The lure of sirens: joint distance and velocity measurements with third generation detectors

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ABSTRACT

The next generation of detectors will detect gravitational waves from binary neutron stars at cosmological distances, for which around a thousand electromagnetic follow-ups may be observed per year. So far, most work devoted to the expected cosmological impact of these standard sirens employed them only as distance indicators. Only recently their use as tracers of clustering, similar to what already proposed for supernovae, has been studied. Focusing on the expected specifications of the Einstein Telescope (ET), we forecast here the performance on cosmological parameters of future standard sirens as both distance and density indicators, with emphasis on the linear perturbation growth index and on spatial curvature. We improve upon previous studies in a number of ways: a more detailed analysis of available telescope time, the inclusion of more cosmological and nuisance parameters, the Alcock-Paczynski correction, the use of sirens also as both velocity and density tracers, and a more accurate estimation of the distance posterior. We find that the analysis of the clustering of sirens improves the constraints on H_0 by 30% and on Ω_{k0} by over an order of magnitude, with respect to their use merely as distance indicators. With 5 years of joint ET and Rubin Observatory follow-ups we could reach precision of 0.1 km/s/Mpc in H_0 and 0.02 in Ω_{k0} using only data in the range 0 < z < 0.5. We also find that the use of sirens as tracers of density, and not only velocity, yields good improvements on the growth of structure constraints.

Key words: cosmological parameters – large-scale structure of Universe – cosmology: observations – methods: data analysis – techniques: radial velocities

1 INTRODUCTION

Since the breakthrough of gravitational wave (GW) astronomy with the first direct detection in 2015 (Abbott et al. 2016) by the LIGO/Virgo collaboration (LVC) (Aasi et al. 2015; Acernese et al. 2015), several events have been reported. The total number is changing fast, and currently approaches one hundred detections (Abbott et al. 2019, 2021b,c). Among those, of special importance was the binary neutron star (BNS) GW170817 (Abbott et al. 2017a), which was accompanied by an electromagnetic (EM) counterpart (Abbott et al. 2017d). This breakthrough *multimessenger* observation provided a precise determination of the GWs redshift, allowing its use for the first time as standard siren, and in particular to measure the Hubble constant H_0 independently (Abbott et al. 2017b).

However, the constraint on the luminosity distance d_L of the single event GW170817 translates into a large uncertainty on H_0 . Better constraints will be achieved with ad-

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ditional GW detections, but only one additional EM counterpart candidate has been reported so far, to wit for the binary black hole (BBH) coalescence GW190521 (Abbott et al. 2021b; Graham et al. 2020), and moreover this association is still uncertain (Ashton et al. 2021; Bustillo et al. 2021). Though this slightly improves the error on H_0 (Mukherjee et al. 2020; Chen et al. 2022; Bustillo et al. 2021), the latter remains much larger than the one obtained indirectly with the cosmological microwave background (CMB) (Aghanim et al. 2020) or clustering (Philcox et al. 2020), or directly with Cepheids (Riess et al. 2021) measurements and the Tip of the Red Giant Branch methods (Freedman et al. 2019).

The need for strong and independent constraints on H_0 nevertheless cannot be overestimated. As well-known, the long standing discrepancy between some (mostly low-z) estimates of H_0 and other (mostly high-z) estimates, did not disappear, and in fact intensified, with more and more precise measurements (for a recent review, see Abdalla et al. 2022).

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This tension might well be the first indication of a failure of the standard cosmological model, and deserves the most intense scrutiny. Moreover, with future upgrades and observing runs of the LVC instruments there are great expectations that standard sirens might clarify the H_0 conundrum.

Other methods to measure cosmological parameters with GWs have been proposed which preclude the need of an electromagnetic counterpart (Abbott et al. 2021a). For instance, one can use the existence of sharp features in the mass distribution of BBHs such as a characteristic mass scale or a mass-gap (Chernoff & Finn 1993; Taylor et al. 2012; Farr et al. 2019; Leyde et al. 2022; Ezquiaga & Holz 2022), correlate GW measurements with catalogues of galaxies in order to constrain the GW redshifts (Schutz 1986; Del Pozzo 2012; Chen et al. 2018), or if thousands of BBHs are available, one can rely also purely on the redshift distribution of sources in order to constrain cosmology (Leandro et al. 2022). But it is still unclear how competitive these methods will be in the future when compared to using BNS mergers, and we will not consider these possibilities in this work.

In the next decade, third generation GW detectors, such as the Einstein Telescope (ET) (Hild et al. 2011; Punturo et al. 2010; Ballmer & Mandic 2015) and Cosmic Explorer (CE) (Abbott et al. 2017c; Reitze et al. 2019), will detect BNS mergers at high redshifts and most BBH coalescences in the Universe with total masses $M \sim 2-2000 M_{\odot}$ (Maggiore et al. 2020). This will greatly increase the number of candidates for multimessenger observations, allowing us to perform precise measurement of cosmological parameters.

Traditionally, standard candles such as type Ia supernovae (SN) and standard sirens have been primarily considered as tools to build a Hubble-Lemaître diagram, and thus to constrain the background expansion of the universe. More recently, it has been realized that standard candles can also be powerful tools to measure cosmological perturbation parameters, which model the amount of structure in the universe and its growth. For higher redshift sources, Quartin et al. (2014) showed that the non-Gaussianities in the Hubble diagram residuals introduced by weak-lensing can yield precise constraints with the upcoming LSST survey. For low and intermediate redshifts, Gordon et al. (2007) proposed instead the measurements of the correlations between the supernova Hubble residual points as a probe of the peculiar velocity field. Both techniques have since been refined (Johnson et al. 2014; Macaulay et al. 2017; Howlett et al. 2017), and low-precision constraints from real data already been established (Castro et al. 2016; Qin et al. 2019; Macaulay et al. 2020).

Since the energy in EM waves obeys the inverse-square law but GW amplitudes decay instead with a single power of the distance, in relative terms it is simpler to detect highredshift BNS GWs than to observe their EM counter-parts. It is thus expected that standard sirens observed with the help of third-generation GW facilities will have higher completeness at lower redshifts. This is precisely the range in which peculiar velocity effects are more relevant, and therefore in this work we will neglect lensing and instead focus on the peculiar velocity measurements made possible by this next generation of GW observatories.

Palmese & Kim (2021) made the first forecasts for this science case. In this work we will revisit this idea improving it in a number of ways. In particular, we will: (i) make more

detailed estimates of the amount of telescope time it takes to perform the EM follow-ups with different current and upcoming telescopes following Chen et al. (2021); (ii) investigate the benefits of the 6 × 2pt method recently proposed by Quartin et al. (2022), which makes use of the standard sirens as both velocity and density tracers; (iii) include a larger number of cosmological and nuisance parameters to account for both linear bias and non-linear redshift-space distortions (RSD); (iv) include the Alcock-Paczynski (AP) corrections (Alcock & Paczynski 1979); (v) make joint forecasts for clustering and traditional distance measurements; (vi) perform a more accurate determination of the distance posterior by adopting the analytical expressions in Chassande-Mottin et al. (2019).

As discussed in Torres-Orjuela et al. (2019, 2021), the analogue of the "beaming effect" for GWs, due to the peculiar velocity of the source relative to the observer, affects the response the response of the detector to the + and \times polarizations in different manners and in addition leads to a mixing between the harmonics of the GW signal. This translates into a phase shift, which would potentially be detectable and would allow to break the degeneracy between source-frame mass and redshift and to measure the peculiar velocity of the source. However, for low peculiar velocities and nearly equal mass ratio systems as we are interested here, this effect should be subdominant and we will therefore neglect it in what follows.

This paper is divided as follows. In Section 2 we discuss the construction of our simulated GW catalogs. This is followed by calculations of the required amount of telescope time in order to perform the EM follow-ups, in Section 3. The final cosmological forecasts are presented in Section 4, and the conclusions are discussed in Section 5. Details on the distance posterior approximations, on the effect of simultaneous BNS mergers, and on how the results depend on the distance uncertainties are covered in three appendices.

