0.348	0.728
	Ethnicity (Pacific)	-0.014	0.060	-0.239	0.811
	Ethnicity (Asian)	-0.076	0.029	-2.654	0.008
	Education	0.002	0.002	0.903	0.367
	SES	-0.002	0.004	-0.581	0.561
	Deprivation	0.018	0.008	2.355	0.019
	Religious (Yes = 1)	0.013	0.013	1.019	0.308
	Punishment	SDO	0.003	0.002	1.767
RWA		0.001	0.001	0.847	0.397
Age		0.000	0.000	1.457	0.145
Gender (Male = 1)		0.000	0.003	-0.111	0.912
Ethnicity (Maori)		-0.001	0.004	-0.225	0.822
Ethnicity (Pacific)		-0.008	0.011	-0.804	0.422
Ethnicity (Asian)		-0.006	0.008	-0.810	0.418
Education		-0.001	0.001	-0.858	0.391
SES		0.000	0.001	-0.390	0.697
Deprivation		0.000	0.002	-0.058	0.954
Religious (Yes = 1)		0.001	0.003	0.369	0.712

Table S2

*Summary of the regression coefficients predicting preference for income redistribution (1-7) in the full structural equation model. Full item: “Redistributing money and wealth more evenly among a larger percentage of the people in New Zealand through heavy taxes on the rich.” Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	1.066	0.335	3.180	0.001
Punishment	-2.890	1.506	-1.919	0.055
Age	-0.005	0.003	-1.604	0.109
Gender (Male = 1)	-0.204	0.076	-2.671	0.008
Ethnicity (Maori)	0.325	0.102	3.187	0.001
Ethnicity (Pacific)	0.163	0.322	0.505	0.614
Ethnicity (Asian)	0.085	0.192	0.444	0.657
Education	0.055	0.015	3.673	0.000
SES	0.031	0.026	1.188	0.235
Deprivation	0.111	0.057	1.958	0.050
Religious (Yes = 1)	-0.183	0.076	-2.399	0.016

Table S3

*Summary of the regression coefficients predicting income attribution (1-7) in the full structural equation model.*

*Full item: "Income equality reduces motivation to work hard." Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	-1.497	0.344	-4.359	0.000
Punishment	2.034	1.427	1.425	0.154
Age	-0.012	0.003	-3.666	0.000
Gender (Male = 1)	0.462	0.077	5.966	0.000
Ethnicity (Maori)	-0.103	0.108	-0.959	0.338
Ethnicity (Pacific)	-0.377	0.336	-1.121	0.262
Ethnicity (Asian)	0.574	0.175	3.288	0.001
Education	-0.042	0.015	-2.815	0.005
SES	-0.067	0.026	-2.558	0.011
Deprivation	-0.033	0.053	-0.635	0.525
Religious (Yes = 1)	0.165	0.075	2.204	0.028

Table S4

*Summary of the regression coefficients predicting willingness to sacrifice for the environment (1-7) in the full structural equation model. Full item: “Are you willing to make sacrifices to your standard of living (e.g., accept higher prices, drive less, conserve energy) in order to protect the environment?” Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	1.205	0.383	3.145	0.002
Punishment	-2.911	1.465	-1.987	0.047
Age	-0.002	0.003	-0.759	0.448
Gender (Male = 1)	-0.304	0.077	-3.950	0.000
Ethnicity (Maori)	0.266	0.110	2.415	0.016
Ethnicity (Pacific)	-0.407	0.294	-1.385	0.166
Ethnicity (Asian)	-0.331	0.204	-1.625	0.104
Education	0.053	0.015	3.500	0.000
SES	0.046	0.026	1.802	0.072
Deprivation	-0.076	0.052	-1.454	0.146
Religious (Yes = 1)	-0.279	0.076	-3.661	0.000