2 GRAVITATIONAL WAVE OBSERVATIONS

2.1 Catalogue generation

In this work we will focus on GWs observations accompanied with EM counterparts, that is called "bright sirens". For simplicity, we will consider that the bright sirens are generated only by the BNS mergers, ignoring the contribution of NS-BH binaries since the estimate of its merger rate is much lower than that found for BNS (Abbott et al. 2021d). In order to construct a mock catalog of the BNS mergers observed by ET, we first assume that the BNS merger population is distributed within the cosmological volume through the redshift distribution p(z) that is written as a function of the merger rate per redshift, in the observer frame $\mathcal{R}(z) (\equiv \frac{dN}{dt_{obs}dz}, \text{ i.e. is}$ the number of mergers per redshift and observer-frame time):

$$p(z) = A\mathcal{R}(z) = A\frac{\mathrm{d}V_c}{\mathrm{d}z}\frac{\mathcal{R}_m(z)}{1+z},\qquad(1)$$

where A is a normalization constant ensuring that the integration of p(z) goes to unity over $0 < z < z_{\text{max}}$, dV/dz is the comoving volume element, and $\mathcal{R}_m(z)$ is the merger rate per comoving volume in the source frame. We assume that \mathcal{R}_m follow the Madau-Dickinson star formation rate (Madau & Dickinson 2014):

$$\mathcal{R}_m(z) = \mathcal{R}_0 \frac{1.00257(1+z)^{2.7}}{\left(1 + \left(1 + z/2.9\right)^{5.6}\right)},\tag{2}$$

with a local volumetric rate of $\mathcal{R}_0 = 300 \,\mathrm{Gpc}^{-3}\mathrm{yr}^{-1}$, that agrees with the latest results of the LVC (Abbott et al. 2021d).¹ To reduce computational time, we truncate the volume integral at z = 5, which is enough to cover the ET detection range as we will show.

We draw NS masses m_1 and m_2 from a Gaussian distribution with mean $1.4\,\mathrm{M}_{\odot}$ and standard deviation $0.2\,\mathrm{M}_{\odot}$, restricted to the range 1–3 M_{\odot} , which is in agreement with the limit inferred by the third observing run of LIGO-Virgo (see section V.B of Abbott et al. 2021d). Due to propagation in a Universe described by the Friedmann-Lemaître-Robertson-Walker metric, the observed masses in the detector frame are redshifted relative to the source-frame ones: $m_{1,2} \to m_{1,2}(1+z)$. We recall that the chirp mass is defined as $\mathcal{M} = (m_1^3 m_2^3 / (m_1 + m_2))^{1/5}$, this is the parameter that drives the evolution of the binary at leading post-Newtonian order (Blanchet 2014). It is redshifted in the same way as the individual masses. Spins are isotropically oriented with magnitudes < 0.05 and tidal deformabilities are distributed according to a uniform distribution U(0, 5000). The events are assumed to be isotropically distributed and randomly oriented. Finally, the coalescence phase (φ_c) and the polarization angle (ψ) are both distributed according to a uniform distribution $U(0, 2\pi)$.

The total number of BNS mergers N can be simply computed using the definition $N = \Delta_{\rm obs} t_{\rm obs} \mathcal{R}_{\rm max}$, where $\mathcal{R}_{\rm max} = \int_0^{z_{\rm max}} \mathcal{R}(z) dz$, $\Delta_{\rm obs}$ is the duty cycle, and $t_{\rm obs}$ is the observation time. Assuming the flat Λ CDM model with $\Omega_{m0} = 0.3$ and $H_0 \equiv 100 h \text{ km/s/Mpc} = 70 \text{ km/s/Mpc}$ as our fiducial cosmological model, we find that the total number of BNS mergers is $N \approx 8 \times 10^5$ per year of observation, assuming $\Delta_{\rm obs} = 0.8$ for the ET (Belgacem et al. 2019).

2.2 Detection

Here we will consider the prospects for detecting BNS mergers with third generation GW detectors, which are expected to take over Advanced Virgo in the 2030's. We will use the Europe-based ET specifications as our baseline third generation configuration, but similar results are expected for the USA-based CE. Both the number of events and the GW parameter estimation could be further improved by combining both detectors, or a combination of second and third generation facilities, but we will conservatively focus only on ET forecasts here.

ET consists of three Michelson interferometers arranged in a triangular shape, with 10-km-long arms and 60° opening angle, and can be seen as a combination of three noiseuncorrelated detectors. ET is sensitive to GWs in the $1-10^4$ Hz band, with a level of noise an order of magnitude lower than current detectors. This improvement will be achieved thanks to the longer arms, but also to the use of cryogenic technologies to reduce thermal noise and quantum technology to reduce the high-frequency quantum shot noise. As a result, ET will observe BNS mergers for much longer, up to tens of hours, improving on the parameter estimation. Moreover, combining the three noise-uncorrelated detectors will allow us to triangularise the signal, providing an accurate sky localization (Mills et al. 2018; Chan et al. 2018).

For each BNS merger in our mock catalogue, we simulate its GW signal using the frequency domain approximant IM-RPhenomPv2 (Dietrich et al. 2017), which includes a 5PN modification to the phase due to tidal effects (Damour 1984; Flanagan & Hinderer 2008; Wade et al. 2014). The signal in each ET detector is given by a combination of the two GW polarisations, weighted by the antenna pattern functions (Schutz & Tinto 1987; Tinto 1987):

$$h_i = F_i^+ h_+ + F_i^{\times} h_{\times}. \tag{3}$$

The expressions for $F_i^{+,\times}$ can be found in Regimbau et al. (2012). The total network signal-to-noise-ratio (SNR) is defined as

$$\rho_{\rm net} = \left[\sum_{i=1}^{3} \rho_i^2\right]^{1/2} , \qquad (4)$$

where

$$\rho_i^2 = 4 \int_0^\infty \frac{|h_i(f)|^2}{S_n(f)} \mathrm{d}f$$
 (5)

is the SNR in the *i*-th ET detector. The power spectral density, $S_n(f)$, measures the level of noise (assumed to be stationary and Gaussian) in the detector at each frequency. We consider an event to be detectable if $\rho_{\text{net}} \ge 12$. We compute SNRs using the Bilby package (Ashton et al. 2019).

2.3 Parameter estimation

The parameters of a source are estimated using Bayes' theorem to obtain their posterior distribution from observed data. For our analysis, we are mostly interested in the accuracy of the sky localization for each event, in order to determine if there should be an EM follow-up, as well as the measurement error on the luminosity distance, which will be combined with EM measurement of redshift to infer cosmological parameters. Therefore, instead of performing computationally expensive Bayesian analyses, we use the analytic approach of Cutler & Flanagan (1994); Chassande-Mottin et al. (2019) to estimate the error on d_L . Assuming the sky localization is known from an EM counterpart and that intrinsic parameters (masses and spins) are not strongly correlated with inclination and distance, it provides the posterior distribution for the latter marginalised over polarization and phase. As illustrated in Chassande-Mottin et al. (2019), this is a very good approximation in most cases, except for very few events where the GW polarization becomes degenerate. These expressions are very convenient to perform forecasts, while at the same time being more accurate than simpler fits which relied exclusively on the relative distance errors of a given GW event (Zhao et al. 2011). In Appendix A we review BNS distance estimations and illustrate the performance of these expressions for the case of the ET.

Moreover, the observation of an EM counterpart provides an additional constraint on the inclination angle, which helps

¹ This merger rate is however still poorly constrained if one accounts for model systematics. Its 95% confidence interval is very broad: $10 - 1700 \text{ Gpc}^{-3} \text{ yr}^{-1}$.

to break the typical degeneracy with the luminosity distance, reducing in turn the uncertainty on d_L (Dhawan et al. 2019). Therefore, we consider a scenario where this information is incorporated via a Gaussian prior on the inclination. The Gaussian is centred at the true value and with a standard deviation of 10° (Dhawan et al. 2019). As for the sky localization, we assume $\Delta\Omega \propto 1/\rho_{\rm net}^2$, using the full Bayesian results for observations of BNSs with LIGO/Virgo network (Del Pozzo et al. 2018) to calibrate the proportionality constant. Since triangularisation can be performed using the three ET detectors, we expect such a scaling with the SNR should hold.

2.4 Comparisons between kilo and supernovae distances

It is interesting to compare the capabilities of both supernovae (SN) and kilonovae as standard candles. The former has a much higher expected rate, to wit $\mathbf{r}_{Ia} = 21000 (1 +$ $z)^{1.95}/({\rm yr~Gpc^3})$ (Cappellaro et al. 2015; Amendola & Quartin 2021). On the other hand, by themselves SN cannot measure H_0 due to their unknown absolute magnitudes M_B , and instead rely on calibration methods with external data such as Cepheids. Moreover, SN surveys indicate the presence of an intrinsic scatter which in the best cases, with infrared observations, is around 0.1 mag (Avelino et al. 2019). KN on the other hand, as far as currently known, have precision limited only by the GW SNR. This means that as GW detectors improve, we should find some high SNR events with distances which are more precise than those of SN. Moreover, both the large number of already identified possible sources of systematic effects on SN (see, e.g., Howell 2011) and in general the empirical nature of the SN light-curve fitting mean that it will be difficult to keep systematic effects subdominant in upcoming SN data.