Table S5

*Summary of the regression coefficients predicting support for same-sex marriage (1-7) in the full structural equation model. Full item: “Support for same-sex marriage in NZ (The Marriage Amendment Act 2013).” Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	0.472	0.411	1.148	0.251
Punishment	-0.162	1.488	-0.109	0.913
Age	-0.012	0.004	-3.398	0.001
Gender (Male = 1)	-0.422	0.086	-4.923	0.000
Ethnicity (Maori)	-0.020	0.122	-0.164	0.870
Ethnicity (Pacific)	-0.859	0.266	-3.223	0.001
Ethnicity (Asian)	-0.268	0.213	-1.259	0.208
Education	0.048	0.017	2.831	0.005
SES	0.030	0.028	1.044	0.297
Deprivation	-0.107	0.057	-1.882	0.060
Religious (Yes = 1)	-0.939	0.085	-11.110	0.000

Table S6

*Summary of the regression coefficients predicting support for euthanasia (1-7) in the full structural equation model.*

*Full item: "Suppose a person has a painful incurable disease. Do you think that doctors should be allowed by law to end the patient's life if the patient requests it?"*

*Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	0.173	0.379	0.456	0.648
Punishment	1.224	1.419	0.863	0.388
Age	-0.005	0.003	-1.330	0.184
Gender (Male = 1)	-0.020	0.086	-0.231	0.818
Ethnicity (Maori)	0.068	0.118	0.582	0.561
Ethnicity (Pacific)	-0.332	0.262	-1.264	0.206
Ethnicity (Asian)	-0.166	0.198	-0.835	0.404
Education	-0.025	0.017	-1.490	0.136
SES	-0.046	0.028	-1.660	0.097
Deprivation	-0.055	0.054	-1.018	0.309
Religious (Yes = 1)	-0.794	0.080	-9.980	0.000

Table S7

*Summary of the regression coefficients predicting support for abortion (1-7) in the full structural equation model.*

*Full item: "Support for legalized abortion for women, regardless of the reason." Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	0.736	0.393	1.876	0.061
Punishment	-0.371	1.465	-0.253	0.800
Age	-0.008	0.003	-2.416	0.016
Gender (Male = 1)	-0.189	0.084	-2.261	0.024
Ethnicity (Maori)	0.053	0.115	0.461	0.645
Ethnicity (Pacific)	-0.287	0.285	-1.007	0.314
Ethnicity (Asian)	-0.201	0.204	-0.985	0.324
Education	0.044	0.016	2.742	0.006
SES	0.015	0.028	0.539	0.590
Deprivation	-0.103	0.055	-1.873	0.061
Religious (Yes = 1)	-0.898	0.079	-11.322	0.000

Table S8

*Summary of the regression coefficients predicting support for increased payments for those receiving Jobseeker Support (1-7) in the full structural equation model. Full item: "Increase payments for those receiving Jobseeker Support (formerly the Unemployment Benefit)."*

*Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	1.360	0.388	3.508	0.000
Punishment	-0.693	1.406	-0.493	0.622
Age	0.005	0.003	1.482	0.138
Gender (Male = 1)	-0.129	0.078	-1.646	0.100
Ethnicity (Maori)	0.210	0.102	2.062	0.039
Ethnicity (Pacific)	-0.358	0.363	-0.985	0.325
Ethnicity (Asian)	-0.577	0.219	-2.634	0.008
Education	0.056	0.014	3.931	0.000
SES	0.016	0.026	0.609	0.543
Deprivation	0.110	0.054	2.038	0.042
Religious (Yes = 1)	-0.015	0.078	-0.191	0.849

Table S9

*Summary of the regression coefficients predicting support for increased payments for those receiving Sole Parent Support (1-7) in the full structural equation model. Full item: “Increase payments for those receiving Sole Parent Support (formerly the Domestic Purposes Benefit).” Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	1.419	0.386	3.673	0.000
Punishment	-0.940	1.440	-0.653	0.514
Age	-0.005	0.003	-1.570	0.116
Gender (Male = 1)	-0.188	0.078	-2.409	0.016
Ethnicity (Maori)	0.149	0.106	1.401	0.161
Ethnicity (Pacific)	-0.144	0.261	-0.550	0.582
Ethnicity (Asian)	-0.607	0.221	-2.749	0.006
Education	0.071	0.015	4.746	0.000
SES	0.021	0.026	0.800	0.424
Deprivation	0.119	0.054	2.213	0.027
Religious (Yes = 1)	-0.049	0.077	-0.634	0.526