In Section 4 below we make quantitative comparisons between KN and SN as cosmological probes when combining GW and EM measurements, under the assumption of no relevant systematic sources of uncertainties in either one.

3 ELECTROMAGNETIC OBSERVATIONS

BNS coalescences are expected to originate various types of EM counterparts across the spectrum. The most promising counterparts for obtaining redshifts for the number of BNS mergers expected to be detected by the ET are the kilonovae (KN). These are optical/near-IR emissions caused by the decay of heavy ions via the r-process, which occurs in the mass ejecta of the BNS merger (Li & Paczynski 1998; Metzger et al. 2010; Barnes & Kasen 2013). This emission can be viewed up to days-weeks after their production. Besides the optical/near-IR radiation, the ejected KN can still interact with the interstellar medium, producing a forward shock that emits radiation in the radio waveband. Another expected counterpart are gamma ray bursts (GRBs). However, those emissions are highly beamed along the binary's orbital axis. As discussed in Chen et al. (2021), this high collimation together with the typical brightness of these GRBs mean that the observable number of these multimessenger events are considerably smaller than those involving the kilonovae. We will therefore for simplicity only consider the latter in this work.

Current and forthcoming wide-field telescopes, such as the Rubin Observatory (Rubin)², Wide Field Survey Telescope (WFST)³, Multi-channel Photometric Survey Telescope (Mephisto)⁴ and the Zwicky Transient Factory (ZTF) (Masci et al. 2019)⁵, will be able to detect a considerable fraction of kilonovae counterparts to BNS events which are well localised by the ET and/or CE. Since these GW detections are still more than a decade away, other telescopes and cameras may still come online in the meantime.

In our analysis we will start by comparing the capabilities of all four telescopes: Rubin, WFST, Mephisto and ZTF. Using the published sensitivities of each one, we compute the required exposure time making use of the power-law fits derived in Zhu et al. (2021) in order to use a fixed amount of the total available telescope time. This calculation follows the one performed for the case of Rubin by Chen et al. (2021). Fixing a priori a fraction of telescope time dedicated to BNS GW follow-up allows a better comparison of all four facilities. Moreover, Rubin scans mostly the southern hemisphere whereas the other telescopes cover the northern region, so their data will complement Rubin's. We quote results for two different fraction of telescope time (F_{time}) needed to observe all the events. To wit, we assume either $F_{\text{time}} = 10\%$ or 50%, with the latter being our baseline case. Although dedicating up to half of all usable telescope time for kilonova follow-ups may sound excessive, the third generation of GW detectors here considered are not expected to come online before the second half of the 2030s, therefore after the completion of the original proposed surveys for these instruments.

For each scenario, following Chen et al. (2021) we select events with sky localization $\Delta \Omega \leq 20 \text{ deg}^2$ to ensure a small number of pointings needed to identify that event in the sky. We also make the conservative simplifying assumption that the average number of pointings p for these events will be given by

$$p = \frac{40 \, \deg^2}{\text{FoV}} \,, \tag{6}$$

where FoV is the telescope effective field of view in squared degrees. The FoV values for each telescope used here are shown in Table 2. We also assume that the observable sky area for all four telescopes are equal in size, and that the fraction $f_{\rm obs}$ of events followed-up is given in all cases by $f_{\rm obs} = 0.4$. Although the observable sky area in each night is larger than this for all four telescopes, this lower number is supposed to already take into account unfavorable meteorological conditions and instrument downtime. We remark that this is a conservative choice, and that other authors assume a more aggressive choice of $p = \Delta\Omega/\text{FoV}$ (McGee et al. 2020). Nevertheless this also takes into account the fact that the GW search area may be elongated which leads to wasted area in the borders of the FoV.

Finally, we consider two changes to the analysis of Chen et al. (2021): (i) we do not adopt a sharp cutoff in the detection horizon $d_{L,\text{lim}}$, and instead consider the full extension of the efficiency curve without ignoring its smooth tail; (ii) for each efficiency curve we propose that the exposure time is

⁵ https://www.ztf.caltech.edu/

² https://www.lsst.org/

³ https://wfst.ustc.edu.cn/

⁴ http://www.mephisto.ynu.edu.cn/site/



Figure 1. Kilonova efficiency ε as a function of d_L for the telescopes used in this work (Rubin, WFST, Mephisto and ZTF), and two exposure time: $t_{\text{exp.}} = 300 \text{ s}$ (solid curve) and $t_{\text{exp.}} = 600 \text{ s}$ (dashed curve).

adjusted so that less time is spent in closer events, which are brighter and easier to spot. One can thus maximize the number and distance of observable kilonovae while maintaining maximum completeness for the closer events. This advanced strategy will be better described in the next subsection.

For the kilonovae program, we use the results presented in Zhu et al. (2021). They assume that all the simulated kilonovae are AT2017gfo-like, adopt a total ejecta mass of $M_{\rm ei} = 0.04 M_{\odot}$, a half-opening angle equal to $\Phi = 60^{\circ}$, and ignore the viewing-angle dependence. In our program, we assume that the electromagnetic counterparts will be identified by three epochs of observations in two filters (gr, that aretypically used by the surveys we consider). Although other filters will be available that could be added to the analysis, we choose not to include them since the information added on the magnitudes PDFs would not compensate the increase of the observation time for the identification of the kilonova. The KN efficiency curves were found using the PDFs for the apparent magnitude and the absolute magnitude presented in Zhu et al. (2021). For each band we convolve these PDFs to find the PDF P_{μ} of distance modulus μ , and then find the luminosity distance PDF P_{d_L} :

$$P_{\mu} \equiv P_{m_{\rm app}} \otimes P_{(-m_{\rm abs})} \,, \tag{7}$$

$$P_{d_L} = \frac{5}{(d_L/\text{pc})\ln 10} P_{\mu} \,. \tag{8}$$

The kilonova detectability will basically depend on the filter selection and the survey sensitivity that is characterized through the limiting magnitude. All the limiting magnitude values were calculated using their relation to the exposure time presented in Table 2 of Zhu et al. (2021). The KN efficiency curves are shown in Figure 1.

Our simulations predict 5498 BNS merger detections (events with $\rho_{\text{net}} \geq 12$) per year with ET. This number represents only ~ 1% of our whole mock catalog ($z \leq 5$). After imposing the cutoff for detection of an EM counterpart, this number is reduced according to the kilonova observation scenario (as will be shown below in Tables 2 and 3). As an example, in Figure 2 we show the distribution of event parameters before and after imposing GW and EM selection cuts in a given scenario. The GW selection curves show that the detections occur in a reduced redshift range, as expected, selecting sources up to $z_{\rm lim} \sim 1.06$; this value shows that the choice of $z_{\rm max} = 5$ as the threshold of the BNS GW distribution is enough to ensure that all relevant events are considered. This selection effect also impact the distribution of \mathcal{M}_z , that prefers events with low redshifted chirp mass. The distribution of $\rho_{\rm net}$ for all injections has a mean equal to ≈ 2.7 , and we can see that most of the events are distributed in $\rho_{\rm net} < 10$ which explain the computed low fraction of BNS GW which are detected.

3.1 Exposure time as a function of luminosity distance

BNS coalescences that are closer to us require shorter exposure times to have their electromagnetic counterpart detected, when compared to similar events at larger distances. This implies that the dedicated exposure time should be a function that grows with the luminosity distance. Often, however, a simplified analysis disregards this fact by considering that the exposure time is constant for all kilonovae, regardless of their distances (e.g. Chen et al. 2021). Here we will propose instead that since BNS GW events themselves already constrain the luminosity distance, dedicated follow-up programs could use this information to adjust exposure times to be a function of the luminosity distance.

A simple estimation indicates that to maintain a constant signal-to-noise ratio, the exposure time should be proportional to d_L^4 . The reason is that the flux (i.e. the signal) decreases with the square of the distance, and that the instrumental noise decreases roughly with the inverse square root of the exposure time. In practice, instrumental noise may not follow exactly this simple rule, but in any case the limiting magnitude of a given telescope as a function of exposure time is well understood. For the four telescopes here considered, Zhu et al. (2021) approximates the limiting magnitude in different bands as a power law of the exposure time. Making use of this we adjusted iteratively the exposure time in order to stretch the efficiency curves of the instruments to the maximum possible distance d_L^{\max} , defined as the luminosity distance where the detection efficiency is larger than 99%. This guarantees that basically all kilonovae up to that distance will be detected.

We then fit a power law in the form:

$$t_{\rm exp} = a \left(\frac{\mathrm{d}_L}{400\,\mathrm{Mpc}}\right)^n,\tag{9}$$

where the constants $\{a, n\}$ provide the best fit to the data. Figure 3 illustrates the result, showing that in all cases a power law is a very good fit. Clearly a power-law exposure time program is problematic when the distances involved are too low or too high. When d_L is too low, the total telescope time becomes dominated by the slewing time when changing pointings. For very far away sources, on the other hand, the time to be spent on a single object becomes prohibitive, not to mention the inherent challenges of performing very long exposures in astronomy. We therefore set a minimum and maximum exposure times: the former is fixed to be 30s, a standard exposure time in surveys, while the latter is set by



Figure 2. Distribution of a subset of GW parameters from our original catalog (blue curves) and after the GWs selection criteria adopted ($\rho_{\text{net}} \geq 12$ - beige curves) and EM counterpart selection considering $F_{\text{time}} = 0.5$ for the Rubin survey (see Table 3 - olive green curves). Here \mathcal{M}_z is the redshifted chirp mass defined as $\mathcal{M}_z = (1+z)\mathcal{M}$.



Figure 3. Required telescope exposure time t_{exp} as a function of the minimum luminosity distance with 99% telescope efficiency for the surveys here considered. The data are well fitted by power-laws with exponents not too different from the naive expectation of 4 [see text]. We assume a minimum t_{exp} of 30s.

the allocated fraction of yearly telescope time. We then have

$$t_{\exp}\left(d_L, t_{\exp}^{\max}\right) = \begin{cases} 30 \, s, & ad_L^n < 30 \, s \\ ad_L^n, & 30 \, s < ad_L^n < t_{\exp}^{\max} \\ t_{\exp}^{\max}, & ad_L^n > t_{\exp}^{\max} \end{cases}$$
(10)

where t_{\exp}^{\max} is the maximum exposure time.

Table 1 presents the values of $\{a, n\}$ that represent the best fit for each of the surveys. The exponents n are not too far from the value 4, the simple expectation discussed above which assumes the limiting magnitudes scale as $1/\sqrt{t_{exp}}$.

As expected, we find that adjusting the follow-up exposure time using the distance estimate from GW observations allows the telescope to spend less time in finding nearby (and brighter) events and longer times in those farther away. The total observing time t_{source} required per source was computed using:

$$t_{\rm source} = p t_{\rm exp} N_{\rm epochs} N_{\rm filters} + 120s \,, \tag{11}$$

where t_{exp} is the exposure time, N_{epochs} is the number of epochs, which we assume to be 3, $N_{filters}$ is the required number of filters, which we assume to be 2 and p is calculated by equation (6). The last term is a two minutes overhead re-

Telescope	a(s)	n
Rubin	1.4	4.6
WFST	45	4.7
Mephisto	78	3.2
\mathbf{ZTF}	5000	4.4

Table 1. The optimized exposure time parameters of Eq. (9) for the different telescopes considered in this work.

Telescope	FoV	$f_{\rm obs}$	scenario	N_{SS}/yr	F_{time}
Rubin	9.6	0.4	$t_{\rm exp} = 300s$	1194	0.66
			$t_{\rm exp} \propto d_L^{4.0}$ $t_{\rm exp} = 300 {\rm e}$		0.47
WFST	6.6	0.4	$t_{\rm exp} = 500s$ $t_{\rm exp} \propto d_L^{4.7}$	307	0.20
Menhisto	3.1	04	$t_{\rm exp} = 300 {\rm s}$	298	0.54
mephibto	0.1	0.1	$t_{\rm exp} \propto d_L^{3.2}$	200	0.46
\mathbf{ZTF}	47	0.4	$t_{\rm exp} = 300 {\rm s}$	11	0.0015
		0.1	$t_{ m exp} \propto d_L^{4.4}$		0.0012

Table 2. Comparison of different wide field telescope properties and of different exposure time strategies. FoV is the field of view in squared degrees, $f_{\rm obs}$ is the assumed fraction of the sky that observed, $N_{SS}/{\rm yr}$ is the number of gravitational waves that will have their counterparts observed per year, and $F_{\rm time}$ is the fraction of telescope time needed to observe all the events. As can be seen, adjusting the exposure time to be a power of the estimated distance decreases the needed amount of survey time.

quired for slewing and for filter changes, following Chen et al. (2021). Table 2 compares a scenario in which fixed 300s exposure times are used with our proposal, fixing the maximum exposure in each case to ensure the observation of the same amount of kilonovae. We assume that all surveys will have a total of 3600 hours available per year. In all cases, the same number of events are observed while consuming in between 20 and 30% less telescope time. In what follows we will therefore assume the observational program described in Eq. (10).

With the observational method established, the next interesting forecast is on the total number of expected observed events for a fixed fraction of the total telescope time available each year. We compare two cases: using 10% and 50% of each telescope time. Table 3 shows the results. Clearly Rubin is the performance winner. Nevertheless, the other telescopes combined are still able to observe a good amount of events, albeit at lower redshifts. Since they cover the northern hemisphere

Telescope	$t_{\exp}^{\max}(s)$	z_{\max}	$f_{20{\rm deg}^2}$	$f_{\rm obs}$	$N_{SS}/{ m yr}$	$F_{\rm time}$
Rubin	90	0.49	0.89	0.4	819	0.1
WFST	200	0.27	0.94	0.4	244	0.1
\mathbf{ZTF}	3200	0.17	0.98	0.4	78	0.1
Mephisto	140	0.23	0.96	0.4	157	0.1
Rubin	315	0.61	0.86	0.4	1196	0.5
WFST	550	0.34	0.92	0.4	404	0.5
\mathbf{ZTF}	8500	0.22	0.96	0.4	152	0.5
Mephisto	320	0.30	0.93	0.4	306	0.5

Table 3. Similar to Table 2 but for the same fixed fractions of survey time. Here $f_{20 \text{deg}^2}$ is the fraction of GW sources that have localization area $< 20 \text{ deg}^2$, and $t_{\text{exp}}^{\text{max}}$ is the maximum exposure time.

it shows that they may perform an interesting complementary follow-up to Rubin in this hemisphere.

Dedicating a large fraction of telescope time to observe transients such as kilonovae with long expositions may lead to some events being lost due to their overlap in the time domain. In Appendix B we investigate this and show that with the possible exception of ZTF, in all other cases this would be a negligible effect.

4 COSMOLOGICAL FORECASTS

In this section we will discuss the advantages of using KN measurements as tracers of both density and velocity fields, besides only as distance indicators. Since EM counterparts from BNS are expected mostly in $z \leq 0.5$, as can be seen in Figure 2, we will split our forecasts in lower redshift ($z \leq 0.5$) and intermediate/high redshifts ($z \geq 0.5$). This is a simple yet useful separation in light of the recent framing of the Hubble tension as a tension between lower and higher redshifts (see e.g. the review Abdalla et al. 2022). We will assume as our baseline survey a 5-year observational run of the ET and the follow-up facilities.