Table S10

*Summary of the regression coefficients predicting support for a flat tax (1-7) in the full structural equation model.*

*Full item: "Support for a 'flat' tax rate (everyone pays the same percentage of tax on their income)." Dependent variable is treated as ordinal. Reference ethnicity category is Pakeha.*

IV	Estimate	SE	z	p
Cooperation	-0.967	0.371	-2.606	0.009
Punishment	-0.101	1.329	-0.076	0.940
Age	0.000	0.003	0.030	0.976
Gender (Male = 1)	0.154	0.076	2.029	0.042
Ethnicity (Maori)	0.007	0.110	0.061	0.951
Ethnicity (Pacific)	-0.122	0.369	-0.329	0.742
Ethnicity (Asian)	0.279	0.187	1.493	0.135
Education	-0.024	0.015	-1.582	0.114
SES	-0.086	0.026	-3.304	0.001
Deprivation	0.042	0.052	0.807	0.420
Religious (Yes = 1)	0.190	0.077	2.458	0.014

Table S11

*Summary of the Bayesian multinomial logistic regression coefficients predicting political party support. The multinomial reference group is NZ First. Reference ethnicity category is Asian.*

Group	IV	Estimate	SE	Lower 95% CI	Upper 95% CI	
National	Intercept	1.297	0.274	0.764	1.824	
	Age	0.026	0.013	0.000	0.051	
	Gender (Male = 1)	0.155	0.234	-0.292	0.612	
	Ethnicity (Maori)	0.043	0.323	-0.593	0.658	
	Ethnicity (Pacific)	-0.249	0.427	-1.068	0.591	
	Ethnicity (Pakeha)	0.581	0.278	0.045	1.125	
	Education	-0.029	0.062	-0.148	0.091	
	SES	0.179	0.102	-0.016	0.379	
	Deprivation	-0.320	0.177	-0.663	0.032	
	Religious (Yes = 1)	0.051	0.223	-0.382	0.497	
	Cooperation	-0.739	0.432	-1.587	0.121	
	Punishment	0.142	0.482	-0.824	1.069	
	Labour	Intercept	1.435	0.271	0.899	1.959
		Age	0.009	0.013	-0.016	0.034
Gender (Male = 1)		-0.334	0.234	-0.779	0.132	
Ethnicity (Maori)		0.348	0.313	-0.253	0.980	
Ethnicity (Pacific)		-0.186	0.424	-1.025	0.635	
Ethnicity (Pakeha)		0.751	0.277	0.216	1.307	
Education		0.045	0.062	-0.081	0.167	
SES		0.176	0.100	-0.016	0.375	
Deprivation		-0.004	0.171	-0.339	0.340	
Religious (Yes = 1)		-0.337	0.224	-0.779	0.103	
Cooperation		0.229	0.439	-0.625	1.076	
Punishment		-0.013	0.485	-0.973	0.905	
Greens		Intercept	0.099	0.297	-0.479	0.688
		Age	-0.025	0.014	-0.054	0.002
	Gender (Male = 1)	0.024	0.266	-0.502	0.542	
	Ethnicity (Maori)	-0.101	0.346	-0.773	0.570	
	Ethnicity (Pacific)	0.052	0.454	-0.861	0.922	
	Ethnicity (Pakeha)	0.791	0.296	0.204	1.375	
	Education	0.137	0.076	-0.015	0.287	
	SES	0.383	0.123	0.139	0.621	
	Deprivation	0.067	0.200	-0.318	0.459	
	Religious (Yes = 1)	-0.628	0.268	-1.159	-0.100	
	Cooperation	0.660	0.449	-0.202	1.511	
	Punishment	-0.168	0.506	-1.119	0.795	