We use the symbols g for galaxies and c for standard candles in general. When we need to specify which standard candle we will use s for supernovae and k for kilonovae (assumed to be measured jointly with the associated GW). From standard sirens we extract directly the luminosity distance $d_L(z_t)$. The observed redshift is the sum of a cosmological redshift \bar{z} plus a small redshift induced by peculiar velocity: $1 + z_t = (1 + \bar{z})(1 + z_p) \rightarrow z_t = \bar{z} + (1 + \bar{z})z_p$, with $v_c = z_p$. Then we can write

$$d_L(z_{\rm obs}) \approx d_L(\bar{z}) + \frac{\mathrm{d}d_L}{\mathrm{d}z} (1+\bar{z})v_c \,, \tag{12}$$

where \bar{z} is the cosmological redshift. Therefore

$$v_c = \frac{\Delta d_L}{d_L} \frac{1}{\mathrm{d}\log d_L/\mathrm{d}\log(1+z)} \,. \tag{13}$$

As stated above, we will neglect the specificities of the beaming effect for GWs discussed in (Torres-Orjuela et al. 2019, 2021) and model the velocity effects like that on photons. In the CMB rest frame (i.e. neglecting our own peculiar velocity) we have that (Hui & Greene 2006; Davis et al. 2011)

$$d_L(z_{\rm obs}) \approx d_L(\bar{z})(1-2v_c) + \frac{\mathrm{d}d_L}{\mathrm{d}z}(1+\bar{z})v_c ,$$
 (14)

which leads to

$$v_{\rm c} = \frac{\Delta d_L}{d_L} \left[\frac{\mathrm{d}\log d_L}{\mathrm{d}\log(1+z)} - 2 \right]^{-1} \,. \tag{15}$$

The statistical uncertainty in the velocity field is thus (Amendola & Quartin 2021):

$$\sigma_{v,\text{eff}}^2 \equiv \left[\frac{\Delta d_L}{d_L}\right]^2 \left[2 - \frac{\mathrm{d}\log d_L}{\mathrm{d}\log(1+z)}\right]^{-2} + \frac{\sigma_{v,\text{nonlin}}^2}{c^2}.$$
 (16)

Since we assume for our fiducial case $\Omega_{k0} = 0$, we can write $d \log d_L/d \log(1+z) = 1 + (1+z)^2/[d_L H(z)].$

We have therefore three random fields: v_c , δ_g and δ_c . We use subscripts v, g, s to refer to the three fields, respectively. These give rise to six linear auto- and cross-power spectra (Quartin et al. 2022)

$$P_{\rm gg}(k,\mu,z) = \Upsilon \left[1 + \beta_{\rm g} \mu^2 \right]^2 b_{\rm g}^2 S_{\rm g}^2 D_+^2 P_{\rm mm}(k) + \frac{1}{n_{\rm g}}, \qquad (17)$$

$$P_{\rm cc}(k,\mu,z) = \Upsilon \left[1 + \beta_{\rm c} \mu^2 \right]^2 b_{\rm c}^2 S_{\rm c}^2 D_+^2 P_{\rm mm}(k) + \frac{1}{n_{\rm c}}, \qquad (18)$$

$$P_{\rm gc}(k,\mu,z) = \Upsilon \left[1 + \beta_{\rm g}\mu^2\right] \left[1 + \beta_{\rm c}\mu^2\right] b_{\rm g} b_{\rm c} S_{\rm g} S_{\rm c} D_+^2 P_{\rm mm}(k) + \frac{n_{\rm gs}}{n_{\rm g} n_{\rm c}}, \qquad (19)$$

$$P_{\rm gv}(k,\mu,z) = \Upsilon \frac{H\mu}{k(1+z)} \Big[1 + \beta_{\rm g} \mu^2 \Big] b_{\rm g} S_{\rm g} S_{\rm v} f D_+^2 P_{\rm mm}(k), \quad (20)$$

$$P_{\rm cv}(k,\mu,z) = \Upsilon \frac{H\mu}{k(1+z)} \Big[1 + \beta_{\rm c} \mu^2 \Big] b_{\rm c} S_{\rm c} S_{\rm v} f D_+^2 P_{\rm mm}(k), \quad (21)$$

$$P_{\rm vv}(k,\mu,z) = \Upsilon \left[\frac{H\mu}{k(1+z)}\right]^2 S_{\rm v}^2 f^2 D_+^2 P_{\rm mm}(k) + \frac{\sigma_{v,\rm eff}^2}{n_{\rm c}}, \quad (22)$$

where $\Upsilon \equiv (Hd_{L,r}^2)/(H_rd_L^2)$, $\beta_i \equiv f/b_i$, $\mu \equiv \hat{k} \cdot \hat{r}$, with i = g, c, $S_{g,c,v}$ are damping terms, D_+ is the growth function, $n_{g,c}$ are the densities of galaxies and sirens, n_{gc} is the fraction of sirens in galaxies belonging to the survey (here assumed to vanish for simplicity), and $P_{\rm mm}$ is the matter power spectrum at z = 0. We also take into account the fact that the values of k and μ depend on the cosmological model (the Alcock-Paczynski effect) through H and d_L (for more details, see Amendola & Quartin 2021). As in Quartin et al. (2022), the smoothing factors are modeled as

$$S_{\rm g,c,v} = \exp\left[-\frac{1}{4}(k\mu\sigma_{\rm g,c,v})^2\right],\tag{23}$$

with fiducial values chosen as $\sigma_{\rm g} = \sigma_{\rm c} = 4.24 \text{ Mpc}/h$ and $\sigma_{\rm v} = 8.5 \text{ Mpc}/h$ (see Koda et al. (2014); Howlett et al. (2017); Dam et al. (2021) for the choice of these values).

The full analysis using all six spectra for supernovae was dubbed the $6 \times 2pt \ g\text{-}s\text{-}s$ method in Quartin et al. (2022). For sirens we can instead refer to it as a $6 \times 2pt \ g\text{-}k\text{-}k$ method, but we will often refer to it as simply $6 \times 2pt$ for short. Likewise, using sirens only as velocity but not density tracers (i.e., dropping the $P_{\rm cc}$, $P_{\rm gc}$ and $P_{\rm cv}$ terms), we have a $3 \times 2pt \ g\text{-}k$ method (or simply $3 \times 2pt$ for short). Finally, using only $P_{\rm gg}$ consists of a traditional full shape galaxy power spectrum approach.

For the sirens, the distance error estimates were obtained with the method discussed in Section 2.3, both including or not EM information on the BNS orbit inclination (through the addition of a prior). For supernovae, which we will also employ to forecast performance as distance estimators for comparison, we assume distances are measured with magnitude uncertainties given by the sum in quadrature of the intrinsic scatter $\sigma_{\rm int} = 0.13$ mag with the lensing-induced scatter of $\sigma_{\rm lens} = 0.052z$ (Quartin et al. 2014). For the number of events for lower redshifts we follow Quartin et al. (2022) assuming 15% SN completeness, for a total of 239k SN. For higher redshifts we use the expected number of SN with the Nancy Grace Roman Space Telescope as computed by Rose et al. (2021). Finally, for the CMB we quote the results found in Quartin et al. (2022), which comes from a combination of Planck 2018 TTTEEE (Aghanim et al. 2020) without lensing and an analysis on σ_8 and γ performed by Mantz et al. (2015).

For galaxies, we produce two different forecasts. For z <0.5 we use the predicted values for the DESI Bright Galaxy Survey presented in Aghamousa et al. (2016), which has an expected linear bias of $b_g = 1.34/D_+(z)$. Converting their numbers to volumetric density, we find $n_a = \{38.50, 17.63,$ 6.439, 1.937, 0.3571} $10^{-3} (h/\text{Mpc})^3$ in the z bins of $\Delta z =$ 0.1 centered on {0.05, 0.15, 0.25, 0.35, 0.45}, respectively. For 0.5 < z < 1.5 we will forecast the performance of the traditional $1 \times 2pt$ using only galaxies. We will use as baseline the DESI Emission Line Galaxy (ELG) survey, which covers well this redshift range. The assumed linear bias in this case is $b_g = 0.84/D_+(z)$ (Aghamousa et al. 2016), and the number densities are {0.1778, 1.099, 0.8130, 0.7940, 0.5007, 0.4384, 0.4097, 0.1532, 0.1316 $10^{-3} (h/Mpc)^3$ for the bins centered on {0.65, 0.75, 0.85, 0.95, 1.05, 1.15, 1.25, 1.35, 1.45}. Finally, since the kilonova bias as density tracers has not been studied in detail, we assume for simplicity that $b_{\rm k} = 1.0/D_+(z)$, and likewise for SN.

The $6 \times 2pt$ Fisher matrix (FM) for a set of parameters θ_{α} , in a survey of volume V and for an interval Δ_k of k-modes, is (Tegmark 1997; Abramo & Amendola 2019)

$$F_{\alpha\beta} = \frac{1}{(2\pi)^3} 2\pi k^2 \Delta_k V \bar{F}_{\alpha\beta} = V V_k \bar{F}_{\alpha\beta} , \qquad (24)$$

where $V_k = (2\pi)^{-3} 2\pi k^2 \Delta_k$ is the volume of the Fourier space integrated over the azimuthal angle but not over the polar angle, and where

$$\bar{F}_{\alpha\beta} = \frac{1}{2} \int_{-1}^{+1} \mathrm{d}\mu \, \frac{\partial P_{ab}}{\partial \theta_{\alpha}} P_{ad}^{-1} \frac{\partial P_{cd}}{\partial \theta_{\beta}} P_{bc}^{-1} \,, \tag{25}$$

to be evaluated at the fiducial value. The elements of the data covariance matrix P_{ab} , with a, b standing for g, c, v, are the six power spectra (17)–(22). Denoting with V(z) the volume of the z-shell, the k-cells are chosen with size $\Delta_k = 2\pi/V(z)^{1/3}$ between $k_{\min}(z)$ and k_{\max} . Following Garcia et al. (2020), we take $k_{\min} = 2\pi/V(z)^{1/3}$.

The choice of $k_{\rm max}$ is more delicate. The amount of information grows rapidly with $k_{\rm max}$ but so do the modeling uncertainties due to non-linear effects. Ideally, one would like to select the highest $k_{\rm max}$ that does not introduce significant non-linear effects: this, however, depends clearly on the cosmological model. In Amendola et al. (2022), the dependence of Fisher matrix forecasts for a non-linear power spectrum on $k_{\rm max}$ has been explored. The conclusion was that, if one does not employ a specific cosmological model, one needs to have independent strong prior constraints on the linear and non-linear bias parameters to decide which $k_{\rm max}$ is safe to use. In this paper, however, we restrict ourselves to ΛCDM (but with a free growth index γ) and in this case the standard choice $k_{\rm max} = 0.1 \ h/\text{Mpc}$ is likely to be sufficient to ensure we remain within the linear regime. This is also a value close to 0.08 h/Mpc, which was found in retrospect to be the one for which Fisher matrix precision calculations best matched that of current real data analysis (Foroozan et al. 2021). Note also that since non-linearities are stronger at lower redshifts, one is expected to be able to reach higher k_{max} at higher redshifts (Nishimichi et al. 2009; Tomlinson & Jeong 2022), but here we adopt a fixed value for simplicity.

The fiducial values of the parameters that are varied in the 6×2 pt Fisher matrix are

$$\{\sigma_8, \gamma, H_0, \Omega_{m0}, \Omega_{k0}\} = \{0.83, 0.545, 70 \,\mathrm{km/s/Mpc}, 0.3, 0\}.$$
(26)

We also fix $n_s = 0.96$ and $\tau = 0.066$. When analysing the combined constraints of clustering and distance methods, we simply sum the 6 × 2pt FM with the usual distance FM for KN or SN (Amendola & Tsujikawa 2015). For KN, the distance parameters are just the background ones, to wit Ω_{m0} , Ω_{k0} , H_0 , while for SN we have an extra parameter M_B to account for their unknown absolute magnitude. We marginalize over this when combining with the clustering FM, but for the distance constraints alone this makes the SN FM degenerate. So, for this case alone, instead of M_B and H_0 separately we use one single parameter which accounts for the combination $M_B - 5 \log H_0$ and marginalize over it. In practice, this can be achieved by simply fixing M_B and marginalizing over H_0 (or vice-versa).

Following Quartin et al. (2022) we also employ three global nuisance parameters to account for the non-linear RSD $(\sigma_g, \sigma_c, \sigma_v)$ and two bias nuisance parameters in each redshift bin $(b_g^{z_i} \text{ and } b_c^{z_i})$. We adopt a Gaussian prior with 0.5 uncertainties for Ω_{m0} , Ω_{k0} , and with 50% errors in all bias and RSD damping parameters. For σ_8 , γ and H_0 we use uninformative priors.

Table 4 summarizes our forecasts for 5 years of observation. All uncertainties are fully marginalized over all other parameters. Results are divided into three redshift ranges. For lower redshifts, where sirens will be useful, we compare the traditional full-shape (linear) power spectrum analysis with those of clustering combining galaxies and sirens, and with standard candle distance estimates. Additionally, we also separate the results to include (i) the contribution of the ι prior on the measurement of d_L , and (ii) assuming the scenario in which a second Rubin-like survey operate simultaneously in the northern hemisphere, doubling the observed sky area and consequently the number of galaxies and standard sirens measured. For the former, we find that the inclusion of a 10° ι prior on the 6 \times 2pt has only a very small effect. In fact, the uncertainty on γ decreases by only 1.5%, even though such a prior results in improvements on their distance measurements by over 15%. In Appendix C we explore why this is the case. Nevertheless for their use as traditional distance indicators, such a prior does lead to important improvements, reducing uncertainties by 20-25% in each of the three background parameters.

Due to the expected low volume density of these objects, compared to traditional galaxy (full shape) power spectrum measurements, the performance gains achieved when combining them with density and velocity spectra measurements from KN are only modest. Peculiar velocity from sirens reduce modestly uncertainties in σ_8 and $\gamma \simeq 5\%$ and 16%, respectively). The addition of sirens as density tracers further decreases the uncertainties in γ and σ_8 by almost 10%.

1σ uncertainties in:	σ_8	γ	H_0	Ω_{m0}	Ω_{k0}		
Low $z \ (0 \le z \le 0.5)$							
DESI BGS gg	0.081	0.165	2.1	0.0095	0.171		
Rubin $3 \times 2pt$ g–k	0.079	0.137	2.1	0.0094	0.168		
Rubin 6 \times 2pt g–k–k	0.070	0.129	2.1	0.0093	0.167		
Rubin $6 \times 2pt + \iota$ prior	0.070	0.127	2.1	0.0093	0.166		
Rubin full-sky ("FS") $6\times 2 \mathrm{pt}$	0.057	0.105	1.6	0.0067	0.127		
Rubin BNS distances	-	-	0.12	0.24	0.41		
Rubin BNS dist + ι prior	-	-	0.097	0.18	0.31		
Rubin BNS dist $+ 6 \times 2pt$	0.069	0.128	0.085	0.0063	0.018		
Rubin as above $+ \iota$ prior	0.069	0.126	0.071	0.0062	0.015		
Rubin FS BNS dist $+6 \times 2pt$	0.056	0.104	0.060	0.0046	0.013		

Table 4. Marginalized absolute forecast uncertainties in each cosmological parameter for different instrument capabilities and 5 years of observation. The combination of clustering measurements using BNS to their distances shrinks H_0 uncertainties by around 30%, and constrains curvature to within 2%. Rubin full-sky assumes a second equivalent Rubin coverage in the north, for a total coverage of 80% of the sky. All results assume $k_{\rm max} = 0.10 \, h^{-1} {\rm Mpc}$. H_0 is given in units of km/s/Mpc.

Since sirens provide absolute distance indicators, they can also measure H_0 as is well known. The use of sirens simply as distance measurements results in better precision in H_0 than the 6 × 2pt method. Nevertheless, the combined measurements of spectra and distances is able to reduce the uncertainties in H_0 by a significant 30%. Moreover, such a combination improves the measurements on curvature by an order of magnitude, and reaches the level of 2% in Ω_{k0} .

Table 5 compares the performance of sirens with other probes, in particular with those of type Ia supernovae at the same redshift range, with supernovae and clustering at higher redshifts, and with the CMB constraints. At intermediate/high redshifts as discussed before we rely on DESI ELG forecasts for clustering, and the Roman Telescope for supernovae. Finally, we also show the CMB constraints. As can be seen, when combining clustering and distances, KN can outperform SN in H_0 due to their capabilities of measuring absolute distances. For the other cosmological parameters, SN perform similarly, with their higher number densities compensating for the extra degeneracies with H_0 . In any case it will be very useful to have two independent distance probes with comparable performance potential, as this will be of great use in highlighting any possible systematics.

Figure 4 illustrates the forecasts above by showing the 1σ and 2σ confidence regions for different parameter combinations. We depict different methods: $1 \times 2pt$ applied to galaxies as matter tracers, $3 \times 2pt$, $6 \times 2pt$, the use of the siren distances alone (the usual Hubble diagram fit), the combination of distances and $6 \times 2pt$ and Planck TTTEEE CMB results. We see that the influence of peculiar velocities is greater on γ and σ_8 , while in the other parameters ($\Omega_{m0}, \Omega_{\kappa0}, H_0$) the effect is minimal. The same occurs when comparing the $3 \times 2pt$ and $6 \times 2pt$ methods, where we see a modest improvement over γ and σ_8 (the 2D figure-of-merit in the $\sigma_8 - \gamma$ plane increases by 12%), while the other parameters have almost identical results.

It is also interesting to note that although siren distances alone produce highly degenerate constraints in the $\Omega_{m0} - \Omega_{k0}$ plane, this is distinctively different from the degeneracies on the clustering measurements, and this high complementarity result in very tight constraints in the background parame-

1σ uncertainties in:	σ_8	γ	H_0	Ω_{m0}	Ω_{k0}		
Low $z \ (0 \le z \le 0.5)$							
Rubin BNS distances	-	-	0.12	0.24	0.41		
Rubin BNS dist $+6{\times}2ptg{-}k{-}k$	0.069	0.128	0.085	0.0063	0.018		
Rubin SN distances	-	-	-	0.050	0.092		
Rubin SN dist $+6\!\times\!2\mathrm{pt}\mathrm{g}\!-\!\mathrm{s}\!-\!\mathrm{s}$	0.065	0.115	1.5	0.0068	0.013		
Intermediate and high z ($0.5 \le z \le 1.5$)							
DESI ELG gg	0.032	0.181	1.4	0.0057	0.066		
Roman SN distances	-	-	-	0.084	0.232		
DESI ELG gg + Roman SN	0.031	0.181	0.98	0.0048	0.022		
Last scattering surface							
Last scatte	ering su	ırface					

Table 5. Similar to Table 4, but comparing the performance with BNS with SN and with probes in other redshift ranges. SN forecasts assume a spectroscopic catalog in the same area of the sky. The higher number densities make supernovae better probes of relative distance and clustering, but KN are capable of measuring H_0 much better.

ters when combining both methods. This in turn leads also to modest improvements in the perturbation forecasts from the 6 × 2pt baseline, and to the 30% improvements in H_0 from the distance-based results mentioned above. We also remark the almost orthogonal degeneracies exhibited by both the distance method and the CMB, in the $H_0 - \Omega_{k0}$ plane, and by the 6 × 2pt method and the CMB in the $\sigma_8 - \gamma$ plane. The latter was first noticed by Quartin et al. (2022) in the case of supernovae, but the former is particular to standard sirens.

Finally, Figure 5 illustrates the effect of including external EM information on the inclination ι (as a 10° prior) on the cosmological results. The ι prior reflects this possible additional information from the electromagnetic counterpart, which helps to break the inclination angle–distance degeneracy, and consequently reduces the uncertainty of d_L . This induces improvements on the uncertainties of H_0 , Ω_{m0} , and $\Omega_{\kappa 0}$ found with the distance method of 19%, 25%, and 24%, respectively. The effect on the 6 × 2pt method alone is almost negligible, as discussed above, so this is not depicted. But still we get interesting improvements for the combined 6 × 2pt + distances case, as can be seen.

5 DISCUSSION

The next generation of ground-based GW detectors predict to detect gravitational waves from BNS at cosmological distances. Here we revisited their use as observables both in the traditional distance measurements as well as in the use as tracers of the velocity and density field. We consider that there are three main results of our work.

First, we propose a method to optimize the use of telescope time in BNS GW follow-ups, which uses the preliminary GW information on the distance to the source to allocate telescope time accordingly. This means longer integration times on further objects, and can result in a reduction of needed telescope time by 20–30% for the same number of kilonova observations in a given redshift range. This is relevant as by the next generation of GW facilities the number of observed standard sirens will be limited by telescope time.

Second, we find that using clustering in the full-shape



Figure 4. Marginalised contours for $\{\sigma_8, \gamma, \Omega_{m0}, \Omega_{m0}, h\}$, assuming $k_{max} = 0.1 \text{ h/Mpc}$ and $z_{max} = 0.5$. Results are separated according to the method used. Purple: $1 \times 2pt$ of galaxies; red: $3 \times 2pt$; green: $6 \times 2pt$; orange: siren distances only; blue: $6 \times 2pt$ + distances.

power spectrum method, using sirens both as velocity and density tracers, their expected low number density results for the sirens means that the gains are smaller than those obtained when using supernovae as tracers. Nevertheless, we obtain modest gains on the growth of structure precision when compared to the constraints arising purely from galaxy clustering. To wit, we obtain a 14% (24%) reduction on the uncertainties in σ_8 and γ .

Finally, we show that the combination of distance and clustering of standard sirens (using the full $6 \times 2pt$ method) results in massive improvements in the amount of background cosmological information that can be inferred. For the curvature density parameter Ω_{k0} , we find that percent-level precision is possible using only siren and galaxy data in $z \leq 0.5$. Given the limited available volume observed at lower redshifts, it is crucial to try to explore the data there to the fullest. For H_0 , in the same redshift range the inclusion of clustering information considerably improves the results arising from distance measurements only, resulting in 30% smaller uncertainties. Given the current efforts to better understand the reasons behind the Hubble tension, this is an important possible gain, and has the advantage of only relying on data in this redshift range. Moreover, sirens should then also outperform the precision on H_0 from either galaxy power spectrum measurements at higher redshifts or the CMB by an order of magnitude. And together with higher redshift measurements, we should be able to perform precise measurements of this parameter at different redshift ranges.

We also find that the inclusion of priors on the inclination of the orbital plane from electromagnetic observations have little impact on the clustering performance alone, but do lead to important improvements on the inferred KN distances and thus also on the combined $6 \times 2pt$ + distance approach, which fully explores the cosmological information from KN.

Our forecasts for σ_{γ} are less promising than what was found in Palmese & Kim (2021), to wit $\sigma_{\gamma} \sim 0.02 - 0.03$, even though they only assume a 3×2 pt method. This discrepancy seems to be due to the analysis of the EM follow-ups adopted, as our predictions for the capabilities of EM counterpart observations with future telescopes resulted in a much lower number of joint GW+EM measurements.

We remark that in our forecasts we have made a number of conservative assumptions. We have assumed that for each event the EM follow-up search area is 40 \deg^2 , instead of sim-



Figure 5. Possible improvements due to the addition of external information on the BNS inclination.

ply the smaller $\Delta\Omega$, and also required three epochs of followup observations for each KN. We also limited the follow-up at any given time to only 40% of the sky. The total accessible area in the sky at any given night is larger, specially if follow-ups could be conducted at lower altitudes. Finally, we restricted ourselves to the Einstein Telescope alone, but the Cosmic Explorer or other concurrent GW facilities may be operational at the same time, which would increase both the amount of events observed and the precision and the precision in the area determination. Together, these assumptions may result in similar constraints being obtained with less than the 5-year observational period here considered. Nevertheless, the still very large uncertainty in the rate of BNS mergers means that the number of years needed to reach our forecast precision could change significantly when this quantity is better understood in the course of the next observational runs of the LVC.

In any case, the $6 \times 2pt$ method applied to standard sirens can provide competitive clustering measurements when compared to SN, being better than the configuration labeled as conservative in Quartin et al. (2022), especially when it comes to H_0 . The results are also more sensitive to the number density of BNS, so if we improve further on our follow-up capabilities even more cosmological information can be obtained.

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DATA AVAILABILITY

The data underlying this article will be shared on reasonable request to the corresponding author.

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APPENDIX A: SHORT REVIEW ON DISTANCE POSTERIOR APPROXIMATIONS

In this appendix we briefly review the Cutler and Flanagan (CF) approximation Cutler & Flanagan (1994) for the posterior distribution on luminosity distance, which was rederived and corrected in Chassande-Mottin et al. (2019).

Under the following assumptions that:

• the noise is Gaussian, stationary and uncorrelated between detectors,

- we work at leading post-Newtonian order,
- the sky localisation is known from an EM counterpart,

• redshifted chirp mass and time to coalescence are very precisely measured and do not impact the measurement of distance, inclination ι , polarization and phase at coalescence,

the posterior distribution on distance and inclination,

marginalised over polarization and phase, can be written as:

$$p(d', v'|\vec{\theta}) \propto \begin{cases} \pi(d', v') \exp\left[-\frac{\rho_0^2 \sigma_d}{2} \left[\left(1 - \epsilon_d \cos 4\overline{\Psi}\right) \left(v - d'^{-1}v'\right)^2 + \left(1 + \epsilon_d \cos 4\overline{\Psi}\right) \left(\chi_+ - d'^{-1}\chi'_+\right)^2 \right] \right], & \epsilon_d \neq 0 \\ \pi(d', v') I_0(z_+) I_0(z_-) \\ \exp\left[-\frac{\rho_0^2 \sigma_d}{2} \left[\left(\chi_+^2 + v^2\right) + d'^{-2} \left(\chi_+'^2 + v'^2\right) \right] \right], & \epsilon_d = 0 \end{cases}$$
(A1)

where $\pi (d', v')$ is the prior, $d = d_L/d_L^{\text{true}}$ with d_L^{true} being the true luminosity distance, $v = \cos \iota$, $\chi_+ = 0.5 (1 + v^2)$, $z_{\pm} = \frac{\rho_0^2 \sigma_d}{2} (\chi_+ \pm v) (\chi'_+ \pm v')$, and $\overline{\Psi} = \Psi + \arctan \left[2\Theta^{+\times}/\Theta^{++} - \Theta^{\times\times}\right]/4$. I_0 is the modified Bessel function. Unprimed quantities refer to the true values, and primed ones to where the posterior is evaluated at. We made use of the tensor Θ^{AB} , defined for $A, B = +, \times$ over a network of detectors as

$$\Theta^{AB} = \frac{F_a^A F_b^B \delta_{ab} \int_0^\infty f^{-7/3} S_{n,a}^{-1}(f) \, df}{\int_0^\infty f^{-7/3} S_{n,\text{aver}}^{-1}(f) \, df}, \qquad (A2)$$

where $S_{n,a}$ is the power spectral density in detector a and $S_{n,\text{aver}}^{-1}$ is the average of the inverse of the power spectral densities over all detectors, in our case the three detectors that make ET. We define

$$\rho_0^2 = \frac{5G\mathcal{M}^{5/3}}{6\pi^{4/3}c^3} \int_0^\infty df \, \frac{f^{-7/3}}{S_{\text{n,aver}}(f)} \,, \tag{A3}$$

which can be seen as the SNR squared for an overhead faceon source in the case of a single detector. From the tensor Θ^{AB} , ϵ_d and σ_d are defined as

$$\sigma_d = 0.5 \,\mathrm{Tr}\Theta^{AB},\tag{A4}$$

$$\epsilon_d = \sqrt{\frac{2 \operatorname{Tr} (\Theta^{AB})^2}{(\operatorname{Tr} \Theta^{AB})^2 - 1}}.$$
 (A5)

The quantity ϵ_d takes values in the range from 0 to 1 and measures the detector's ability to distinguish the two polarization amplitudes. In the limit $\epsilon_d \rightarrow 1$ the detector can only estimate one of the polarization tensor components, which makes it difficult to estimate the parameters of the gravitational wave (see discussion in section IV.A of Chassande-Mottin et al. (2019)). σ_d is called the amplitude sensitivity and is related to the loudness of the signal in the network of detectors. large values being favorable. Figure A1 shows how σ_d and ϵ_d changes with the sky position for the ET detectors. Comparing this figure with Figures 1, 2 of Chassande-Mottin et al. (2019) we can observe the improvement in the sensitivity achieved by the ET over the detectors currently in operation.

We validate our results produced with the CF approximation by comparing with those found by a Bayesian analysis performed with the Bilby package (Ashton et al. 2019). For the Bayesian inference, we use emcee as our sampler, with 100 steps, 200 walkers, and the IMRPhenomPv2 waveform as our signal template. The MCMC was performed varying $\{\mathcal{M}_z, q = m_2/m_1, d_L, \varphi_c, \psi, \iota\}$ and fixing the remaining parameters at their injection values. The analysis shows that the results of the CF approximation are equivalent to the one obtained with MCMC when $\epsilon_d \leq 0.85$, which is in agreement with what was found in Chassande-Mottin et al. (2019)



Figure A1. Amplitude sensitivity σ_d and polarization sensitivity ϵ_d over the sky for the ET detectors.

Survey	Overlapping Fraction					
	$\delta t = 12$ h	$\deltat=48$ h				
LSST	$16.1\pm0.5\%$	0	0			
Mephisto	$36.1\pm1.2\%$	$3.8\pm0.5\%$	0			
WFST	$30.6\pm1.1\%$	$1.4\pm0.3\%$	0			
\mathbf{ZTF}	$69.6 \pm 1.5\%$	$63.7\pm1.4\%$	0			

Table B1. The fraction of overlapping events estimated generating100 random time-series for each configuration assuming kilonovaeare detectable δt hours after the event.

 $(\epsilon_d \lesssim 0.8)$. To correctly describe the uncertainty on d_L we imposed this cut in our catalog. The fraction of events with $\epsilon_d \lesssim 0.85$ is approximately 0.85 of the BNS mergers, but for the detected events (i.e., with $\rho_{\rm net} > 12$) this fraction is 0.98. Thus, the selection cut ϵ_d has an almost negligible impact the number of events in our analysis.

APPENDIX B: FRACTION OF OVERLAPPING EVENTS

We assessed the impact of overlapping events on the number of observed objects presented in Tables 2 and 3 by Monte-Carlo sampling a time series of events and assuming that the events are observed serially and that kilonovae are detectable during a total of δt hours after the event. The fraction of overlapping events are depicted in Table B1. The overlapping fraction variance was estimated to generate 100 random time series for each configuration. The number of expected kilonovae would be only strongly affected in all surveys if



Figure C1. Dependency of the $3 \times 2pt$ method on the standard candle distance uncertainty. The Figure of Merit (FoM) used is the inverse area of the ellipses in the $\sigma_8 - \gamma$ plane. Solid (dashed) lines are the results for our BNS (SN) forecasts. Also shown are dotted gray lines with example power laws that roughly fit the BNS and SN curves for 0.2 < z < 0.4. As can be seen, the low number density of BNS means that the influence of P_{vv} is smaller and thus its contribution to $3 \times 2pt$ and $6 \times 2pt$ are also smaller, which explains the lower sensitivity to σ_{int} . The uncertainties due to P_{vv} alone scale instead roughly as σ_{int}^{-2} .

kilonovae happened to be detectable during only a fraction of a day. Only ZTF would still be significantly affected if kilonovae lasted only for 24 hours.

In our simulations we did not find any scenario where an event could not be observed if kilonovae lasted at least two days. The detectability and best strategy to observe kilonovae have been already studied (see, for instance, Cowperthwaite & Berger 2015; Mochkovitch et al. 2021; Chase et al. 2022). The exact details for the best observation strategy go beyond the scope of our work; still, we can neglect the impact of overlapping events for two reasons. Due to the slow decaying of kilonovae in certain magnitudes, the detectability of a large fraction of the sample will be larger than two days. Secondly, our simulation assumed that the observation strategy followed the events serially; any strategy optimization will reduce the overlapping fraction presented in Table B1.

APPENDIX C: DEPENDENCE ON THE DISTANCE MEASUREMENT UNCERTAINTIES

The inclusion of information on the inclination of the BNS orbit, which can be independently obtained from EM observations, is capable of improving the precision of the sirens distance measurements. A 10° constraint on this inclination leads to improvements on the relative distance of a given source by between 15–25%. Nevertheless as discussed in Section 4, this leads only to very minor improvements on the clustering results from BNS. This is also in opposition to what was found for the case of supernovae by Amendola & Quartin (2021), where a much stronger dependency of the $3 \times 2pt$ and $6 \times 2pt$ methods on σ_{int} was found. And neglecting the small extra lensing scatter, $\sigma_{int} = (5/\ln 10)\sigma d_L/d_L$.

The reason for this small effect is basically the low expected number density of BNS, when compared to that of

LSST supernovae. To understand this, we remark first that the uncertainties due to $P_{\rm vv}$ alone scale instead roughly as $\sigma_{\rm int}^{-2}$. However, the weight of the velocities on the full $3 \times 2pt$ and $6 \times 2pt$ methods depend on the number density of velocity tracers. The low number density of BNS means that the influence of $P_{\rm vv}$ is small and thus its contribution to $3 \times 2 {\rm pt}$ and $6 \times 2pt$ are also smaller, which explains the lower sensitivity to σ_{int} . Figure C1 illustrates this, and compares the inverse area of the ellipses in the $\sigma_8 - \gamma$ plane, which is our Figure of Merit (FoM), as a function of σ_{int} , assumed constant and the same for all sources (contrary to our baseline KN analysis in Section 4). For illustration purposes we normalize all results to the one obtained for $\sigma_{int} = 0.1$ mag. For very low redshifts both SN and KN clustering exhibit small dependencies on σ_{int} , but for 0.2 < z < 0.4 the SN Figure of Merit scale as $\sigma_{\rm int}^{-0.27}$, whereas the KN clustering one only as $\sigma_{\rm int}^{-0.06}$.